B Additional Reduced Form Results

B.1 Robustness of Results to Transformations

In the main text, I use the inverse hyperbolic sine transformation. Bellemare and Wichman (2020) show that for a model with continuous variables x and y and specification $\sinh^{-1}(y) = \beta x + \varepsilon$, the elasticity of y with respect to x is $(\beta x/y)\sqrt{y^2 + 1} \approx \beta x$ whenever $y \geq 2$. This may also refine estimates using the more common $\log(y+1)$ transformation. In this section, I show robustness of these results to other transformations. Throughout, these robustness checks use simple TWFE regressions (rather than the LP-DID approach).

Table B.2 demonstrates that results are robust to two standard transformations for skewed spending variables: the inverse hyperbolic sine transform, as reported in the main text, and the $\log(y+1)$ transformation.

B.2 Discussion of Appropriate Control Groups

In reduced-form estimation, I include both not-yet-treated and never-treated households in the control group for each cohort. Doing so allows for separate identification of dynamic treatment effects from time fixed-effects, but may come at the cost of introducing violations in the parallel trends assumption: namely, it may be the case that in the absence of major health events, treated households and never-treated households may have had differing spending and utilization trajectories.

This is less likely to be true in my setting than in other contexts, for a number of reasons. First, the diagnoses included here span a large range of chronic conditions, including those that do not ultimately affect lifetime health or spending for diagnosed individuals (e.g., asthma or major depressive disorder). As spending and quality of life are unlikely to be meaningfully affected by these diagnoses, it is unlikely that in the absence of a diagnosis, these groups would have differed on some other measure affecting health utilization. Second—and perhaps more important—these diagnoses are largely heritable and unaffected by health behaviors such as diet, exercise, or engaging in other risky behaviors such as smoking. Given this, there is unlikely to be selection of less healthy households (whose health may have deteriorated for other, unobserved, reasons in the absence of diagnosis) into the treatment versus control groups. Importantly, my results are robust to considering only not-yet-treated households in the control group (Figure B.1).

	OOP, chror	OOP, chronic diagnosis	OOP, acut	OOP, acute diagnosis	Wellness	Wellness spending	Low-value	Low-value spending
	$sinh^{-1}(y)$	log(y+1)	$ sinh^{-1}(y) $	log(y+1)	$sinh^{-1}(y)$	log(y+1)	$sinh^{-1}(y)$	log(y+1)
t-5	-0.02	-0.02	-0.11	-0.10	**60.0-	-0.08**	+90.0-	-0.05*
	(0.028)	(0.026)	(0.070)	(0.064)	(0.031)	(0.028)	(0.033)	(0.03)
t-4	0.02	0.01	-0.11	-0.10	-0.03	-0.03	-0.04	-0.03
	(0.024)	(0.022)	(0.059)	(0.055)	(0.026)	(0.024)	(0.028)	(0.024)
t-3	0.00	0.00	-0.02	-0.02	-0.02	-0.02	-0.03	-0.02
	(0.020)	(0.018)	(0.052)	(0.048)	(0.022)	(0.020)	(0.023)	(0.021)
t-2	-0.00	-0.00	-0.07	-0.00	-0.03	-0.03	-0.01	-0.01
	(0.017)	(0.015)	(0.045)	(0.042)	(0.019)	(0.017)	(0.020)	(0.018)
t-1	I	1	1	ı	ı	1	1	1
t	***80.0	0.07***	-0.01	-0.01	0.12***	0.11***	0.05*	0.04*
	(0.014)	(0.013)	(0.041)	(0.037)	(0.016)	(0.015)	(0.018)	(0.016)
t+1	0.10***	0.10***	0.10*	*60.0	0.09	0.08***	0.05	0.04**
	(0.016)	(0.014)	(0.047)	(0.043)	(0.017)	(0.016)	(0.019)	(0.017)
t+2	0.10***	0.09***	90.0	0.07	0.10***	0.10***	0.05*	0.04*
	(0.018)	(0.017)	(0.055)	(0.050)	(0.020)	(0.018)	(0.021)	(0.019)
t+3	0.09***	0.08***	0.10	0.09	0.11***	0.10***	0.04	0.04
	(0.018)	(0.019)	(0.062)	(0.057)	(0.022)	(0.020)	(0.024)	(0.021)
t+4	0.08	0.08***	0.14	0.13	0.13***	0.12***	**60.0	0.08**
	(0.025)	(0.022)	(0.074)	(0.068)	(0.025)	(0.023)	(0.028)	(0.024)
t + 5	0.07***	*90.0	0.12	0.12	0.10***	0.09***	0.12***	0.11***
	(0.030)	(0.028)	(0.088)	(0.081)	(0.030)	(0.027)	(0.033)	(0.029)
R^2	0.51	0.52	0.50	0.51	0.43	0.44	0.20	0.20
N	1,538,161	1,538,161	1,374,359	1,374,359	1,538,161	1,538,161	1,538,161	1,538,161

