ARCH models for Value at Risk forecasting in Latin American stock and Forex markets

Alejandro Cermeño

Pontificia Universidad Católica del Perú

October 19, 2021

Estructura

- Introduction
- 2 Literature Review
- Methodology
- 4 Empirical evidence
 - Data
 - Results
- 5 Conclusions (for stock markets)
- 6 Agenda

Introduction

Motivation

An adequate VaR forecast can help agents and institutions to reduce financial uncertainty, prevent liquidity problems, and optimize provisioning mechanisms.

Introduction

Goals

Identify the optimal ARCH-type model to forecast the one-day ahead Value at Risk (VaR) at 1% and 5% confidence level for a group of Latin American countrys

Literature Review

Stylized facts of financial returns: Franses and van Dijk (2000), Engle (2004) and Teräsvirta (2009)

- Clustering of extreme returns
- Non normality in distribution
- Leverage effect: high volatility after extreme negative returns
- ullet Long memory in r_t^δ or other transformations

Literature Review

- Engle (1982) proposes the ARCH model and Bollerslev (1986) generalizes it to GARCH.
- Arch-type models have been widely developed. Key references are Glosten el at. (1993), Ding et al. (1993) and Baillie et al. (1996).
- Volatility as a key input in Value at Risk: Angelidis et al. (2004) for the major stock indices, Fan et al. (2008) for crude oil price, Liu and Hung (2010) for S&P 100, Brooks and Persand (2003), Ardia and Hoogerheide (2014), among others.

Methodology

A set of ARCH models is used to forecast the one-day ahead VaR at 1% and 5% confidence level for a group of Latin American countrys.

- An ARCH-type model parameters are estimated for a given rolling-window of 250 observations.
- 2 With the fitted model the one-day-ahead volatility.
- VaR is computed from forecasted volatility and distribution parameters.
- The accuracy of the VaR forecasts is analyzed by backtest tests.

Additionally, as a previous step, an in-sample analysis is performed where the models are specified and its parameters estimated with the full available series.

Methodology

A complete ARCH model is divided into three components:

- a mean model: zero, constant, AR(1)
- a volatility process: ARCH, GARCH, GJR, FIGARCH, APARCH.
- ullet a distribution for the standardized residuals: Normal, \mathcal{S} -Student, Skeweed, and generalized error distribution (GED).

The ARCH model of Engle (1982)

$$\sigma_t^2 = \omega + \sum_{\rho=1}^P \alpha_\rho \epsilon_{t-P}^2 \tag{1}$$

The GARCH model of Bollerslev (1986)

$$\sigma_t^2 = \omega + \sum_{i=1}^q \alpha_i \epsilon_{t-i}^2 + \sum_{j=1}^p \beta_j \sigma_{t-j}^2$$
 (2)

The GJR (1993) model

$$\sigma_t^2 = \omega + \sum_{i=1}^q \alpha_i \varepsilon_{t-i}^2 + \sum_{i=1}^q \gamma_i I_{[\varepsilon_{t-i} < 0]} \varepsilon_{t-i}^2 + \sum_{j=1}^p \beta_j \sigma_{t-j}^2$$
 (3)

Methodology ARCH models

The FIGARCH of Baillie et al. (1986)

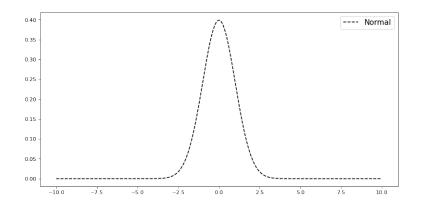
$$\sigma_t^2 = \omega + \left[1 - \beta L - \phi L (1 - L)^d\right] \varepsilon_t^2 + \beta \sigma_{t-1}^2 \tag{4}$$

The APARCH of Ding et al. (1993)

$$\sigma_t^{\delta} = \omega + \sum_{i=1}^{q} \alpha_i \left(|\varepsilon_{t-i}| - \gamma_i I_{[o \ge i]} \varepsilon_{t-i} \right)^{\delta} + \sum_{i=1}^{p} \beta_i \sigma_{t-i}^{\delta}$$
 (5)

