

Hidden Markov Models and Sequential Monte Carlo

A study based on Fast Filtering with Large Option Panels

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2 SVCJ Model

3 Bootstrap Particle Filter

4 PMMH

5 Orthogonal MCMC

6 Discussion

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Objective of the Study

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Methodological Overview

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Objective of the Study

- Study and replicate a stochastic volatility state-space model
- Model studied in the paper:
 - SVCJ
- Focus on inference using Sequential Monte Carlo methods
- Compare particle-based approaches for parameter estimation

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Data Description

- Daily S&P 500 index prices
- Period: January 1996 – December 2015
- Risk-free rate: 3-month Treasury Bill
- Time discretization:

$$\Delta t = \frac{1}{252}$$

- Log-returns used as observations:

$$R_t = \log \left(\frac{S_t}{S_{t-1}} \right)$$

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Methodological Overview

- SVCJ Model
- Bootstrap Particle Filter
- Particle Marginal Metropolis–Hastings (PMMH)
- Orthogonal MCMC

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Observed and Latent Variables (SVCJ)

1. Observation Equation (log-returns)

$$R_{t+1} = r_t - \delta_t - \frac{V_t}{2} + \eta_s V_t - \lambda \bar{\mu}_s + \sqrt{V_t} z_{t+1} + J_{t+1}^s B_{t+1}$$

- R_{t+1} : observed log-return
- $\sqrt{V_t} z_{t+1}$: continuous diffusion, $z_{t+1} \sim \mathcal{N}(0, 1)$
- $J_{t+1}^s B_{t+1}$: price jump, $B_{t+1} \sim \text{Bernoulli}(\lambda \Delta t)$

2. State Equation (variance dynamics)

$$V_{t+1} - V_t = \kappa(\theta - V_t) + \sigma \sqrt{V_t} w_{t+1} + J_{t+1}^v B_{t+1}$$

- mean reversion: $\kappa(\theta - V_t)$
- volatility of volatility: $\sigma \sqrt{V_t} w_{t+1}$
- variance jumps: $J_{t+1}^v B_{t+1}$

Model Parameters (SVCJ)

Diffusion parameters

- κ : mean reversion speed
- θ : long-run variance
- σ : volatility of volatility
- ρ : correlation between price and variance shocks
- η_S : volatility risk premium

Jump parameters

- λ : jump intensity
- μ_S, σ_S : mean and standard deviation of price jumps
- μ_V : mean variance jump
- ρ_J : correlation between price and variance jumps

Fixed parameter

- $\delta = 0$: dividend yield

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Why Particle Filtering?

Implementation and Validation

Marginal Likelihood

Volatility Filtering: SVCJ vs Realized Volatility

Impact of the Number of Particles

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Why Particle Filtering?

- Bootstrap Particle Filter applied to the SVCJ model
- Number of particles: $N = 20,000$
- Parameters fixed to empirical estimates (Table A1, paper)
- Time discretization: $\Delta t = 1/252$
- Resampling method: **systematic resampling**
- Resampling threshold:

$$ESS_{\min} = 0.5 N$$

- Goal:
 - estimate latent states
 - estimate the marginal log-likelihood

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Fixed SVCJ Parameters (Implementation and Validation)

- Parameters fixed at the values reported in the article (Table A1)

- $\kappa = 7.4387$
- $\theta = 0.0244$
- $\sigma = 0.4387$
- $\rho = -0.8232$
- $\eta_s = 3.2508$
- $\delta = 0.0$
- $\lambda = 0.8128$
- $\mu_s = -0.0261$
- $\sigma_s = 0.0221$
- $\mu_v = 0.0822$
- $\rho_J = -0.0960$
- Number of particles: $N = 30,000$
- Time step: $\Delta t = 1/252$

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Marginal Likelihood

- Bootstrap filter provides an unbiased estimate of

$$p_{\theta}(R_{1:T})$$

- Estimated marginal log-likelihood:

$$\log \hat{p}(R_{1:T}) = 16,160.95$$

- Very close to the paper benchmark (Table A1)

