

Exchange Rate Misalignments, Interdependence, Crises, and Currency Wars: An Empirical Assessment*

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Contents: 1. Introduction; 2. Motivation; 3. Literature on exchange rate misalignment; 4. Econometric methodology; 5. Database description; 6. Results; 7. Comparison with the results of existing literature; 8. Conclusions; 9. Appendix A; 10. Appendix B: Complementary Tables.

Keywords: Real effective exchange rate, Cointegration, Exchange rate misalignment

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This study aims to compare two different methodologies of calculating exchange rate misalignment and test whether there is interdependence among countries in determining the real effective exchange rate. Two different econometric approaches are used to achieve these goals. The first one involves estimating a multivariate time series model that contains only country-specific variables and evaluating if this basic model can be improved by adding other countries' variables. The study uses the algorithm suggested by Hendry and Krolzig (2005) to select the best model specification. The second strategy involves estimating long panel data with the real effective exchange rate and fundamentals for a group of countries and explicitly testing the interdependence hypothesis. The results suggest that the long-run exchange rate is mainly driven by its own fundamentals for most countries. The existence of interdependence is restricted to short-run dynamics.

Este trabalho tem por objetivo comparar metodologias distintas para cálculo de desalinhamento cambial além de testar a hipótese se as taxas de câmbio dos diversos países sofrem influencia apenas dos seus próprios fundamentos ou também da taxa de câmbio e dos fundamentos de outros países. Estas hipóteses consistem, respectivamente, na ausência ou na existência de interdependência entre os diversos países. Para realizar tal tarefa utilizam-se duas estratégias empíricas. A primeira baseia-se em avaliar se

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um modelo multivariado de séries de tempo usualmente utilizada na literatura de desalinhamento cambial com dados apenas do próprio país em análise pode ser melhorado através da adição de variáveis relacionadas a outros países usando o algoritmo proposto por Hendry and Krolzig (2005). A segunda estratégia consiste em estimar um panel longo com as variáveis utilizadas para estimar desalinhamento cambial e testar formalmente a hipótese de ausência de interdependências. Os resultados sugerem que em ambas estratégias existe evidência de existência de interdependência. Esta ocorreria mais por conta de fatores ligados ao curto prazo, ou seja, o que explicaria o valor da taxa de câmbio de um país no longo prazo seriam seus próprios fundamentos enquanto no curto prazo fatores externos poderiam causar desvios.

1. INTRODUCTION

The exchange rate is an important macroeconomic variable and its variation affects global and sectorial economic activity, prices and interest rates, and trade flows. Large fluctuations of the exchange rate exacerbate debates regarding whether such movements are “excessive,” reflect “fundamentals,” or are “rational.”

Empirical literature on the exchange rate has advanced toward building models for assessing the long-term determinants of real exchange rates. Empirical strategies can be formulated based on models that use the doctrine of Purchasing Power Parity or an analysis based on fundamentals.¹

Given this second approach, in this study, I estimate the long-term determinants of the real exchange rate of the G-20 as well as certain other selected countries. From these models, I construct the estimates of exchange rate misalignment for each country and test whether there is interdependence between various currency misalignments. The results suggest that the long-run exchange rate is mainly driven by its own fundamentals for most countries. The existence of interdependence is restricted to short-run dynamics.

The remainder of this paper is organized in the following manner. In the second section, I discuss the importance of exchange misalignment. In the third section, I conduct a non-exhaustive review of existing literature. In the fourth section, I discuss the econometric methodology employed in greater detail. In the fifth section, I present the results and compare them with those of similar studies in existing literature. In the last section, I present some conclusions.

2. MOTIVATION

The severity of the crisis in the US economy in 2008 brought the fear of a strong negative contagion in the rest of the world. Among other measures, the US authorities adopted an aggressive monetary policy with a strong reduction in nominal interest rates and monetary expansion. According to analysts, such measures generated strong pressure for US dollar depreciation against currencies of countries whose governments did not follow a similar monetary policy of lowering their domestic interest rates. If a country chooses not to follow such a reduction and opts to keep interest rates unchanged and accumulate reserves to prevent appreciation of its currency, it will eventually have to face inflation pressures. Some authors argue that the Federal Bank, after the sub-prime crisis, has been trying to use

¹Clark and MacDonald (1998). Exchange rates and economic fundamentals: a methodological comparison of BEERs and FEERs, International Monetary Fund.

monetary policy to generate a depreciation of the dollar rate to create an aggregate demand stimulus to speed up economic recovery. This policy has generated worldwide repercussions, and the possible disadvantages of such effects are a subject of debate.

One of the channels through which the monetary policy operates is real exchange rate depreciation. If the effect on other countries is deleterious, these countries could attempt to use available instruments to avoid the appreciation of their local currencies. This might configure a policy known in the literature as “beggar thy neighbor”.² Countries can adopt measures to prevent the appreciation of their currencies to avoid internalizing the costs of a deliberate expansionary monetary policy (“quantitative easing”).³

A cost-benefit analysis of this policy should be conducted not only for ascertaining its short-term effects but also for its long-term effect on output, inflation, and the welfare of countries. The countries’ strategic choice should aim to obtain short-term economic gains through an expansionary monetary policy despite being exposed to suffering some sort of retaliation from partners. It is also important to analyze the circumstances under which a country’s partners do not retaliate and accept the costs of appreciation as the best strategy.

For a small country, the strategy of attempting to retaliate is perhaps not the best as it is very difficult to cause significant damage to a leader country, thereby rendering such a policy ineffective. The main mechanism by which monetary policy would function is via devaluation of the exchange rate; however, for a small country, it would be difficult to counteract the effects of exchange rate appreciation by changing its own monetary policy. Turnovsky and Basar (1987) is an example of a study in which this strategic interaction is formally analyzed.

