

Supervised Classification of Multidimensional and Irregularly Sampled Signals

Application to Earth Observation

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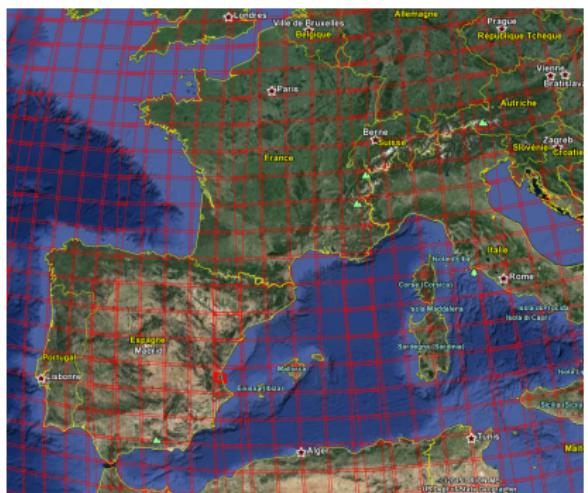
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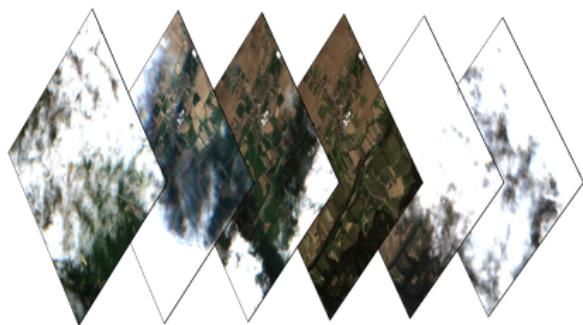
Satellite Image Time Series - SITS

Sentinel 2



Sentinel 2 Tiles

Width: 290km, altitude: 786km, ground resolution: 10m (R, G, B, near IR) - 20m (6 IR).



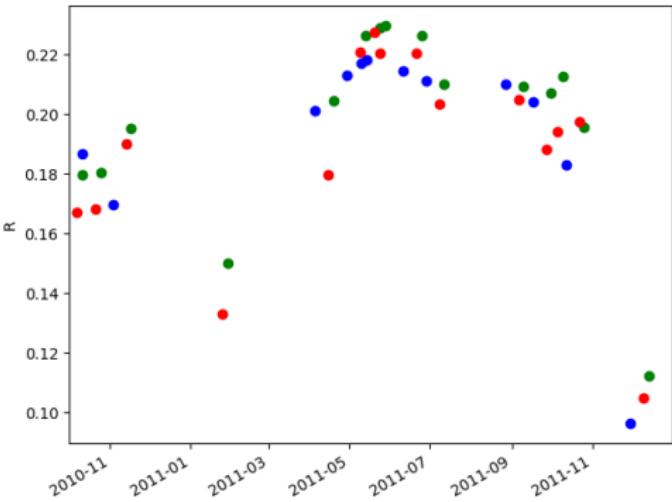
Time Series¹.

¹From *Sentinel Hub Blog* 2019.

Satellite Image Time Series - SITS

Temporal Features

- ▶ High temporal coverage
 - ▶ Large amount of data
(~ 70TB for Sentinel-2)
 - ▶ Noisy data
(clouds/shadows/..)



3 sampled SITS (Broad-Leaved, 2010)

Supervised Classification

For each SITS we want to assign a class c in $\{1, \dots, C\}$ among

- Artificial Areas (urban fabrics, road surfaces, ..),
 - Agricultural Areas (annual winter crops, vineyards, ..),
 - Forests and Semi-natural Areas (dunes, glaciers, ..),
 - Water Bodies (swamp, water, ..).

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Originality of our model

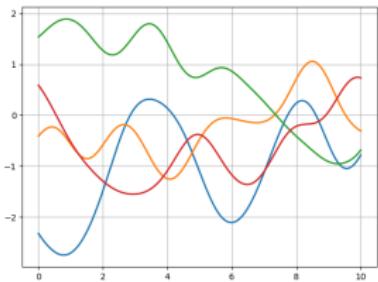
- Handle noisy data and **Irregular** sampling.
 - SITS might be observed at any time.

Standard approaches use temporal resampling, such as [Inglada et al. 2017] and use SITS classifier.

Gaussian Processes

Definition

A **Gaussian process**¹ (GP) is a stochastic process such that any finite-dimensional marginal follows a multivariate Gaussian distribution.



4 samples from a GP.

We write $X \sim \mathcal{GP}(m, K)$, where

$$m(t) = \mathbb{E}[X_t]$$

$$K(t, t') = \mathbb{E}[(X_t - m(t))(X_{t'} - m(t'))]$$

X_t is real-valued and t, t' are times.

¹[Williams and Rasmussen 2006]

Irregularly Sampled GP (IrSGP) Model

Y_b is a SITS and $b \in \{1, \dots, p\}$.

