

Experimental quantum characterization

- Bayesian inference and numerical integration

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Presentation outline

1. Context

- Problem statement
- Bayesian inference
- A simple example

2. Open problems

- Experimental design (adaptivity)
- Integration (Monte Carlo)

3. Characterization of a quantum device

What to characterize?

We consider the characterization of a **time-dependent function** describing a system's dynamics (e.g. wavefunction coefficient norms, or elements of a density operator)

$$\varphi(t) = a_0(t)|0\rangle + a_1(t)|1\rangle$$

$$P(0|t) = |\langle 0|\varphi(t)\rangle|^2 = |a_0(t)|^2 = ?$$

Characterization is useful for certification, tuning sensing, and predictions.

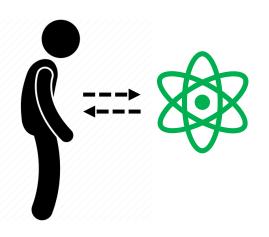
(the generalization to state or process tomography, phase estimation, etc. is straightforward)

$$P(0|t) = ?$$

We derive empirical knowledge via observation, i.e. by **measuring the system**.

The outcome of the measurement(s) should depend on the entity to be characterized. In our case, this dependence is probabilistic.

This adds a layer of complexity to metrology – one which can be dealt with using **statistical inference**



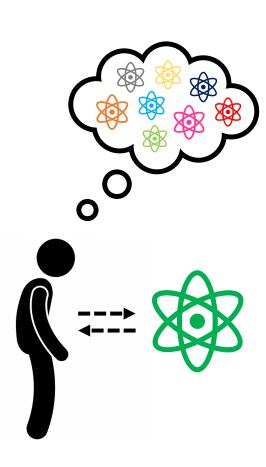
$$P(0|t) = ?$$

Bayesian statistics provides a flexible paradigm for statistical inference.

This strategy ultimately relies on **comparing the**behaviour of the black-box system with that of

potential replicas.

Having chosen these speculative replicas, the onus is on the **experimental data collection** process, which settles the predictive power



$$P(0|t) = ?$$

We can exploit some structured knowledge:

- Parametrized model for the probability (can be exploratory)
- Pre-existing information about the parameters (can be insubstantial)

- Ability to measure the system
- Ability to simulate the system (given the model)
- Classical processing

$$P(0|t) = ?$$

We can exploit some structured knowledge:

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- Ability to **simulate** the system **for some** θ , t
- Classical processing elementary operations, eventually optimization

$$P(0|t) = ? \longrightarrow \theta = ?$$

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- Classical processing elementary operations, eventually optimization

$$\theta = ?$$

Bayesian **probabilities** represent **degrees of belief**. Thus, in Bayesian inference we attribute a *probability* to parameter values, which are seen as random variables:

$$P(\theta)$$

This randomness can be ascribed to our own imperfect knowledge: the object of Bayesian statistics is our reality, rather than an idealized version of it.

To **assess the merit of a parameter** instance, we will assign it a degree of belief **according to our observations**

$$\theta = ?$$

• We start with some **prior probability** distribution over the parameter θ .

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- We then collect an experimental datum D, and re-weight the prior probability depending on the likelihood $P(D|\theta)$ that it would have generated D.

$$P(D|\theta)P(\theta)$$

$$\theta = ?$$

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- Apart from a normalizing constant, we obtain the posterior probability, representing our updated knowledge (after observing D).

$$P(\theta|D) \propto P(D|\theta)P(\theta)$$

$$\theta = ?$$

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Each datum should bring our knowledge slightly closer to reality

$$P(\theta|D) \propto P(D|\theta)P(\theta)$$

$$\theta = ?$$

$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)}$$

Bayes' rule

- $P(\theta|D)$: the **posterior** probability (our updated knowledge)
- $P(D|\theta)$: the **likelihood** (a generative model for the data)
- $P(\theta)$: the **prior** probability (can be flat over Θ)
- P(D): the marginal probability (for a fixed model, this is a normalizing constant)

The Bayes' rule assesses the **relative merit** of any given **parameter instance**, taken as **the probability of its having generated the observed system behaviour** (up to previous knowledge)

$$\theta = ?$$

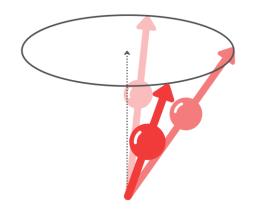
$$P(\theta|D) = \frac{P(D|\theta)P(\theta)}{P(D)}$$

Bayes' rule

Bayesian inference refers to the systematic use of Bayes' rule to **learn** system properties using empirical evidence. In the quantum case:

- The data $D \in \{0,1\}$ are binary experimental outcomes
- The likelihood $P(D|\theta)$ can be calculated using analog **quantum** simulation
- Experimental controls may offer additional degrees of freedom

A simple example: spin precession



Learn the model Hamiltonian:

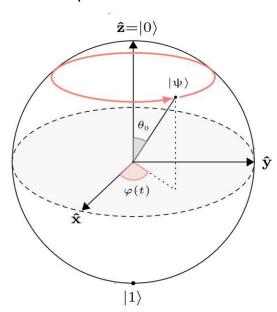
$$\widehat{H} = \frac{\boldsymbol{\omega}}{2} \, \widehat{\sigma}_{z} \qquad \boldsymbol{\omega} = ?$$

This type of oscillation describes several phenomena of interest - namely **Larmor** precession, **Rabi** flopping, **Ramsey** interferometry, and **phase estimation**.

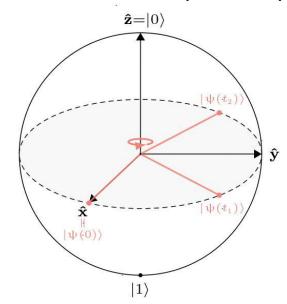
In this context, frequency estimation (learning ω) has applications as important as **magnetic field sensing** and resonant **gate tuning** in superconducting qubits

A simple example: spin precession

In general, this Hamiltonian induces a **precession** about the Bloch sphere's \hat{z} axis.



For simplicity, we consider $|+\rangle$ initialization at t=0, and thus oscillations contained in the **equatorial plane**.



To collect the data, we **measure the qubit on the z basis** for some chosen **evolution times** t_i (counting from initialization).

This evolution time is an experimental control.

A simple example: spin precession

In this case, as in many others, the **likelihood** is physically motivated.

$$P(0|\omega;t) = \cos^2(\omega t/2)$$

$$P(1|\omega;t) = 1 - P(0|\omega;t)$$

The **prior** can be chosen to barely influence the results. Generally, we establish a **frequency range**, over which we can lay out a flat prior:

$$P(\omega) = \begin{cases} 1/V_{\Omega} & \text{if } \omega \in \Omega \\ 0 & \text{otherwise} \end{cases}$$

All that's left is the marginal, which can be regarded as a normalization constant. Thus, we can evaluate the posterior:

$$P(\omega|D;t) = \frac{P(D|\omega;t)P(\omega)}{P(D;t)}$$

$$P(\omega|D;t) = \frac{P(D|\omega;t)P(\omega)}{P(D;t)}$$

The posterior is a **continuous probability distribution** representing our posterior (updated) knowledge on the model parameter(s). From a Bayesian viewpoint, **it holds** *all* **the** available information.

Conceptually, this solves the problem; but in practice, it creates another one. How do we **extract the information** we want?