An expansionary monetary policy operates via the mechanism of excessive depreciation of currency. The degree and persistence of exchange rate misalignment is the mechanism by which net exports react in the short term and generate an increase in aggregate demand and smooth output fluctuations around its potential level. Some empirical questions can be formulated in this regard:

- a) What is the extent to which monetary policy can generate a significant currency misalignment?
- b) How long it will last?
- c) Can significant misalignments of a country’s monetary policy be quickly offset by similar policies of relevant trading partners of the leader country?
- d) What are the effects of such policies on the economic welfare of a country Turnovsky and Basar (1987)?
- e) To what extent would the exchange rate policies of countries be the genesis of currency crises Glick and Rose (1998)?

3. LITERATURE ON EXCHANGE RATE MISALIGNMENT

The existing literature on the real exchange rate is vast. The classical doctrine and perhaps one of the oldest one is Purchasing Power Parity (PPP). Reference to this theory can be found in classical economics. Recently, some studies have collected evidence in favor of the validity of PPP for tradable goods, although the adjustment toward equilibrium can be very slow Froot and K. Rogoff (1995). Ahmad and Craighead (2010) analyzed monthly consumer price indexes for a secular American and British data set. The study showed strong evidence in favor of mean reversion, but with a high half-life. Their work follows Taylor’s guidelines Taylor (2001).

²Eichengreen and Sachs (1986). Exchange rates and economic recovery in the 1930s, National Bureau of Economic Research Cambridge, Mass., USA.

³Ugai (2007). “Effects of the quantitative easing policy: A survey of empirical analyses.” MONETARY AND ECONOMIC STUDIES-BANK OF JAPAN 25(1): 1.



3.1. The economics of exchange rate misalignment

There is a theoretical discussion on the variables that drive the exchange rate in the long run. Two old but very important references are Edwards (1987) and Dornbusch (1976). The first analyzes economic misalignment, its causes, and consequences. The second is the classical model of flexible exchange rates. They analyze how monetary policy shocks can cause deviations from long-term PPP and the phenomena of under and over shooting.

Bilson (1979) and Mussa (1976) are also related classical studies. They are in keeping with the monetary approach to exchange rate. According to this approach, the exchange rate is linked primarily to the relative evolution of output and money supply among countries on the assumption that PPP and uncovered interest parity hold continuously and countries have stable money demand functions. However, Meese and Rogoff (1983) casts doubt on the explanatory power of this theory. The study showed that exchange rate forecasts that are constructed using this approach are not superior to a naive model such as a pure random walk with drift.

Stein (1995) proposes the natural rate of exchange approach (NATREX). According to the study, the equilibrium exchange rate level is determined at the point at which investment and savings generated by economic fundamentals are equal.

A classical reference to exchange rate misalignment is made by Williamson (1994). He defines equilibrium exchange rate as the level at which the country is able to maintain a certain desired level of deficit or surplus (seen as sustainable) in its external accounts. This is called the fundamental approach of the real exchange rate (FRER). Cline (2008) is another reference in this regard.

One source of criticism comes from the fact that there is a high degree of arbitrariness in choosing the desirable level of external accounts. Another source of criticism comes from the focus of the approach on flows and not on stocks.

Faruquee (1995) seeks to incorporate issues related to the evolution of stocks and builds a model in which there is an interaction between flows and stocks. Therefore, he shows that there must be a stable relationship between real exchange and the net foreign investment position between residents and non-residents. This is called the behavioral approach of the real exchange rate (BRER). The model was extended by Alberola and Cervero (1999). Further, Kubota (2009) developed a model with intertemporal consumption choice and capital accumulation. The model suggests that the real exchange rate is a function of relative terms of trade, net external position, and relative productivity of tradable and non-tradable sectors. This is the approach used in the current study.

The advantage of this approach is the low degree of subjectivity in the exchange rate misalignment estimate. The model connects the level of real exchange rate to a group of fundamentals obtained from a theoretical model. This approach is empirically implemented using an econometric model to decompose the series of real exchange rates into permanent and transitory components.

4. ECONOMETRIC METHODOLOGY

In this study, we propose to test whether the models estimated using a traditional approach to calculate currency misalignment from a list of fundamentals can be improved upon by modeling with long panel data that take into account a possible interdependence among different countries. In order to this, two testing strategies are used.

4.1. Testing Strategies

4.1.1. Strategy 1

The starting point is Larsson and Lyhagen's model Larsson and Lyhagen (2007). It is a generalization of Johansen and Juselius' procedure (Johansen (1988), Johansen (1995) and Juselius (2009)).

Let $x_{it} = [RER_{it} \text{ Fundamentals}_{it}]$ be the vector that contains the real exchange rate and its fundamentals at time t for country i , and let $X_t = [x'_{1t} \dots x'_{Nt}]'$ be the vector in which country data are stacked.

The general model is given by the following equation:

$$\Delta X_t = \Gamma_1 \Delta X_{t-1} + \dots + \Gamma_k \Delta X_{t-k} + \Phi D_t + A B X_{t-1} + \varepsilon_t \quad (1)$$

where ε_t represents random errors, and Ω is its variance and covariance matrix of the shocks. Shocks are assumed to be uncorrelated in time and cross-sectional dimensions. The vector X_t contains the data series from all countries, D_t contains deterministic terms, and $\theta = \Gamma_1, \dots, \Gamma_k, \Phi, A, B$ denotes the set of parameters to be estimated. In the general case, all the variables could cointegrate with each other,⁴ but Larsson and Lyhagen (2007), for example, restrict their analysis to the case in which there can be cointegration only between variables of the same country. Then,

$$B = \begin{bmatrix} \beta_1 & 0 & 0 & \dots & 0 \\ 0 & \beta_2 & 0 & \dots & 0 \\ 0 & 0 & \ddots & \ddots & \vdots \\ \vdots & \vdots & \dots & \beta_{N-1} & 0 \\ 0 & 0 & \dots & 0 & \beta_N \end{bmatrix}, \quad (2)$$

where β is a vector of dimension pxr_i .