Our **main assumptions** are, conditionally to class Z (discrete)

- ## 1. It is a GP

$$Y_b | Z = c \sim \mathcal{GP}(m_{b,c}, K_{b,c})$$

2. Y_1, \dots, Y_p are independent.

The i th sample Y^i associated with Z^i at times $\{t_1, \dots, t_{T_i}\}$, then

$$Y_b^i | Z^i = c \sim \mathcal{N}_{T_i}(\mu_{b,c}^i, \Sigma_{b,c}^i)$$

where $\mu_{b,c}^i = [m_{b,c}(t_1), \dots, m_{b,c}(t_{T_i})]$ and $(\Sigma_{b,c}^i)_{k,l} = K_{b,c}(t_k, t_l)$.

IrSGP Model

Estimation of the b th wavelength

- m is supposed to belong to a subspace of J basis functions $\{\varphi_j\}_{j \geq 1}$

$$m_{b,c}(t) = \sum_{j=1}^J \alpha_{b,c,j} \varphi_j(t), \quad t \in \mathcal{T}.$$

Then $\mu_{b,c}^i = \mathbf{B}^i \alpha_{b,c}$ where

- $(\mathbf{B}^i)_{k,j} = \varphi_j(t_k^i)$ the design matrix;
 - $\alpha_{b,c} = [\alpha_{b,c,1}, \dots, \alpha_{b,c,p}]^\top$ projection coefficients.

- K is chosen in a set of parametric kernels indexed by θ , we write $\Sigma_{b,c}^i = \Sigma^i(\theta_{b,c})$.

Finally

$$Y_b^i | Z^i = c \sim \mathcal{N}_{T_i}(\mathbf{B}^i \color{blue}{\alpha_{b,c}}, \boldsymbol{\Sigma}^i(\color{red}{\theta_{b,c}}))$$

with parameters $\{\alpha_{b,c}, \theta_{b,c}\}$.

Optimizing hyperparameters

Maximizing Likelihood

Find $\{\hat{\alpha}_{b,c}, \hat{\theta}_{b,c}\}$ that maximizes the log-marginal likelihood.

$$LML(\boldsymbol{\alpha}_{b,c}, \boldsymbol{\theta}_{b,c}) \propto - \sum_{i|Z_i=c} \log |\boldsymbol{\Sigma}^i(\boldsymbol{\theta}_{b,c})| - \sum_{i|Z_i=c} (\boldsymbol{Y}_b^i - \mathbf{B}^i \boldsymbol{\alpha}_{b,c})^T \boldsymbol{\Sigma}^i(\boldsymbol{\theta}_{b,c})^{-1} (\boldsymbol{Y}_b^i - \mathbf{B}^i \boldsymbol{\alpha}_{b,c})$$

Alternate¹ minimization of:

- α : Explicit formula in $\nabla_{\alpha}(LML) = 0$.
 - θ : No explicit formulation: **Gradient descent** algorithm.

¹More details can be found in [Constantin et al. 2019] (French).

Optimizing hyperparameters

Optimizing w.r.t. α and θ

$$\bullet \hat{\alpha}_{b,c} = \left[\sum_{i|Z_i=c} (\mathbf{B}^i)^\top \boldsymbol{\Sigma}^i (\theta_{b,c})^{-1} \mathbf{B}^i \right]^{-1} \left[\sum_{i|Z_i=c} (\mathbf{B}^i)^\top \boldsymbol{\Sigma}^i (\theta_{b,c})^{-1} Y_b^i \right],$$

$$\bullet \frac{\partial}{\partial \theta_k} LML_{b,c} = \sum_{i|Z_i=c} \text{tr} \left(([\Sigma^i(\theta_{b,c})]^{-1} - \beta_{b,c}^i \beta_{b,c}^{i\top}) \frac{\partial \Sigma^i(\theta_{b,c})}{\partial \theta_k} \right),$$

where $\beta_{b,c}^i = \Sigma^i(\theta_{b,c})^{-1}(\mathbf{Y}_b^i - \mathbf{B}^i\hat{\alpha}_{b,c})$.

We compute $p \times C$ independent likelihood optimizations.

Classification

For a new signal \tilde{Y} , Bayes' rule yields to

$$P(Z = c | \tilde{Y}) \propto P(Z = c) \prod_{b=1}^p P(\tilde{Y}_b | Z = c)$$

- $P(Z = c)$ is estimated by the proportion of obs. in class c.
 - $P(\tilde{Y}_b | Z = c)$ is the density of the Gaussian vector evaluated with parameters $\hat{\alpha}_{b,c}$ and $\hat{\theta}_{b,c}$ at times $[\tilde{t}_1, \dots, \tilde{t}_{\tilde{T}}]$.
 - **Select** the class with highest probability (MAP).