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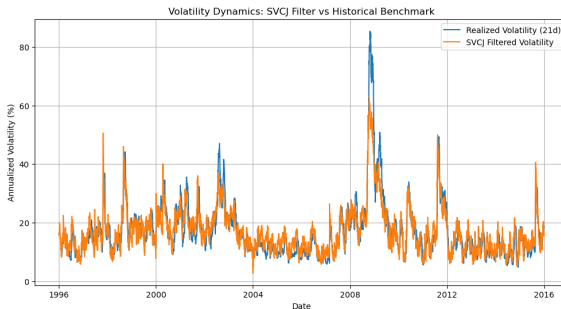
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Volatility Filtering: SVCJ vs Realized Volatility

- Latent volatility estimated as:

$$\mathbb{E}[V_t \mid R_{1:t}]$$

- Benchmark: 21-day realized volatility
- Correct timing of volatility spikes
- Extreme peaks slightly underestimated



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Number of Particles (1/2)

Why this can be a problem

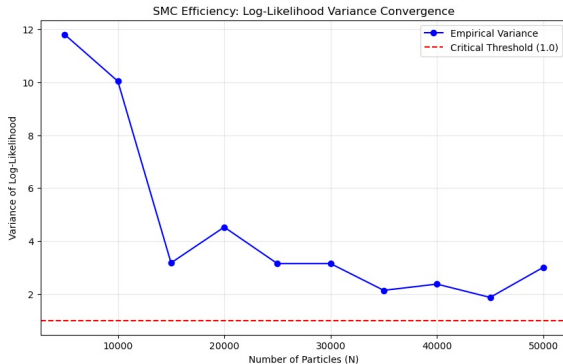
- PMMH efficiency strongly depends on the variance of the log-likelihood estimator.
- Optimal performance is achieved when:

$$\text{Var}(\log \hat{p}_{\theta}(y_{1:T})) \ll 1$$

- If the variance is too large: low acceptance rates and poor mixing.
- Increasing the number of particles indefinitely is not sufficient: diminishing variance reduction and prohibitive computational cost.

Number of Particles (2/2)

What we observe



- Variance decreases rapidly for small N , then plateaus.
- Beyond $N \approx 30,000$, gains are marginal and the variance remains > 1 .

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Why Multiple PMMH Runs?

PMMH on SV Model

PMMH on Jump Parameters

PMMH on Full SVCJ Model

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Why Multiple PMMH Runs?

- Full SVCJ model is high-dimensional and noisy
- Likelihood estimation is costly and unstable
- Strategy: estimate parameters progressively
- Each PMMH targets a different model specification

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PMMH on SV Model

- Model: stochastic volatility without jumps
- Parameters estimated:

$$\kappa, \theta, \sigma, \rho, \eta_s$$

- Jump components fixed to zero
- Goal: validate PMMH on diffusion dynamics

PMMH Methodology: SV Model

- Configuration: 30,000 particles and 500 iterations to balance numerical noise with execution speed.
- Prior distributions: Uniform for $\kappa, \theta, \sigma, \eta_s$ and Logit-Normal for $\rho \in (-1, 1)$.
- Initial variance: $(0.1 \times \text{width})^2$ for Uniforms and $(0.1 \times \text{scale})^2$ for ρ .
- Adaptation: learned parameter correlation structure and automated scaling to target the optimal acceptance rate.

PMMH Results: SV Model vs. Table A1 Benchmark

Numerical Performance

Param.	Table A1 (SV)	Our Run
Mean Log-L	16,133.41	16,113.51
Variance	3.5106	1.0660
Acc. Rate	–	29.66%

*Based on 20 Bootstrap runs ($N = 30,000$)

- **Stability:** Our variance is **3x lower** than the paper, ensuring significantly more stable posterior exploration.
- **Fit:** Log-likelihood is close to the benchmark.