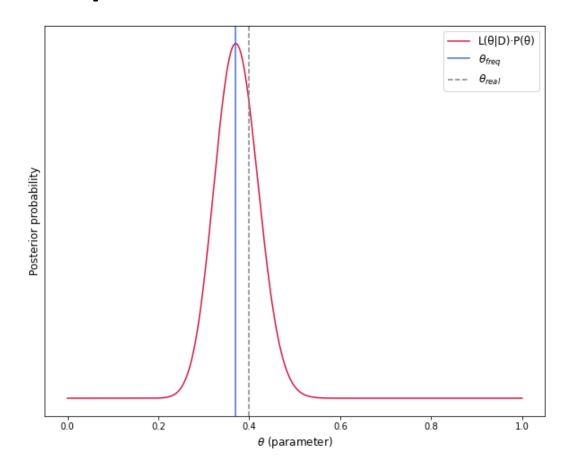
Taking the single parameter case, the standard **frequentist** approach would take **multiple trials** with **unchanged experimental controls** and **invert the likelihood** to produce a **parameter estimate**.

$$P(0|\omega;t) \approx \frac{\# \text{ 0 outcomes}}{\# \text{ trials}} = \cos^2(\omega t/2) \leftrightarrow \omega = \omega_X$$

By contrast, the Bayesian approach offers *only* a way to compute the probability of parameter estimates of our choice.

$$P(\omega|D;t) = \frac{P(D|\omega;t)P(\omega)}{P(D;t)} \leftrightarrow P(\omega|D;t) = p_X$$

We must then use these **pointwise evaluations** to produce our estimates.

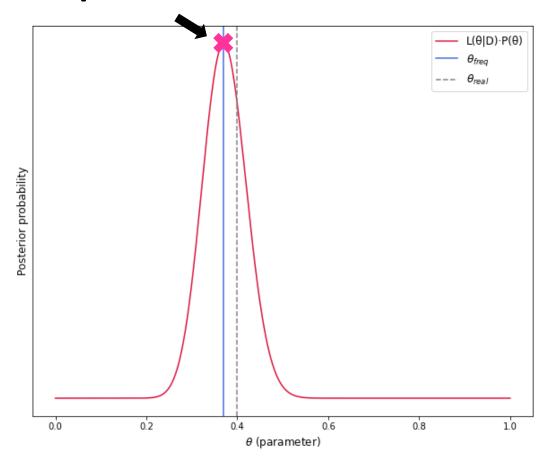


frequentist point estimate:

 $\omega = \omega_X$

Bayesian posterior:

$$P(\omega|D;t) = p_X$$



frequentist point estimate:

 $\omega = \omega_X$

Bayesian posterior:

$$P(\omega|D;t) = p_X$$

Querying the posterior

There are two known treatment options for the posterior:

Optimization

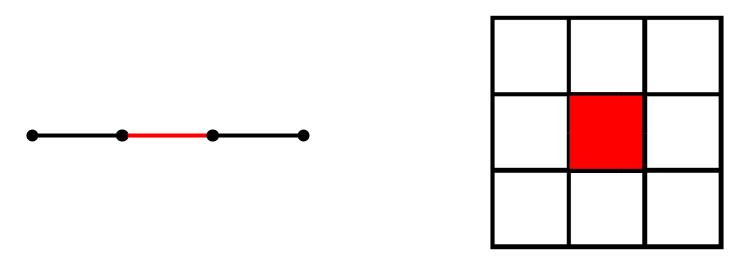
$$\widehat{\theta} = \underset{\theta}{\operatorname{argmax}} P(\theta|D)$$

Integration

$$\mathbb{E}[f(\theta)] = \int f(\theta) P(\theta|D) d\theta$$

The behaviour of a probability distribution is dictated by **probability mass**, $P(\theta|D) d\theta$, whereas optimization considers the **density** $P(\theta|D)$ without heeding the **volume** $d\theta$. This blatant disregard is based on a naive, unscalable intuition.

Querying the posterior



As dimensionality increases, the fraction of volume contained within any given region shrinks exponentially. Likewise, the volume occupied by the highest density points becomes negligible.

As a consequence, **integration is the only sensible option** for anything other than simple (low-dimensional, convex) models.

Querying the posterior

$$\mathbb{E}[f(\theta)] = \int f(\theta)P(\theta|D) d\theta$$

We can then **retrieve any estimate** we want by **integrating over the posterior**. For instance, we can take the **mean** as our **parameter** estimator:

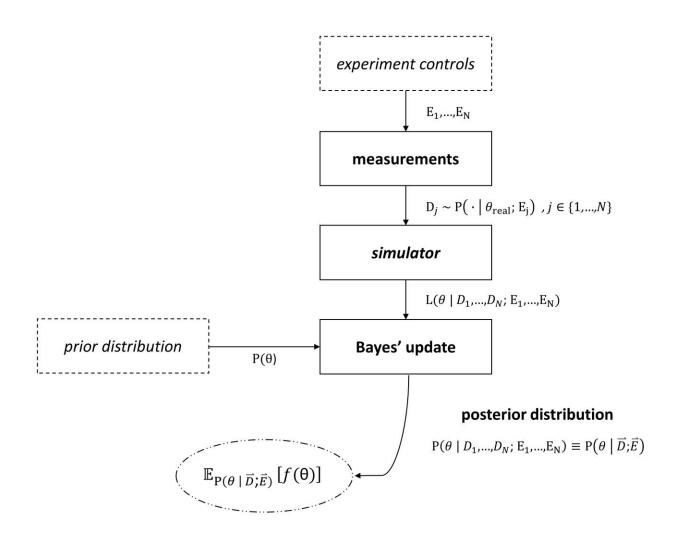
$$\widehat{\theta} = \overline{\theta} \equiv \mathbb{E}[\theta]$$

And the **standard deviation** as our **uncertainty** estimator:

$$\sigma^2 = \mathbb{E}[(\theta - \bar{\theta})^2]$$

We can compute many other expectations, be they **statistical assessments** (e.g. interquartile ranges, credible regions) or **empirical predictions** (e.g. probability of finding the qubit at state X when measuring on some basis at a time Y).

Summary: the algorithm



Why Bayesian inference?

This strategy may seem more cumbersome than a simple frequentist guess, but **the framework is also much richer**.

- It is very **general**: it can be applied whenever one can measure a system and simulate parametrized descriptions of it.
- It naturally **estimates the uncertainty** and any other well-defined quantities.
- It is remarkably **robust** and **scalable**.
- It is capable of online estimation, and of learning from small datasets.

Open problems

querying the system

How to efficiently design informative experiments ?

data → posterior distribution

querying the posterior

How to efficiently capture the information ?

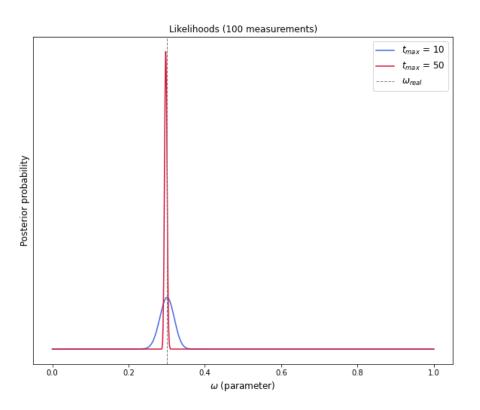
posterior distribution \rightarrow predictions

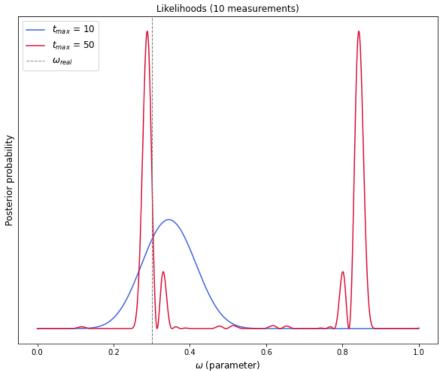
Finding an efficient optimization strategy for the experimental controls

Finding an efficient integration strategy for the posterior

The experiments ultimately determine the **sharpness** of the posterior

...and even its **ambiguity**, or lack thereof





The Bayesian posterior allows for calculating any statistical expectation. This includes the **expected utility** of any given **experiment**(s).

