

Event history analysis

An introduction

Alessandro Di Nallo

Università Bocconi

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PART II

Estimation using STATA

Estimation using STATA

- The aim of this workshop is to illustrate how to use Stata to estimate multivariate continuous & discrete time survival time models.
- These include the parametric models (e.g., Weibull hazard and piecewise constant exponential hazard functions).
- Stata provides an extensive suite of estimators. Parametric regression survival-time models are estimated by maximum likelihood

Estimation using STATA

We will take a few steps

- 1 Prepare the data for survival analysis
- 2 Understand the 'survival dynamics' of the phenomenon under investigation (Kaplan-Meier estimates)
- 3 Estimate the phenomenon with a continuous/discrete time model and time-invariant explanatory variables
- 4 Include time-varying explanatory variables

Continuous time models in STATA

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 - ▶ **Survival time** and **censoring variables** already exist (we have already derived them from start and end dates)

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 - ▶ A single record per 'subject'
 - ▶ **No** complications arising from **left censoring**, **gaps**, **left truncation** ('delayed entry'), or **multiple events**
 - ▶ There are **no missing values** for simplicity's sake
 - ▶ Data do not need to be weighted
 - ▶ **Survival time** and **censoring variables** already exist (we have already derived them from start and end dates)
 - ▶ **No time-varying covariates**, for the moment.
- Here I assume that there is a Stata format **data set ready to use**.

Data structure of continuous time models (an example)

```
. use cancer,clear
(Patient Survival in Drug Trial)
. de
Contains data from cancer.dta
obs:          48          Patient Survival in Drug Trial
vars:          4          16 Nov 1998 11:49
size:         576 (100.0% of memory free)
```

| | | | |
|-------------|-----|-------|--------------------------------|
| 1. studytim | int | %8.0g | Months to death or end of exp. |
| 2. died | int | %8.0g | 1 if patient died |
| 3. drug | int | %8.0g | Drug type (1=placebo) |
| 4. age | int | %8.0g | Patient's age at start of exp. |

- studytim is the survival time
- died is the censoring indicator;
- drug and age are potential explanatory variables.

Data structure of continuous time models (an example)

| studytime | died | drug | age | _st | _d | _t | _t0 |
|-----------|------|---------|-----|-----|----|----|-----|
| 11 | Yes | Placebo | 55 | 1 | 1 | 11 | 0 |
| 12 | Yes | Placebo | 49 | 1 | 1 | 12 | 0 |
| 12 | Yes | Placebo | 62 | 1 | 1 | 12 | 0 |
| 15 | Yes | Placebo | 51 | 1 | 1 | 15 | 0 |
| 17 | Yes | Placebo | 49 | 1 | 1 | 17 | 0 |
| 22 | Yes | Placebo | 57 | 1 | 1 | 22 | 0 |
| 23 | Yes | Placebo | 52 | 1 | 1 | 23 | 0 |
| 6 | Yes | Other | 67 | 1 | 1 | 6 | 0 |
| 6 | No | Other | 65 | 1 | 0 | 6 | 0 |
| 7 | Yes | Other | 58 | 1 | 1 | 7 | 0 |
| 9 | No | Other | 56 | 1 | 0 | 9 | 0 |
| 10 | No | Other | 49 | 1 | 0 | 10 | 0 |
| 11 | No | Other | 61 | 1 | 0 | 11 | 0 |

- `studytim` is the survival time
- `died` is the censoring indicator;
- `drug` and `age` are potential explanatory variables.

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| | 15 | Yes | Placebo | 51 | 1 | 1 | 15 | 0 |
| | 17 | Yes | Placebo | 49 | 1 | 1 | 17 | 0 |
| | 22 | Yes | Placebo | 57 | 1 | 1 | 22 | 0 |
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- One row in the data set for each 'subject' (e.g. person or firm).
- Columns in the data set (variables) contain at least two types of information for each subject:
 - 1 the length of time in the state (the **survival time** = the length of time the subject was exposed to the risk of experiencing a 'failure');
 - 2 **censoring status** (a variable typically equal to 1 if the person experienced a 'failure', and equal to 0 otherwise).
 - 3 At least one (time-invariant, for now) variable used as **regressor** in estimation of multivariate hazard models.

Data structure of continuous time models (an example)

`stset` is the only command required to organise the data.

```
stset timevar , failure(censvar)
```

- `stset timevar` identifies the phenomenon at issue (`studytime`)
- `failure(censvar)` specifies the failure event (`died`)– there is a failure whenever the variable is not equal to zero and not missing

```
. stset studytim , failure(died)
```

```
      failure event:  died != 0 & died < .  
obs. time interval:  (0, studytim]  
exit on or before:  failure
```

```
48  total obs.
```

```
0   exclusions
```

```
48  obs. remaining, representing
```

```
31  failures in single record/single failure data
```

```
744 total analysis time at risk, at risk from t =           0
```

```
      earliest observed entry t =           0
```

```
      last observed exit t =          39
```

Data structure of continuous time models (an example)

`stset` creates a set of new variables in the data

```
. de _*
```

| variable name | storage type | display format | value label | variable label |
|---------------|--------------|----------------|-------------|----------------|
| _____ | | | | |
| _st | byte | %8.0g | | |
| _d | byte | %8.0g | | |
| _t | byte | %10.0g | | |
| _t0 | byte | %10.0g | | |

```
. su _*
```

| Variable | Obs | Mean | Std. Dev. | Min | Max |
|----------|-----|----------|-----------|-----|-----|
| _____ | | | | | |
| _st | 48 | 1 | 0 | 1 | 1 |
| _d | 48 | .6458333 | .4833211 | 0 | 1 |
| _t | 48 | 15.5 | 10.25629 | 1 | 39 |
| _t0 | 48 | 0 | 0 | 0 | 0 |

- `_st` is a 0/1 variable, equal to 1 for observations whose data has been `stset`;
- `_d` variable is the binary censoring indicator (died in this case);
- `_t` is the duration variable (`studytim` in this case);
- `_t0` records the date of entry to the study for each case, i.e. when they were first observed at risk of the event (i.e. $t=0$);

Data structure of continuous time models (an example)

stsum shows how the data are currently set

```
. stsum
```

```
      failure _d:  died  
analysis time _t:  studytim
```

| | | time at risk | incidence rate | no. of subjects | ----- Survival time ----- | | | |
|-------|--|--------------|-------------------|--------------------|---------------------------|-----|-----|-----|
| | | | | | | 25% | 50% | 75% |
| total | | 744 | .0416667 | 48 | | 8 | 17 | 33 |

