

Introduction to statistics

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Problems on Confidence intervals

Summary

- $(1 - \alpha) \cdot 100\%$ confidence intervals for the mean μ :

case 1:

when

1. $X \rightarrow N(\mu, \sigma^2)$
2. we **know** σ^2 then CI for the estimation of $\hat{\mu}$

CI for μ :

$$(l, u) = (\bar{x} - z_{\alpha/2} \sigma / \sqrt{n}, \bar{x} + z_{\alpha/2} \sigma / \sqrt{n})$$

case 2:

when

1. $X \rightarrow N(\mu, \sigma^2)$
2. and we **don't know** σ^2 then

CI for μ :

$$(l, u) = (\bar{x} - t_{\alpha/2, n-1} s / \sqrt{n}, \bar{x} + t_{\alpha/2, n-1} s / \sqrt{n})$$

- $(1 - \alpha) \cdot 100\%$ confidence intervals for the proportion p :

case 3:

when

1. $X \rightarrow \text{Bernoulli}(p)$ 2 and np and $n(1 - p) > 5$

CI for p :

$$(l, u) = (\bar{x} - z_{\alpha/2} \left[\frac{\bar{x}(1 - \bar{x})}{n} \right]^{1/2}, \bar{x} + z_{\alpha/2} \left[\frac{\bar{x}(1 - \bar{x})}{n} \right]^{1/2})$$

- $(1 - \alpha) \cdot 100\%$ confidence intervals for the estimation of the variance σ^2 :

case 4:

when

$$1. X \rightarrow N(\mu, \sigma^2)$$

CI for σ :

$$(l, u) = \left(\frac{s^2(n-1)}{\chi_{\alpha/2, n-1}^2}, \frac{s^2(n-1)}{\chi_{1-\alpha/2, n-1}^2} \right)$$

Remember for the formulas:

- In cases 1 and 3: $z_{\alpha/2} = \text{qnorm}(1-\alpha/2)$
- In case 2: $t_{\alpha/2, n-1} = \text{qt}(1-\alpha/2, n-1)$
- In case 4: $\chi_{\alpha/2, n-1}^2 = \text{qchisq}(1-\alpha/2, n-1)$

Problem 1

Consider:

- $n = 5$
 - $\alpha = 1 - 0.95 = 0.05$
 - $CI = (229.7, 233.5)$
- a. $P(\mu \in (229.7, 233.5)) = 0.95?$

No. μ is not a random variable it is a parameter of a probability function. Probabilities are defined only for random variables.

- b. compute \bar{x} and s

we are given 95% confidence interval

$$CI = (l, u) = (229.7, 233.5)$$

Let's remember the definition for **case 2**

$$(l, u) = (\bar{x} - t_{\alpha/2, n-1} s / \sqrt{n}, \bar{x} + t_{\alpha/2, n-1} s / \sqrt{n})$$

for $\alpha = 0.05$: $t_{0.025, n-1}$

then

$$(l, u) = (\bar{x} - t_{0.025, n-1} s / \sqrt{n}, \bar{x} + t_{0.025, n-1} s / \sqrt{n})$$

where

$$t_{0.025, n-1} = F_{t, n-1}^{-1}(0.975) = \text{qt}(0.975, 4) = 2.77$$

Therefore we are given the confidence interval:

$$(l, u) = (\bar{x} - 2.77s / \sqrt{5}, \bar{x} + 2.77s / \sqrt{5}) = (229.7, 233.5)$$

and two equations to solve for \bar{x} and s :

- $\bar{x} - 1.23s = 229.7$
- $\bar{x} + 1.23s = 233.5$

With solutions:

$$\bar{x} = (229.7 + 233.5)/2 = 231.6 \text{ and } s = (233.5 - \bar{x})/1.23 = 1.53$$

c. compute 99% CI:

we have:

$$(l, u) = (\bar{x} - t_{\alpha/2, n-1} s / \sqrt{n}, \bar{x} + t_{\alpha/2, n-1} s / \sqrt{n})$$

We leave out a total of $\alpha = 0.01$.