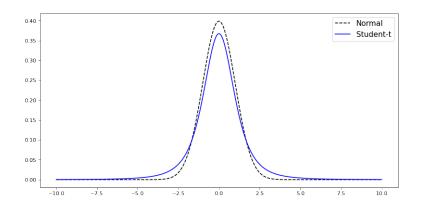
Methodology

Distributions

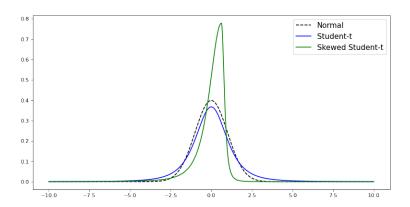


${\sf Methodology}$

Distributions

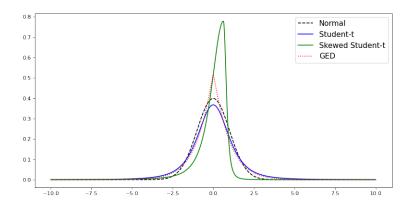


Methodology Distributions



${\sf Methodology}$

Distributions



Value at Risk and backtest

$$\widehat{VaR}_{t+1}^{\alpha} = -\hat{\mu}_{t+1} - \widehat{\sigma}_{t+1} q_{\alpha} \tag{6}$$

backtest procedures:

- Unconditional coverage (UC) test of Kupiec (1995)
- Conditional Coverage (CC) test of Christoffersen (1998)
- Dynamic quantile (DQ) test of Engle and Manganelli (2004).

Empirical Evidence: Data

- Countries: Argentina, Brazil, Chile, Colombia, Mexico and Peru
- Markets: stocks and Forex
- Frecuency: daily
- Source: Bloomberg Financial Data
- Returns are calculated as $y_t = 100 * [\log{(p_t)} \log{(p_{t-1})}]$

Table 1: Descriptive Statistics for Stock and Forex Markets Returns

| Country | Security ID | Start Date | End Date | Obs. | Mean | Std, | Min | Max | Skew | Kurt | | |
|-----------|-------------|------------|------------|------|------|------|--------|-------|-------|--------|--|--|
| (a) Stock | | | | | | | | | | | | |
| Argentina | MERVAL | 1991-12-26 | 2021-08-31 | 7744 | 0.06 | 2.30 | -47.69 | 16.12 | -1.43 | 27.89 | | |
| Brasil | IBOV | 1995-03-16 | 2021-08-31 | 6904 | 0.05 | 1.96 | -17.23 | 28.82 | 0.09 | 14.05 | | |
| Chile | IPSA | 1990-08-09 | 2021-08-31 | 8104 | 0.05 | 1.14 | -15.22 | 11.80 | -0.32 | 12.83 | | |
| Colombia | IGBC | 2001-07-26 | 2021-08-31 | 5244 | 0.05 | 1.14 | -11.05 | 14.69 | -0.18 | 15.73 | | |
| México | MEXBOL | 1994-03-31 | 2021-08-31 | 7154 | 0.04 | 1.40 | -14.31 | 12.15 | 0.03 | 7.56 | | |
| Perú | SPBLPGPT | 2002-02-07 | 2021-08-31 | 5104 | 0.05 | 1.35 | -13.29 | 12.82 | -0.59 | 12.69 | | |
| (b) Forex | | | | | | | | | | | | |
| Argentina | ARS | 2014-03-06 | 2021-08-31 | 1954 | 0.13 | 1.14 | -6.03 | 30.80 | 12.32 | 294.71 | | |
| Brasil | BRL | 1999-06-03 | 2021-08-31 | 5804 | 0.02 | 1.04 | -10.34 | 7.11 | 0.07 | 5.90 | | |
| Chile | CLP | 1990-01-18 | 2021-08-31 | 8249 | 0.01 | 0.60 | -4.33 | 4.68 | 0.25 | 7.54 | | |
| Colombia | COP | 1992-09-03 | 2021-08-31 | 7564 | 0.02 | 0.66 | -7.60 | 6.02 | 0.17 | 10.31 | | |
| México | MXN | 1996-05-09 | 2021-08-31 | 6604 | 0.01 | 0.71 | -6.65 | 7.98 | 0.85 | 11.43 | | |
| Perú | PEN | 1995-05-25 | 2021-08-31 | 6854 | 0.01 | 0.30 | -2.85 | 3.55 | 0.08 | 14.54 | | |