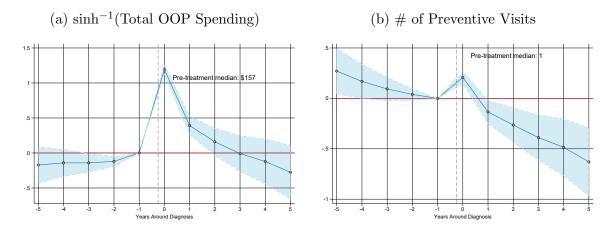
Notes: This table presents estimates for the main event study regression results reported in the paper. The first column of each pair of results are the results shown graphically in the text, while the second column uses the log transformation. Standard errors are clustered at the household level. *p < 0.05,** p < 0.01,*** p < 0.001

Table B.2. Robustness: Inverse Hyperbolic Sine & Log Transformations

B.3 Household Response to Major Medical Events

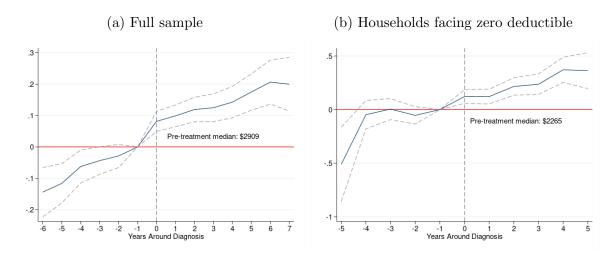
In this section, I include additional results from a suite of two-way fixed effects models estimating the causal effect of major medical events on health behaviors. Figures B.3 and B.4 illustrate the estimated effect on billed spending for both chronic and acute medical events.

Figure B.2. Effect of Chronic Diagnoses on OWN Healthcare Utilization



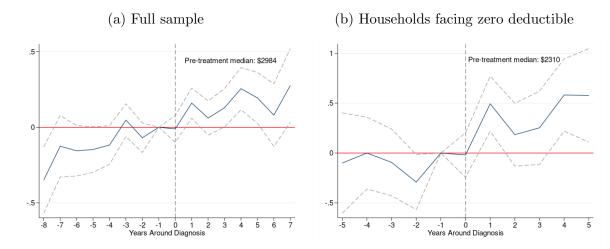
Notes: Figures show regression coefficients from "stacked" TWFE regressions, with 95% confidence intervals. Regressions estimate the effect of a new chronic diagnosis on the medical utilization of the diagnosed individual. In panel (a), the dependent variable is the inverse hyperbolic sine of total OOP spending; panel (b) estimates the effect on the number of household preventive services per year using Poisson regression. Coefficients are presented relative to the year prior to diagnosis. Spending is measured in 2020 USD. Standard errors are clustered at the household level.

Figure B.3. Estimated Effect of a Chronic Diagnosis on Billed Non-Diagnosed Spending



Note: Dependent variable is the inverse hyperbolic sine of total billed spending for all non-diagnosed individuals in a household. Coefficients are presented relative to the year prior to diagnosis. Spending is measured in 2020 USD. Standard errors are clustered at the household level.

Figure B.4. Estimated Effect of an Acute Health Event on Billed Non-Diagnosed Spending



Note: Dependent variable is the inverse hyperbolic sine of total billed spending for all non-diagnosed individuals in a household. Coefficients are presented relative to the year prior to diagnosis. Spending is measured in 2020 USD. Standard errors are clustered at the household level.