or

$$(l, u) = (\bar{x} - t_{0.005, n-1} s / \sqrt{n}, \bar{x} + t_{0.005, n-1} s / \sqrt{n})$$

where

$$t_{0.005, n-1} = F_{t, n-1}^{-1}(0.995) = \text{qt}(0.995, 4) = 4.60$$

$$\text{then: } (\bar{x} - 4.60s/\sqrt{5}, \bar{x} + 4.60s/\sqrt{5}) = (228.45, 234.75)$$

Problem 2

Consider:

- $n = 1000$
- $x = 17$

Where $X \rightarrow \text{Bernoulli}(p)$

Remember: the sample mean of n Bernoulli trials \bar{X} is an estimator of p :

- $\bar{X} = \sum_{i=1}^n X_i$ where $E(\bar{x}) = E(X) = p$
- $S^2 = \bar{x}(1 - \bar{x})$ where $E(S^2) = \sigma^2 = p(1 - p)$

The $(1 - \alpha)100\%$ CI of sample is (case 3)

$$CI = (l, u) = (\bar{x} - z_{\alpha/2} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2}, \bar{x} + z_{\alpha/2} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2})$$

when np and $n(p - 1) > 5$

a. Then we have: $\bar{x} = \hat{p} = 0.017$

$n\bar{x} = 17$ and $n(1 - \bar{x}) = 983 > 5$ and we can use the approximation (TCL):

$$Z = \frac{\bar{x} - p}{[p(1-p)/n]^{1/2}} \rightarrow N(0, 1)$$

b. compute the 99% confidence interval then $\alpha = 0.01$

$$CI = (l, u) = (\bar{x} - z_{0.005} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2}, \bar{x} + z_{0.005} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2})$$

Since: $z_{0.005} = \phi^{-1}(0.995) = \text{qnorm}(0.995) = 2.575829$

Then

$$CI = (0.006474, 0.027526)$$

or

$$\bar{x} = 0.017 \pm 0.01$$

c. the estimate $\bar{x} \leq 0.02752$ we cannot guarantee the conditions of the client with 99% confidence.

Problem 3

Consider:

- $P(Y = 0) = a$
- $P(Y = 1) = 1 - a$
- $na(1 - a) \geq 8$

a. Compute n such that for the CI (l, u) , $D = u - l$ is maximum.

Y is a Bernoulli variable with $p = 1 - a$ and as $n(1 - p)p \geq 8$ then

$$CI = (\bar{y} - z_{\alpha/2} \left[\frac{\bar{y}(1-\bar{y})}{n} \right]^{1/2}, \bar{y} + z_{\alpha/2} \left[\frac{\bar{y}(1-\bar{y})}{n} \right]^{1/2})$$

$$D = 2m = 2z_{\alpha/2} \left[\frac{\bar{y}(1-\bar{y})}{n} \right]^{1/2} \text{ is maximum when } \bar{y}(1 - \bar{y}) \text{ is maximum.}$$

$$\frac{d(\bar{y}(1-\bar{y}))}{d\bar{y}} = (1 - 2\bar{y}) = 0 \text{ then } D \text{ is maximum when } \bar{y} = 1/2. \text{ That is } D_{max} = \frac{z_{\alpha/2}}{\sqrt{n}}$$

b. Consider $D = 0.02$ and $\alpha = 0.10$ for 90% confidence. Compute minimum n .

$$z_{\alpha/2} = z_{0.05} = \phi^{-1}(0.95) = \text{qnorm}(0.95) = 1.644854$$

$$D_{max} = 0.02 = \frac{1.644854}{\sqrt{n_{min}}} \text{ then}$$

$$n_{min} = (1.644854/0.02)^2 = 6763.862$$

or

$$n \geq 6764 \text{ (differences with results are because from tables } z_{0.05} \sim 1.65)$$

Problem 4

Consider:

- 682, 553, 555, 666, 657, 649, 522, 568, 700, 558
- $X \rightarrow N(\mu, \sigma_X^2)$

a. Compute 95% CI

- $\bar{x} = 611$
- $s = 65.51$
- $n = 10$
- As we don't know the variance of X then we use s

the observed CI is:

$$(l, u) = (\bar{x} - t_{0.025, n-1} s / \sqrt{n}, \bar{x} + t_{0.025, n-1} s / \sqrt{n})$$

$$t_{0.025, n-1} = F_{t, n-1}^{-1}(0.975) = \text{qt}(0.975, 9) = 2.26$$

$$\text{then: } (l, u) = (\bar{x} - 2.26s / \sqrt{n}, \bar{x} + 2.26s / \sqrt{n}) = (564.135, 657.865)$$