Table 2a: Estimated Parameters for daily Latin American stock Markets Return

| Model | μ | ω | α | β | γ | δ | ν | λ | BIC | log-lik |
|-------------------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---|-------------|-------------|
| | | | | Argentin | a (MERVA | .L) | | | | |
| μ-APARCH-GED | 0.0462ª | 0.1573 ^a | 0.1120 ^a | 0.8564ª | 0.2148ª | 1.9847ª | 1.0919 ^a | | 31888.194 | -15912.755 |
| μ -GJR-GED | 0.0462 ^a | 0.1592 ^a | 0.0692 ^a | 0.8558 ^a | 0.0955ª | | 1.0921 ^a | | 31879,24598 | -15912,758 |
| | | | | Brazi | l (IBOV) | | | | | |
| μ-APARCH-GED | 0.0503ª | 0.0701 ^a | 0.0759 ^a | 0.8980ª | 0.4245 ^a | 1.7129ª | 1.3592ª | | 26273.472 | -13105.796 |
| μ -GJR-GED | 0.0510 ^a | 0.0814 ^a | 0.0275 ^a | 0.8933 ^a | 0.1067ª | | 1.3594ª | | 26266,91876 | -13106,939 |
| | | | | Chile | e (IPSA) | | | | | |
| μ-APARCH-GED | 0.0304ª | 0.0295 ^a | 0.1411 ^a | 0.8475 ^a | 0.1485ª | 1.7939ª | 1.3146ª | | 21853.744 | -10895.372 |
| $\mu	ext{-}FIGARCH	ext{-}GED$ | 0.0372 ^a | 0.0400 ^a | 0.2548 ^a | 0.5357 ^a | | 0.4905 ^a | 1.3183 ^a | | 21845,49687 | -10895,7481 |
| | | | | Colomi | bia (IGBC) | ı | | | | |
| APARCH-GED | | 0.0045 ^a | 0.0943 ^a | 0.9203 ^a | 0.2276 ^a | 0.3190 ^a | 1.0853 ^a | | 5889.822 | -2919.216 |
| μ -APARCH-GED | -0.0000 | 0.0000 | 0.1932 | 0.8068 | 0.0230 | 1.8244 | 1.0100 | | 6023.314 | -2981.680 |
| | | | | Mexico | (MEXBOL | .) | | | | |
| μ-APARCH-GED | 0.0316 ^a | 0.0160 ^a | 0.0786 ^a | 0.9214 ^a | 0.4290 ^a | 1.4718 ^a | 1.3037 ^a | | 22136.973 | -11037.422 |
| μ -GJR-GED | 0.0331 ^a | 0.0167 ^a | 0.0312 ^a | 0.9131^{a} | 0.0988ª | | 1.3015 ^a | | 22138,60164 | -11042,674 |
| | | | | Peru (S | PBLPGPT | .) | | | | |
| μ -FIGARCH-GED | 0.0397 ^a | 0.0846 ^a | 0.0867 | 0.3316 ^a | | 0.4235 ^a | 1.1342 ^a | | 14482.021 | -7215.397 |
| μ-APARCH-GED | 0.0375 ^b | 0.0394 ^a | 0.1473 ^a | 0.8337 ^a | 0.0743 ^b | 1.9825 ^a | 1.1281 ^a | | 14524,71567 | -7232,475 |