We can then use our knowledge to choose the best possible set of experiments.

However, this has two problems:

- We incur an exponential cost in the number of experiments.
- The knowledge we start out with is limited, and may not be very informative.

These two issues can be solved at once, by applying Bayes' rule sequentially:

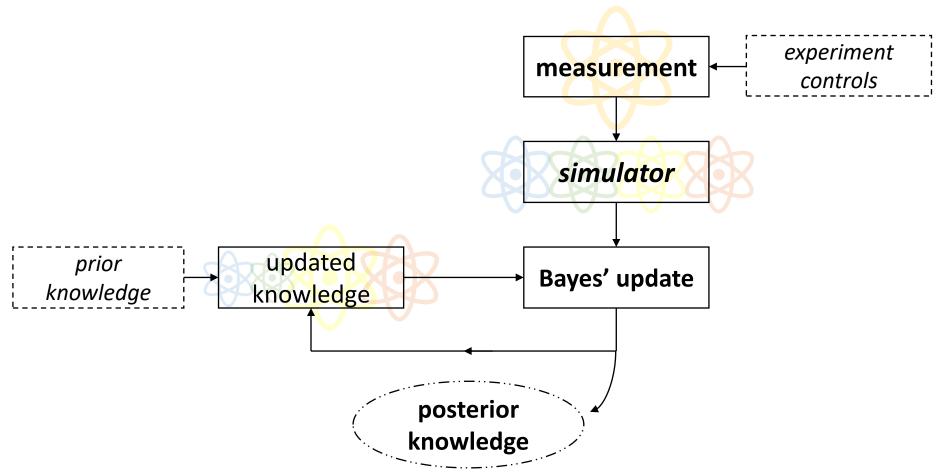
$$P(\theta|D_1) \propto P(D_1|\theta)P(\theta)$$

$$P(\theta|D) \propto P(D|\theta)P(\theta)$$

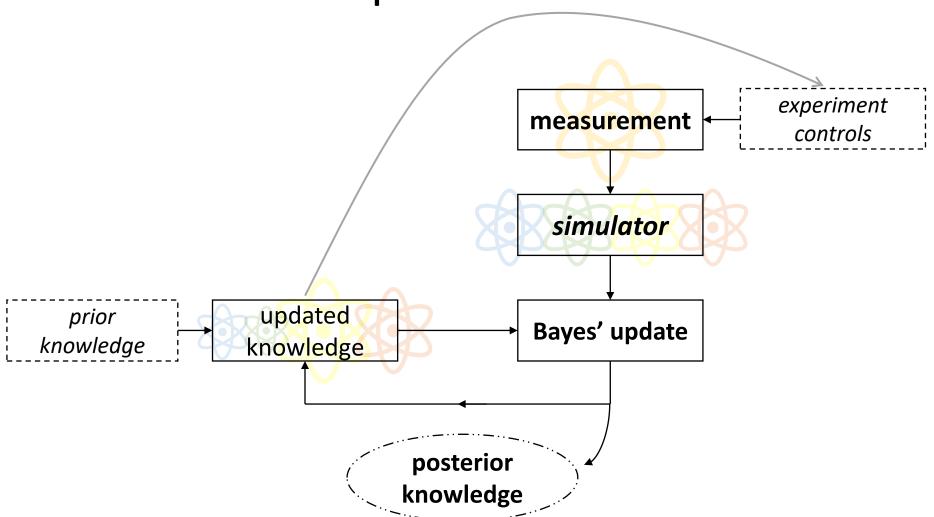
$$P(\theta|D_1) \propto P(D_1|\theta)P(\theta)$$

$$P(\theta|D_1) \propto P(D_2|\theta)P(\theta|D_1)$$

35



How to efficiently design informative experiments?



How to efficiently design informative experiments?

Adaptivity bring two major advantages:

- It brings the cost down to linear in the number of experiments.
- The experiments can be chosen according to the full extent of our growing knowledge

This greatly simplifies the problem, after which we can employ:

- (Local) optimization
- Heuristics (can be analytically/empirically motivated, or use machine learning)
- An in-between strategy

$$\mathbb{E}[f(\theta)] = \int f(\theta)P(\theta|D) d\theta$$

We have seen that sound statistical inquiries directed at a probability distribution correspond to integrals.

In general, the main difficulty of scaling up Bayesian inference is integration.

$$\mathbb{E}[f(\theta)] = \int f(\theta)P(\theta|D) d\theta$$

There are two major options for integration over probability distributions:

- Variational inference
- Numerical integration

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$$\mathbb{E}[f(\theta)] = \int f(\theta)P(\theta|D) d\theta$$

$$\approx \sum_{i} f(\theta_i) \cdot w_i$$

, with
$$w_i \propto \frac{P(\theta_i|D)}{\pi(\theta_i)}$$
 , $\theta_i \sim \pi(\cdot)$

Numerical integration

- Monte Carlo

$$\mathbb{E}[f(\theta)] = \int f(\theta)P(\theta|D) d\theta$$

The problem of numerical integration is to choose the evaluation sites

$$\theta_i \sim \pi(\cdot)$$

Numerical integration

- Monte Carlo

Two Monte Carlo solutions stand out:

- Markov Chain Monte Carlo (MCMC)
 - Random walk Metropolis
 - Hamiltonian Monte Carlo
- Sequential Monte Carlo (SMC)
 - Sequential importance resampling
 - Tempered likelihood estimation

$$\{\theta_i\}_{i=1}^N = ?$$

It is simple to build a Markov Chain that samples from the posterior, using random walk Metropolis:

- 1. Choose a starting point $heta_{
 m curr}$
- 2. For *N* iterations:
 - 1. Propose θ_{new} at random.
 - 2. Replace θ_{curr} with θ_{new} with probability

$$\frac{P(\theta_{\text{new}}|D)}{P(\theta_{\text{curr}}|D)}$$

The sequence of θ is distributed according to $P(\cdot | D)$ as $N \to \infty$



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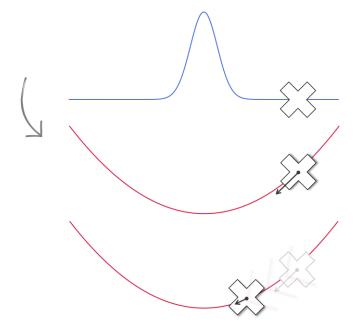
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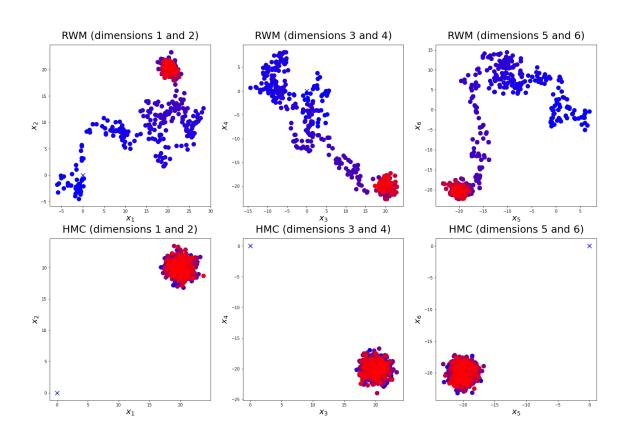
$$\{\theta_i\}_{i=1}^N = ?$$