Imputation of Missing Values

Let t^* be the time where the value is missing, then

$(Y_b^i(t^*), \mathbf{y}_{i,b}^\top)_{Z_i=c}^\top$ follows a Gaussian distribution

$$\begin{pmatrix} Y_b^i(t^*) \\ \mathbf{y}_{i,b} \end{pmatrix}_{Z_i=c} \sim \mathcal{N}_{T_i+1} \left(\begin{bmatrix} m_{b,c}(t^*) \\ \mathbf{B}^i \boldsymbol{\alpha}_{b,c} \end{bmatrix}, \begin{bmatrix} K^i(\theta_{b,c}|t^*, t^*) & \mathbf{K}^i(\theta_{b,c}|t_{1:T_i}, t^*) \\ \mathbf{K}^i(\theta_{b,c}|t^*, t_{1:T_i}) & \Sigma^i(\theta_{b,c}) \end{bmatrix} \right).$$

where $\mathbf{K}^i(\theta_{b,c}|t^*, t_{1:T_i}) = \{K(\theta_{b,c}|t^*, t_s)\}_{s \in T_i}^\top$.

Thanks to the properties of conditional expectations,

$$Y_b^i(t^*) | \mathbf{y}_{i,b}; Z_i = c \sim \mathcal{N}_1(\hat{Y}_{b,c}^i(t^*), \text{cov}(\hat{Y}_{b,c}^i(t^*))).$$

Imputation of Missing Values

Global and Conditional

When class is known to be c (Conditionally to $Z_j = c$)

$$\left\{ \begin{array}{l} \hat{Y}_{b,c}^i(t^*) = B^i(t^*) \hat{\alpha}_{b,c} \\ \quad + \mathbf{K}^i(\hat{\theta}_{b,c}|t^*, t_{1:T_i})^\top \Sigma^i(\hat{\theta}_{b,c})^{-1} (\mathbf{y}_{i,b} - \mathbf{B}^i \hat{\alpha}_{b,c}) \\ \text{cov}(\hat{Y}_{b,c}^i(t^*)) = K^i(\hat{\theta}_{b,c}|t^*, t^*) \\ \quad - \mathbf{K}^i(\hat{\theta}_{b,c}|t^*, t_{1:T_i}) \Sigma^i(\hat{\theta}_{b,c})^{-1} \mathbf{K}^i(\hat{\theta}_{b,c}|t_{1:T_i}, t^*) \end{array} \right.$$

When class is unknown

$$\left\{ \begin{array}{l} \hat{Y}_b^i(t^*) = \sum_{c=1}^C P(Z_i = c | \mathbf{y}_{i,b}) \hat{Y}_{b,c}^i(t^*), \\ \text{cov}(\hat{Y}_b^i(t^*)) = \sum_{c=1}^C P(Z_i = c | \mathbf{y}_{i,b}) \text{cov}(\hat{Y}_{b,c}^i(t^*) | \mathbf{y}_{i,b}, Z_i = c) \\ \quad + \sum_{c=1}^C P(Z_i = c | \mathbf{y}_{i,b}) \hat{Y}_{b,c}^i(t^*) \hat{Y}_{b,c}^i(t^*)^\top - \hat{Y}_b^i(t^*) \hat{Y}_b^i(t^*)^\top. \end{array} \right.$$

Results

Formosat 2011

Formosat 2011 SITS: $\mathcal{C} = 17$ classes and $p = 6$ dimensions, including 4 wavelenghts and 2 spectral indices (*NDVI* and *Brightness*).

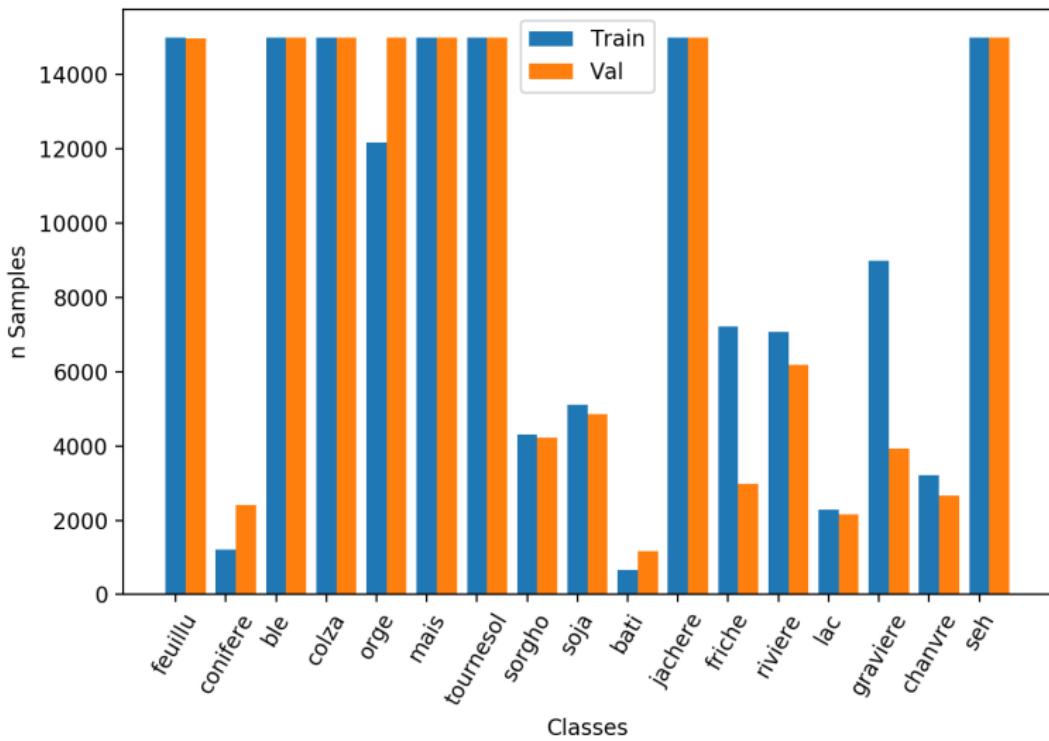
Groundtruth is divided into two spatially disjoint sets (train/val), with a maximum of 15k samples (per class).