a, b, c denote significance level at 1%, 5% and 10% respectively

Table 2a: Estimated Parameters for daily Latin American Forex Markets Return

| Model | μ | ω | α | β | γ | δ | ν | λ | BIC | log-lik |
|--|-------------------------------|--|--|--|---|--|--|---------------------|------------------------|-------------------------|
| | | | | Arge | ntina (ARS |) | | | | |
| μ -FIGARCH-sk ${\cal S}$ μ -FIGARCH- ${\cal S}$ | 0.0539 ^a 0.0501 | 0.0003 0.0003 | 0.0541 0.0768 | 0.6814 ^b 0.6077 ^a | | 0.8918 ^a 0.8463 ^a | 3.0927 ^a 3.0500 ^a | 0.0909 ^b | 500.480 505,101 | -223.718 -229,8176 |
| | | | | Bra | azil (BRL) | | | | | |
| μ-APARCH-GED GJR-GED | 0.0022 | 0.0139ª 0.0101ª | 0.0968 ^a 0.1438 ^a | 0.9032 ^a 0.8940 ^a | -0.3064ª -0.0872ª | 1.4421 ª | 1.3863ª 1.3768ª | | 14806.536 14800,927 | -7372.936 -7378,798 |
| | | | | Ch | ile (CLP) | | | | | |
| APARCH- $\mathcal S$ APARCH-sk $\mathcal S$ | | 0.0270 ^a 0.0177 ^a | 0.1300° 0.0884° | 0.8686ª 0.9099ª | -0.4153ª 0.2023ª | 0.2499 ^a 0.3396 ^a | 2.8537 ^a 3.0470 ^a | 0.0971 ^a | 8009.367 8837.625 | -3977.630 -4387.2502 |
| | | | | Colo | mbia (COP |) | | | | |
| μ -FIGARCH-GED FIGARCH-GED | 0.0000 | 0.0024 ^c 0.0024 ^c | 0.2664 ^a 0.2662 ^a | 0.5728 ^a 0.5730 ^a | | 0.4672 ^a 0.4675 ^a | 1.0100 ^a 1.0100 ^a | | 10252.889 10244,165 | -5099.651 -5099,754 |
| | | | | Mex | ico (MXN) | | | | | |
| μ -APARCH-sk ${\cal S}$ μ -GJR-GED | 0.0117ª 0.0331ª | 0.0120ª 0.0167ª | 0.0966 ^a 0.0312 ^a | 0.9034 ^a 0.9131 ^a | -0.4849 ^a 0.0988 ^a | 1.1410ª | 7.3379 ^a 1.3015 ^a | 0.1336ª | 11378.144 22138,601 | -5653.890 -11042,674 |
| | | | | Pe | ru (PEN) | | | | | |
| μ -FIGARCH-GED μ -APARCH-GED | -0.0000 -0.0000 | 0.0002 0.0004 | 0.1766 ^a 0.1618 ^a | 0.6098 ^a 0.8382 ^a | -0.1393 | 0.6468 ^a 2.0107 | 1.0100° 1.0100° | | -2873.629 -2844,943 | 1463.312 1453,385 |

a, b, c denote significance level at $1\,\%,\,5\,\%$ and $10\,\%$ respectively

Empirical Evidence

Results: Out-of-sample analysis

Figure: One-step ahead VaR forecast for Peru (SPBLPGPT)