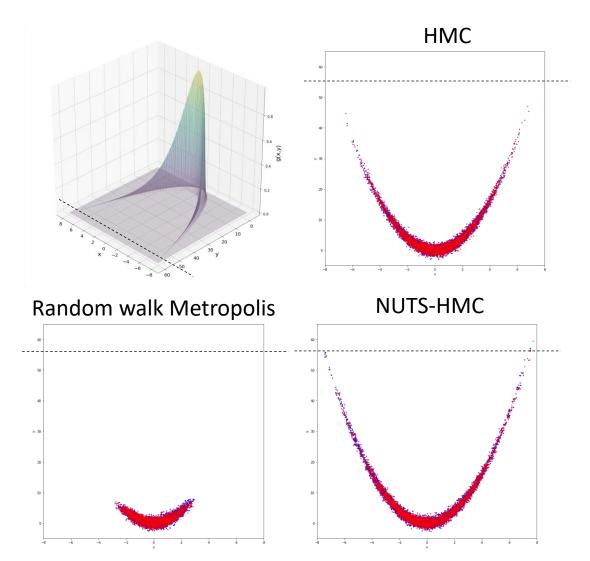
A more efficient approach must be informed by the differential geometry of the target distribution. This is the idea behind

Hamiltonian Monte Carlo



Random walk Metropolis (top) vs. Hamiltonian Monte Carlo (bottom) for a Gaussian:





Random walk Metropolis vs. Hamiltonian Monte Carlo for a Rosenbrock "smile" function

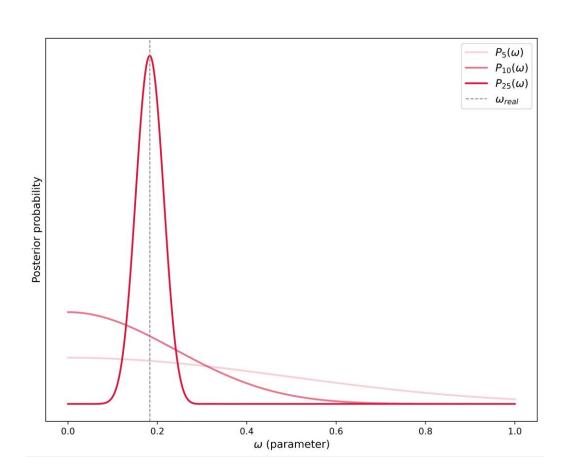
$$\{\theta_i\}_{i=1}^N = ?$$

HMC enables a highly efficient exploration for a wide class of functions. However, there are still two problems:

- MCMC is incompatible with sequential processing, and thus online sensing and adaptivity.
- Markov chains are unfit to capture multimodality.

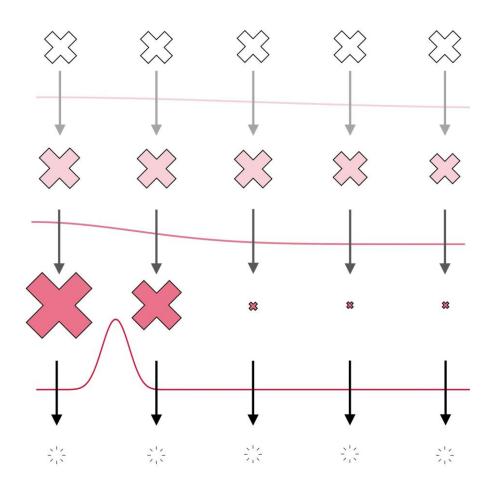
Often, these two problems can again be solved at once using adaptivity.

$$\{\theta_i\}_{i=1}^N = ?$$



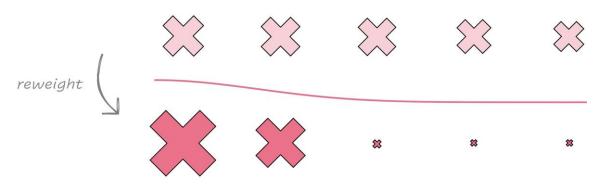
Just as the experimental design, integration may benefit from adaptivity

$$\{\boldsymbol{\theta_i}\}_{i=1}^N = ?$$



A grid-based strategy can be made sequential, but it leads to poor results and rigid exponential scaling.

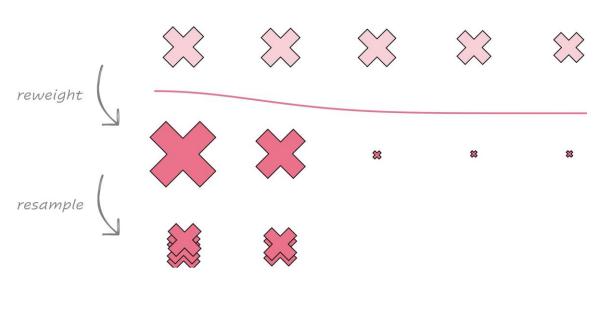
$$\{\boldsymbol{\theta_i}\}_{i=1}^N = ?$$



We can correct this by again exploiting the available knowledge, this time to refocus the computational resources.

Sequential Monte Carlo guides the construction of such a dynamic grid

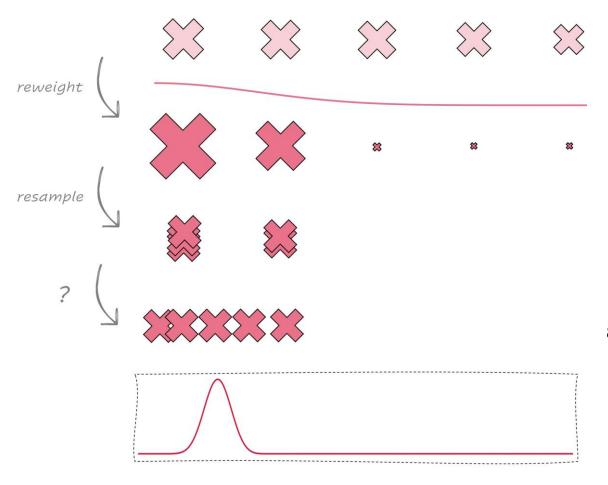
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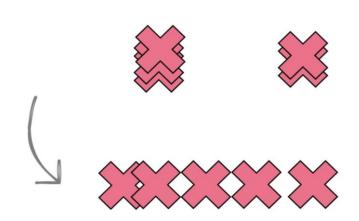
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$$\{\theta_i\}_{i=1}^N = ?$$

The move step is the most crucial.
One possibility is to induce random perturbations, but we incur information loss. To compensate, we can pull the points towards the center.

This is called kernel smoothing and shrinkage, and is done in e.g. the Liu-West particle filter, the most (only?) used option in SMC for quantum characterization

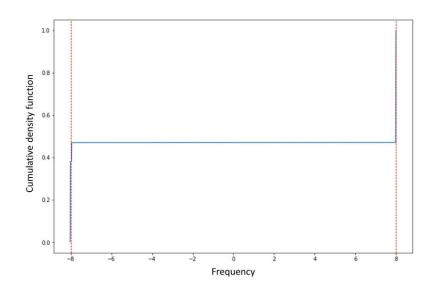


$$\{\boldsymbol{\theta_i}\}_{i=1}^N = ?$$

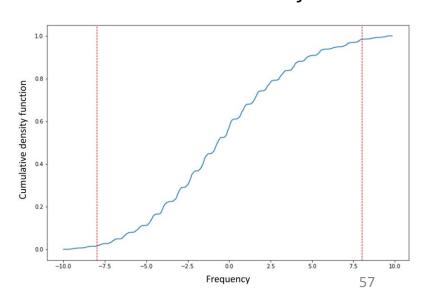
...but it only works perfectly for a Gaussian, and assumes unimodality. It can bias expectations, and get trapped in local minima.

