

(Non)-Impact of Transmission Type on Fuel Consumption for 1973-74 Car Models

Alexander Lemm

October 25, 2015

Executive summary

This report examines the relationship between a set of variables and fuel consumption measured in miles per gallon (MPG). Looking only at the transmission type, our findings suggest that cars with an automatic transmission consume more fuel than cars with a manual transmission. However, this effect can no longer be observed when including other variables in our final regression model. Based on our findings MPG can be best expressed as linear combination of *Weight*, *Weight*² and *Quarter mile time*. The entire report including the respective code can be found on [GitHub](#).

Examining the Effect of Transmission Type on Fuel Consumption

We assume that the data was randomly sampled from the *1974 Motor Trend* magazine. Moreover, we deal with non-paired data which means that the independence condition is also satisfied between groups. Since both samples have $n < 30$ observations and are not strongly skewed we can apply a two sample t-test ($H_0 : \mu_{\text{automatic}} = \mu_{\text{manual}}$. $H_A : \mu_{\text{automatic}} \neq \mu_{\text{manual}}$):

estimate	estimate1	estimate2	statistic	p.value	parameter	conf.low	conf.high
-7.244939	17.14737	24.39231	-3.767123	0.0013736	18.33225	-11.28019	-3.209684

Because the p-value $0.0014 < 0.05$, we reject the null hypothesis. The data provides convincing evidence that fuel consumption indeed differs between transmission types. We are 95% confident that cars from the 1973/74 population with an automatic transmission consume between 11.28 and 3.21 less fuel than cars with a manual transmission.

Examining the Effect of Multiple Variables on Fuel Consumption

In this section we will determine whether the difference in fuel consumption by transmission type holds, if we take additional variables and their effect on fuel consumption into account.

The first step is to perform a linear regression of *MPG* on all other variables in the *mtcars* data set leveraging the *step()* function: This ensures that we will perform a variable selection leaving only the most important variables in the model. This model we refer to as **Model 1** includes 3 significant variables:

term	estimate	std.error	statistic	p.value
(Intercept)	9.617781	6.9595930	1.381946	0.1779152
wt	-3.916504	0.7112016	-5.506882	0.0000070
qsec	1.225886	0.2886696	4.246676	0.0002162
am	2.935837	1.4109045	2.080819	0.0467155

When performing model diagnostics we observe a non-linear pattern in the residual plot for Model 1 (Left side of Figure 2). This is a problem because all of the conclusions that we draw from the fit are suspect. Our

findings suggest that it is best to add $Weight^2$ to Model 1 to accommodate this non-linear relationship. We call this new model **Model 2**. When adding $Weight^2$ only a slight pattern can be observed in the residuals (Right side of Figure 2).

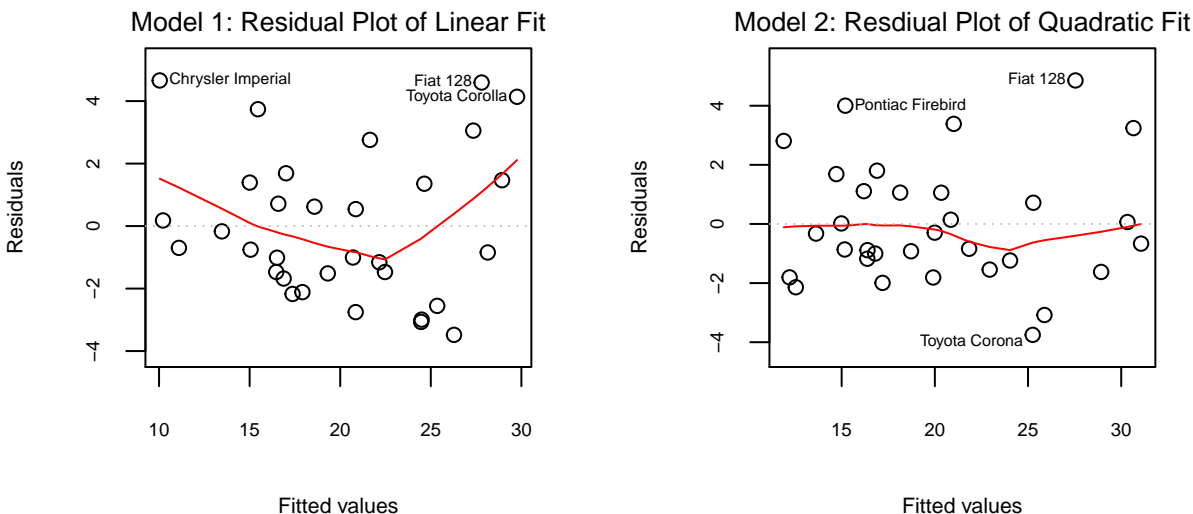


Figure 1: Plots of residuals versus fitted values for the *mtcars* data set. **Left:** A linear regression of *MPG* on *Weight*, *Quarter mile time* and *Transmission Type*. A pattern in the residuals indicates non-linearity in the data. **Right:** A linear regression of *MPG* on the same variables plus $Weight^2$. The former pattern nearly vanished.

Performing a hypothesis test comparing the two models with the `anova()` function reveals the following:

Model	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
Model 1	28	169.2859	NA	NA	NA	NA
Model 2	27	130.8573	1	38.42859	7.929032	0.0089778

Here the F-statistic is 7.93 and the associated p-value is 0.009. This provides evidence that the model containing the predictors *Weight* and $Weight^2$ is superior to the model that only contains the predictor *Weight*.

Model 2 is superior to Model 1 in terms of R^2 (0.8666 vs. 0.8336) and *RSE* (2.2015 vs. 2.4588). However, examining the individual p-values from the predictors of Model 2 reveals that *Transmission Type (AM)* is no longer significant:

term	estimate	std.error	statistic	p.value
(Intercept)	27.7376147	8.9574469	3.0965983	0.0045282
wt	-11.2490694	2.6807530	-4.1962349	0.0002629
I(wt^2)	0.9581950	0.3402858	2.8158538	0.0089778
qsec	0.9705112	0.2739061	3.5432263	0.0014614
am	1.0215185	1.4345500	0.7120828	0.4825217

This suggests that we might drop *Transmission Type (AM)* from the quadratic model. Dropping this predictor results in Model 3 with the respective summary information below:

term	estimate	std.error	statistic	p.value
(Intercept)	32.6418325	5.6767588	5.750083	0.0000036
wt	-12.4330965	2.0841792	-5.965464	0.0000020
I(wt^2)	1.0730270	0.2969987	3.612901	0.0011739
qsec	0.8598587	0.2235665	3.846099	0.0006339

Now, all included variables are highly significant again. Like Model 2, Model 3 does not show any pattern in the residual plot (not shown here).

Comparing all three models we can clearly see the superiority of Model 3 in terms of *Adjusted R²* and *RSE*:

Model	Adj.r.squared	RSE
Model 1	0.8335561	2.458846
Model 2	0.8665743	2.201492
Model 3	0.8689233	2.182028

Model 3 would be our final choice when modeling the relationship of specific car variables and fuel consumption. The 95% confidence intervals are as follows: (-16.702, -8.164) for *Weight*, (0.465, 1.681) for *Weight²*, and (0.402, 1.318) for *Quarter mile time*.