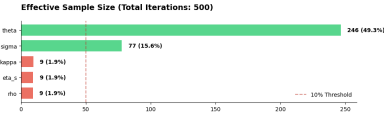
Parameter Estimates

Param.	Benchmark	Posterior Mean
κ	6.4802	7.0878
θ	0.0339	0.0336
σ	0.5121	0.4789
ρ	-0.7886	-0.6465
η_s	2.3818	4.7806

- **Efficiency:** The 29.7% acceptance rate validates the proposal scaling in the 5D space.
- **Alignment:** κ , θ and σ match closely; η_s and ρ reflect local data adaptations.

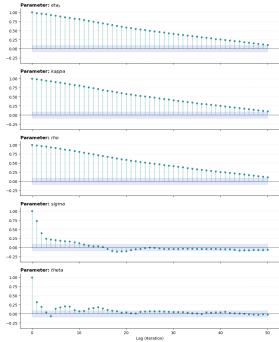
MCMC Diagnostics: SV Model

ESS Efficiency



- **High Identification:** θ (49%) and σ (15%) exceed the 10% threshold, ensuring reliable scale estimates.
- **Critical Persistence:** κ, ρ, η_s at 1.9% (9 samples). $N = 500$ is too short to decouple these dynamics.

ACF Mixing



- **Fast:** θ, σ mix instantly.
- **Persistent:** κ, ρ, η_s remain significant at lag 50.

Verdict: Chain too short ($N = 500$) for structural parameters (κ, ρ, η_s).

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PMMH on Jump Parameters

- Diffusion parameters fixed to SV posterior means
- Parameters estimated:

$$\lambda, \mu_s, \sigma_s, \mu_v, \rho_j$$

- Methodology: Same configuration as the diffusion stage ($N = 30,000$, 500 iterations, adaptive mechanism).
- Priors: Uniform for jump sizes (μ_s, σ_s, μ_v) and Logit-Normal for intensity (λ) and correlation (ρ_j) .
- Goal: isolate the contribution of jumps

PMMH Results: Jump Parameters vs. Table A1

Numerical Performance

Params	Table A1 (SVCJ)	Our Run
Mean Log-L	16,160.82	16,128.72
Variance	5.0186	2.0912
Acc. Rate	—	28.06%

*Based on 20 Bootstrap runs ($N = 30,000$)

- **Model Fit:** +16 log-points over SV; jumps are essential for capturing tail behavior.
- **Numerical Stability:** Variance (2.1) is lower than benchmark, reducing filter noise.

Parameter Estimates

Param.	Benchmark	Posterior Mean
λ	0.8128	0.6509
μ_S	-0.0261	-0.0089
σ_S	0.0221	0.0348
μ_V	0.0822	0.1134
ρ_J	-0.0960	-0.0706

- **Robustness:** Current estimates provide a more stable likelihood surface than Table A1.
- **Efficiency:** 28% acceptance rate validates the adaptive proposal calibration.

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PMMH on Full SVCJ Model

- Full parameter vector (10 parameters)
- Warm-start: Initialized with posterior means from previous partial runs.
- Covariance matrix initialized with SV and Jump-only sub-matrices.
- Increased particle count: $N = 40,000$
- Adaptive refinement of joint correlations over 500 iterations.

Numerical Performance

*Based on 20 Bootstrap runs ($N = 30,000$)

- ## Parameter Estimates

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Orthogonal MCMC (O-MCMC): Concept

- **Problem:** Standard PMMH chains often get stuck in local modes due to the high dimensionality of the SVCJ model.
- **Solution:** Run a population of parallel chains that interact to share information.
- **Algorithm Structure:**
 - ① **Vertical Step (Local Exploration):** Each chain runs independent PMMH steps (M_v iterations) to explore its local neighborhood using the Particle Filter.
 - ② **Horizontal Step (Global Interaction):** Chains exchange information using a **Differential Evolution** proposal:

$$\theta_{\text{prop}} = \theta_i + \gamma(\theta_a - \theta_b) + \epsilon$$

Allows "bold jumps" across the parameter space based on population dispersion.