A more reliable solution is to use MCMC within SMC:

with Hamiltonian Monte Carlo kernels



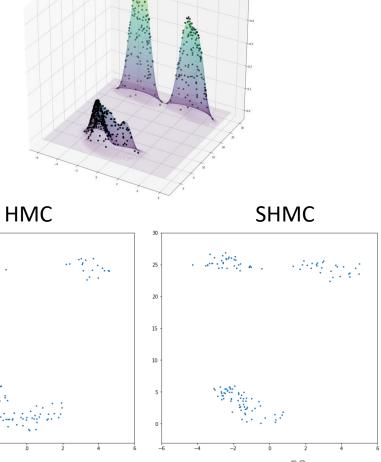
with the Liu-West filter



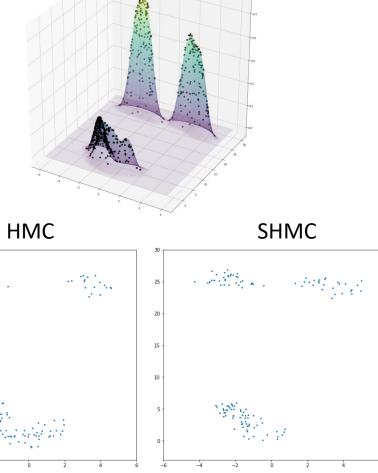
Random walk Metropolis vs. Hamiltonian Monte Carlo vs. their SMC counterparts for a"smiley" function

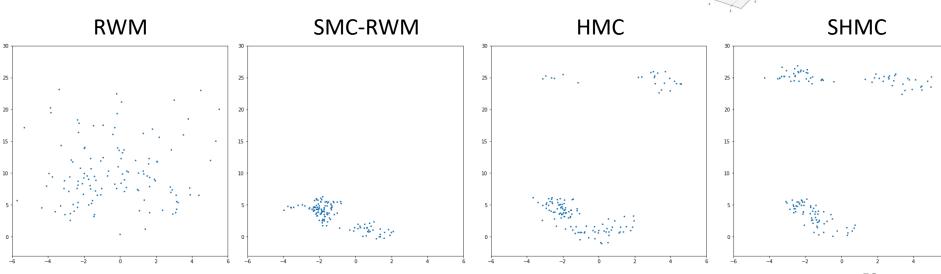
RWM

SMC-RWM



MCMC within SMC conciles online processing with robustness and correctness

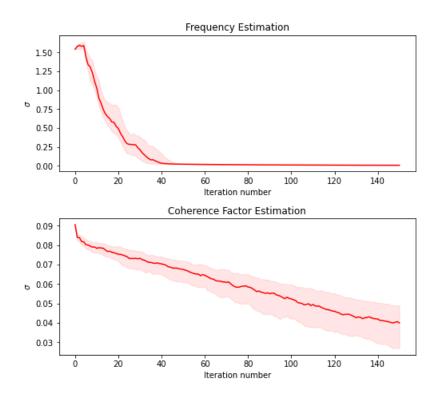




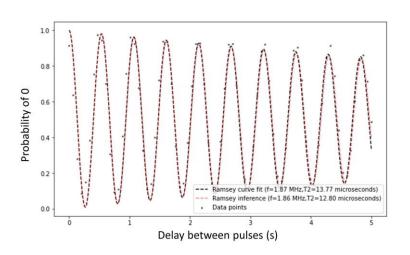
Ramsey experiment on IBMQ device $ibmq_armonk$: joint estimation of ω and T_2^*

$$P(t) = e^{-t/T_2^*} \cos^2(\omega t/2) + \frac{1}{2} (1 - e^{-t/T_2^*})$$

- Data collection with Qiskit Pulse
- Representation via Sequential Monte Carlo with Markov kernels
- Each iteration corresponds to one datum
- Median values over 100 runs
- Matched against the default curve fitters

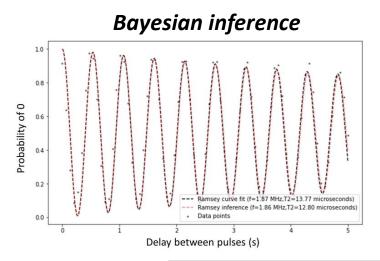


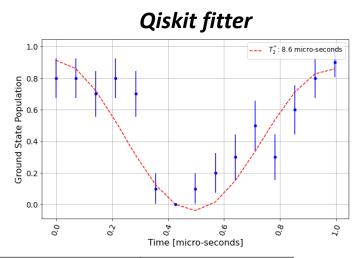
Evolution of the inference's **uncertainty** with the iterations/data



- Data points: 512 measurements per time
- Dashed curve: obtained by regression on data points
- Red curve: Bayesian inference using less than
 0.4% of the data

Allowing the fitter as many data as the inference and the same measurement scheme resulted in an **infinite uncertainty** in T_2^* . Finite values could be achieved for fits on shorter intervals, but the best observed value was still **10 times worse**, whereas the frequency estimation performed roughly **40 times worse** (0.4% vs. 15%)

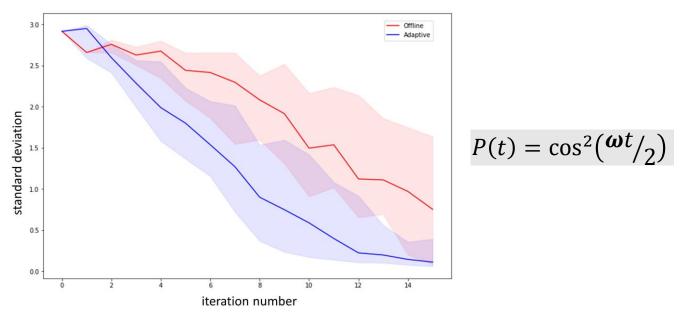




Note: direct comparisons for T_2^* are not viable due to fast fluctuations

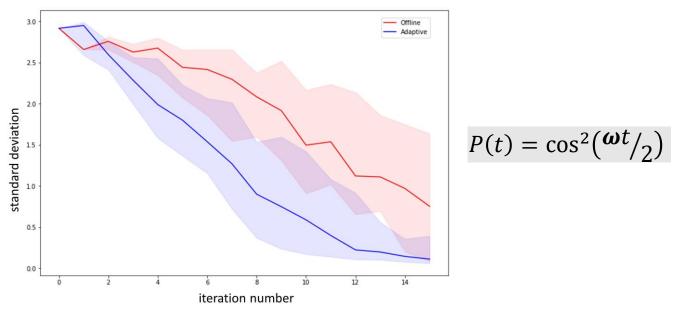
	T_2^* (μs)	Standard deviation	Total shot count
Bayesian inference	9	3	$75 \cdot 2 = 150$
Qiskit fitter	21	1	$75 \cdot 512 = 38400$
Qiskit fitter	9	30	$15 \cdot 10 = 150$

Adaptive *Hahn-Ramsey* experiment on IBMQ device **ibmq_armonk**



	δ (MHz)	Standard deviation	Total shot count	Precision ($\sigma^2 \Delta t_{ m acc}$)
Adaptive	1.8	0.1	15	0.38
Offline	2.3	0.8	15	9.3

Adaptive Hahn-Ramsey experiment on IBMQ device ibmq_armonk



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Main bibliographic references

- Sequential Monte Carlo: Doucet et al. An Introduction to Sequential Monte Carlo Methods (Springer NY, 2001).
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 - Gunawan et al. Subsampling sequential Monte Carlo for static Bayesian models (Springer Science and Business Media LCC, 2020).