- There are 31 failures (deaths in this case);
- The incidence rate, $0.04167 = 31/744$;
- The median survival time since the start of the study is 17 months

Estimation using `streg`

The basic syntax is

```
streg [varlist], dist(distname) nohr time tr nolog
```

- specifies the survival model to be estimated;
- is one of the following: exponential, **weibull**, gompertz, lognormal, loglogistic or gamma

Estimation of a Weibull model using `streg`

```
. streg drug age, dist(weibull) nolog nohr
```

```
      failure _d:  died  
analysis time _t:  studytim
```

Weibull regression -- log relative-hazard form

```
No. of subjects =          48          Number of obs   =          48  
No. of failures =          31  
Time at risk    =          744  
  
Log likelihood   =  -42.931335          LR chi2(2)       =          35.39  
                                          Prob > chi2      =          0.0000
```

| _t | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-------|-----------|-----------|--------|-------|----------------------|-----------|
| drug | -2.196936 | .4087791 | -5.374 | 0.000 | -2.998129 | -1.395744 |
| age | .1202027 | .0371599 | 3.235 | 0.001 | .0473707 | .1930348 |
| _cons | -10.58396 | 2.326271 | -4.550 | 0.000 | -15.14337 | -6.024553 |
| /ln_p | .5204297 | .1389037 | 3.747 | 0.000 | .2481834 | .792676 |
| p | 1.682751 | .2337403 | | | 1.281695 | 2.209301 |
| 1/p | .5942651 | .0825456 | | | .452632 | .7802168 |

The `nohr` option means that **coefficient estimates** are shown.

Estimation of a Weibull model using `streg`

```
. streg, hr
```

```
Weibull regression -- log relative-hazard form
```

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No. of failures =           31
Time at risk    =           744
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                                                Prob > chi2        =           0.0000
```

| ----- | | | | | | |
|-------|------------|-----------|--------|-------|----------------------|----------|
| _t | Haz. Ratio | Std. Err. | z | P> z | [95% Conf. Interval] | |
| ----- | | | | | | |
| drug | .1111431 | .045433 | -5.374 | 0.000 | .0498803 | .2476487 |
| age | 1.127725 | .0419062 | 3.235 | 0.001 | 1.048511 | 1.212925 |
| ----- | | | | | | |
| /ln_p | .5204297 | .1389037 | 3.747 | 0.000 | .2481834 | .792676 |
| ----- | | | | | | |
| p | 1.682751 | .2337403 | | | 1.281695 | 2.209301 |
| 1/p | .5942651 | .0825456 | | | .452632 | .7802168 |
| ----- | | | | | | |

The `hr` option means that **hazard ratios** are shown.

Interpretation of coefficient & hazard ratio of drug

$$\beta_{drug} \simeq -2.2$$

The **coefficient estimates** indicate that those receiving the drug (drug = 1) have lower [coeff. < 0] hazard rates *ceteris paribus* (i.e. lower conditional death rates and hence longer survival times).

$$\exp(\beta_{drug}) \simeq 0.11$$

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The **hazard ratio** compares the hazard rates of two categories (in this case): those who receive the drug and those who do not



The hazard *rate* for those who received the drug is only 11% of the hazard rate for those who received the placebo.

Interpretation of coefficient & hazard ratio of age

$$\beta_{(age)} \simeq 0.12$$

The **coefficient estimates** indicate that older people (age increasing) have higher [coeff. > 0] hazard rates *ceteris paribus* (i.e. higher conditional death rates and hence shorter survival times).

$$\exp(\beta_{(age)}) \simeq 1.13$$

The **hazard ratio** "compares" the hazard rates of people with age (say) X to those with age $X + 1$



The hazard *rate* of a one year rise in age ($\Delta X = 1$) is associated with a 13% higher hazard rate.

Generalizing the interpretation of β s & hazard ratios

A **dummy** (or categorical) variable (ex. drug)

β (coeff. estimate) and hazard *ratio* contrast the hazard rate of group **G** to that of the baseline group **B**

If $\beta > 0$ or hazard ratio $> 1 \rightarrow \text{hazard (G)} > \text{hazard (B)}$

If $\beta < 0$ or hazard ratio $< 1 \rightarrow \text{hazard (G)} < \text{hazard (B)}$

A **continuous** variable (ex. age)

β (coeff. estimate) and hazard *ratio* capture the **elasticity** of the hazard with respect to a **one unit** change in the explanatory variable X

If $\beta > 0$ or hazard ratio $> 1 \Rightarrow \Delta X = 1 \rightarrow \uparrow \text{hazard}$

If $\beta < 0$ or hazard ratio $< 1 \Rightarrow \Delta X = 1 \rightarrow \downarrow \text{hazard}$

More on the interpretation of β s & hazard ratios of a continuous variable

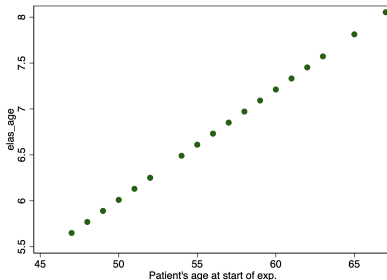
The elasticity of the hazard rate with respect to a one unit change ($\Delta X_k = 1$) in the k^{th} explanatory variable is given by $\beta_k X_{ik}$

For age, it is $0.12 * age_j$.

Let age_j vary and calculate the elasticity of the hazard wrt age:

```
ge elas_age = _b[age]*age
```

Let's plot 'elasticity of age' (y-axis) vs. 'age' (x-axis)



More on the interpretation of β s & hazard ratios of a continuous variable

More generally, hazard rate ratios at each survival time are related to absolute differences in characteristics (exploiting the properties of exponential functions):

$$h(t, X_1)/h(t, X_2) = \exp[\beta X_1]/\exp[\beta X_2] \equiv \exp[\beta(X_1 - X_2)]$$

Example 1. A 10-year difference in age, other things equal, is associated with a hazard rate ratio of some 3.3:

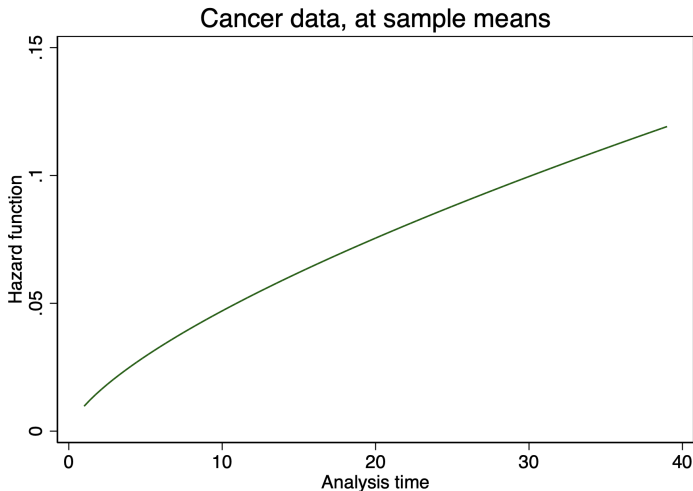
$$\frac{h(t; \text{age} = y + 10, \text{drug} = x)}{h(t; \text{age} = y, \text{drug} = x)} = \exp(\beta_{\text{age}} * 10) \approx 3.3$$

Example 2. Some one aged $y+10$ and who is receiving the drug has a hazard ratio that is 37% of some one aged y who gets the placebo:

$$\frac{h(t; \text{age} = y + 10, \text{drug} = 1)}{h(t; \text{age} = y, \text{drug} = 0)} = \exp(\beta_{\text{age}} * 10 + \beta_{\text{drug}} * 1) \approx 0.37$$

Estimated hazard function

```
stcurv, hazard title("Cancer data, at sample means")
```



Estimated hazard function

Recall the general formula

$$h(t; \mathbf{X}) = h_0(t) \exp(\beta \mathbf{X})$$

In this case we calculate the function $h_0(t)$ keeping the covariates at their means:

$$h(t; \overline{age}, \overline{drug}) = h_0(t) \exp(\beta \overline{\mathbf{X}})$$

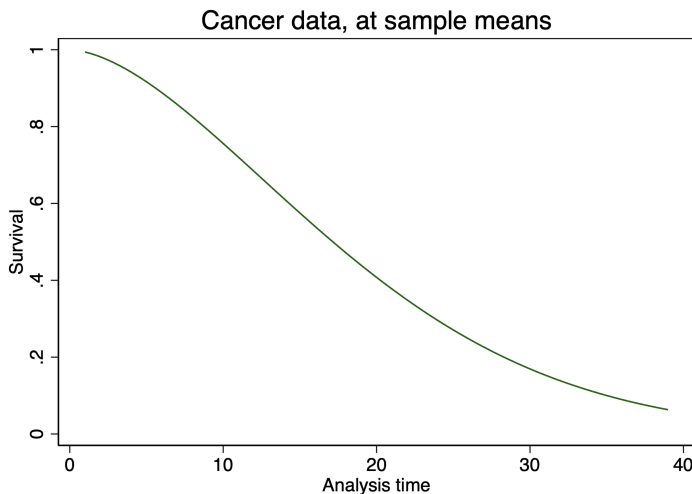
The *baseline* hazard function is a function of time, and is identical for all the individuals (or firms, etc...), regardless of their characteristics.

The curve is monotonically increasing.

The risk of death increases over time (time since treatment began).

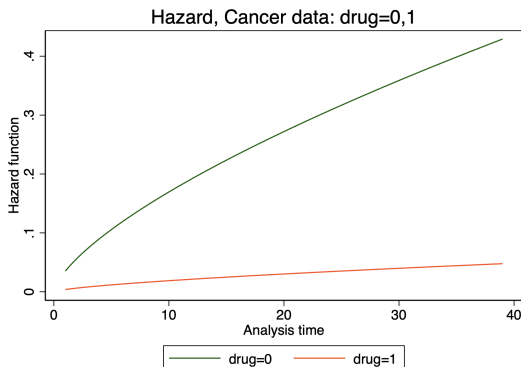
Estimated survival function

```
stcurv, survival title("Cancer data, at sample  
means")
```



Estimated hazard function

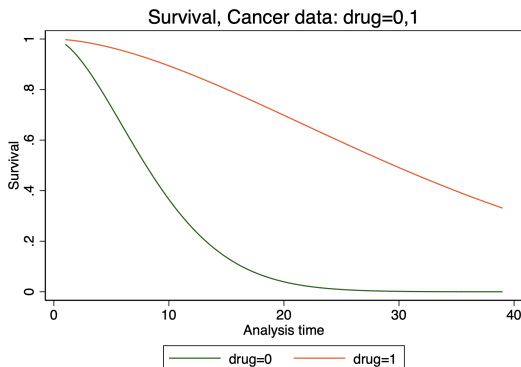
```
stcurv, hazard title("Cancer data, at sample means")
```



- Patients with the placebo ($\text{drug} = 0$): $h_0(t) \exp(\beta_{\text{age}} \overline{\text{age}})$
- Patients with the treatment ($\text{drug} = 1$): $h_0(t) \exp(\beta_{\text{age}} \overline{\text{age}} + \beta_{\text{drug}})$

Estimated survival function