I also explore the effect of acute health events on household out-of-pocket spending, similar to Figure 1 in the text. In general, acute events do not generate the same household response that chronic diagnoses do.

Finally, I find a strong extensive margin response among household members who experience major medical events in their families. Table B.4 shows that individuals are more likely to spend any positive amount (billed and OOP) on medical care, use any outpatient

visits or preventive care, or fill any prescriptions. This effect is strongest in the year of the diagnosis and decays slightly over time, but remains significant for five years following the health event.

Table B.4. Estimated Extensive Margin Health Effects of Family Diagnosis

	Year of Event $(t=0)$	Following Years $(t > 0, \text{ averaged})$
Any Billed Spending	1.54***	0.60***
	(0.08)	(0.13)
Any OOP Spending	2.62***	1.41***
	(0.11)	(0.18)
Any Outpatient Visits	2.20***	0.65***
	(0.09)	(0.15)
Any Preventive Care	3.23***	0.90***
	(0.15)	(0.22)
Any Prescription Fills	4.74***	2.45***
	(0.41)	(0.53)

Notes: Table shows estimated difference-in-difference regression coefficients for the effect of a new chronic diagnosis (N=1,538,161). Outcome variables are dummy variables indicating the likelihood of each outcome, scaled from 1 to 100. Standard errors clustered at the household level are reported in parentheses.

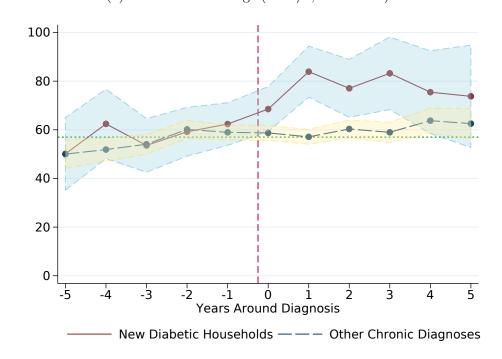
B.4 Disease-Specific Screenings: Additional Results

Figure B.5 illustrates, in the raw data, the central takeaway from Section 3.2 in the paper: that households respond to a new diagnosis by selecting into services related to that diagnosis. The figure shows that households with a newly diagnosed diabetic takeup more diabetes screenings than households affected by a non-diabetes chronic diagnosis.

^{*} p < 0.05, ** p < 0.01, *** p < 0.001.

Figure B.5. Rate of Diabetes Screenings Around Time of Diagnoses

(a) Diabetes Screenings (Rate/1,000 Adults)



Notes: Figure plots re-centered time series that depict the associations between household diagnoses and the takeup of diabetes screenings for adults within a household. Utilization rates of diabetes screenings for non-diagnosed household members 18 years of age and older, measured in rates per 1,000 adults, are shown, including averages and 95% confidence intervals. The top (solid maroon) line indicates average rates for households who experience a diabetes diagnosis, and the bottom (dashed navy) line indicates rates for those affected by other chronic diagnoses. The horizontal, dotted green line indicates the average utilization rate for all other households in the sample who do not experience a diagnosis, about 59 screenings per 1,000 adults. Individuals whose family members are diagnosed with conditions other than diabetes do not appear to significantly alter their screening behaviors from unaffected households (whose average is depicted in the horizontal, dotted green line). On the other hand, household members of those diagnosed with diabetes increase screenings in the first three years following the diagnosis, being about 36% more likely to be screened for diabetes than unaffected individuals.

B.5 Intra-Familial Relationships

For example, while a diabetes diagnosis is most likely to affect adult household members with similar lifestyles to the original diagnosed individual, a mental health diagnosis may have a stronger genetic component. Hence, households where an adult was diagnosed with diabetes may choose to screen other adults, such as spouses, while households where someone received a mental health diagnosis may choose to screen children or siblings of the affected individual.