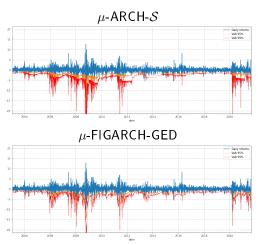


Table 3: Accuracy of VaR predictions for Stock and Forex Markets Returns

| | VaR 5% | risk level | | | | VaR 1% risk | level | | |
|----------------------|---------------------------|------------|---------|-----------|---------------------|-------------|---------|---------|--------|
| Model | % Failures UC CC DQ Model | | | | Model | % Failures | UC | CC | DQ |
| | | | | (a) Stock | Market Returns | | | | |
| | | | | Argenti | na (MERVAL) | | | | |
| AR(1)-GARCH-S | 0.0653 | 0.0373* | 0.0078* | 0.0000* | AR(1)-FIGARCH-GED | 0,0179 | 0.0272* | 0.0518 | 0.0000 |
| AR(1)-GJR-√ | 0.0674 | 0.0185* | 0.0061* | 0.0011* | GARCH-√ | 0,0189 | 0.0133* | 0.0331* | 0.0000 |
| | | | | Bra | zil (IBOV) | | | | |
| μ-GJR-GED | 0.055285 | 0.0749 | 0.0066* | 0.0155* | APARCH-S | 0.0167 | 0.0000* | 0.0000* | 0.0000 |
| μ-APARCH-GED | 0.055286 | 0.0748 | 0.0067* | 0.0155* | APARCH-GED | 0.0167 | 0.0000* | 0.0000* | 0.0000 |
| | | | | Chi | ile (IPSA) | | | | |
| GJR-S | 0,0568 | 0.0058* | 0.0000* | 0.0000* | FIGARCH-sk <i>S</i> | 0,0166 | 0.0000* | 0.0000* | 0.0000 |
| APARCH- $\mathcal S$ | 0,0569 | 0.0058* | 0.0000* | 0.0000* | GARCH-GED | 0,0166 | 0.0000* | 0.0000* | 0.0000 |
| | | | | Mexico | (MEXBOL) | | | | |
| FIGARCH-S | 0, 044 | 0.3030 | 0.0047* | 0.0121* | GJR- <i>N</i> | 0,0146 | 0.0690 | 0.0003* | 0.0000 |
| FIGARCH-GED | 0, 04 5 | 0.3030 | 0.0003* | 0.0002* | APARCH-√ | 0,0147 | 0.0690 | 0.0002* | 0.0000 |
| | | | | Peru (| SPBLPGPT) | | | | |
| FIGARCH-GED | 0,0569 | 0.0312* | 0.0000* | 0.0000* | FIGARCH-skS | 0,0165* | 0.0000* | 0.0000* | 0.0000 |
| μ -GARCH- S | 0,0571 | 0.0266* | 0.0000* | 0.0000* | FIGARCH-GED | 0,0173* | 0.0000* | 0.0000* | 0.0000 |

(b) Forex Market Returns

Conclusions (for stock markets)

- In general, FIGARCH volatility process and leptokurtic distributions are able to produce better one-step-ahead VaR forecasts
- The models that best fit the full series in-sample are not necessarily the ones that obtain the most accurate VaR forecasts out-of-sample
- The models producing the most accurate forecasts vary by market and country.

Agenda'

- Obtain the results also for the Forex market
- Compare the results for the stock market and the Forex market.

References

- Baillie, R., T. Bollerslev, and H. O. Mikkelsen (1996). Fractionally integrated generalized autoregressive conditional heteroskedasticity. Journal of Econometrics 74 (1), 3-30.
- Bollerslev, T. (1986). Generalized autoregressive conditional heteroskedasticity. Journal of Econometrics 31 (3), 307–327.
- Christoffersen, P. (1998). Evaluating interval forecasts. International Economic Review 39 (4), 841–62.
- Ding, Z., C. Granger, and R. Engle (1993). A long memory property of stock market returns and a new model. Journal of Empirical Finance 1 (1), 83–106.
- Engle, R. (1982). Autoregressive conditional heteroscedasticity with estimates of the variance of united kingdom inflation. Econometrica 50 (4), 987–1007.

References

- Engle, R. F. and S. Manganelli (2004). Caviar: Conditional autoregressive value at risk by regression quantiles. Journal of Business Economic Statistics 22 (4), 367–381.
- Glosten, L. R., R. Jagannathan, and D. E. Runkle (1993). On the relation between the expected value and the volatility of the nominal excess return on stocks. The Journal of Finance 48 (5), 1779–1801.
- Kupiec, P. H. (1995). Techniques for verifying the accuracy of risk measurement models. The Journal of Derivatives 3 (2), 73–84.
- Nelson, D. B. (1991). Conditional heteroskedasticity in asset returns:
 A new approach. Econometrica 59 (2), 347–70.