Problem 5

Consider:

- $n = 9$
- the 90% CI is (118.25, 123.55)

- $X \rightarrow N(\mu, \sigma^2)$
- We know σ

the observed CI is: $(l, u) = (\bar{x} - z_{0.05}\sigma/\sqrt{n}, \bar{x} + z_{0.05}\sigma/\sqrt{n})$

where

$$z_{0.05, n-1} = \phi^{-1}(0.95) = \text{qnorm}(0.95) = 1.644$$

then we have two equations

$$\text{i. } l = \bar{x} - 1.644 * \sigma / \sqrt{9} = 118.25$$

$$\text{ii. } u = \bar{x} + 1.644 * \sigma / \sqrt{9} = 123.55$$

$$\text{a. y b. Compute } \bar{x} \text{ and } \sigma_X^2$$

solving i and ii for \bar{x} and σ_X^2

$$\bar{x} = (118.25 + 123.55)/2 = 120.9 \text{ and } \sigma = 3 * (123.55 - \bar{x})/1.644 = 4.83 \text{ or } \sigma^2 = 23.36$$

$$\text{c. Compute } 97\% \text{ confidence interval, then } \alpha = 0.03$$

Remember:

$$(l, u) = (\bar{x} - z_{\alpha/2}\sigma/\sqrt{n}, \bar{x} + z_{\alpha/2}\sigma/\sqrt{n})$$

$$\text{as } z_{\alpha/2} = z_{0.015} = \phi^{-1}(0.985) = \text{qnorm}(0.985) = 2.170 \text{ then}$$

$$(l, u) = (120.9 - 2.170 * 4.83/\sqrt{9}, 120.9 + 2.170 * 4.83/\sqrt{9})$$

$$= (117.4063, 124.3937)$$

$$\text{d. for } d = 123.55 - 118.25 = 5.3 = (u - l) \text{ compute } n$$

$$d = u - l = 2 * 2.170 * 4.83/\sqrt{n} = 5.3$$

then

$$n = \left(\frac{2 * 2.170 * 4.83}{5.3} \right)^2 = 15.64 \sim 16$$

Problem 6

consider:

- $\bar{x} = 0.5354$
- $s = 0.3479$
- $n = 51$
- $\alpha = 0.05$

a. the 95% CI for the variance

$$(l, u) = \left(\frac{s^2(n-1)}{\chi_{\alpha/2, n-1}^2}, \frac{s^2(n-1)}{\chi_{1-\alpha/2, n-1}^2} \right)$$

$$l = \frac{s^2(n-1)}{\chi_{\alpha/2, n-1}^2}$$

$$\chi_{\alpha/2, n-1}^2 = \chi_{0.025, 50}^2 = \text{qchisq}(1-0.025, 50) = 71.42$$

$$\text{then } l = 0.3479^2 * 50 / 71.42 = 0.0847$$

$$u = \frac{s^2(n-1)}{\chi_{1-\alpha/2, n-1}^2}$$

$$\chi_{1-\alpha/2, n-1}^2 = \chi_{0.975, 50}^2 = \text{qchisq}(1-0.975, 50) = 32.35$$

$$\text{then } u = 0.3479^2 * 50 / 32.35 = 0.1870$$

$$CI = (l, u) = (0.0847, 0.1870)$$

consider:

- $X = 17$, number of fisheries with concentrations greater than 0.700 ppm
- $\alpha = 0.01$
- then $\bar{x} = \bar{x} = 1/3$ where $X \rightarrow \text{Bernoulli}(p)$.

since $n\bar{x} = 17$ and $n(1 - \bar{x}) = 34 > 5$ then CI is

$$(l, u) = (\bar{x} - z_{\alpha/2} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2}, \bar{x} + z_{\alpha/2} \left[\frac{\bar{x}(1-\bar{x})}{n} \right]^{1/2})$$

since:

$$z_{\alpha/2} = \text{qnorm}(1-0.005) = 2.575$$

then

$$(l, u) = (1/3 + 2.575 * \sqrt{\frac{1/3 * 2/3}{51}}, 1/3 + 2.575 * \sqrt{\frac{1/3 * 2/3}{51}}) = (0.1634, 0.5033)$$