The matrix of load (A) has dimension $(PXN)x(r1 + r2 + \dots + rN)$ and no *a priori* constraint and enables all sorts of configuration, provided that the system is stable. One cointegrated vector that contains the variables of one country can generate not only part of the equation of that country but also other countries' equations. It is precisely this possibility that makes this structure attractive for analyzing the problem considered in this work.

For example, it is possible to test if the cointegration vector, which represents the fundamentals of a country (e.g., US), is present in exchange rates and fundamentals variables' equations. This hypothesis could be tested under the null hypothesis that certain elements of matrix A are zero. Larsson and Lyhagen (2007) showed that after the rank of a long-run matrix is determined, the test can be implemented using a likelihood ratio test with an asymptotic chi-squared distribution. The authors also derived a test to determine the rank of cointegration; this test has no standard asymptotic distribution, and this distribution was tabulated. Among other interesting hypotheses, it is also possible to test jointly, for example, whether the net external investment position is one of the fundamentals for all analyzed countries.

One limitation of this structure is the untested assumption that B is a diagonal matrix. Except for cases in which the temporal dimension of the database is sufficiently large, it is difficult to assess the validity of this hypothesis due to the curse of dimensionality. In order to circumvent this problem, we chose two different strategies. We estimated models with small sizes in relation to time and cross-sectional dimensions. We decided to estimate models for all combinations of countries taken as pairs for a sample with data from 1981 to 2011, and a model of higher dimension with five developed countries and Brazil using a larger sample (beginning in 1970).

The estimation of the model is done by using the algorithm of generalized reduced rank regression (GRRR) developed by Hansen (2003). All restrictions imposed on the general model can be tested using a likelihood ratio test with an asymptotic chi-squared distribution.

⁴The concept of cointegration has become widely used in empirical research with the seminal work of Engle e Granger (1987). "Co-integration and Error Correction: Representation, Estimation and Testing." *Econometrica* 55: 251-276.



The likelihood ratio procedure used in the study suffers from a serious size distortion. This encourages analysts to use a bootstrap technique to obtain appropriate critical values. Studies that discuss similar problems are Johansen (1999), Johansen et alii (2000), Gredenhoff and Jacobson (2001), and Swensen (2006). We chose to perform a bootstrap to simulate the appropriate distribution for the test statistic. The results suggest that the size distortion is actually severe.

4.1.2. Strategy 2

The second strategy for testing the existence of interdependence involves estimating multivariate models. First, a country-by-country analysis is done to obtain the long-term relationships that will be used as input for the construction of the exchange rate misalignment estimate. Then, the error correction mechanisms obtained for all countries are included in each individual country model. Thereafter, their relevance is tested using the algorithm suggested by Santos and Hendry (2008) and Hendry and Krolzig (2005). This procedure was implemented using software Oxmetrics. If the final selected model contains only the series of the individual country, it can be concluded that there is no evidence for the interdependence hypothesis, whereas in the event that the final country model contains variables related to other countries, it can be concluded that there is some degree of interdependence.

4.2. Gonzalo and Granger's decomposition

Many decomposition methods have been suggested for the decomposition of the process between permanent and transitory components.

The data generating process for a specific country is assumed to be a multivariate VECM model given by:

$$\Delta x_{it} = \psi_{i1}\Delta x_{i,t-1} + \dots + \psi_{i1}\Delta x_{i,t-k+1} + \phi_{i1}d_{it} + \alpha_i\beta x_{i,t-1} + v_{i,t} \quad (3)$$

where $v_{i,t}$ represents random errors, and θ_i is its variance and covariance matrix of the shocks. Shocks are assumed to be uncorrelated in time. The vector $x_{i,t}$ contains the data series for country i , $d_{i,t}$ contains deterministic terms, and $\theta_i = \psi_{i1}, \dots, \psi_{i1}, \phi_{i1}, \alpha_i, \beta_i$ denotes the set of parameters to be estimated.

In general, the decomposition has the following form:⁵

$$x_{i,t} = \beta_{i,\perp}(\beta_{i,\perp}'\beta_{i,\perp})^{-1}\beta_{i,\perp}'x_{i,t} + c_{\perp}(\beta_{i,\perp}'c_{\perp})^{-1}\beta_{i,\perp}'x_{i,t} \quad (4)$$

The decompositions vary according the choice of vector c . A condition for the existence of the decomposition is that the matrix $(\beta_{i,\perp}'c_{\perp})$ has full rank. However, this condition is not always satisfied.

Gonzalo and Granger (1995) propose that $c = \alpha_{i,\perp}$.⁶ This representation always exists for the case of a VECM of zero order. Johansen (1995) proposes that . This decomposition exists in the system since there are variables whose order of integration is at most 1. Kasa (1992) proposes that . Another possibility is to generate forecasts from the VECM estimated for each point. The values for which the series converge are the fundamentals. In this study, I use the decomposition given by Gonzalo and Granger (1995).⁷

⁵ β_{\perp} denotes the vector orthogonal to β in such a manner that $\beta' \perp \beta = 0$

⁶Gonzalo and Granger's decomposition-Gonzalo, J. and C. W. J. Granger (1995). "Estimation of Common Long-Memory Components in Cointegrated Systems." Journal of business and Economics Statistics 13(1).was implemented in Matlab software.

