# Lecture 1

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### 1 Introduction

Finna learn which analyses to undertake and how to interpret them

#### Divided into 4 parts:

- 1. Basic probability theory
- 2. Statistical inference
- 3. Linear (regression) models
- 4. Time series analysis

Lecture notes are the primary course material

Exercises are heavily useful  $\rightarrow$  do these every week

#### 1.1 Assessment

- Part 1 40% midterm (in groups of up to 3)
- Part 2-60% final take-home exams

#### 1.2 Contact

- Office Hours: by apt
- Contact: Adam Lee
- for god's sake use email

# 2 Probability

$$S: \text{Sample space} \to \{x_1,\dots,x_n\} \to \text{cannot have duplicate numbers}$$
 
$$\mathcal{S}: \sigma\text{-algebra set of sets}$$
 
$$S = \{1,2,3\}$$
 
$$A \subset S$$
 
$$A = \{1,2\} \quad A = \{1\} \quad A = \{\} \to \text{This is the empty set } \emptyset$$
 
$$P: \mathcal{S} \to \mathbb{R} \text{ is a probability IF}$$
 
$$i)P(A) \geq 0 \quad \forall A \in \mathcal{S}$$
 
$$ii)P(\mathcal{S}) = 1$$
 
$$iii)A_1,\dots,A_n \text{ s.t. } i \neq j = )A_i \cap A_j = \emptyset \to P(u_{1,2}^\infty) = \sum_{i=1}^\infty P(A_i)$$

### 2.1 2.1.1 Events and probabilities

Prop: if 
$$P:\mathcal{S}\to\mathbb{R}$$
 is a probability s.t.  $A\in\mathcal{S}$  then: 
$$i)P(\emptyset)=0$$
 
$$\mathcal{S}\cap\emptyset=\emptyset$$
 
$$ii)P(A)\leq 1$$
 
$$A^C=\{x\in\mathcal{S}:x\notin A\}$$
 
$$S=A\cap A^C$$
 
$$A\cup A^C=\emptyset$$
 
$$1=P(\mathcal{S})=P(A)+P(A^C)$$
 
$$iii)P(A^C)=1-P(A)$$

$$A_n \uparrow A = U_{n=1}^{\infty} A \to P(A_n) \uparrow P(A)$$

$$A_n \subset A_{n+1}$$

$$B_1 = A_1$$

$$B_k = A_k \setminus A_{k-1}$$

$$U_{n=1}^{\infty} A_i = \{ x \in \mathcal{S} : \exists i : x \in A_i \}$$

$$P: \mathcal{S} \rightarrow [0,1]$$
 
$$A, B \in \mathcal{S}$$
 then: 
$$i)$$
 
$$P(A \cap A^C) = P(B) - P(A \cap B)$$
 
$$B = (B \cap A) \cup (B \cap A^C)$$
 B and A complement have no commen elements 
$$P(B) = P(B \cap A) + P(B \cap A^C)$$
 
$$ii)$$
 
$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$
 
$$A = \{1, 2, 3\} \quad B = \{3, 4, 5\}$$
 
$$A \cup B = A \cup (B \cap A^C)$$
 
$$P(A \cup B) = P(A) + P(B \cap A^C) = P(A) + P(B) - P(A \cap B)$$
 
$$iii)$$
 
$$A \subset B \rightarrow P(A) \leq P(B)$$
 
$$P(B) = P(A) + P(B \cap A^C)$$
 
$$iv)$$
 
$$(A_n)$$
 
$$A \subset U_{n=1}^{\infty} A_n$$
 
$$A \subset U_{n=1}^{\infty} A_n$$
 
$$A'_n = A_n \cap A$$
 
$$B_n = A'_n \setminus U_{n=1}^{n-1} A'_m$$
 
$$U_{n=1}^{\infty} B_n = A \rightarrow P(A) = \sum_{n=1}^{\infty} P(B_n) \leq \sum_{n=1}^{\infty} P(A_n)$$

$$X: S \to \mathbb{R}$$
 (R is the sample space)

$$\mathcal{X} \subset \mathbb{R}$$

 $\mathcal{B}$  (collection of subsets)

 $B \in \mathcal{L}$ 

$$P_x(B) = P(\{s \in S : X(s) \in B\}) \to P(X^{-1}(B)) \to P(X \in B) = P(\{x \in B\})$$

$$S = \{(1,1), (1,2), \dots, (6,5), (6,6)\}$$
 (throwing two dice)