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Problems on hypothesis testing

Summary

For hypothesis testing, follow steps from 1 to 7:

1. From the problem context, identify the parameter of interest: μ , p or σ^2

2. State the hypothesis contrast, for instance for μ

- two tailed: a. $H_0 : \mu = \mu_0$, b. $H_1 : \mu \neq \mu_0$
- or, upper tailed: a. $H_0 : \mu \leq \mu_0$, b. $H_1 : \mu > \mu_0$
- or, lower tailed: a. $H_0 : \mu \geq \mu_0$, b. $H_1 : \mu < \mu_0$

Note: the type of tail refers to H_1 .

3. Choose a significance level: α (for example $\alpha = 0.05$ when we want 95% confidence).

4. Define the statistic

case 1

Hypothesis for μ

i. $X \rightarrow N(\mu, \sigma^2)$

ii. we **know** σ^2

then

$$Z = \frac{\bar{X} - \mu_0}{\sigma / \sqrt{n}}$$

is standard. In R: `pnorm(z)`

case 2

Hypothesis for μ

i. $X \rightarrow N(\mu, \sigma^2)$

ii. and we **don't know** σ^2 then

then

$$T = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$$

is t-distributed with $n - 1$ degrees of freedom. In R: `pt(t,n-1)`

case 3

Hypothesis for p

- i. $X \rightarrow \text{Bernoulli}(p)$
- ii. and np and $n(1 - p) > 5$

$$Z = \frac{\bar{X} - p_0}{\left[\frac{p_0(1-p_0)}{n}\right]^{1/2}}$$

is standard. In R: `pnorm(z)` .

case 4

Hypothesis for σ^2

- i. $X \rightarrow N(\mu, \sigma^2)$

then

$$W = \frac{(n - 1)S^2}{\sigma_0^2}$$

is χ^2 -distributed with $n - 1$ degrees of freedom. In R: `pchisq(w,n-1)` .

5. Test the hypothesis

- a. compute the confidence interval at $(1 - \alpha)100\%$ confidence assuming H_0 is true
- b. or, define the acceptance **region** for the statistic assuming H_0 is true with probability of α at the edges
- c. or, compute the P -value:
 - two tail: $P\text{-value} = 2(1 - F(|z|))$,
 - upper tail: $P\text{-value} = (1 - F(z))$,
 - lower tail: $P\text{-value} = F(z)$

7. Decide

Reject H_0 (accept H_1) if:

- a. the $(1 - \alpha)100\%$ confidence interval does not contain μ_0 (σ_0 , p_0 , etc...)
- b. or, the observed statistic falls outside the acceptance region
- c. or, if the $P\text{-value} < \alpha$

Otherwise **do not** reject H_0 and accept the null.

R code for testing hypotheses:

case 1

```
library(BSDA)

x <- c(...)

z.test(x, mu = 0, alternative = , sigma.x = , conf.level = )
```

case 2

```
x <- c(...)

t.test(x, mu = , alternative = , conf.level = )
```

case 3

```
x <- c(...)

prop.test(x, n = , p = , alternative = , conf.level = , correct = FALSE )
```

case 4

```
library(EnvStats)

x <- c(...)

varTest(x, sigma.squared = , alternative = , conf.level = , )
```

Problem 1

- consider the measurements:

204.999, 206.149, 202.150, 207.048, 203.496, 206.343, 203.496, 206.676, 205.831

- $\mu \geq 206.5$
- a. At 90% confidence ($\alpha = 0.1$)
- Test for the mean where we do not know σ

Lower tail: $H_0 : \mu \geq 206.5, H_1 : \mu < 206.5$

- Statistic: $T = \frac{\bar{X} - \mu_0}{S/\sqrt{9}}$, t-distribution with $n - 1$ degrees of freedom
- Observed value: $t = \frac{205.132 - 206.5}{1.707/\sqrt{9}} = -2.404$

P -value for lower tail:

$$P\text{-value} = F(T) = P(T < -2.404) = \text{pt}(-2.404, 8)$$

$$= 0.021 < \alpha = 0.10$$

Since: $P < \alpha$ we reject the null hypothesis

- b. if $\sigma_X^2 = 4$. Test the hypothesis $\mu_0 = 206.5$ at 95% confidence

- Test for the mean where know σ

Two tail: $H_0 : \mu = 206.5, H_1 : \mu \neq 206.5$

- Statistic: $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{9}}$, standard distribution