The model is tested according to

- Estimation of the processes (m and K).
 - The imputation of missing values.
 - The classification scores.

Results

Formosat 2011 - Samples per Class



Model Settings

3 different basis functions:

- *Fourier*: α represents the Fourier coefficients

$$m(t) = \alpha_0 + \sum_{j=1}^J \alpha_j \cos(2\pi * j * f_0 t) + \sum_{j=1}^J \alpha_j \sin(2\pi * j * f_0 t),$$

- *Exp*: equidistant Gaussian densities

$$m(t) = \sum_{j=0}^J \alpha_j \exp\left(-\frac{(t-t_j)^2}{d}\right),$$

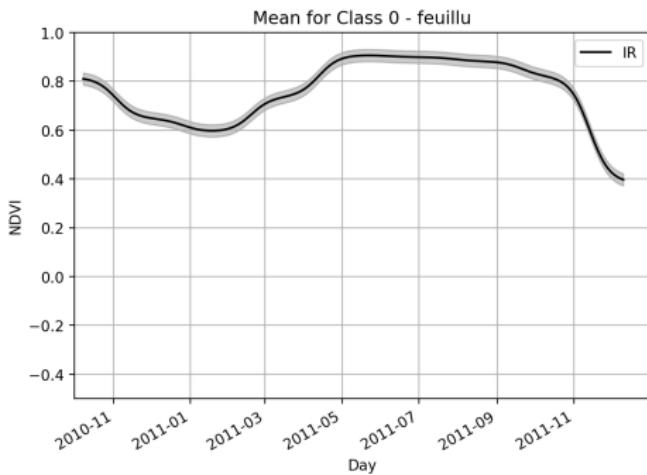
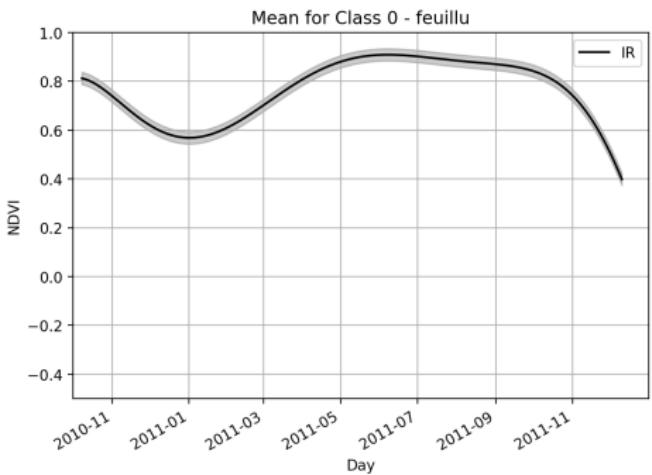
- *Polynomial*: Polynomial function of t

$$m(t) = \alpha_0 + \sum_{j=1}^J \alpha_j t^j.$$

K is parametrized by $K(t, t' | \theta_{b,c}) = \sigma_{b,c} \exp\left(-\frac{(t-t')^2}{h_{c,b}}\right) + \gamma_{b,c} \delta_{tt'}.$

