

# Mode of Multivariate Gaussian Distribution

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May 11, 2016

*Proof.* The pdf of  $\mathbf{x} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  is given by

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\} \quad (1)$$

where  $\mathbf{x}$  is  $d$  dimensional.

To find the mode i.e. the maximum of the Gaussian distribution, we differentiate the pdf with respect to  $\mathbf{x}$  and equate it to 0 to find the critical point where the function is maximum or minimum and then we use the second derivative test to ascertain that the function is maximized at that point. *(The second derivative test for a function of more than one variable generalizes to a test based on the eigenvalues of the function's Hessian matrix at the critical point. Assuming that all the second order partial derivatives of the function are continuous on a neighbourhood of a critical point  $\mathbf{x}$ , then if the eigenvalues of the Hessian at  $\mathbf{x}$  are all **positive**, then  $\mathbf{x}$  is a local **minimum**, if the eigenvalues are all **negative**, then  $\mathbf{x}$  is a local **maximum**, and if some are positive and some negative, then the point is a **saddle point**. If the Hessian matrix is **singular**, then the second derivative test is **inconclusive**.)*

Differentiating (1) with respect to  $\mathbf{x}$ ,

$$\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}} = \frac{1}{(2\pi)^{D/2}} \frac{1}{|\boldsymbol{\Sigma}|^{1/2}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\} \{ -\boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \} \quad (2)$$

$$= -f(\mathbf{x}) \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \quad (3)$$

Equating (2) to zero, we get the critical point

$$\mathbf{x} = \boldsymbol{\mu} \quad (4)$$

To verify that the pdf is maximized at (4), we evaluate the Hessian matrix at  $\boldsymbol{\mu}$  and check that it is negative definite.

Differentiating (3) with respect to  $\mathbf{x}$ ,

$$\frac{\partial^2 f(\mathbf{x})}{\partial \mathbf{x} \partial \mathbf{x}^T} = \frac{\partial}{\partial \mathbf{x}} \left( \frac{\partial f(\mathbf{x})}{\partial \mathbf{x}^T} \right) \quad (5)$$

$$= \frac{\partial}{\partial \mathbf{x}} \left( -f(\mathbf{x}) (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} \right) \quad (6)$$

$$= f(\mathbf{x}) \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} - f(\mathbf{x}) \boldsymbol{\Sigma}^{-1} \quad (7)$$

$$= f(\mathbf{x}) \left( \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} - \boldsymbol{\Sigma}^{-1} \right) \quad (8)$$

Evaluating (8) at  $\mathbf{x} = \boldsymbol{\mu}$  we get,

$$\left. \frac{\partial^2 f(\mathbf{x})}{\partial \mathbf{x} \partial \mathbf{x}^T} \right|_{\mathbf{x}=\boldsymbol{\mu}} = -f(\boldsymbol{\mu}) \boldsymbol{\Sigma}^{-1} \quad (9)$$

We know that the covariance matrix  $\boldsymbol{\Sigma}$  is positive definite, hence its inverse  $\boldsymbol{\Sigma}^{-1}$  is also positive definite. *(since if  $\lambda$  is an eigenvalue of  $\boldsymbol{\Sigma}$ , then  $1/\lambda$  is the corresponding eigenvalue of  $\boldsymbol{\Sigma}^{-1}$ . As  $\lambda$  is positive, so is  $1/\lambda$  which implies that all the eigenvalues of  $\boldsymbol{\Sigma}^{-1}$  are positive.)*

We know that  $f(\boldsymbol{\mu})$  is positive everywhere, since the pdf is positive, and hence  $-\boldsymbol{\Sigma}^{-1}$  is negative definite which proves that the pdf is maximized at  $\mathbf{x} = \boldsymbol{\mu}$ . *(if  $\mathbf{Ax} = \lambda\mathbf{x}$ , then  $-\mathbf{Ax} = -\lambda\mathbf{x}$ )*

**Hence the mode of the multivariate Gaussian distribution is  $\boldsymbol{\mu}$ .** □

## References

[1] Christopher M. Bishop. *Pattern Recognition and Machine Learning*. **Exercise 1.9**