⁷In this case, the components of the deterministic model as constant and time trends should be restricted to the space of cointegration.

In Gonzalo and Granger's decomposition, there is no Granger causality between transitory and permanent components in that the former do not⁸ cause a change in permanent components in the long run; in other words, the exchange rate misalignment, defined as the transitory component of the real exchange rate equation in a multivariate system, does not contain information that is relevant for predicting the change in the permanent components in the long run.⁹ Gonzalo and Granger's decomposition is widely used in the literature of exchange rate misalignment. See, for example, Alberola and Cervero (1999). Their decomposition can be easily adapted if the data generating process is given by (1).

5. DATABASE DESCRIPTION

The data for this study were collected from the International Financial Statistics of the International Monetary Fund (IMF). The real exchange rate for each country is calculated from a basket of currencies that uses data from consumer price indexes. The data on net foreign investment position and reserves have been collected by the IMF from 2000 onward and those for previous years were collected on the basis of those given in Lane and Milesi-Ferretti (2007). The gross domestic product data were collected from the World Bank World Development Indicators available online.

The codes were run in Matlab and Oxmetrics software, and the frequency is annual. In many cases, data begins in 1970 and ends in 2010; in other cases, the sample begins in 1980 and ends in 2010. The countries that are analyzed are Australia, Brazil, Canada, Colombia, Denmark, Finland, France, Germany, Greece, India, Italy, Portugal, Ireland, Japan, Mexico, Netherlands, South Korea, South Africa, Singapore, Spain, Sweden, United Kingdom, Uruguay, Turkey, and the United States. Data from 1970 until 2012 is not available for all countries. However, all listed countries at least have information from 1980 onward.

6. RESULTS

As a first step, analysis tests were conducted using multivariate cointegration Johansen (1988). For countries that showed significant relationships, the error correction mechanisms were calculated. Table 20 contains the results of cointegration tests Johansen (1988), Shin (1994), and Engle e Granger (1987). The estimated error correction mechanism was used as input for strategy 2. Moreover, the error correction mechanisms were again estimated for strategy 1.

6.1. Results of Strategy 1

6.1.1. Long panel data (1970-2010): Australia, Brazil, Canada, the United States, Japan, and the United Kingdom

In order to investigate the existence of interdependence among exchange rate misalignments in the above countries, I estimated the model of equation (1). The model under the null hypothesis is given by

⁸For a rigorous definition of Granger causality, see Hendry (Hendry (1995). *Dynamic Econometrics*. Oxford, Oxford University Press).

⁹If the most appropriate VECM contains only the erroneous mechanism and deterministic terms, then Gonzalo and Granger's decomposition Gonzalo and Granger (1995) and C. W. J. Granger (1995). "Estimation of Common Long-Memory Components in Cointegrated Systems." *Journal of Business and Economics Statistics* 13(1) satisfies the condition to be a decomposition of the class proposed by Beveridge-Nelson (Beveridge e Nelson (1981). "A new approach to decomposition of economic time series into permanent and transitory components with particular attention to measurement of the business cycle." *Journal of Monetary Economics* 7(2): 151-174.).



$$(2)B = \begin{bmatrix} \beta_1 & 0 & 0 & \dots & 0 \\ 0 & \beta_2 & 0 & \dots & 0 \\ 0 & 0 & \ddots & \ddots & \vdots \\ \vdots & \vdots & \dots & \beta_{N-1} & 0 \\ 0 & 0 & \dots & 0 & \beta_N \end{bmatrix} \quad \text{and} \quad (2)A = \begin{bmatrix} \alpha_1 & 0 & 0 & \dots & 0 \\ 0 & \alpha_2 & 0 & \dots & 0 \\ 0 & 0 & \ddots & \ddots & \vdots \\ \vdots & \vdots & \dots & \alpha_{N-1} & 0 \\ 0 & 0 & \dots & 0 & \alpha_N \end{bmatrix}$$

Except for the variance and covariance matrices of the shocks, there is no source of interdependence. In the alternative hypothesis, matrices A and B are completely unrestricted.

The results of this test are presented in Table 1. Based on the asymptotic value, the null hypothesis is easily rejected, but it is accepted if bootstrap critical values are used. Since the tests can suffer from size distortions, I tend to rely on the option of not rejecting the null.

Table 1: Results of the test for assessing the existence of interdependence based on misalignments in exchange rates

Likelihood Ratio Test:	
Test Statistic	689,67
Number of restrictions	180
asymptotical <i>p</i> -value	0,00%
bootstrap <i>p</i> -value*	28,10%

* 300 replications.

6.1.2. Peer analysis: Using data from 1980-2011 and $N = 2$

The second group of estimates was conducted using pairs of countries. The results are presented in Table 2. Again, the difference between the asymptotic critical and simulated critical values is large. As a general rule, the simulated critical values were consistently larger than the asymptotic critical values.

The analysis in Table 2 suggests that there is an important interdependence across countries, and these findings are not in line with the evidence presented in section 6.1.1. For example, in the Brazilian case, there is evidence of interdependence with the United Kingdom, the United States, and Australia. The lack of interdependence between Brazil and Canada as well as Canada and Japan is confirmed. Interdependence was detected for the United States with Brazil, Canada, France, India, South Africa, South Korea, and the United Kingdom. However, there is no evidence of interdependence between the remaining countries and the United States. Further, there is strong evidence of interdependence among European countries, as expected.