```
stcurv, hazard title("Cancer data, at sample means")
```



- Patients with the placebo ($\text{drug} = 0$): $S(t, \overline{age}, \text{drug} = 0)$
- Patients with the treatment ($\text{drug} = 1$): $S(t, \overline{age}, \text{drug} = 1)$

Additional commands

`predict xb, xb`

- Generates a new variable equal to the estimate βX_i for each person i
- Calculates the shape parameter (α , in the case of Weibul) that is stored in the 'estimation class' $e(aux_p)$
- Provides the parameters to estimate, for instance, mean and median survival time for groups of people
 - ▶ Mean survival time for the 50 years old with a placebo
 - ▶ Median survival time for the 35 years old with a treatment
- More details in the lecture notes

Discrete time models in STATA

Preparing survival discrete time data for analysis and estimation

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- All one has to do is **re-organise the data set**, define some new variables (to specify the baseline hazard function in particular), and then **apply logit or cloglog regression**.
 - ▶ For each unit of analysis (e.g., person), there are as many data rows as there are time intervals at risk of the event occurring for each unit of analysis.
 - ▶ "1 row - 1 person" (continuous models) \Rightarrow "Each person contributes T_i rows" (discrete models), where T_i is the number of time periods (e.g. months) i was at risk of the event.
 - ▶ A unique identifier variable for each subject is needed, plus a spell month (or semester, or year) identifier variable for each subject.

Preparing survival discrete time data for analysis and estimation

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- All one has to do is **re-organise the data set**, define some new variables (to specify the baseline hazard function in particular), and then **apply logit or cloglog regression**.
- The binary dependent variable also needs to be created.
 - ▶ If subject i 's survival time is **censored**, the binary dependent variable is equal to 0 for all of i 's spell months;
 - ▶ If subject i 's survival time is **not censored**, the binary dependent variable is equal to 0 for all but the last of i 's spell months (month 1,..., T_{i-1}) and equal to 1 for the last month (month T_i).
- Full tutorial on the lecture notes.

Preparing survival discrete time data for analysis and estimation. An example.

Data Editor (Browse) — cancer.dta

| | id | j | studytime | died | dead | drug | age | _st | _d | _t | _t0 |
|----|----|---|-----------|------|------|---------|-----|-----|----|----|-----|
| 13 | 6 | 2 | 4 | Yes | 0 | placebo | 67 | 1 | 1 | 4 | 0 |
| 14 | 6 | 3 | 4 | Yes | 0 | placebo | 67 | 1 | 1 | 4 | 0 |
| 15 | 6 | 4 | 4 | Yes | 1 | placebo | 67 | 1 | 1 | 4 | 0 |
| 16 | 7 | 1 | 5 | Yes | 0 | placebo | 63 | 1 | 1 | 5 | 0 |
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| 27 | 9 | 2 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 28 | 9 | 3 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 29 | 9 | 4 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 30 | 9 | 5 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 31 | 9 | 6 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 32 | 9 | 7 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
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| 35 | 10 | 2 | 8 | No | 0 | placebo | 58 | 1 | 0 | 8 | 0 |

id: the unit of analysis (persons)

Preparing survival discrete time data for analysis and estimation. An example.

Data Editor (Browse) — cancer.dta

| id[1] | 1 | | | | | | | | | | | |
|-------|----|---|-----------|------|------|---------|-----|-----|----|----|-----|--|
| | id | j | studytime | died | dead | drug | age | _st | _d | _t | _t0 | |
| 13 | 6 | 2 | 4 | Yes | 0 | placebo | 67 | 1 | 1 | 4 | 0 | |
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| 32 | 9 | 7 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 | |
| 33 | 9 | 8 | 8 | Yes | 1 | placebo | 56 | 1 | 1 | 8 | 0 | |
| 34 | 10 | 1 | 8 | No | 0 | placebo | 58 | 1 | 0 | 8 | 0 | |
| 35 | 10 | 2 | 8 | No | 0 | placebo | 58 | 1 | 0 | 8 | 0 | |

id: the unit of analysis (persons)

j: time unit (months)

Preparing survival discrete time data for analysis and estimation. An example.

Data Editor (Browse) — cancer.dta

| | id | j | studytime | died | dead | drug | age | _st | _d | _t | _t0 |
|----|----|---|-----------|------|------|---------|-----|-----|----|----|-----|
| 13 | 6 | 2 | 4 | Yes | 0 | placebo | 67 | 1 | 1 | 4 | 0 |
| 14 | 6 | 3 | 4 | Yes | 0 | placebo | 67 | 1 | 1 | 4 | 0 |
| 15 | 6 | 4 | 4 | Yes | 1 | placebo | 67 | 1 | 1 | 4 | 0 |
| 16 | 7 | 1 | 5 | Yes | 0 | placebo | 63 | 1 | 1 | 5 | 0 |
| 17 | 7 | 2 | 5 | Yes | 0 | placebo | 63 | 1 | 1 | 5 | 0 |
| 18 | 7 | 3 | 5 | Yes | 0 | placebo | 63 | 1 | 1 | 5 | 0 |
| 19 | 7 | 4 | 5 | Yes | 0 | placebo | 63 | 1 | 1 | 5 | 0 |
| 20 | 7 | 5 | 5 | Yes | 1 | placebo | 63 | 1 | 1 | 5 | 0 |
| 21 | 8 | 1 | 5 | Yes | 0 | placebo | 58 | 1 | 1 | 5 | 0 |
| 22 | 8 | 2 | 5 | Yes | 0 | placebo | 58 | 1 | 1 | 5 | 0 |
| 23 | 8 | 3 | 5 | Yes | 0 | placebo | 58 | 1 | 1 | 5 | 0 |
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| 26 | 9 | 1 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 27 | 9 | 2 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 28 | 9 | 3 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 29 | 9 | 4 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 30 | 9 | 5 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 31 | 9 | 6 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 32 | 9 | 7 | 8 | Yes | 0 | placebo | 56 | 1 | 1 | 8 | 0 |
| 33 | 9 | 8 | 8 | Yes | 1 | placebo | 56 | 1 | 1 | 8 | 0 |
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Preparing survival discrete time data for analysis and estimation. An example.

Data Editor (Browse) — cancer.dta

| id[1] | 1 | | | | | | | | | | |
|-------|----|---|-----------|------|------|---------|-----|-----|----|----|-----|
| | id | j | studytime | died | dead | drug | age | _st | _d | _t | _t0 |
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Preparing survival discrete time data for analysis and estimation. An example.

Data Editor (Browse) — cancer.dta

| id[1] | | | | | | | | | | | |
|-------|----|---|-----------|------|------|---------|-----|-----|----|----|-----|
| | id | j | studytime | died | dead | drug | age | _st | _d | _t | _t0 |
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Preparing survival discrete time data for analysis and estimation. An example.

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Choose the functional form for the baseline hazard function

- We define new time-varying covariates which are functions of survival time t per person.
- Different alternatives are possible
 - ▶ $\log(\text{time})$
 - ▶ polynomial in time (e.g., $\alpha t + \gamma t^2 + \delta t^3 \dots$)
 - ▶ piece-wise constant
 - ▶ fully non-parametric (e.g., $\gamma_1 * j_1 + \gamma_2 * j_2 + \gamma_3 * j_3 + \dots + \gamma_k * j_k$, where j_k are duration-interval-specific dummy variables). *Caveat*: check whether events occur at each value of j_k .

Choose the functional form for the baseline hazard function

| | id | j | studytime | lnj | j2 | j3 | dur1 | dur2 | dur3 | dur4 | dur5 | dur6 | d1 | d2 | d3 | d4 | d5 | d6 | d7 | d8 | d9 | d10 |
|----|----|----|-----------|----------|----|-----|------|------|------|------|------|------|----|----|----|----|----|----|----|----|----|-----|
