

# Estimation

How to get information from samples

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# ESTIMATION

1 THE PROBLEM : Estimate a System Parameter

2 MODELLING : Stochastic Model

3 Synthesis

# PROTOCOL VALIDATION

Typical random access protocol to a common channel (CSMA family)

Protocol-Send ( $M$ )

**while Message is not sent**

    Send( $M$ )

**if Collision**

$W = \text{Random}(I_n)$

$n = n + 1$

$I_{n+1} = g(n, I_n)$

        Wait ( $W$ )

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  while Message is not sent
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    if Collision
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      n = n + 1
      In+1 = g(n, In)
      Wait (W)
```

What should be the *amount of time* ?

## Protocol dimensioning

Waiting time :

- ▶ Random
- ▶ Uniformly distributed on an interval  $[0, I_n]$
- ▶ Length of the interval depends on the number of collisions
- ▶ Adaptive scheme  $I_{n+1} = 2 \times I_n$ ,
- ▶  $I_0$  fixed, characteristic of the protocol

# PROTOCOL HISTORY

University of Hawai'i 1970

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Ancestor of CSMA/CD (ethernet), CSMA/CA (WiFi)...

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## Decision

How could you conclude on the validity of the implementation of the protocol ?

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## SAMPLES

**The observations :** a sequence of  $n$  waiting times.

We suppose that the experiment has been well driven (no outliers, ...) and the observed values denoted by

$$\{x_1, \dots, x_n\}.$$

**The stochastic model :** the observations are considered to be realizations of independent random variables with the same probability law  $F$

$$\{X_1, \dots, X_n\}$$

Question : What could be said on  $F$  from the observations  $\{x_1, \dots, x_n\}$  ?

### A priori knowledge :

- The shape of the law is known and depends on some parameter(s) unknown : **parametric estimation**

ex:  $X$  follows a uniform distribution on  $[0, \theta[$  with  $\theta$  unknown

$$F_\theta(x) = \begin{cases} 0 & x \leq 0; \\ \frac{1}{\theta}x & x \in [0, \theta[; \\ 1 & \theta \leq x. \end{cases}$$

- The shape of the law is unknown and some parameters are under study (expectation, variance, moments,...) : **non-parametric estimation**

## BASIC CONCEPTS

- ▶ A **statistic** is a function of the observations :  $t_n(x_1, \dots, x_n)$ , it usually summarize some parameter of the distribution.
- ▶ An **estimator** is a random variable  $T_n = t(X_1, \dots, X_n)$  (model of the statistic)

### Example

$t(x_1, \dots, x_n) = \max_{1 \leq i \leq n} x_i$  is a statistic on the samples;

and  $T_n = \max_{1 \leq i \leq n} X_i$  the corresponding estimator.

Law of  $T_n$  under the hypothesis  $X_i$  uniformly distributed on  $[0, \theta]$  :

$$F_n^\theta(x) = \mathbb{P}\left(\max_{1 \leq i \leq n} X_i \leq x\right) = \frac{1}{\theta^n} x^n \text{ using independence and uniformity law of } X_i;$$

and density

$$f_n^\theta(x) = \frac{1}{\theta^n} n x^{n-1}.$$

# BIAS

An estimator  $T_n$  of some parameter  $\theta$  is **unbiased** iff

$$\mathbb{E}T_n = \theta$$

## Example

Bias of estimator  $T_n$

$$\mathbb{E}T_n = \int_0^\theta xf_n^\theta(x)dx = \int_0^\theta x \frac{1}{\theta^n} nx^{n-1} dx = \frac{1}{\theta^n} n \left[ \frac{x^{n+1}}{n+1} \right]_0^\theta = \frac{n}{n+1}\theta \neq \theta.$$