Table 2: Dependence tests with models from pairs of countries: Short Panel

	Australia	Brazil	Canada	Colombia	Denmark	Finland	France	Germany	Greece	India	Ireland	Japan	Mexico	Netherlands	South Korea	South Africa	Singapore	Spain	Sweden	UK	Uruguay	United States		
Austria	Statistic	38.9	43.4	23.4	25.3	27.5	31.4	32.0	29.1	42.6	34.2	33.7	52.8	25.4	55.1	26.9	31.9	46.9	62.1	24.2	29.9	29.3	29.8	
	Asymptotic p-value	0.0%	0.0%	0.9%	0.5%	0.2%	0.1%	0.0%	0.1%	0.0%	0.0%	0.0%	0.0%	0.5%	0.0%	0.3%	0.0%	0.0%	0.0%	0.7%	0.1%	0.1%	0.1%	
	Bootstrap p-value	8.0%	4.0%	22.0%	48.0%	54.0%	30.0%	28.0%	36.0%	24.0%	10.0%	18.0%	0.0%	44.4%	4.0%	48.0%	22.0%	2.0%	8.0%	36.0%	18.0%	38.0%	46.0%	
Australia	Statistic	58.9	47.0	36.4	40.6	44.8	13.1	21.8	18.9	34.0	21.8	24.0	31.8	24.0	31.8	24.2	23.6	40.6	33.4	22.5	35.3	29.9	16.5	
	Asymptotic p-value	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	22.0%	1.6%	4.1%	0.0%	100.0%	1.6%	0.8%	0.0%	0.7%	0.9%	0.0%	0.0%	1.3%	0.0%	0.1%	8.6%	
	Bootstrap p-value	0.0%	6.0%	2.0%	10.0%	10.0%	0.0%	42.0%	12.0%	60.0%	2.0%	38.0%	20.0%	2.0%	6.0%	24.0%	24.0%	0.0%	4.0%	48.0%	2.0%	26.0%	46.0%	
Brazil	Statistic	24.2	23.6	0.7%	0.9%	0.1%	34.4%	1.7%	0.0%	1.1%	4.9%	0.0%	0.2%	1.7%	0.6%	7.7%	0.3%	0.0%	1.5%	38.8%	0.0%	0.2%	0.0%	
	Asymptotic p-value	0.7%	0.9%	0.0%	0.1%	0.0%	34.4%	1.7%	0.0%	1.1%	4.9%	0.0%	0.2%	1.7%	0.6%	7.7%	0.3%	0.0%	1.5%	38.8%	0.0%	0.2%	0.0%	
	Bootstrap p-value	14.0%	52.0%	0.0%	10.0%	18.0%	68.0%	30.0%	12.0%	56.0%	64.0%	34.0%	30.0%	24.0%	14.0%	44.0%	14.0%	2.0%	56.0%	84.0%	4.0%	6.0%	2.0%	
Canada	Statistic	39.3	21.7	48.7	45.8	32.8	23.8	39.0	23.5	15.7	24.7	44.0	12.5	25.8	19.4	25.8	19.4	22.1	21.1	28.9	21.4	31.0	31.0	
	Asymptotic p-value	0.0%	1.7%	0.0%	0.0%	0.0%	0.0%	0.8%	0.0%	0.9%	10.9%	0.6%	0.0%	25.4%	0.0%	60.0%	12.0%	3.6%	1.4%	2.0%	0.1%	1.9%	0.1%	
	Bootstrap p-value	10.0%	0.0%	0.0%	0.0%	0.0%	0.0%	16.0%	32.0%	0.0%	38.0%	70.0%	4.0%	2.0%	2.0%	60.0%	12.0%	50.0%	26.0%	18.0%	28.0%	42.0%	4.0%	
Colombia	Statistic	28.2	36.9	32.5	36.7	39.2	20.3	26.2	31.5	32.8	36.7	19.2	87.2	33.5	46.4	21.9	36.2	33.5	46.4	21.9	36.2	29.9	34.2	
	Asymptotic p-value	0.2%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.26%	0.3%	0.0%	0.0%	0.0%	0.0%	0.0%	3.8%	0.0%	0.0%	0.0%	1.5%	0.0%	0.1%	0.0%	
	Bootstrap p-value	16.0%	0.0%	8.0%	12.0%	12.0%	24.0%	20.0%	22.0%	2.0%	12.0%	26.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	28.0%	6.0%	18.0%	8.0%	
Denmark	Statistic	29.8	25.2	29.9	46.2	25.8	34.7	48.2	22.9	37.8	47.1	44.8	49.6	53.5	38.3	31.9	26.5	41.2	49.6	53.5	38.3	31.9	26.5	41.2
	Asymptotic p-value	0.1%	0.5%	0.1%	0.0%	0.4%	0.0%	0.0%	0.0%	1.1%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.3%	0.0%	
	Bootstrap p-value	28.0%	44.0%	24.0%	24.0%	26.0%	52.0%	46.0%	12.0%	48.0%	24.0%	8.0%	6.0%	6.0%	6.0%	6.0%	6.0%	2.0%	6.0%	6.0%	26.0%	52.0%	22.0%	
Finland	Statistic	47.7	24.8	53.7	56.6	35.9	45.1	23.8	26.4	26.7	39.5	30.6	43.6	41.3	24.5	28.1	36.5	30.6	43.6	41.3	24.5	28.1	36.5	
	Asymptotic p-value	0.0%	0.6%	0.0%	0.0%	0.0%	0.0%	0.0%	0.8%	0.3%	0.3%	0.0%	0.0%	0.0%	0.1%	0.0%	0.0%	0.0%	0.0%	0.0%	0.6%	0.2%	0.0%	