Estimation¹ of the Process

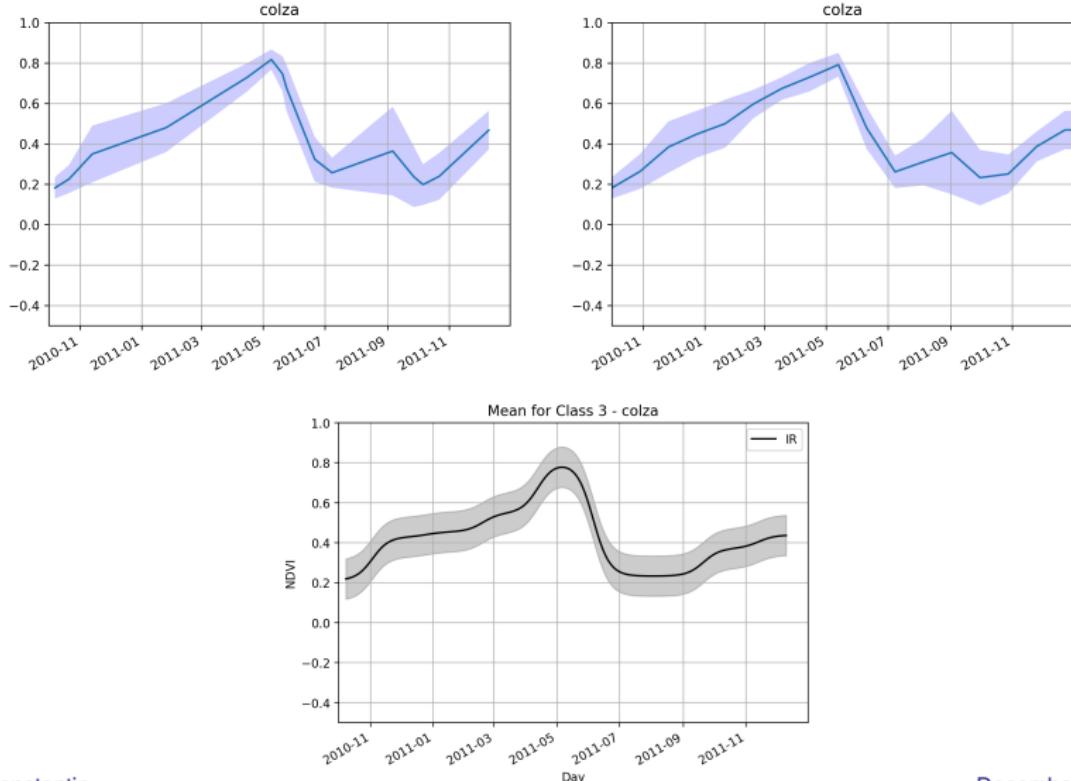
Polynomial and exp-basis



¹The standard deviation is the square root of the diagonal from the covariance matrix.

Estimation of the Process - Colza with Exp-basis

Figure: Emp. distrib. of training set (Top) - GP Estimation (Bottom)



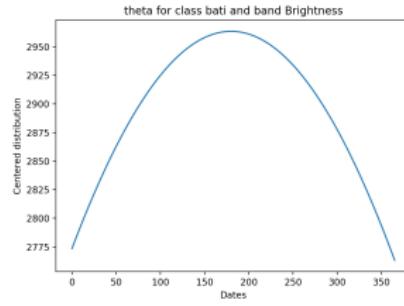
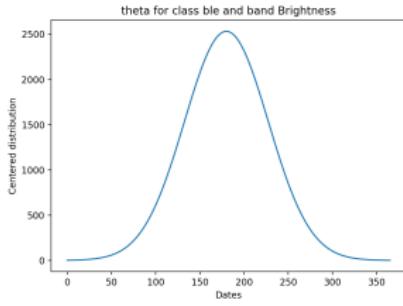
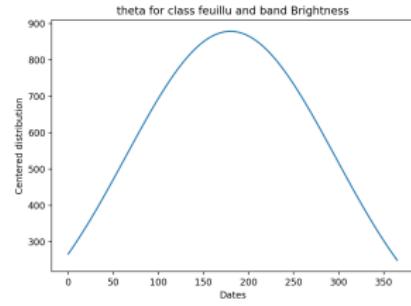
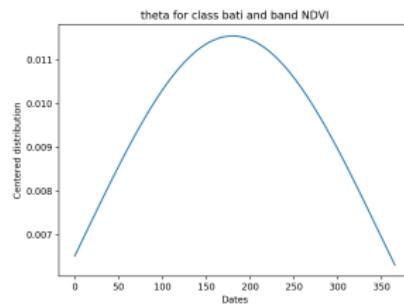
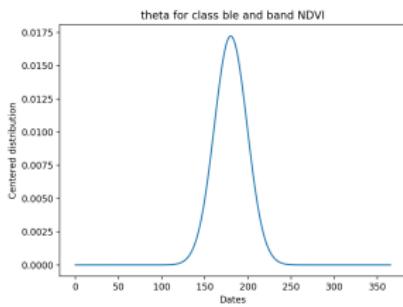
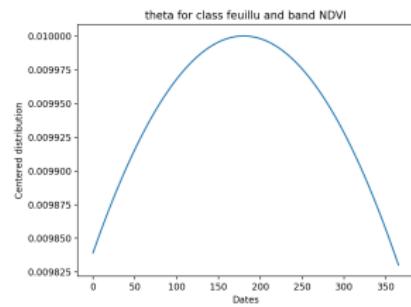
Motivation 000