| 1 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 2 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 3 | 1 | 2 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 2 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1 | 3 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 2 | 3 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 4 | 3 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 5 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 5 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 5 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 6 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 6 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 6 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 6 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 16 | 7 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 7 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 18 | 7 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 7 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 20 | 7 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 21 | 8 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 8 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 8 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 8 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 25 | 8 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 26 | 9 | 1 | 8 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27 | 9 | 2 | 8 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 28 | 9 | 3 | 8 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 29 | 9 | 4 | 8 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 30 | 9 | 5 | 8 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 31 | 9 | 6 | 8 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 32 | 9 | 7 | 8 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 33 | 9 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| 34 | 10 | 1 | 8 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 35 | 10 | 2 | 8 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 36 | 10 | 3 | 8 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 37 | 10 | 4 | 8 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 38 | 10 | 5 | 8 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 39 | 10 | 6 | 8 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 40 | 10 | 7 | 8 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 41 | 10 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| 42 | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- | -- |

Choose the functional form for the baseline hazard function

| id | studytime | j | lnj | j2 | j3 | dur1 | dur2 | dur3 | dur4 | dur5 | dur6 | d1 | d2 | d3 | d4 | d5 | d6 | d7 | d8 | d9 | d10 |
|----|-----------|---|----------|----|-----|------|------|------|------|------|------|----|----|----|----|----|----|----|----|----|-----|
| 1 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 2 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 2 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 2 | 1.609438 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 3 | 2.079442 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 4 | 2.795388 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 8 | 2 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 8 | 2 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 2 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 2 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 9 | 3 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 9 | 3 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 3 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 3 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 10 | 8 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 10 | 8 | 6 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 10 | 8 | 7 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 10 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |
| 10 | 8 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 10 | 8 | 6 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 10 | 8 | 7 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 10 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |

log(time)

Choose the functional form for the baseline hazard function

| id | studytime | lnj | j | j2 | j3 | dur1 | dur2 | dur3 | dur4 | dur5 | dur6 | d1 | d2 | d3 | d4 | d5 | d6 | d7 | d8 | d9 | d10 |
|----|-----------|----------|---|----|-----|------|------|------|------|------|------|----|----|----|----|----|----|----|----|----|-----|
| 1 | 1 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 1 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 2 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 2 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 1.098612 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1.098612 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1.386294 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 1.098612 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 1.386294 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 0 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | .6931472 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 1.386294 | 5 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 5 | 2.079442 | 6 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 1.386294 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 2.079442 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 9 | 8 | 0 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 9 | 8 | .6931472 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 9 | 8 | 1.386294 | 5 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 9 | 8 | 2.079442 | 6 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 10 | 8 | 0 | 1 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | .6931472 | 2 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 1.098612 | 3 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 1.386294 | 4 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 10 | 8 | 1.609438 | 5 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 10 | 8 | 1.791759 | 6 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 10 | 8 | 1.94591 | 7 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 10 | 8 | 2.079442 | 8 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 |

Polynomial in time (e.g., $\alpha t + \gamma t^2 + \delta t^3 \dots$)

Choose the functional form for the baseline hazard function

| id | j | studytime | lnj | j2 | j3 | dur1 | dur2 | dur3 | dur4 | dur5 | dur6 | d1 | d2 | d3 | d4 | d5 | d6 | d7 | d8 | d9 | d10 |
|----|----|-----------|-----|----------|----|------|------|------|------|------|------|----|----|----|----|----|----|----|----|----|-----|