	Bootstrap p-value	2.0%	26.0%	6.0%	0.0%	0.0%	26.0%	10.0%	62.0%	46.0%	38.0%	10.0%	10.0%	32.0%	14.0%	78.0%	34.0%	32.0%	6.0%	14.0%	78.0%	34.0%	18.0%	
France	Statistic	28.2	18.2	27.9	21.6	50.9	38.9	18.4	24.3	26.1	16.9	52.9	35.5	38.9	24.3	53.4	16.9	52.9	35.5	38.9	24.3	53.4	53.4	
	Asymptotic p-value	0.2%	5.1%	0.2%	1.7%	0.0%	0.0%	4.9%	0.7%	0.4%	7.8%	0.0%	0.0%	0.7%	0.0%	0.0%	0.4%	7.8%	0.0%	0.0%	0.0%	0.7%	0.0%	
	Bootstrap p-value	16%	84%	56.0%	36.0%	4.0%	4.0%	64.0%	64.0%	28.0%	28.0%	70.0%	0.0%	8.0%	12.0%	50.0%	28.0%	70.0%	0.0%	8.0%	12.0%	50.0%	4.0%	
Germany	Statistic	42.7	55.1	38.1	38.3	48.3	40.8	35.1	16.1	16.1	26.6	32.7	33.8	30.0	45.8	41.2	26.6	32.7	33.8	30.0	45.8	41.2	41.2	
	Asymptotic p-value	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.3%	0.0%	0.0%	0.0%	0.1%	0.0%	0.0%	
	Bootstrap p-value	12.0%	2.0%	28.0%	24.0%	0.0%	0.0%	8.0%	8.0%	8.0%	72.0%	34.0%	14.0%	38.0%	2.0%	16.0%	34.0%	0.0%	0.0%	14.0%	38.0%	2.0%	16.0%	
Greece	Statistic	33.9	57.5	32.7	44.9	39.9	36.0	34.2	36.0	67.5	35.3	39.4	41.3	36.0	67.5	35.3	39.4	36.0	67.5	35.3	39.4	41.3	41.3	
	Asymptotic p-value	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	
	Bootstrap p-value	44.0%	0.0%	64.0%	40.0%	20.0%	16.0%	48.0%	28.0%	0.0%	22.0%	16.0%	30.0%	22.0%	0.0%	0.0%	28.0%	0.0%	0.0%	22.0%	16.0%	30.0%	22.0%	
India	Statistic	42.2	28.5	27.0	76.7	45.8	100.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	45.8	53.9	30.2	54.6	40.5	71.6	71.6	
	Asymptotic p-value	0.0%	0.1%	0.3%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.1%	0.0%	0.0%	0.0%	
	Bootstrap p-value	18.0%	30.0%	60.0%	2.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	4.0%	62.0%	0.0%	10.0%	0.0%	
Ireland	Statistic	31.6	27.6	19.1	21.1	2.0%	36.6	26.7	31.9	33.4	54.1	41.8	22.9	2.0%	0.0%	0.0%	0.3%	26.7	31.9	33.4	54.1	41.8	22.9	
	Asymptotic p-value	0.0%	0.2%	3.9%	2.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.3%	0.0%	0.0%	0.0%	0.0%	1.1%	
	Bootstrap p-value	40.0%	52.0%	68.0%	62.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	20.0%	56.0%	26.0%	24.0%	2.0%	18.0%	30.0%	
Japan	Statistic	28.5	27.0	0.1%	0.1%	0.3%	76.7	45.8	100.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	44.5	44.5	31.9	40.7	47.5	45.1	45.1	
	Asymptotic p-value	0.1%	0.1%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	
	Bootstrap p-value	48.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	
Mexico	Statistic	25.4	32.0	33.9	33.6	34.5	23.8	33.6	34.5	23.8	33.6	34.5	23.8	33.6	34.5	23.8	33.6	34.5	23.8	33.6	34.5	23.8	33.6	
	Asymptotic p-value	0.5%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.8%	0.0%	0.0%	
	Bootstrap p-value	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	14.0%	24.0%	24.0%	12.0%	46.0%	46.0%	12.0%	
Netherlands	Statistic	44.7	36.1	36.5	46.6	19.7	42.5	30.3	44.7	36.1	36.5	46.6	19.7	42.5	30.3	44.7	36.1	36.5	46.6	19.7	42.5	30.3	44.7	
	Asymptotic p-value	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.0%	0.2%	0.0%	0.0%	
	Bootstrap p-value	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	12.0%	8.0%	4.0%	58.0%	2.0%	38.0%	12.0%	