Model of Gaussian Processes

Results

Conclusion

Estimation of Thetas

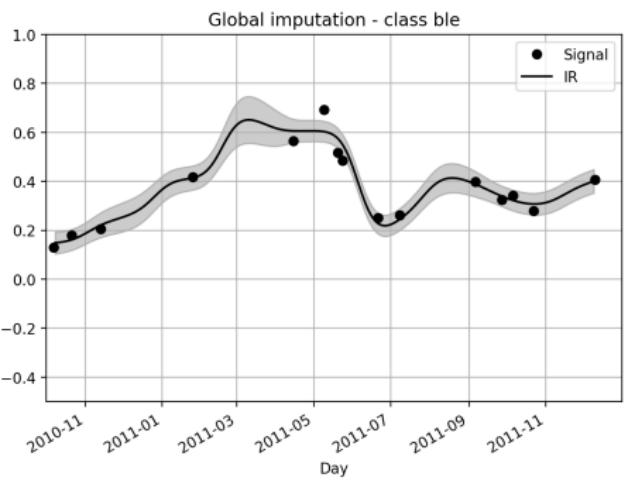


Motivation
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Model of Gaussian Processes
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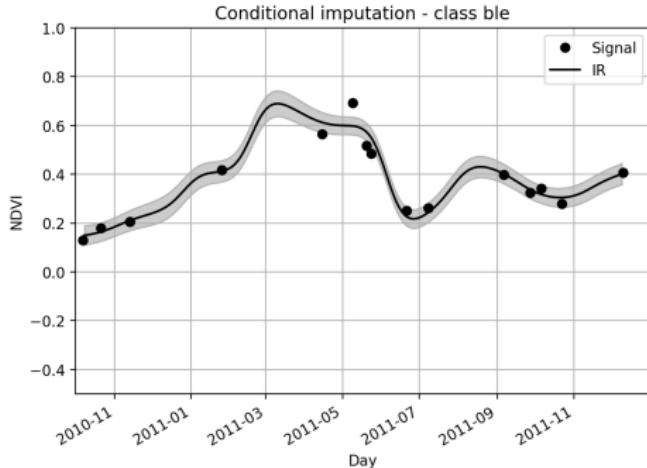
Reconstruction