| 1 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 2 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 3 | 1 | 2 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 2 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1 | 3 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 2 | 3 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 4 | 3 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 5 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 5 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 5 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 6 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 6 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 6 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 6 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 7 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 7 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 18 | 7 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 7 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 20 | 7 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 21 | 8 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 8 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 8 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 8 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 25 | 8 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 26 | 9 | 1 | 8 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27 | 9 | 2 | 8 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 28 | 9 | 3 | 8 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 29 | 9 | 4 | 8 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 30 | 9 | 5 | 8 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 31 | 9 | 6 | 8 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 32 | 9 | 7 | 8 | 1.94591 | 49 | 343 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 33 | 9 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 34 | 10 | 1 | 8 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 35 | 10 | 2 | 8 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 36 | 10 | 3 | 8 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 37 | 10 | 4 | 8 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 38 | 10 | 5 | 8 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 39 | 10 | 6 | 8 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 40 | 10 | 7 | 8 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 41 | 10 | 8 | 8 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |

Piece-wise constant

Choose the functional form for the baseline hazard function

| | id | j | studytime | lnj | j2 | j3 | dur1 | dur2 | dur3 | dur4 | dur5 | dur6 | d1 | d2 | d3 | d4 | d5 | d6 | d7 | d8 | d9 | d10 |
|----|----|---|-----------|----------|----|-----|------|------|------|------|------|------|----|----|----|----|----|----|----|----|----|-----|
| 1 | 1 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 2 | 2 | 1 | 1 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 3 | 3 | 1 | 2 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 4 | 3 | 2 | 2 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 5 | 4 | 1 | 3 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 6 | 4 | 2 | 3 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 7 | 4 | 3 | 3 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 8 | 5 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 9 | 5 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 10 | 5 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 11 | 5 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 12 | 6 | 1 | 4 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 13 | 6 | 2 | 4 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 14 | 6 | 3 | 4 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 15 | 6 | 4 | 4 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 16 | 7 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 17 | 7 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 18 | 7 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 19 | 7 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 20 | 7 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 21 | 8 | 1 | 5 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 22 | 8 | 2 | 5 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 23 | 8 | 3 | 5 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 24 | 8 | 4 | 5 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 25 | 8 | 5 | 5 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 26 | 9 | 1 | 6 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 27 | 9 | 2 | 6 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 28 | 9 | 3 | 6 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 29 | 9 | 4 | 6 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 30 | 9 | 5 | 6 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 31 | 9 | 6 | 6 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 32 | 9 | 7 | 6 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 33 | 9 | 8 | 6 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |
| 34 | 10 | 1 | 6 | 0 | 1 | 1 | 1 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 35 | 10 | 2 | 6 | .6931472 | 4 | 8 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 36 | 10 | 3 | 6 | 1.098612 | 9 | 27 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 |
| 37 | 10 | 4 | 6 | 1.386294 | 16 | 64 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 |
| 38 | 10 | 5 | 6 | 1.609438 | 25 | 125 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 | 0 |
| 39 | 10 | 6 | 6 | 1.791759 | 36 | 216 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 | 0 |
| 40 | 10 | 7 | 6 | 1.94591 | 49 | 343 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 | 0 |
| 41 | 10 | 8 | 6 | 2.079442 | 64 | 512 | 0 | 1 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 0 | 1 | 0 | 0 |

Fully non-parametric (e.g., $\gamma_1 * j_1 + \gamma_2 * j_2 + \gamma_3 * j_3 + \dots + \gamma_k * j_k$)

Estimation of a discrete-time model

Logistic model

`logit depvar varlist, [or noconstant]`

`'depvar'` is the (new) event variable – dead in the illustration – and `'varlist'` refers to the explanatory variables (covariates) together with the variables used to summarise the baseline hazard function.

The `'noconstant'` option means estimate a model without a constant term – we mainly use this for estimating models with a fully non-parametric baseline hazard.

Estimation of a discrete-time model

Logistic model

Odds ratios of hazard rates refer to ratios of form $\frac{h_1/(1-h_1)}{h_0/(1-h_0)}$ from a one-unit change in an explanatory variable from zero to one ($\Delta X = 1$ as $X = 0 \rightarrow X = 1$).

I personally find these difficult to interpret.

On the other hand, as $h \rightarrow 0$, the odds ratio tends to the hazard ratio $\frac{h_1}{h_0}$, which does have a ready interpretation.

Estimation of a discrete-time logistic (or logit) model

logit dead drug age lnj, nolog

Logistic regression

Log likelihood = **-111.16102**

Number of obs = **744**

LR chi2(3) = **35.41**

Prob > chi2 = **0.0000**

Pseudo R2 = **0.1374**

| dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] | |
|-------|------------------|-----------------|--------------|--------------|----------------------|------------------|
| drug | -2.297714 | .4388526 | -5.24 | 0.000 | -3.157849 | -1.437579 |
| age | .1253644 | .0391588 | 3.20 | 0.001 | .0486145 | .2021143 |
| lnj | .6781536 | .2582365 | 2.63 | 0.009 | .1720194 | 1.184288 |
| _cons | -10.25004 | 2.383314 | -4.30 | 0.000 | -14.92125 | -5.578826 |

- Drug recipients ($\text{drug} = 1$) have lower hazard rates ($\beta_{\text{drug}} < 0$ if $\text{drug} = 0 \rightarrow \text{drug} = 1$)
- Hazard rate increases with age ($\beta_{\text{age}} > 0$ if $\text{age} \uparrow$)
- Baseline hazard rises with elapsed survival time ($\ln j > 0$)

Comparison of discrete and continuous-time models

Discrete time (logistic): `logit dead drug age lnj, nolog`