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Table 2 – Continued

	Australia	Brazil	Canada	Colombia	Denmark	Finland	France	Germany	Greece	India	Ireland	Japan	Mexico	Netherlands	South Korea	South Africa	Singapore	Spain	Sweden	UK	Uruguay	United States
South Korea																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
South Africa																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
Singapore																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
Spain																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
Sweden																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
UK																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						
Uruguay																						
Statistic																						
Asymptotic p-value																						
Bootstrap p-value																						

6.2. Results of Strategy 2

The results of a sequence of tests for assessing the existence of interdependence among different countries are reported in the following tables. Two models are estimated. The first one is a multivariate VECM model for each country, expanded with the cointegration vectors of the remaining countries. Thereafter, I used a simplification algorithm available at Oxmetrics. The second model is a multivariate model that is expanded not only with the error correction mechanisms but also with the first difference of the logarithm of the exchange rate and the ratio of Net Foreign Asset Position (NFA) on GDP, which is the indicator of the Balassa-Samuelson effect. In this case, this second option not only tests the existence of interdependence of misalignments (imbalances) but also tests the short-run dynamics of a specific country in relation to the short-run dynamics of other countries.

6.2.1. Brazil

Table 3 presents the results of the multivariate model for Brazil. In the analysis of the extended model (model 1) using the cointegration vectors of other countries, The algorithm excluded the variables of the Brazilian final model. Thus, the exchange rate misalignments elsewhere do not appear to cause large influences on the dynamics of the Brazilian series. The results from the search algorithm for outliers suggest three points of instability of the Brazilian model in 1999, 2008, and 2009. The last two are related to the recent developments of the international crisis. The first is related to the transition from an almost fixed exchange rate regime after the Real Plan to a floating exchange rate regime in 1999.

Table 3: Final model selected using the Ox algorithm: Brazil

Brazil								
Model 1								
		DLRER_bra_1	I:1999	I:2008	I:2009	Constant	Cla_bra_1	
DLRER_t	Coefficient	0,2122	-0,3592	0,0651	-0,0232	0,0147	-0,0527	
	Standard Error	0,1395	0,1131	0,1125	0,1117	0,0184	0,0180	
	t	1,5200	-3,1800	0,5780	-0,2080	0,7990	-2,9300	
	p-value	13,78%	0,32%	56,69%	83,66%	43,01%	0,62%	
DLNFA_t	Coefficient	-0,013	-0,118	0,242	-0,204	0,000	-0,020	
	Standard Error	0,049	0,039	0,039	0,039	0,006	0,006	
	t	-0,264	-2,980	6,180	-5,250	-0,050	-3,200	
	p-value	79,33%	0,53%	0,00%	0,00%	96,07%	0,31%	
DLBS_t	Coefficient	-0,259	0,088	0,028	-0,006	0,025	0,008	
	Standard Error	0,067	0,055	0,054	0,054	0,009	0,009	
	t	-3,860	1,610	0,525	-0,110	2,850	0,903	
	p-value	0,05%	11,72%	60,32%	91,29%	0,74%	37,32%	
Model 2								
		DLRER_bra_1	Cla_bra_1	DLRER_usa_1	I:1999	I:2008	I:2009	Constant
DLRER_t	Coefficient	0,2090	-0,0497	-0,1766	-0,3490	0,0558	-0,0287	0,0132
	Standard Error	0,1412	0,0191	0,3445	0,1161	0,1152	0,1134	0,0188
	t	1,4800	-2,6000	-0,5130	-3,0100	0,4840	-0,2530	0,7040
	p-value	14,87%	1,39%	61,16%	0,51%	63,16%	80,19%	48,63%
DLNFA_t	Coefficient	-0,015	-0,018	-0,120	-0,111	0,236	-0,208	-0,001
	Standard Error	0,049	0,007	0,119	0,040	0,040	0,039	0,006
	t	-0,309	-2,740	-1,010	-2,770	5,940	-5,320	-0,200
	p-value	75,96%	0,99%	32,11%	0,94%	0,00%	0,00%	84,29%
DLBS_t	Coefficient	-0,251	0,000	0,459	0,061	0,053	0,008	0,029
	Standard Error	0,060	0,008	0,146	0,049	0,049	0,048	0,008
	t	-4,200	0,018	3,150	1,250	1,080	0,173	3,650
	p-value	0,02%	98,58%	0,35%	22,20%	28,83%	86,39%	0,09%

Note: Cla denotes the cointegration vector.



Model 2 was obtained from the general model that contained not only the cointegration vectors of all the countries analyzed, but also the first difference from the logarithm of the variables collected for all countries. The final model suggests that the changes in the effective exchange rate in the United States influence the Brazilian economy. Either way, the dynamics of the Brazilian exchange rate do not appear to have been influenced directly by global factors. The only economy that appears to have influenced the Brazilian data is the US economy, and this was widely expected given the extent of trade and financial integration between the two countries.

6.2.2. The United States

Table 5 presents the results of the tests of interdependence applied to the United States. The results suggest that the misalignment of other countries does not influence the dynamics of the model that contains US series. This result contrasts with the conclusions reported in section 6.1.2. Only the model with an extended information set provides evidence in favor of interdependence. The change in the real exchange rate of the Netherlands appears to be linked to United States data. This result is not intuitive and could have been due to the casual retention of unimportant variables. The results for the United States suggest that the US economy is barely influenced by changes in the economies of the rest of the world.

In the North American model, there were some points of instability-the years 1974, 1986, 2005, 2008, and 2009. The last two are associated with the recent financial crisis, while the first one is related to the oil crisis.

Table 4: Final model obtained from the selection made using the Ox algorithm: United States