Screening Diagnosis	Hypertension Any Chronic	Diabetes Diabetes	Cholesterol Diabetes	High BMI Diabetes	Cancer Cancer	Depression $MDD/Bipolar$
$\operatorname{Post}_t \times \operatorname{Diagnosis}_f$	0.39*** (0.03)	-0.85*** (0.21)	-2.20*** (0.29)	-0.38** (0.12)	2.55*** (0.43)	0.30** (0.10)
$\operatorname{Post}_t \times \operatorname{Diagnosis}_f \times \operatorname{Parent}_j$	-0.34** (0.11)	3.49* (1.71)	3.73 (2.26)	1.73* (0.70)	-1.90 (2.49)	-0.93*** (0.13)
$\operatorname{Post}_t \times \operatorname{Diagnosis}_f \times \operatorname{Spouse}_j$	-0.74*** (0.13)	2.54^{***} (0.45)	5.15*** (0.60)	1.03*** (0.20)	-3.33*** (0.81)	-0.62*** (0.11)
$\operatorname{Post}_t \times \operatorname{Diagnosis}_f \times \operatorname{Sibling}_j$	$0.09 \\ (0.04)$	0.76 (1.09)	2.89 (1.86)	0.16 (0.69)	1.56 (1.55)	0.68^* (0.32)
Observations Adjusted R^2	$4,039,602 \\ 0.024$	3,680,725 0.217	3,680,725 0.388	3,680,725 -0.025	3,671,064 0.473	3,724,608 0.117

Standard errors in parentheses

Notes: Table shows results of a difference-in-differences estimation strategy highlighting the potentially differential effects of chronic illnesses on preventive care utilization by household relationships. The primary outcome variable in each column is a screening or new diagnosis, shown in the top row. The specific chronic illness used as the Diagnosis f dummy is shown in the second row. Standard errors are clustered at the household level.

Table B.5. DDD Estimates: Disease-Specific Spending

To assess these potentially heterogeneous effects, I utilize a simple difference-in-differences framework. In Table B.5, I present estimation results for the same six diagnosis/outcome pairs shown in Table 3. The dependent variable—either a screening or a new diagnosis—is shown in the top row, with the treatment variable—the chronic illness affecting the household—below in italics. I explore the potentially heterogeneous responses for four family relationships: parents, spouses, siblings, and children of the affected individual, with children as the reference group.

Throughout, I find consistent evidence that households respond by not only selecting

^{*} p < 0.05, ** p < 0.01, *** p < 0.001

¹The vast majority of diabetes diagnoses in my sample are for Type 2 Diabetes Mellitus, which generally affects adults and risk of which is increased or decreased based on specific lifestyle choices, such as diet and exercise. The same is not as true for Type 1 DM diagnoses.

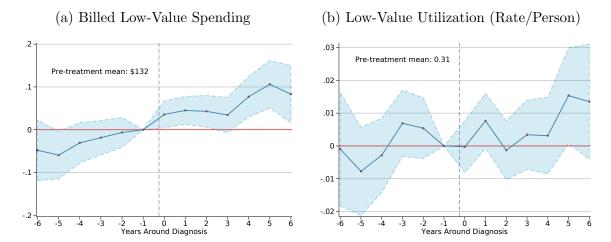
screenings associated with the health events they experienced, but also selecting which individuals to screen based on their associated risk. New hypertension diagnoses following a chronic event are concentrated among children rather than parents and spouses, suggesting that households are identifying previously ignored risks among the previously lower-risk members of their household. Additionally, households affected with diabetes focus screenings on spouses more than on children, consistent with the lifestyle factors that affect diabetes risk. In contrast, households affected with chronic illnesses that communicate a greater level of genetic risk—cancer and mental health conditions—choose instead to screen children and siblings (in the case of mental health conditions) more than parents or spouses.

As an alternative test to salience, I explore how each individual reacts differently to a diagnosis in their home stratified by their pre-event diagnostic risk. This uses the risk proxy described in Section 5 of the main paper.

B.6 Low Value Care

Figure B.6 presents estimates for the effect of new chronic diagnoses on the overall utilization of low-value services, including both total spending and overall utilization rates. Major health events are associated with a small increase in overall low-value spending of about 5 percent. In contrast, the average rate of service use among non-diagnosed household members does not change meaningfully following a diagnosis. Table B.6 depicts the event study regressions discussed in the text.