```
Logistic regression                                Number of obs =   744
                                                    LR chi2(3)      =   35.41
                                                    Prob > chi2     =   0.0000
Log likelihood = -111.16102                        Pseudo R2      =   0.1374
```

| | dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] |
|-------|------|-------------|-----------|-------|-------|----------------------|
| drug | | -2.297714 | .4388526 | -5.24 | 0.000 | -3.157849 -1.437579 |
| age | | .1253644 | .0391588 | 3.20 | 0.001 | .0486145 .2021143 |
| lnj | | .6781536 | .2582365 | 2.63 | 0.009 | .1720194 1.184288 |
| _cons | | -10.25004 | 2.383314 | -4.30 | 0.000 | -14.92125 -5.578826 |

Continuous time (Weibull): `streg drug age, dist(weibull) nolog`

`nohr`

```
Weibull regression -- log relative-hazard form
```

```
No. of subjects =      48                Number of obs   =      48
No. of failures =      31
Time at risk   =      744
Log likelihood  = -42.931335            LR chi2(2)       =      35.39
                                                    Prob > chi2     =      0.0000
```

| _t | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|-------|-----------|-----------|--------|-------|----------------------|
| drug | -2.196936 | .4087791 | -5.374 | 0.000 | -2.998129 -1.395744 |
| age | .1202027 | .0371599 | 3.235 | 0.001 | .0473707 .1930348 |
| _cons | -10.58396 | 2.326271 | -4.550 | 0.000 | -15.14337 -6.024553 |
| /ln_p | .5204297 | .1389037 | 3.747 | 0.000 | .2481834 .792676 |
| p | 1.682751 | .2337403 | | | 1.281695 2.209301 |
| 1/p | .5942651 | .0825456 | | | .452632 .7802168 |

Comparison of discrete and continuous-time models

Discrete time (logistic): `logit dead drug age lnj, nolog`

Logistic regression

Number of obs = **744**

LR chi2(3) = **35.41**

Prob > chi2 = **0.0000**

Pseudo R2 = **0.1374**

Log likelihood = **-111.16102**

| | dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] |
|-------|------|------------------|-----------------|--------------|--------------|----------------------------|
| drug | | -2.297714 | .4388526 | -5.24 | 0.000 | -3.157849 -1.437579 |
| age | | .1253644 | .0391588 | 3.20 | 0.001 | .0486145 .2021143 |
| lnj | | .6781536 | .2582365 | 2.63 | 0.009 | .1720194 1.184288 |
| _cons | | -10.25004 | 2.383314 | -4.30 | 0.000 | -14.92125 -5.578826 |

Continuous time (Weibull): `streg drug age, dist(weibull) nolog` `nohr`

No. of subjects = 48
No. of failures = 31
Time at risk = 744

Number of obs = **48**

LR chi2(2) = 35.39

Prob > chi2 = 0.0000

Log likelihood = -42.931335

| _t | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] | |
|-------|-----------|-----------|--------|-------|----------------------|-----------|
| drug | -2.196936 | .4087791 | -5.374 | 0.000 | -2.998129 | -1.395744 |
| age | .1202027 | .0371599 | 3.235 | 0.001 | .0473707 | .1930348 |
| _cons | -10.58396 | 2.326271 | -4.550 | 0.000 | -15.14337 | -6.024553 |
| ----- | | | | | | |
| /ln_p | .5204297 | .1389037 | 3.747 | 0.000 | .2481834 | .792676 |
| ----- | | | | | | |
| p | 1.682751 | .2337403 | | | 1.281695 | 2.209301 |
| 1/p | .5942651 | .0825456 | | | .452632 | .7802168 |
| ----- | | | | | | |

Comparison of discrete and continuous-time models

Discrete time (logistic): `logit dead drug age lnj, nolog`

Logistic regression

Number of obs = 744

LR chi2(3) = 35.41

Prob > chi2 = 0.0000

Log likelihood = -111.16102

Pseudo R2 = 0.1374

| | dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] |
|-------|------|-------------|-----------|-------|-------|----------------------|
| drug | | -2.297714 | .4388526 | -5.24 | 0.000 | -3.157849 -1.437579 |
| age | | .1253644 | .0391588 | 3.20 | 0.001 | .0486145 .2021143 |
| lnj | | .6781536 | .2582365 | 2.63 | 0.009 | .1720194 1.184288 |
| _cons | | -10.25004 | 2.383314 | -4.30 | 0.000 | -14.92125 -5.578826 |

Continuous time (Weibull): `streg drug age, dist(weibull) nolog`

`nohr`

No. of subjects = 48
No. of failures = 31
Time at risk = 744
Log likelihood = -42.931335

Number of obs = 48
LR chi2(2) = 35.39
Prob > chi2 = 0.0000

| _t | Coef. | Std. Err. | z | P> z | [95% Conf. Interval] |
|-------|-----------|-----------|--------|-------|----------------------|
| drug | -2.196936 | .4087791 | -5.374 | 0.000 | -2.998129 -1.395744 |
| age | .1202027 | .0371599 | 3.235 | 0.001 | .0473707 .1930348 |
| _cons | -10.58396 | 2.326271 | -4.550 | 0.000 | -15.14337 -6.024553 |
| /ln_p | .5204297 | .1389037 | 3.747 | 0.000 | .2481834 .792676 |
| p | 1.682751 | .2337403 | | | 1.281695 2.209301 |
| 1/p | .5942651 | .0825456 | | | .452632 .7802168 |

Comparison of discrete and continuous-time models

Discrete time models

Pros

- Time-varying covariates $X(t)$
- More flexible baseline hazard function

Continuous time models

Pros

- Cost-effective dataset construction
- Time-invariant covariates X
- Long time span with frequent events

Estimation of a discrete-time logit model

logistic dead drug age ln

| dead | Odds ratio | Std. err. | z | P> z | [95% conf. interval] | |
|-------|------------|-----------|-------|-------|----------------------|----------|