Hamiltonian Monte Carlo:

Foundations:

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- Betancourt. A conceptual introduction to Hamiltonian Monte Carlo (Arxiv e-print, 2018).

Variants:

- Hoffman et al. *The No-U-Turn Sampler: Adaptively Setting Path Lengths in Hamiltonian Monte Carlo* (Journal of Machine Learning Research, 2011).
- Betancourt. Adiabatic Monte Carlo (Arxiv e-print, 2015).
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Subsampling:

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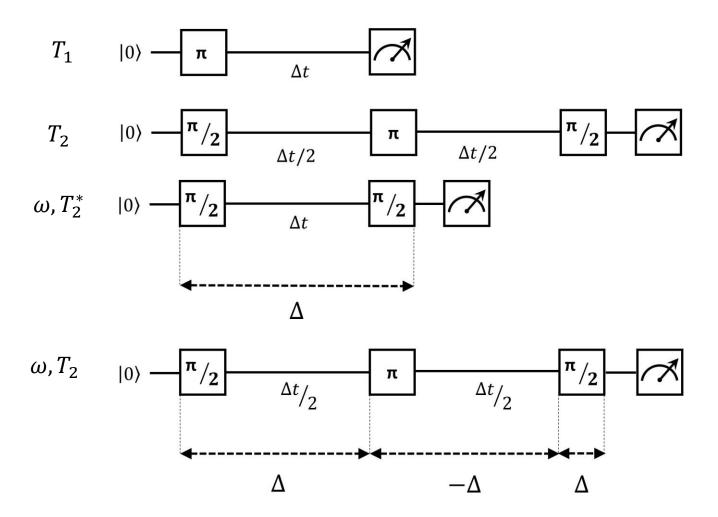
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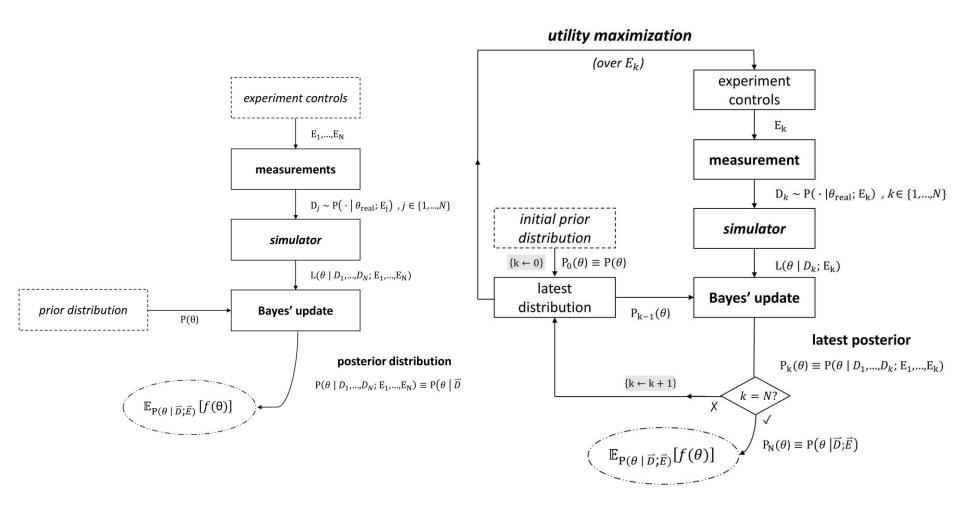
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Other graphs