Conditional imputation of missing values

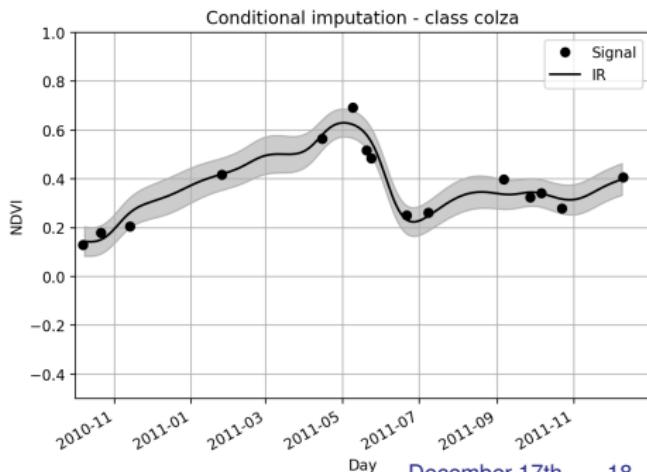


$$P(Z = \text{'wheat'} | Y^i) = 0.796$$
$$P(Z = \text{'colza'} | Y^i) = 0.204$$

Results



Conclusion

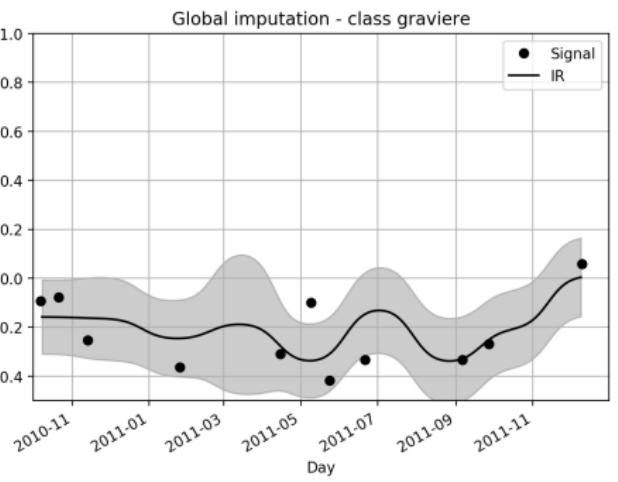


Motivation
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Model of Gaussian Processes
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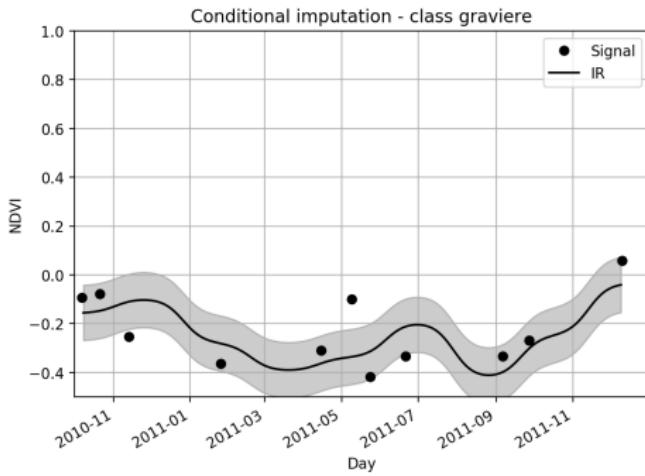
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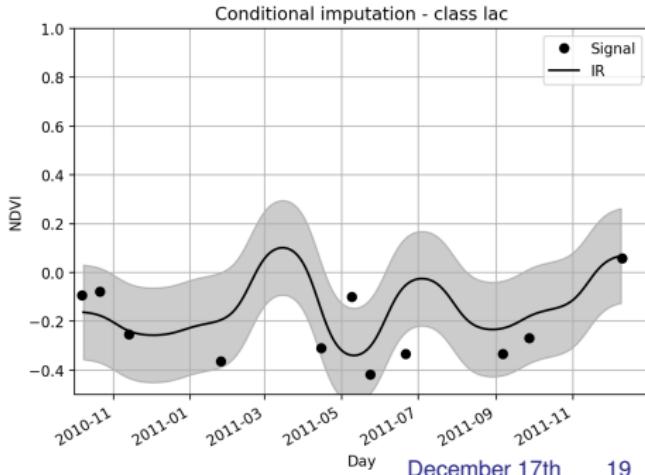


$$P(Z = \text{'gravel'} | Y^i) = 0.592$$
$$P(Z = \text{'lake'} | Y^i) = 0.408$$

Results



Conclusion



Classification Results

Table: Average scores among 10 independent runs.

		Gap-filled		
		QDA	RF	IrSGP
Acc	0.73	0.79	0.71	
κ	0.70	0.77	0.68	
F_1_{mean}	0.70	0.76	0.68	

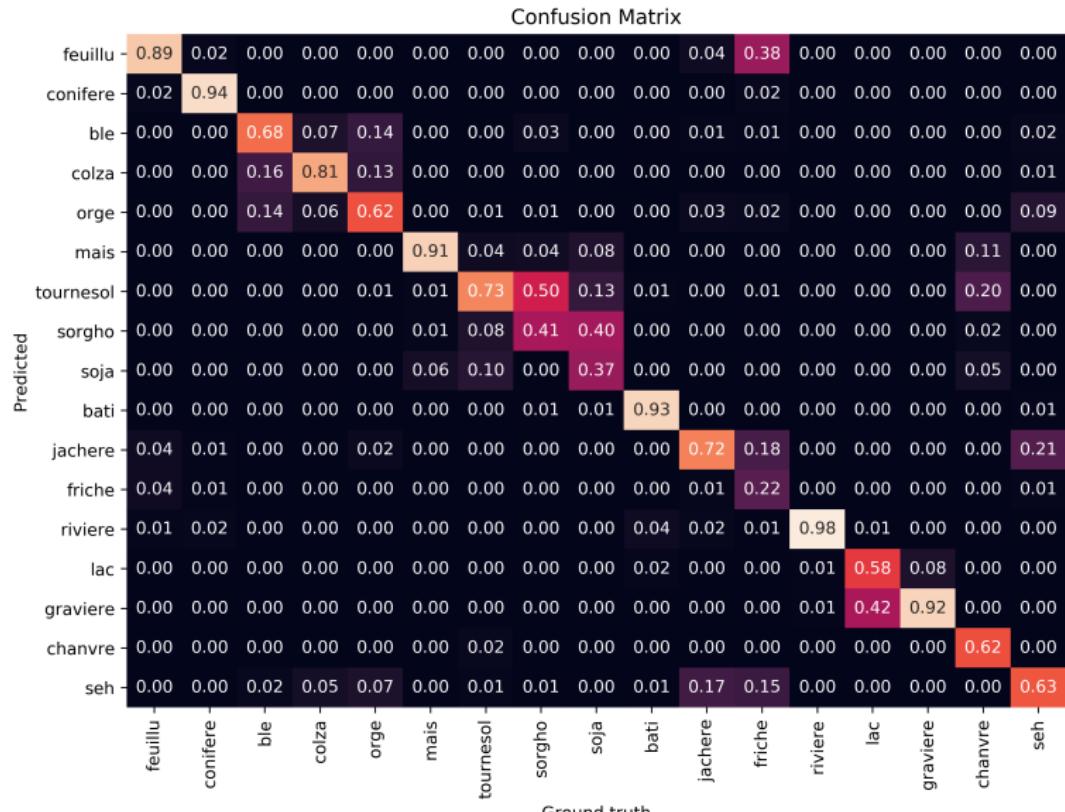
Settings:

QDA (Quadratic Discriminant Analysis), RF (Random Forest)
maxdepth 25, IrSGP (exp basis - 11 centers).

Comment:

IrSGP almost same performances as QDA and RF despite independence assumptions on the bands.