Figure B.6. Chronic Diagnoses Increase Utilization of Low-Value Care



Notes: This figure shows estimated coefficients and 95% confidence intervals for the effect of major health events on the use of low-value services (see Appendix A for definitions). In the first panel, the outcome is the inverse hyperbolic sine of billed spending. In the second panel, the outcome is the number of low-value services used per household member. Spending is measured in 2020 USD. Standard errors are clustered at the household level.

All Pediatric Spending Rai	te Sper	Adult Drugs ding Rate	Adult Imaging Spending Rat	naging Rate	Adult Screening Spending Rate	reening Rate	Adult Surgery Spending Rat	Surgery Rate
0.02***		-0.00	0.03***	0.01***	0.10***	0.03***	-0.10***	-0.04**
$ \begin{array}{c cccc} (0.017) & (0.003) & (0.000) \\ 0.192 & 0.228 & 0.143 \end{array} $		(0.000)	(0.013)	(0.002) 0.141	(0.014) 0.163	(0.005) 0.151	(0.012)	(0.002) 0.255
-0.02*	,	0.00*	0.01	-0.00	-0.10***	***50.0-	0.09***	0.03***
$ \begin{array}{c ccc} (0.014) & (0.008) & (0.003) & (\\ -0.02 & -0.01 & 0.00 & \\ \end{array} $		(0.002) 0.00	(0.016) -0.01	(0.005) -0.01	(0.021) -0.03	(0.011) -0.09	(0.012) $0.04***$	$(0.004) \\ 0.02***$
(0.007) (0.002)		(0.001)	(0.013)	(0.004)	(0.019)	(0.010)	(0.010)	(0.003)
-0.01* 0.00		0.00	0.01	0.00	-0.02	0.00	0.01	0.01**
$(0.010) \qquad (0.005) \qquad (0.002) \qquad ($	<u> </u>	0.001)	(0.016)	(0.004)	(0.016)	(0.010)	(0.000)	(0.002)
I I		ı	ı	ı	ı	I	I	ı
0.008 0.00		0.00	0.01	0.01	0.03*	0.008	-0.03***	-0.01***
(0.004) (0.002)		(0.001)	(0.010)	(0.003)	(0.015)	(0.008)	(0.008)	(0.002)
0.01*** 0.00		0.00	0.03***	0.01***	0.07***	0.04**	-0.07***	-0.02***
(0.005) (0.002)		(0.001)	(0.011)	(0.003)	(0.015)	(0.008)	(0.009)	(0.003)
0.02***		00.0	0.02*	0.01**	0.06***	0.03	-0.08***	-0.03**
(0.005) (0.002)		(00.0	(0.012)	(0.003)	(0.016)	(0.000)	(0.011)	(0.003)
		00.0	0.03**	0.02***	0.07***	0.03**	-0.11***	-0.05***
(0.006) (0.002)		(0.001)	(0.013)	(0.004)	(0.018)	(0.011)	(0.013)	(0.005)
		0.00	0.06***	0.02***	0.10^{***}	0.03*	-0.10***	-0.05***
$(0.013) \qquad (0.007) (0.003) (0.003)$	_	(0.002)	(0.016)	(0.005)	(0.021)	(0.012)	(0.016)	(0.005)
$0.228 \mid 0.143$		0.259	0.123	0.141	0.163	0.151	0.230	0.255
1,538,161 $1,538,161$ $1,538,161$ $1,$		1,538,161	1,538,161	1,538,161	1,538,161	1,538,161	1,538,161	1,538,161

Notes: Table shows estimated difference-in-difference and event study regression coefficients for the effect of a new chronic diagnosis. Two outcome variables are reported for each category: the inverse hyperbolic sine of billed spending and the number of low-value services used per household member. See Appendix A for service definitions. Spending is measured in 2020 USD. Standard errors clustered at the household level are reported in parentheses. * p < 0.05, ** p < 0.01, *** p < 0.001.

Table B.6. Estimated Effects of Chronic Illness on Low-Value Care Utilization, by Category