| drug | .1004883 | .0440996 | -5.24 | 0.000 | .0425171 | .2375022 |
| age | 1.133561 | .044389 | 3.20 | 0.001 | 1.049816 | 1.223988 |
| lnj | 1.970237 | .5087869 | 2.63 | 0.009 | 1.187701 | 3.268358 |
| _cons | .0000354 | .0000843 | -4.30 | 0.000 | 3.31e-07 | .003777 |

logit dead drug age lnj, or

| dead | Odds ratio | Std. err. | z | P> z | [95% conf. interval] | |
|-------|------------|-----------|-------|-------|----------------------|----------|
| drug | .1004883 | .0440996 | -5.24 | 0.000 | .0425171 | .2375022 |
| age | 1.133561 | .044389 | 3.20 | 0.001 | 1.049816 | 1.223988 |
| lnj | 1.970237 | .5087869 | 2.63 | 0.009 | 1.187701 | 3.268358 |
| _cons | .0000354 | .0000843 | -4.30 | 0.000 | 3.31e-07 | .003777 |

Logit and logistic models are equivalent.

logistic → OR

logit → coefficients (log odds)

logit [...], 'or' → OR

Estimation of a discrete-time logit model

Cubic polynomial: `logit dead drug age j j2 j3, nolog`

| dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] | |
|-------|-------------|-----------|-------|-------|----------------------|-----------|
| drug | -2.39327 | .4626735 | -5.17 | 0.000 | -3.300093 | -1.486447 |
| age | .123828 | .039411 | 3.14 | 0.002 | .0465838 | .2010721 |
| j | .0881094 | .1809724 | 0.49 | 0.626 | -.26659 | .4428087 |
| j2 | .0009124 | .0127206 | 0.07 | 0.943 | -.0240196 | .0258444 |
| j3 | -.0000481 | .0002508 | -0.19 | 0.848 | -.0005397 | .0004436 |
| _cons | -9.678605 | 2.441734 | -3.96 | 0.000 | -14.46431 | -4.892895 |

Piece-wise constant baseline: `logit dead drug age dur1-dur6, nocons nolog`

| dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] | |
|------|-------------|-----------|-------|-------|----------------------|-----------|
| drug | -2.280694 | .4565036 | -5.00 | 0.000 | -3.175425 | -1.385963 |
| age | .1190845 | .0389076 | 3.06 | 0.002 | .042827 | .195342 |
| dur1 | -9.098389 | 2.265066 | -4.02 | 0.000 | -13.53784 | -4.658941 |
| dur2 | -8.434023 | 2.178112 | -3.87 | 0.000 | -12.70304 | -4.165002 |
| dur3 | -8.438919 | 2.182175 | -3.87 | 0.000 | -12.7159 | -4.161936 |
| dur4 | -7.596855 | 2.169892 | -3.50 | 0.000 | -11.84977 | -3.343945 |
| dur5 | -7.445229 | 2.273184 | -3.28 | 0.001 | -11.90059 | -2.98987 |
| dur6 | -7.499636 | 2.382459 | -3.15 | 0.002 | -12.16917 | -2.830103 |

Estimation of a discrete-time logit model

Fully non-parametric baseline: logit dead drug age d1-d39, nolog

note: **d38** != 0 predicts failure perfectly;
d38 omitted and 1 obs not used.

note: **d39** != 0 predicts failure perfectly;
d39 omitted and 1 obs not used.

Logistic regression

Number of obs = 573

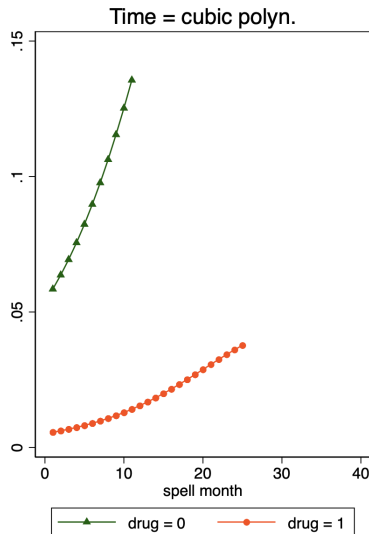
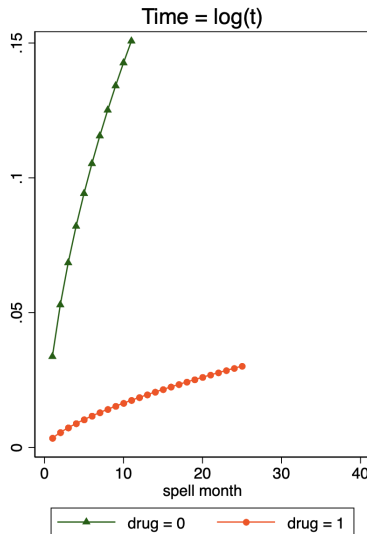
Wald chi2(23) = 159.53

Log likelihood = -96.988418

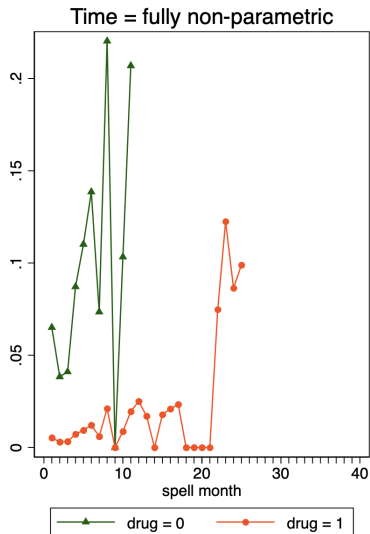
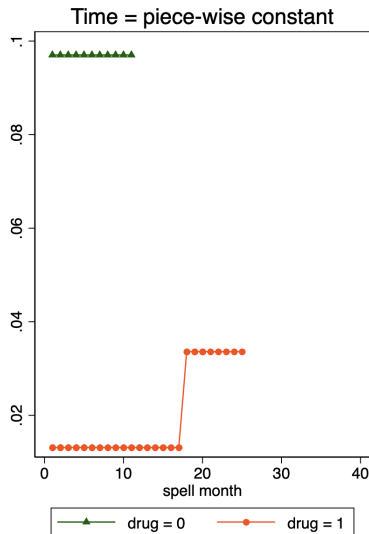
Prob > chi2 = 0.0000

| dead | Coefficient | Std. err. | z | P> z | [95% conf. interval] | |
|------|-------------|-----------|-------|-------|----------------------|-----------|
| drug | -2.572486 | .5104139 | -5.04 | 0.000 | -3.572879 | -1.572093 |
| age | .130051 | .0406452 | 3.20 | 0.001 | .0503879 | .209714 |
| d1 | -9.817929 | 2.507439 | -3.92 | 0.000 | -14.73242 | -4.903439 |
| d2 | -10.37478 | 2.586792 | -4.01 | 0.000 | -15.4448 | -5.304765 |
| d3 | -10.30505 | 2.584837 | -3.99 | 0.000 | -15.37124 | -5.238863 |
| d4 | -9.501784 | 2.473035 | -3.84 | 0.000 | -14.34884 | -4.654724 |
| d5 | -9.242106 | 2.418842 | -3.82 | 0.000 | -13.98295 | -4.501263 |
| d6 | -8.980254 | 2.376713 | -3.78 | 0.000 | -13.63853 | -4.321983 |
| d7 | -9.687424 | 2.476246 | -3.91 | 0.000 | -14.54078 | -4.834071 |
| d8 | -8.416641 | 2.31529 | -3.64 | 0.000 | -12.95453 | -3.878756 |
| d9 | 0 (omitted) | | | | | |
| d10 | -9.313809 | 2.482374 | -3.75 | 0.000 | -14.17917 | -4.448445 |
| d11 | -8.496452 | 2.365123 | -3.59 | 0.000 | -13.13201 | -3.860896 |
| d12 | -8.242216 | 2.375201 | -3.47 | 0.001 | -12.89752 | -3.586907 |
| d13 | -8.635718 | 2.447953 | -3.53 | 0.000 | -13.43362 | -3.837819 |
| d14 | 0 (omitted) | | | | | |
| d15 | -8.588234 | 2.440998 | -3.52 | 0.000 | -13.3725 | -3.803967 |
| d16 | -8.425413 | 2.464191 | -3.42 | 0.001 | -13.25514 | -3.595688 |
| d17 | -8.312966 | 2.4293 | -3.42 | 0.001 | -13.0744 | -3.5515 |

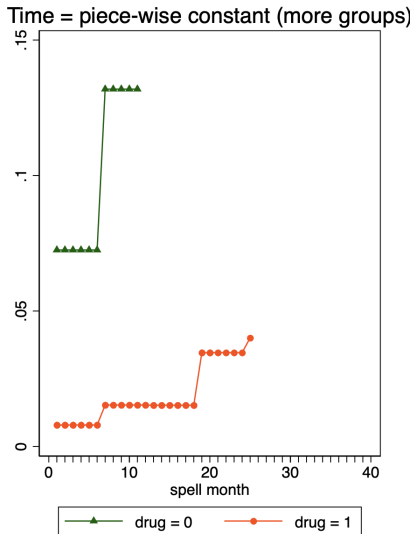
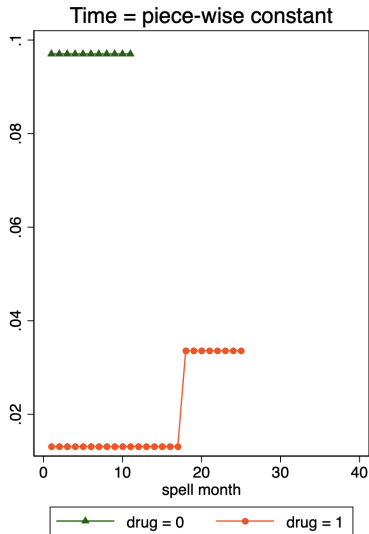
Graphical representation of a discrete-time logit hazard



Graphical representation of a discrete-time logit hazard



Graphical representation of a discrete-time logit hazard



Comparison of baseline hazard functions

Parametric functions (logarithmic, quadratic, cubic, etc.)

Pros

- Smooth hazard
- Few observations (units of analysis or transitions)

Partially or fully parametric functions (piecewise, time-dummy, etc.)

Pros

- Baseline of your choice
- Easy interpretation
- Many observations (units of analysis or transitions)

More on discrete-time models

- Parametric models
 - ▶ Complementary log-log regression
 - ▶ GLM
- Within and out-of-sample prediction
- Survival functions
- Implementation on software STATA

Full display of these topics on the Lecture notes and do-files.