

STA 510

MANDATORY 2

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Student: 236227

Aleksander B. Jakobsen

$$\textcircled{1} \quad \int_{-1}^1 \int_{-1}^1 I_D(x,y) dx dy = \pi$$

where  $I_D(x,y) = \begin{cases} 1, & x^2 + y^2 \leq 1 \\ 0, & \text{otherwise} \end{cases}$

Crude first attempt:  $\Theta_{CMC} = \frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)$

$X_i, Y_i$  independent

a) Rewrite  $\int_{-1}^1 \int_{-1}^1 I_D(x,y) dx dy$  as

$$\begin{aligned} \int_{b_1}^{b_2} \int_{a_1}^{a_2} I_D(x,y) dx dy &= (a_2 - a_1)(b_2 - b_1) \iint_{b_1, a_1}^{b_2, a_2} I_D(x,y) \cdot \frac{1}{(b_2 - b_1)} \cdot \frac{1}{(a_2 - a_1)} \\ &= (a_2 - a_1)(b_2 - b_1) E(I_D(X,Y)) \end{aligned}$$

and a crude estimate is

$$\frac{(a_2 - a_1)(b_2 - b_1)}{n} \sum_{i=1}^n I_D(X_i, Y_i) = \frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)$$

Inserting values for  $a$  and  $b$ , we get

$$\frac{(1 - (-1))(1 - (-1))}{n} \sum_{i=1}^n I_D(X_i, Y_i) = \frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)$$

and thus,  $\frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)$  is a crude monte carlo estimator



Show that  $I_D(X_i, Y_i)$  has a Bernoulli distribution with success probability  $\frac{\pi}{4}$ .

$$I_D(X_i, Y_i) = \begin{cases} 1 & \text{with prob } p \\ 0 & \text{with prob } q = 1-p \end{cases}$$

$I_D(X_i, Y_i)$  is either 0 or 1

The area of the circle is  $\pi$ .

If "success", then  $I_D(X_i, Y_i) = 1$

Assuming success, then the integral becomes

$$\int_{-1}^1 \int_{-1}^1 dx dy = \int_{-1}^1 x|_1 dx = \int_{-1}^1 2 dx = 4$$

Then the success probability is the prob of the circle of area  $\pi$  ~~to~~ ~~the~~

divided by square area 4, hence  $p = \frac{\pi}{4}$ .

Since this is a "either/or" or "success/failure",

and  $p = \frac{\pi}{4}$ , yielding  $q = 1-p = 1 - \frac{\pi}{4}$ , then

this is a Bernoulli distribution



b)  $\sum_{i=1}^n I_D(X_i, Y_i)$  has a binomial distribution,

since it is the repetition of Bernoulli trials, all with "success" =  $\frac{\pi}{4}$  and "failure" =  $1 - \frac{\pi}{4}$ .

Thus  $\sum_{i=1}^n I_D = \text{Binomial}$

Expectation for Binomial:  $np$

Variance for Binomial:  $np(1-p)$

$$\begin{aligned} E\left(\frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)\right) &= \frac{4}{n} E\left(\sum_{i=1}^n I_D(X_i, Y_i)\right) = \frac{4}{n} \cdot np \\ &= 4 \cdot \frac{\pi}{4} = \underline{\underline{\pi}} \end{aligned}$$

$$\begin{aligned} \text{Var}\left(\frac{4}{n} \sum_{i=1}^n I_D(X_i, Y_i)\right) &= \left(\frac{4}{n}\right)^2 \text{Var}\left(\sum_{i=1}^n I_D(X_i, Y_i)\right) \\ &= \frac{16}{n^2} \cdot np(1-p) = \frac{16}{n} \cdot \frac{\pi}{4} \left(1 - \frac{\pi}{4}\right) \\ &= \underline{\underline{\frac{\pi}{n} (4 - \pi)}} \end{aligned}$$

1e) ~~I expect the antithetic variables will reduce the Monte Carlo variance, because the covariances,  $\text{Cov}(X_i, V_i)$  and  $\text{Cov}(Y_i, W_i)$  should be negative, and thus reduce variance.~~

Function is for a circle, and is thereby not monotonic. Thus, variance will not be reduced.

2b)

Find smallest possible  $\lambda_{\max}$ :

$$\lim_{t \rightarrow \infty} \lambda(t) = \lim_{t \rightarrow \infty} \frac{297(1 + \cos(2\pi(t+1/10)))}{10} (1 - \exp(-t/10)/2) + \frac{3}{5}$$

Evaluate the lim where there are t's.

$$\lim_{t \rightarrow \infty} \cos(2\pi(t+1/10)) = 1$$

$$\lim_{t \rightarrow \infty} \exp(-t/10) = \lim_{t \rightarrow \infty} \frac{1}{\exp(t/10)} = 0$$

Thus

$$\lim_{t \rightarrow \infty} \lambda(t) = \frac{297(1+1)}{10} (1-0) + \frac{3}{5} = \underline{\underline{60 = \lambda_{\max}}}$$



$$4a) f_R(r) = \begin{cases} \frac{e^{-\frac{(r-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi}\sigma(1-F(L))} & \text{if } r \geq L \\ 0 & \text{if } r < L \end{cases}$$

where  $F$  is the cdf of a  $N(\mu, \sigma^2)$  r.v.

$$G_R(r) = \int_{-\infty}^r f_R(r) dr = \int_L^r \frac{e^{-\frac{(r-\mu)^2}{2\sigma^2}}}{\sqrt{2\pi}\sigma(1-F(L))} dr$$

$$= \frac{1}{1-F(L)} \underbrace{\int_L^r \frac{1}{\sqrt{2\pi}\sigma} \cdot e^{-\frac{1}{2}\left(\frac{r-\mu}{\sigma}\right)^2} dr}_{\text{This is pdf of a normal r.v.}}$$

This is pdf of a normal r.v.

$$= \frac{1}{1-F(L)} \cdot \underbrace{\frac{1}{2} \left[ 1 + \operatorname{erf}\left(\frac{r-\mu}{\sigma}\right) \right]}_{\text{cdf of a } N(\mu, \sigma^2) \text{ r.v.}} \bigg|_{r=L}^{r=r}$$

cdf of a  $N(\mu, \sigma^2)$  r.v.

$$= \frac{1}{1-F(L)} \cdot F(r) - F(L)$$

$$\underline{G_R(r) = \frac{F(r) - F(L)}{1 - F(L)}}$$

$$G_R^{-1}(u) = ?$$

$$\frac{F(r) - F(L)}{1 - F(L)} = u$$

$$F(r) = u(1 - F(L)) + F(L)$$

$$\underline{F^{-1}(F(r)) = F^{-1}(u(1 - F(L)) + F(L)) = G_R^{-1}(u)}$$

(5)

4c) To simulate  $S$ :

$$\int_{\frac{0.5}{R}}^S \frac{10R}{3} dS = \frac{10RS}{3} \Big|_{\frac{0.5}{R}}^S = U[0,1]$$

$$= \frac{10RS}{3} - \frac{10R \cdot 0.5}{3} = U[0,1]$$

$$\frac{10RS}{3} = U[0,1] + \frac{5}{3}$$

$$S = \frac{3 \cdot U[0,1]}{10} + \frac{1}{2R}$$

This is used in code to  
simulate SIR.

Source as used in the .R file

1. B. Efron and R. Tibishirani. *An introduction to the Bootstrap*. Chapman & Hall/CRC, Boca Raton, FL, 1993.
2. Rizzo, M.L. *Statistical computing with R, 2nd ed.* Chapman & Hall/CRC, Boca Raton, FL, 2019.