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Implementation Details

Configuration ("Fast Prototyping")

- **Population:** $Z = 10$ parallel chains.
- **Particles:** $N_x = 2,000$ (Lower than PMMH for computational speed).
- **Vertical step:** $M_v = 10$ short bursts.
- **Schedule:** 100 cycles of alternating vertical/horizontal steps.
- **Total Effort:** $\approx 10,000$ likelihood evaluations.

Objective

- Test if population-based methods can find high-probability regions faster than a single tuned chain.
- Recover latent volatility dynamics under constraints.

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Model Selection: Naive vs. O-MCMC

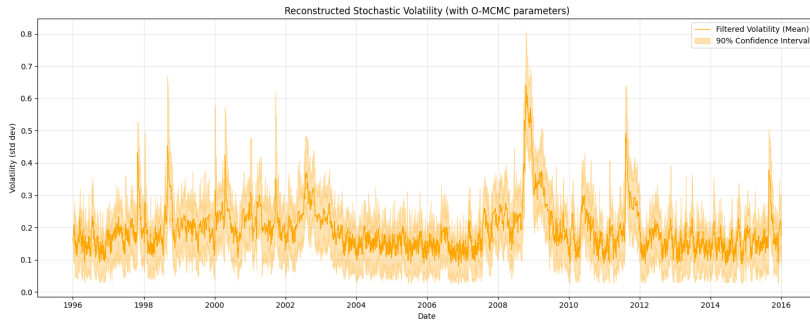
- **Naive Benchmark:** Geometric Brownian Motion (Constant Volatility).
- **O-MCMC:** Stochastic Volatility with Correlated Jumps.

Metric	Naive Model	PMMH	O-MCMC
Log-Likelihood	14,929.86	16,132.44	16,099.53
Fit Gain	–	+1,202.58	+1,169.67

Interpretation:

- The massive gain (> 1000 points) provides "decisive evidence" (Bayes Factor) in favor of the SVCJ specification.
- Successfully captures stylized facts: Fat tails, Volatility Clustering, and Leverage Effect.

Latent State Reconstruction



- **Regime Detection:** Clearly identifies the 2008 crisis ($t \approx 3200$) with volatility peaking at $\sim 80\%$.
- **Uncertainty:** The 90% CI (orange band) naturally expands during turbulent periods, correctly quantifying risk.

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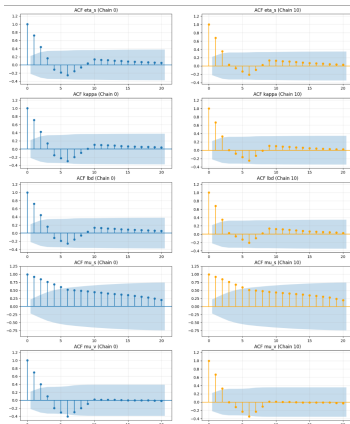
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Diagnostics: The "Sticky Chain" Problem (1/2)



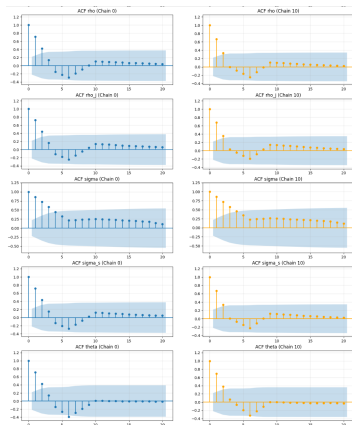
Observations

- **High Autocorrelation:** The visual evidence of slow decay in the ACF plots indicates "sticky" chains.
- **Root Cause:** With only $N = 2000$ particles, the estimator variance is high ($\text{Var}(\log \hat{L})$).

Diagnostics: The "Sticky Chain" Problem (2/2)

Mechanics of Failure

- **Effect:** High variance leads to a noisy acceptance ratio. The Metropolis-Hastings step rejects most proposals, causing chains to stay fixed (flat lines) for many cycles.



Global Conclusion

- The O-MCMC algorithm acted as a powerful **optimizer** (successfully finding the high-probability mode).
- However, it performed as a poor **sampler** (failed to quantify

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Thanks!