S = all subsets of S

$$P(\{(i,j)\}) = \frac{1}{36}$$

$$A = \{(1,1), \dots, (6,6)\} \to A_1 = \{1,1\}, \dots, A_6 = \{6,6\}$$

$$U_{i=1}^6 A_i = A$$

$$P(A) = P(U_{i=1}^{6} A_i) = \sum_{i=1}^{6} P(A_i) = \sum_{i=1}^{6} \frac{i}{36} = \frac{6}{36} = \frac{1}{6}$$

$$X:S\to\mathbb{R}$$

$$X((i,j)) = i + j$$

 $\mathcal{X} = \{\text{all the possible sums you can run through}\} \rightarrow \mathcal{X} = \{(2, 3, 4, \dots, 12)\}$ 

 $\mathcal{B} = \text{subset of } \mathcal{X}$ 

$$P_x(\{3\}) = P(X^{-1}(\{3\})) = P((i,j) \in \mathcal{S} : X((i,j)) \in \{3\})$$

$$\mathcal{X} \supset A = \{2, 4, 6, 8, 10, 12\}$$

$$P_x(A) =$$

$$P(\text{both even}) = P(\text{first even}) \cdot P(\text{second even}) = \frac{3}{6} \cdot \frac{3}{6} = \frac{9}{36}$$

$$P(\text{both odd}) = P(\text{first odd}) \cdot P(\text{second odd}) = \frac{3}{6} \cdot \frac{3}{6} = \frac{9}{36}$$

$$P_x(A) = \frac{9}{36} + \frac{9}{36} = \frac{18}{36} = \frac{1}{2}$$

#### 3 CDF

Cumulative Distribution Function

$$F: \mathbb{R} \to [0,1] \quad \text{cdf of } X_1$$

$$F(x) = P(X \le x) = P(s \in S: X(s) \le x) = P_x ((-\infty, x])$$

$$(P_x(A): A \in B)$$

**Properties of CDF:** if F is the cdf of a random variable X, then:

i) 
$$\lim_{x\to -\infty} F(x)=0$$
 and  $\lim_{x\to \infty} F(x)=1$  
$$ii) F(x) \leq F(y) \text{ whenever } y\geq x$$
 
$$iii) \lim_{x\downarrow z} F(x)=F(z)$$

$$x_n \to -\infty$$

$$F(x_n) = P_x \big( (-\infty, x_n] \big) \big) \downarrow P_x(\emptyset) \quad \text{(converges downwards)}$$

$$(-\infty, x_n] \uparrow (-\infty, \infty) = \mathbb{R}$$

$$ii)(-\infty, x] \subset (-\infty, y] \quad \text{if } y \ge x$$

$$F(x) = P_x \big( (-\infty, x] \big) \le P_x \big( (-\infty, y] \big) = F(y)$$

$$z_n \downarrow x \to (-\infty, z_n] \downarrow (-\infty, x] \to F(z_n) \to F(x)$$

$$z_n \uparrow x \to (-\infty, z_n] \uparrow (-\infty, x)$$

#### Identically distributed:

let x and y be a random variable

 $\forall B \in \mathcal{B}$ 

 $P(x \in B) = P(y \in B)$ 

 $P_x(B) = P_y(B)$ 

This is like flipping two coins, coin one is x and coin two is y (they are not the same coin, hence the same probability) random variables can also be Identically distributed:

if 
$$F_x = F_y \quad \forall x \in \mathbb{R}$$

$$F_x(x) = P(x \in (-\infty, x]) = P_x((-\infty, x])$$

$$= P_y((-\infty, x])$$

$$= P(y \in (-\infty, x])$$

$$= F_y(x)$$

#### 3.1 Continuous and discrete distributions:

#### x is discrete:

x has a pmf:  $f(x) = P(X = x), \quad x \in \mathbb{R}$ 

$$P_x(B) = P(x \in B) = \sum_{t \in B} f(t)$$

P(even) = f(2) + f(4) + f(6) (like the example about the dice)

$$F(x) = P_x \big( (-\infty, x] \big)$$
 untable points that are more than zero)

 $=\sum_{t\leq x}f(t)$  (countable points that are more than zero)

### 4 Three common distributions:

#### 1. Bernoulli:

$$X \sim \text{Bernoulli}(p), \quad p \in [0, 1]$$

$$X = \begin{cases} 1 & \text{if prob. } p \\ 0 & \text{if prob. } 1 - p \end{cases}$$

$$f(x) \begin{cases} p & x = 1 \\ p & x01 \\ p & x \notin \{0, 1\} \end{cases}$$

Ex: we have two different websites and we want to measure website performance of how many people bought stuff on each of the websites  $\rightarrow$  **proportions** 

#### 2. Binomial:

$$X \sim \text{Binom}(n, p)$$
$$f(x) = \binom{n}{x} p^x (1-p)^{n-x}, \quad x = 0, 1, \dots, n$$

#### 3. Poisson:

$$f(x) = \frac{\exp(-\lambda)\lambda^x}{x!}, \quad x = 0, 1, 2, \dots$$

Example: ordered packages, how likely is it that 6 of my packages are going to arrive before lunch? lambda is package arrival intensity