$T_n$  is a biased estimator, on average it underestimate  $\theta$  on average. But, for large samples the bias decreases to 0.

$\lim_{n \rightarrow +\infty} \mathbb{E}T_n = \theta$   $T_n$  is asymptotically unbiased (or consistent)

To compensate the bias,

$$T'_n = \frac{n+1}{n} T_n,$$

which is unbiased.

# RISK

The quality  $T_n$  of an estimator is evaluated by the **risk** function

$$R(T_n) = \mathbb{E}(T_n - \theta)^2 = \mathbb{E}(T_n - \mathbb{E}T_n)^2 + (\mathbb{E}T_n - \theta)^2 = \text{Var}T_n + (\mathbb{E}T_n - \theta)^2$$

- ▶  $\text{Var}T_n$  : the concentration of the distribution
- ▶  $(\mathbb{E}T_n - \theta)^2$  : impact of the bias.

For an unbiased estimator

$$R(T_n) = \text{Var}T_n$$

## Risk of $T'_n$

For the example

$$\begin{aligned}\mathbb{V}ar T_n &= \mathbb{E}(T_n - \mathbb{E}T_n)^2 = \mathbb{E}T_n^2 - (\mathbb{E}T_n)^2 = \int_0^\theta x^2 f_n^\theta(x) dx - \left(\frac{n}{n+1}\theta\right)^2 \\ &= \left(\frac{n}{n+2} - \frac{n^2}{(n+1)^2}\right)\theta^2\end{aligned}$$

$$\begin{aligned}\mathbb{V}ar T'_n &= \mathbb{V}ar \frac{n+1}{n} T_n = \frac{(n+1)^2}{n^2} \mathbb{V}ar T_n = \frac{(n+1)^2}{n^2} \left(\frac{n}{n+2} - \frac{n^2}{(n+1)^2}\right) \theta^2 \\ &= \frac{1}{n(n+2)} \theta^2\end{aligned}$$

## ANOTHER ESTIMATOR

Consider

$$U_n = \frac{2}{n} \sum_{i=1}^n X_i.$$

$$\mathbb{E}U_n = \frac{2}{n} \sum_{i=1}^n \mathbb{E}X_i = \frac{2}{n} \sum_{i=1}^n \frac{\theta}{2} = \theta.$$

$U_n$  is an unbiased estimator of  $\theta$ .

### Risk of $U_n$

For the example

$$\begin{aligned}\mathbb{V}ar U_n &= \mathbb{V}ar \left( \frac{2}{n} \sum_{i=1}^n X_i \right); \\ &= \left( \frac{2}{n} \right)^2 \sum_{i=1}^n \mathbb{V}ar X_i \text{ because of the independence of } X_i; \\ &= \frac{4}{n^2} n \frac{\theta^2}{12} = \frac{\theta^2}{3n}.\end{aligned}$$

The risk associated to  $U_n$  is much larger than the risk of  $T'_n$  so we will prefer  $T'_n$ .

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# SYNTHESIS

Observations are realizations of a same experiment in the same conditions

## Parametric Estimation

- ▶ Assumptions : the results of the experiments are modeled by a sequence of independent identically distributed with a **known probability law**;
- ▶ An **estimator** is a random variable function of the modeled random variables, expecting to approximate a parameter of the law;
- ▶ The **bias** of an estimator is the expected difference between the estimator and the real value  $\mathbb{E}_\theta(T_n - \theta)$
- ▶ The **risk** is the expectation of the error around the real value  $\mathbb{E}_\theta(T_n - \theta)^2$

## Non-Parametric Estimation : estimation of the mean

- ▶ Assumptions : the results of the experiments are modeled by a sequence of independent identically distributed with a **unknown probability law**, but with some properties mean  $m$  and variance  $\sigma^2$ ;
- ▶ The average  $\frac{1}{n} \sum X_i$  is an unbaised estimator of  $m$ ;
- ▶ The error  $\frac{1}{n} \sum X_i - m$  converges almost surely to 0 (strong law of large numbers)
- ▶ The law of errors satisfies the central limit theorem (central limit theorem and confidence intervals)

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{X_i - m}{\sigma} \xrightarrow{d} \mathcal{N}(0, 1)$$