Confusion Matrix



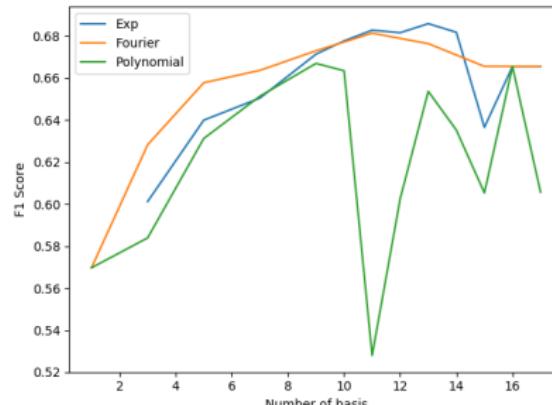
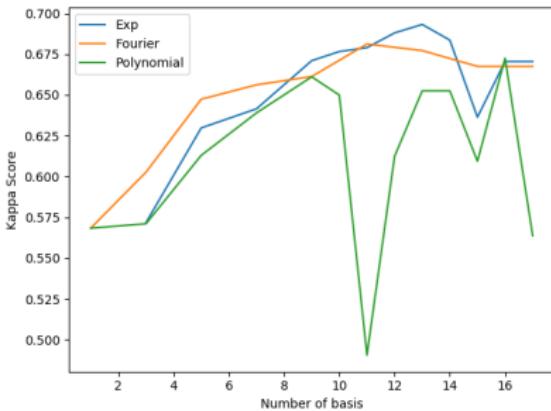
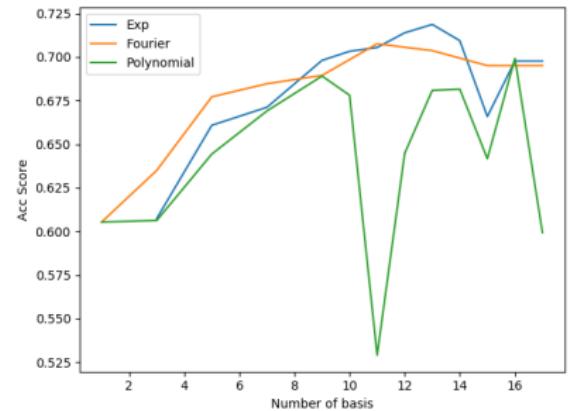
Motivation
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Model of Gaussian Processes
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Results
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Conclusion
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Classification - Sensitivity



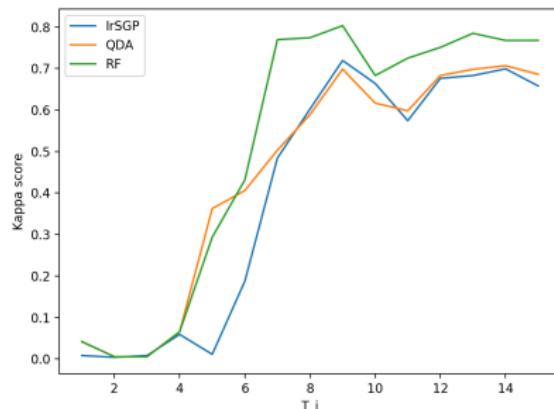
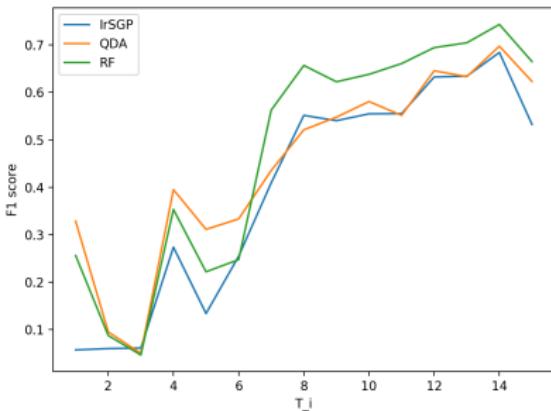
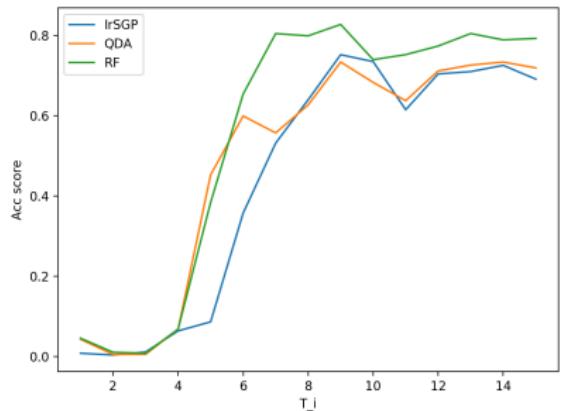
Motivation
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Model of Gaussian Processes
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Results
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Conclusion
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Classification - T_i available acquisitions



Conclusion

Able to handle irregularly sampled signals including

1. Modeling the classes using Gaussian processes.
2. Imputation of missing values.
3. Classify the timeseries.

Future work

1. Scale up to Sentinel-2 Dataset.
2. Consider the dependence between wavelengths.

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