- Observed value: $z = \frac{205.132 - 206.5}{2/\sqrt{9}} = -2.05$

P -value for two tail:

$$P\text{-value} = 2(1 - F(|-2.05|)) = 2*(1 - \text{pnorm}(2.05)) = 0.0403 < \alpha = 0.05$$

Since: $P < \alpha$ we reject the null hypothesis

Problem 2

- consider the measurements:

53700, 55500, 53000, 52400, 51000, 62000, 75000, 53800, 56600

- $\mu_0 = 62000$
- a. At 95% confidence ($\alpha = 0.05$)
- Test for the mean where we do not know σ

Lower tail: $H_0 : \mu \geq 62000, H_1 : \mu < 62000$

- Statistic: $T = \frac{\bar{X} - \mu_0}{s/\sqrt{9}}$, t-distribution with $n - 1 = 8$ degrees of freedom
- Observed value: $t = \frac{57000 - 62000}{7464.08/\sqrt{9}} = -2.01$

- a. critical region

$$P(T < t_{0.95}) = 0.05$$

$$t_{0.95,8} = \text{qt}(0.05, 8) = -1.8595$$

The critical region is $T < -1.8595$

since $t = -2.01 < -1.8595$ we then reject the null hypothesis

- b. P -value for lower tail:

$$P\text{-value} = F(T) = P(T < -2.01) =$$

$$\text{pt}(-2.01, 8) = 0.03966 < \alpha = 0.05$$

Since: $P < \alpha$ we reject the null hypothesis

- c. compute the 99% CI for the if $\sigma^2 = 54760000$

When we know the variance then the CI for an n sample of normal variables is:

$$(l, u) = (\bar{x} - z_{\alpha/2}\sigma/\sqrt{n}, \bar{x} + z_{\alpha/2}\sigma/\sqrt{n})$$

$$\bar{x} = 57000, \sigma = 7400, n = 9 \text{ and}$$

$$z_{\alpha/2} = z_{0.005} = \text{qnorm}(0.995) = 2.5758$$

Putting everything together

$$(l, u) = (50646.29, 63353.71) \text{ (since it contains 62000, we do not reject } H_0 \text{ with 99\% confidence)}$$

Problem 3

consider:

- $\bar{x} = 7750$

- $s = 145$
- $n = 6$
- $\mu_0 = 8000$
- We don't make a claim if $\mu \geq \mu_0$
- We make a claim if $\mu < \mu_0$

a. if $\alpha = 0.1$ should we make a claim?

- Test for the mean where we do not know σ

Lower tail: $H_0 : \mu \geq 8000, H_1 : \mu < 8000$

- Statistic: $T = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$, t-distribution with $n - 1$ degrees of freedom
- Observed value: $t = \frac{7750 - 8000}{145/\sqrt{6}} = -4.22$

P -value for lower tail:

$$P\text{-value} = F(T) = P(T < -4.22) = \text{pt}(-4.22, 5) \\ = 0.004 < \alpha = 0.01$$

Since: $P < \alpha$ we reject the null hypothesis, we make a claim.

b. consider:

- $\sigma_X^2 = 136161$
- $\alpha = 0.05$

Test the same hypothesis.

- Test for the mean where we do not know σ

Lower tail: $H_0 : \mu \geq 8000, H_1 : \mu < 8000$

- Statistic: $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$, is a standard variable
- Observed value: $z = \frac{7750 - 8000}{369/\sqrt{6}} = -1.659$

P -value for lower tail:

$$P\text{-value} = F(Z) = P(Z < -1.659) = \text{pnorm}(-1.659) \\ = 0.004 < \alpha = 0.05$$

Since: $P < \alpha$ we reject the null hypothesis

c.