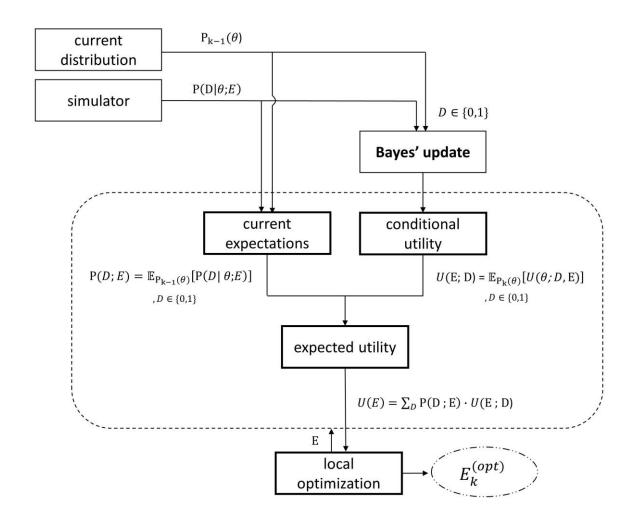
Pulse schedules



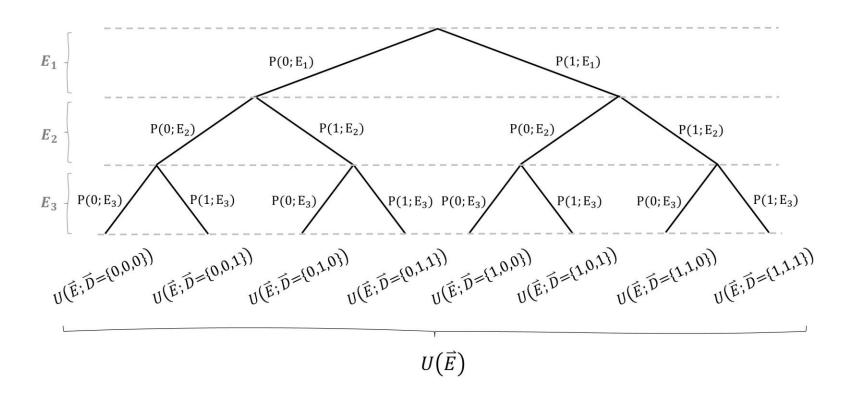
Offline vs. online inference



Offline vs. online inference

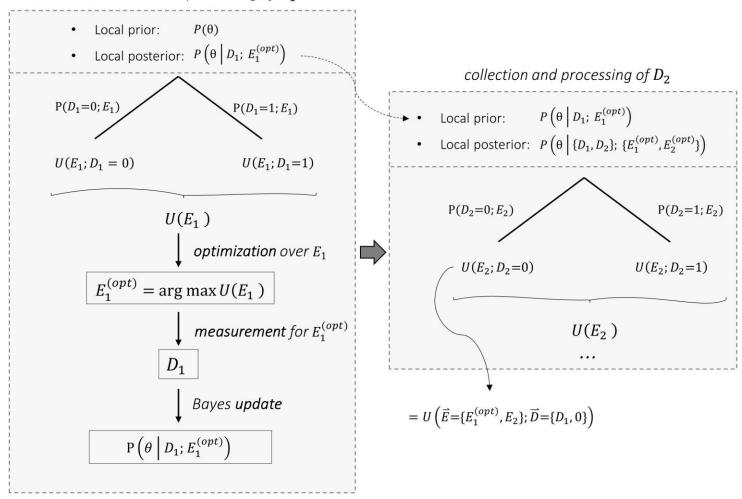


Offline vs. online inference

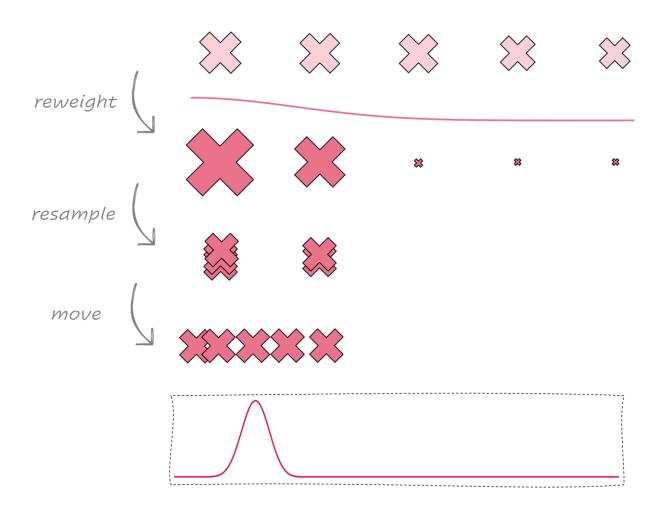


Offline vs. online inference

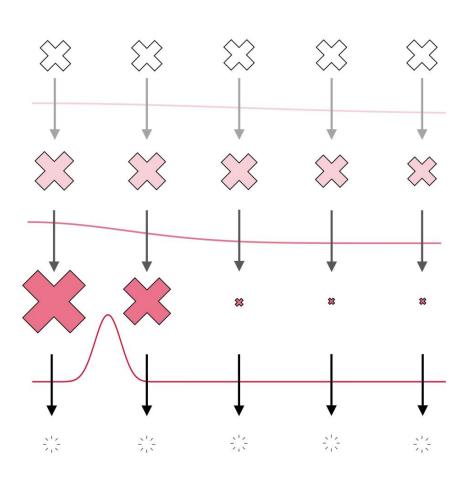
collection and processing of D_1

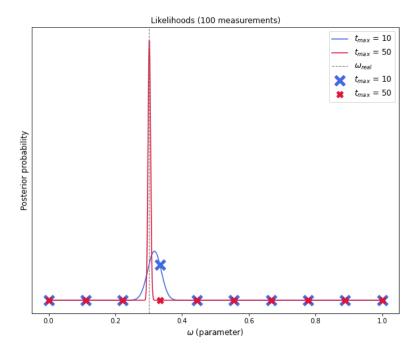


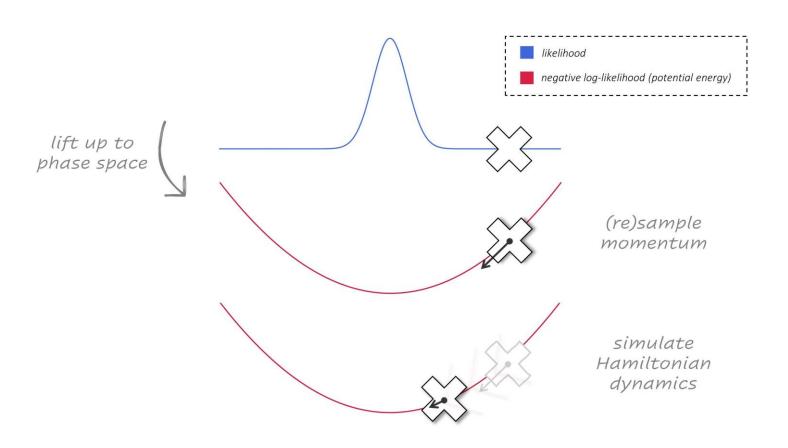
SMC



Particle degeneracy in SMC

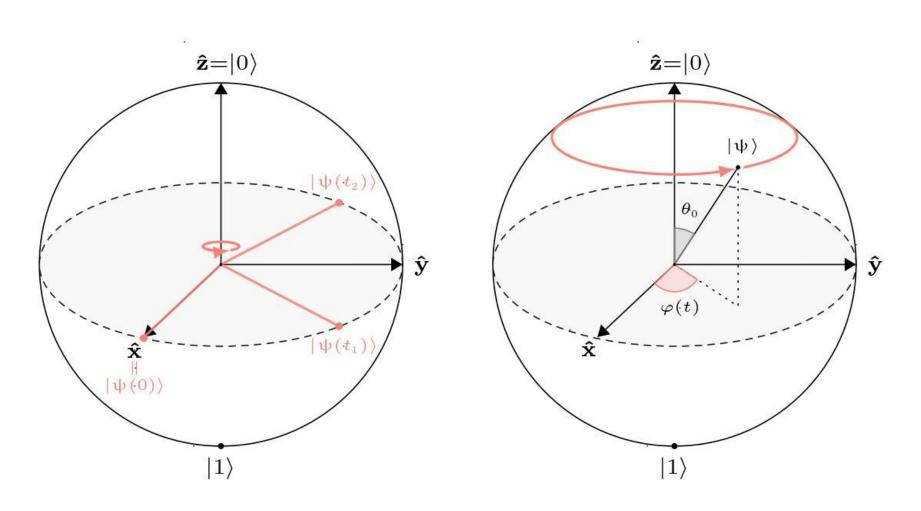




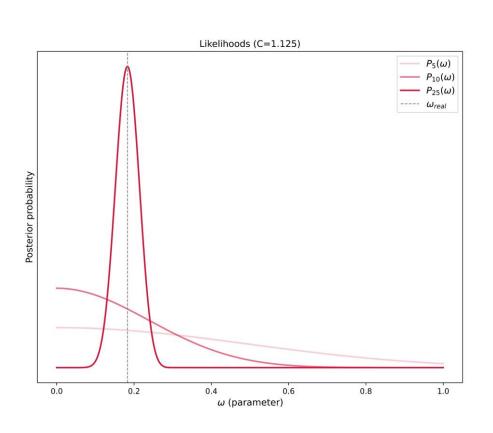


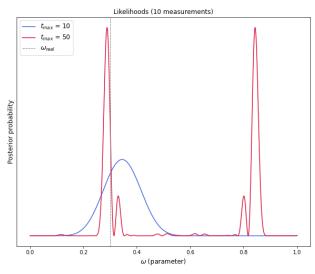
Precession

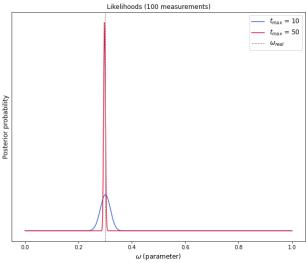