#### 4.1 Continuous random variables

X is continuous

$$f.$$

$$f(x) \neq P(X = x)$$

 $P(X = x) = 0 \quad \forall x \in \mathbb{R}$  (The probability of hitting any specific exact value is essentially zero)

$$P(X \le x) = F(x) = \int_{-\infty}^{x} f(t)dt$$
 
$$P(x \in B) = \int_{-B}^{B} f(t)dt = \int_{-\infty}^{\infty} \mathbb{1}_{B}(t)f(t)dt$$

#### Uniform distribution

 $X \sim \text{uniform}(a, b)$ 

$$f(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

#### Normal distribution:

 $X \sim \text{Normal}(\mu, \sigma^2)$ 

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{-(x-\mu^2)}{2\sigma^2}\right)$$

#### Student's distribution

 $X \sim t(v)$ 

$$f(x) = \frac{\Gamma((v+1)/2)}{\Gamma(v/2)\sqrt{v\pi}} \left(1 + \frac{x^2}{v}\right)^{-(v\pi)/2}$$

#### Chi-squared - related to the N dist

$$X \sim \chi^2(p)$$

Ugly formula – google it

f is a pdf/pmf of r.v. x

$$i)f(x) \ge 0$$

$$ii) \int_{-\infty}^{\infty} f(x)dx = 1; \quad \sum_{x \in \mathcal{X}} f(x) = 0$$

i) continuous

as 
$$F(x) \uparrow F(x) = \int_{-\infty}^{x} f(t)dt$$

ii) discrete

$$\sum_{x \in \mathcal{X}} f(x) = \lim_{t \to \infty} \sum_{x \le t, x \in \mathcal{X}} f(x)$$
$$= \lim_{t \to \infty} F(t)$$
$$= 1$$

$$\int_{-\infty}^{\infty} f(x) dx = \lim_{t \to \infty} \int_{-\infty}^{t} f(x) dx = \lim_{t \to \infty} F(t) = 1$$

# 5 Creating r.vs from an r.v.

$$X:S \to \mathbb{R}$$

$$g:\mathbb{R}\to\mathbb{R}$$

$$Y = g(X): S \to \mathbb{R}$$

$$P(y \in A) = P(g(x) \in A) = P(\{x \in \mathbb{R} : g(x) \in A\}) = P(x \in g^{-1}(A))$$

If this transformation is well '-behaved, we can compute the pdf of g(x)

i) If x is discrete, then Y = g(x) is also discrete:  $f_y(y) = \sum_{x \in g^{-1}(\{y\})} f_x(x)$ 

ii) If x is cont., then the output will be discrete:

$$F_Y(y) = \int_{\{x: g(x) \leq y\}} f_X(x) dx = P(g(x) \leq y) = P\Big(x \in g^{-1}\big(\{y\}\big)\Big) = P(Y \leq y)$$

iii) if g is strictly monotone increasing, then  $h=g^{-1}$  is differentiable, THEN

$$F_Y(y) = F_X(h(y)), \quad f_Y(y) = f_X(h(x))h'(y)$$

$$F_X(y) = \int_{-\infty}^{g^{-1}(y)} f_X(x) dx = F_X(h(y))$$

# 6 Mean variance of an arbitrary distribution

$$\mathbb{E}g(x) = \int_{-\infty}^{\infty} g(x)f(x)dx$$
$$= \sum_{x \in \mathcal{X}} g(x)f(x) \quad \text{(in the discrete case)}$$

$$\mathbb{E}X = \int_{-\infty}^{\infty} x(fx)dx$$
 this technically does not exist, hope this helps

we can always write

$$g(x) = g^{+}(x) - g^{-}(x)$$
 (both funcs are positive)

expectation exists if:

$$\mathbb{E}|g(x)| < \infty$$

in the case that

$$\mathbb{E}|g(x)| \ge \infty$$

we can say that it may not exist

$$i)\mathbb{E}[\alpha g_1(x)] = \alpha \mathbb{E}[g(x)]$$
  
$$\mathbb{E}[g_1(x) + g_2(x)] = \mathbb{E}[g_1(x)] + \mathbb{E}[g_2(x)]$$

$$ii)g_1(x) \ge 0 \quad \forall f(x) > 0 \to \mathbb{E}g_1(x) \ge 0 \quad \text{(proof below)}$$

 $\pm$  inf are implied

$$\mathbb{E}\alpha g_1(x) = \int \alpha g_1(x)f(x)dx = \alpha \int g(x)f(x)dx = \alpha \mathbb{E}g_1(x)$$

$$\mathbb{E}(g_1(x) + g_2(x)) = \int (g_1(x) + g_2(x))f(x)dx = \int g_1(x)f(x) + g_2(x)f(x)dx$$

$$= \int g_1(x)f(x)dx + \int g_2(x)f(x)dx$$

remember to look at ii) from the notes