- if $\mu_1 = 7700$ what is type II error or false negative probability?

$$\beta = P(H_0 : \text{accept} | H_0 : \text{false})$$

We accept H_0 when the observed value z falls in the acceptance region, at $\alpha = 0.05$ that is

$$z > z_c = \text{qnorm}(0.05) = -1.644$$

For this critical z_c , the critical \bar{x}_c is

$$z_c = \frac{\bar{x}_c - 8000}{369/\sqrt{6}} = -1.644 \text{ then } \bar{x}_c = 7752.19$$

For accepting the null hypothesis then we need to observe an average that is greater than $\bar{x} = 7752.19$

What is the probability that if $\mu = 7700$, we accept H_0 , that is:

$$\beta = P(\text{accept} | H_0 : \text{false}) = P(\bar{X} > 7752.19 | \mu = 7700)$$

$$P(\bar{X} > 7752.19 | \mu = 7700) = P(Z > \frac{7752.19 - 7700}{369/\sqrt{6}})$$

$$= 1 - \phi(0.35) = 1 - \text{pnorm}(0.35) = 0.36316$$

Problem 4

- $\mu_0 = 14$
- $\sigma = 4.8$
- $n = 26$
- $\bar{x} = 12.5$
- $s = 2.7$
- Test for the mean where we do not know σ

a.

Lower tail: $H_0 : \mu \geq 14$, $H_1 : \mu < 14$

reject H_0 when H_0 is true is a **false positive** or type I error.

false positive probability:

$$\alpha = P(\text{reject} | H_0 : \text{true})$$

The probability is what we leave out from the rejection zone when H_0 is true.

consider:

- $\alpha = 0.01$ test the hypothesis
- Test for the mean where we do not know σ
- Statistic: $T = \frac{\bar{X} - \mu_0}{s/\sqrt{n}}$, t-distribution with $n - 1$ degrees of freedom
- Observed value: $t = \frac{12.5 - 14}{2.7/\sqrt{26}} = -2.832$

P-value for lower tail:

$$P\text{-value} = F(T) = P(T < -2.832) = \text{pt}(-2.832, 25)$$

$$= 0.0045 < \alpha = 0.001$$

Since: $P < \alpha$ we reject the null hypothesis

c. consider:

- $\sigma = 4.8$
- $\alpha = 0.05$

Test the same hypothesis.

- Test for the mean where we do not know σ

Lower tail: $H_0 : \mu \geq 14$, $H_1 : \mu < 14$

- Statistic: $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$, is a standard variable
- Observed value: $z = \frac{12.5 - 14}{4.8/\sqrt{26}} = -1.593$

P -value for lower tail:

$$P\text{-value} = F(Z) = P(Z < -1.593) = \text{pnorm}(-1.593)$$

$$= 0.0555 > \alpha = 0.05$$

Since: $P > \alpha$ we **do not** reject the null hypothesis

Problem 5

consider:

- $\bar{x} = 10$
- $s = 1.5$
- $n = 41$

a. What is the CI at 87.886% confidence $\alpha = 1 - 0.87886 = 0.12114$

while we do not **know** σ^2 n is big and then we can use the standard distribution:

$$(l, u) = (\bar{x} - z_{\alpha/2}\sigma/\sqrt{n}, \bar{x} + z_{\alpha/2}\sigma/\sqrt{n})$$

$$(l, u) = (\bar{x} - z_{\alpha/2}\sigma/\sqrt{n}, \bar{x} + z_{\alpha/2}\sigma/\sqrt{n})$$

$$\text{where } z_{\alpha/2} = \text{qnorm}(1 - 0.12114/2) = 1.55$$

then

$$(l, u) = (10 - 1.55 * 1.5/\sqrt{41}, 10 + 1.55 * 1.5/\sqrt{41}) = (9.63, 10.36)$$

b. consider:

- $\mu_0 = 10.5$
- $\alpha = 0.05$

Test the hypothesis.

Lower tail: $H_0 : \mu \geq 10.5, H_1 : \mu < 10.5$

- Statistic: $Z = \frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}}$, is a standard variable
- Observed value: $z = \frac{10 - 10.5}{1.5/\sqrt{41}} = -2.134$

P -value for lower tail:

$$P\text{-value} = F(Z) = P(Z < -2.134) = \text{pnorm}(-2.134)$$

$$= 0.016 < \alpha = 0.05$$

Since: $P < \alpha$ we reject the null hypothesis

Problem 6

consider:

- measurements: 515, 464, 558, 491

- $\sigma^2 = 10000$
- $\alpha = 0.1$

Lower tail: $H_0 : \sigma^2 \geq 10000, H_1 : \sigma^2 < 10000$

- the sample variance is

$$s^2 = \text{sd}(c(515, 464, 558, 491))^2 = 1590$$

- Statistic: $W = \frac{(n-1)S^2}{\sigma_0^2}$, is a random variable that follows a χ^2 with $n - 1$ degrees of freedom.
- Observed value: $w_{obs} = \frac{(4-1)1590^2}{10000} = 0.477$

P -value for lower tail:

$$P\text{-value} = F(W) = P(W < 0.477)$$

$$= \text{pchisq}(0.477, 3) = 0.076 < \alpha = 0.1$$

Since: $P < \alpha$ we reject the null hypothesis

consider

- $X \rightarrow N(\mu, \sigma^2)$
- $\mu = 500$
- $\sigma = 50$

b. what is n such that $P(|\bar{X} - \mu| < 10) = 0.95$?

$$P(-10 < \bar{X} - \mu < 10) = 0.95$$

$$P\left(\frac{-10}{\sigma/\sqrt{n}} < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < \frac{10}{\sigma/\sqrt{n}}\right) = 0.95$$

$$\Phi\left(\frac{10}{\sigma/\sqrt{n}}\right) - \Phi\left(\frac{-10}{\sigma/\sqrt{n}}\right) = 0.95$$

$$\Phi\left(\frac{10}{\sigma/\sqrt{n}}\right) - (1 - \Phi\left(\frac{10}{\sigma/\sqrt{n}}\right)) = 0.95$$

$$\Phi\left(\frac{10}{\sigma/\sqrt{n}}\right) = (1 + 0.95)/2 = 0.975$$

$$\frac{10}{\sigma/\sqrt{n}} = \text{qnorm}(0.975) = 1.959$$

$$\frac{10}{\sigma/\sqrt{n}} = 1.959$$

solving for n then $n_{min} = 96.4, n > 97$

Problem 7

consider

- $X \rightarrow N(\mu, \sigma^2)$
- $\sigma = 5$
- Upper tail: $H_0 : \mu \leq 80, H_1 : \mu > 80$
- $n = 100$
- $\alpha = 0.0505$

compute type II error, false negative:

$$\beta = P(\text{accept} | H_0 : \text{false})$$

We accept H_0 when the observed value z falls in the acceptance region, at $\alpha = 0.0505$

Acceptance region:

$$z < z_c = \text{qnorm}(1-0.0505) = 1.640$$

For this critical value z_c we obtain the critical value for \bar{x}_c

$$z = \frac{\bar{x}_c - 80}{5/\sqrt{100}} = 1.640 \text{ then } \bar{x}_c = (5/10) * 1.64 + 80 = 80.82$$

For accepting the null hypothesis H_0 we need to observe an average that is lower than $\bar{x} = 80.82$

What is the probability that if $\mu = 81$, we accept H_0 ?

that is:

$$\beta = P(\text{accept} | H_0 : \text{false}) = P(\bar{X} < 80.82 | \mu = 81)$$

$$P(\bar{X} < 80.82 | \mu = 81) = P(Z < \frac{80.82 - 80}{5/\sqrt{100}})$$

$$= \phi(-0.36) = \text{pnorm}(-0.36) = 0.3594236$$

Problem 8

Consider:

- X is a Bernoulli variable $X \rightarrow \text{Bernoulli}(p)$
- The measurements give

$$\bar{x} = \bar{x} = 926/1225 = 0.7559$$

a.

upper tail: $H_0 : p \leq p_0 = 0.75$, $H_1 : p > p_0 = 0.75$

- Statistic:

$$Z_{obs} = \frac{\bar{X} - p_0}{\left[\frac{p_0(1-p_0)}{n} \right]^{1/2}}$$

is a standard variable because $np_0, n(1-p_0) > 5$ by the CLT

- Observed value: $z = \frac{0.7559 - 0.75}{\sqrt{0.75(1-0.75)/\sqrt{1225}}} = 0.47$

P -value for upper tail:

$$P\text{-value} = 1 - F(Z) = 1 - P(Z < 0.47) = 1 - \phi(0.47)$$

$$= 1 - \text{pnorm}(0.47) = 0.31$$

only with a level of significance of 0.31 we can reject the null hypothesis, and conclude that the method produces the results within the tolerance limits.