Bloch sphere representation



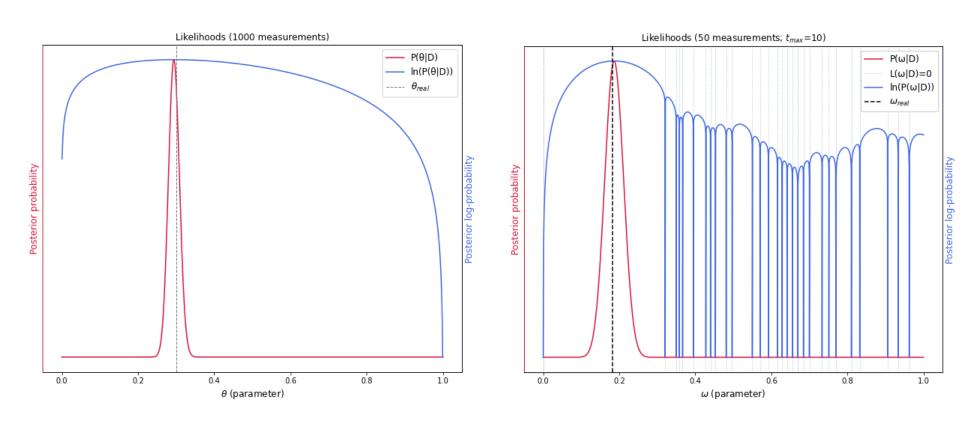
Impact of the evolution times



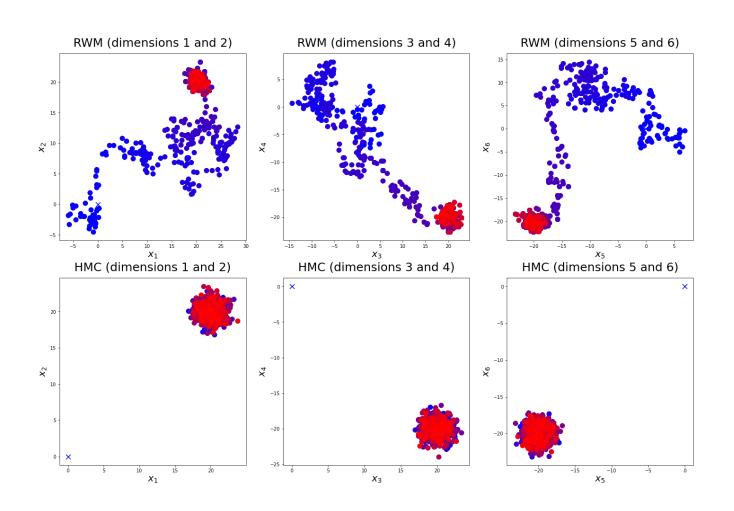


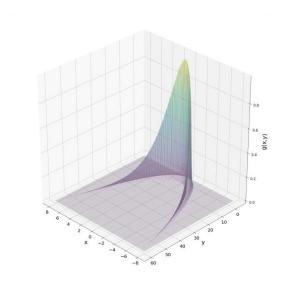


Log-density

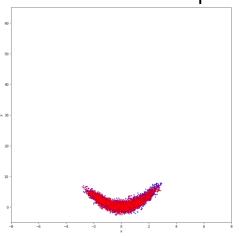


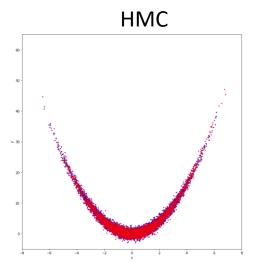
HMC/SHMC/NUTS vs. RWM/GRF/LW

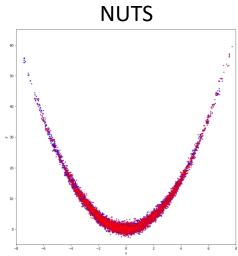


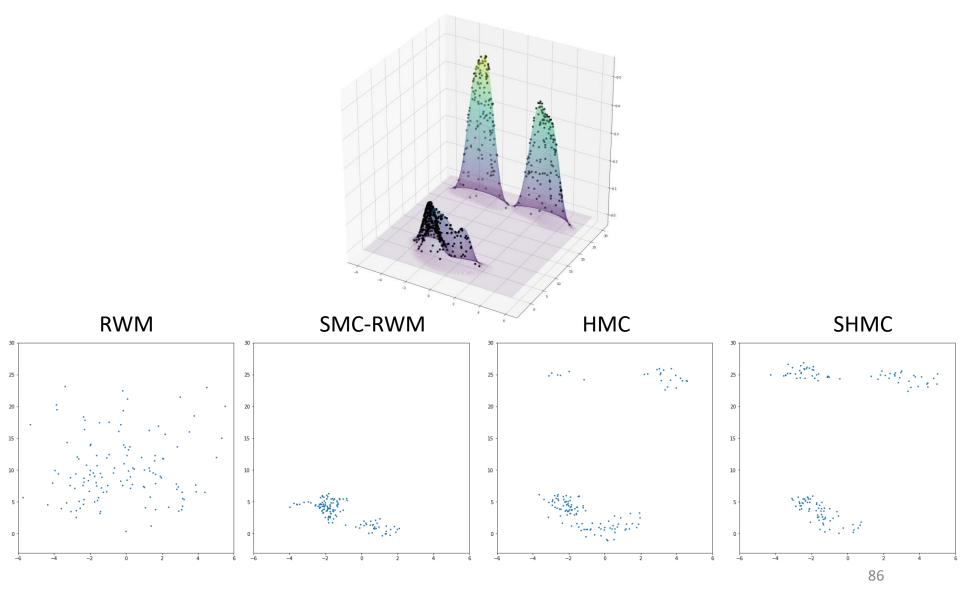


Random walk Metropolis

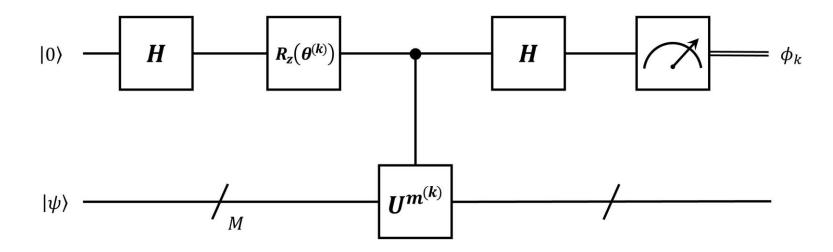




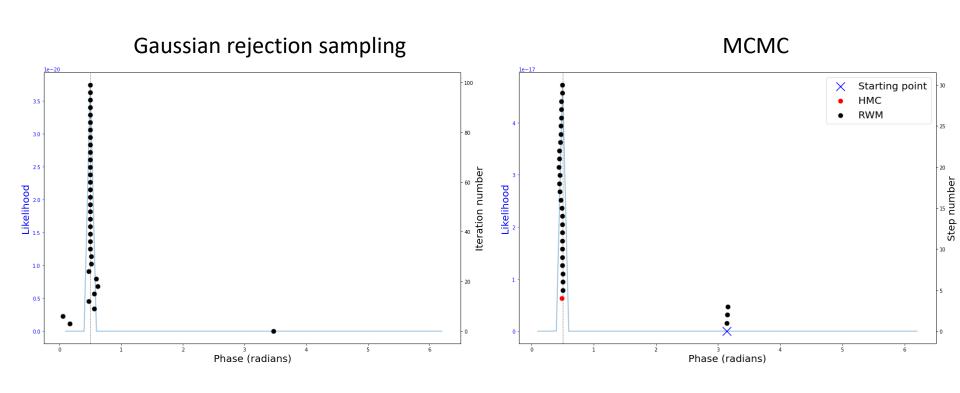




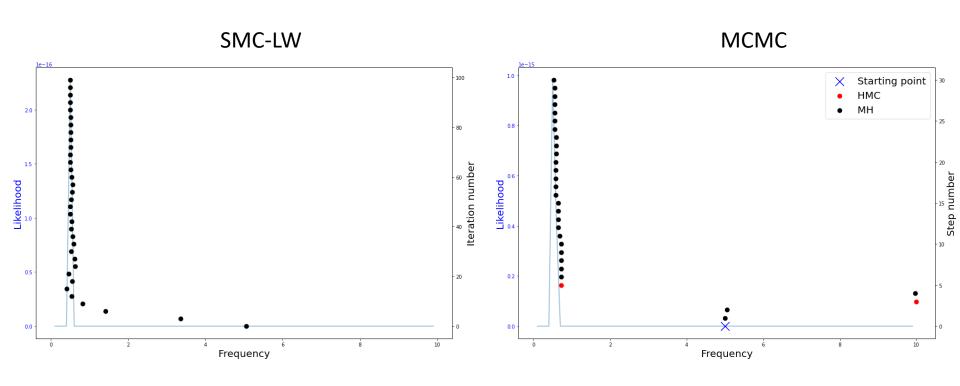
Bayesian phase estimation



Bayesian phase estimation



Bayesian phase estimation



Step by step tempered estimation