United States								
Model 1								
		DLRER_usa_1	I:1974	I:1986	I:2005	I:2008	I:2009	Constant
DLRER_t	Coefficient	0,5442	0,0502	-0,1537	0,0146	-0,0201	0,0513	-0,2004
	Standard Error	0,1269	0,0446	0,0469	0,0429	0,0431	0,0436	0,1036
	t	4,2900	1,1300	-3,2800	0,3390	-0,4660	1,1800	-1,9300
	p-value	0,02%	26,89%	0,26%	73,68%	64,48%	24,85%	6,22%
DLNFA_t	Coefficient	-0,011	-0,014	0,011	0,078	-0,095	0,061	-0,043
	Standard Error	0,070	0,025	0,026	0,024	0,024	0,024	0,057
	t	-0,152	-0,551	0,423	3,270	-3,980	2,530	-0,744
	p-value	87,98%	58,53%	67,50%	0,26%	0,04%	1,68%	46,24%
DLBS_t	Coefficient	-0,175	0,358	0,109	0,038	0,020	-0,060	-0,203
	Standard Error	0,151	0,053	0,056	0,051	0,051	0,052	0,124
	t	-1,150	6,720	1,950	0,745	0,386	-1,150	-1,640
	p-value	25,75%	0,00%	5,98%	46,21%	70,20%	26,01%	11,04%
Model 2								
		DLRER_net_1	I:1974	I:1986	I:2008	Constant	Cla_eua_1	
DLRER_t	Coefficient	-0,9557	0,0370	-0,1691	-0,0418	-0,1534	0,0095	
	Standard Error	0,2112	0,0424	0,0458	0,0413	0,0986	0,0060	
	t	-4,5300	0,8720	-3,7000	-1,0100	-1,5600	1,5900	
	p-value	0,01%	38,92%	0,08%	31,91%	12,93%	12,18%	
DLNFA_t	Coefficient	0,301	-0,026	0,022	-0,100	-0,076	0,004	
	Standard Error	0,136	0,027	0,029	0,027	0,064	0,004	
	t	2,210	-0,939	0,749	-3,760	-1,200	1,140	
	p-value	3,39%	35,47%	45,89%	0,07%	23,73%	26,30%	
DLBS_t	Coefficient	0,085	0,368	0,103	0,028	-0,188	0,011	
	Standard Error	0,265	0,053	0,057	0,052	0,124	0,008	
	t	0,322	6,920	1,790	0,548	-1,520	1,430	
	p-value	74,96%	0,00%	8,31%	58,75%	13,75%	16,30%	

Note: Cla denotes the cointegration vector.

6.2.3. Other Countries

Detailed results for all countries are presented in the appendix (Table 9-Table 19). Table 5 presents the results of the evaluation of the hypothesis that the exchange rate misalignments of a country do not influence the variables in a different country. The country-specific model was expanded to include the misalignment of other countries and then subjected to a simplification algorithm using Oxmetrics. The results suggest that there is little evidence of interdependence via exchange rate misalignment. In other words, misalignments in a country should not generate, as a rule, long-term effects in other countries.

Table 5: Summary of the results of evaluating interdependence using exchange rate misalignment

Test Interdependence of exchange rate misalignments													
Country ij	Austria	Brazil	Canada	Japan	UK	Sweden	Colombia	India	Ireland	Netherlands	Spain	United States	Uruguay
austria	n.a.						Yes				Yes		
Brazil		n.a.											
Canada			n.a.									Yes	
Japan				n.a.		Yes							
UK					n.a.								
Sweden	Yes				Yes	n.a.							
India							n.a.						
Ireland							Yes	n.a.					
Colombia									n.a.				
USA										n.a.			
Uruguay										Yes	n.a.		
Spain												n.a.	
Netherlands						Yes							n.a.

n.a. - not applicable.

Table 6 presents the results of the strategy that contains the results of a data set that is expanded with not only misalignments but also the first differences of the various variables of other countries. In this case, the evidence for interdependence increases substantially. This finding suggests that in the long run, exchange rates would be simply explained by domestic fundamentals, while in the short run, events in larger economies like the United States or United Kingdom may generate a certain impact on other countries. There is also evidence of a strong regional interconnection among European countries, which was highly expected.

A competitive depreciation of the exchange rate of a country may generate long-term effects on the exchange rate of remaining countries if it permanently affects the level of fundamentals in one country. For models reported in Table Table Table and Table significant coefficients of the changes in the exchange rates of other countries in the international investment position (NFA) and productivity indicator (BS) equations can be interpreted as evidence that a competitive depreciation in one country would lead to long-term effects in certain countries. One of the few countries that appear to generate such an influence is the United States.

Table 6: Summary of the results of evaluating interdependence among countries using the expanded information data set

Country <i>i</i>	Evaluating interdependence using the expanded data set																
	Austria	Brazil	Canada	Japan	United Kingdom	Ireland	India	Sweden	United States	Uruguay	Colombia	France	Germany	Spain	Netherlands	Australia	Denmark
Austria	n.a.	n.a.	n.a.	n.a.	Yes	n.a.	n.a.	n.a.	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Brazil	Yes	n.a.	n.a.	n.a.	Yes	n.a.	n.a.	n.a.	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Canada	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Japan	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
UK	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Sweden	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
India	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Ireland	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Colombia	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
USA	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Uruguay	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Spain	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Netherlands	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

n.a. - not applicable.

Finally, the algorithm used permits the identification of the points of instability in the structure. Table 7 presents the points of instability in the expanded models obtained for each country. The findings reinforce the idea that the events in the oil market in the 1970s and the consequent large price increase of oil in international markets had a global impact. Recent events related to the US and European economic crisis also appear to have a global impact.

Table 7: Summary of identified points of instability in the estimated models using exchange rate misalignment

Points of Instability in the global model extended by exchange rate misalignment																								
Country i - Dummy	1973	1974	1975	1976	1977	1979	1980	1981	1983	1984	1985	1986	1993	1994	1999	2000	2001	2002	2003	2005	2006	2008	2009	2010
austria	X										X	X												X
Brazil															X							X	X	
Canada																						X		
Japan			X												X					X		X	X	X
UK	X							X		X													X	
Sweden						X							X				X	X						
India																								
Ireland		X				X										X			X					X
Colombia		X			X							X				X								
USA		X										X								X	X	X	X	
Uruguay		X	X					X	X	X														
Spain						X	X			X			X											X
Netherlands										X	X				X				X					
X - means significance.																								

X - means significance.

Table 8 presents the results for the expanded data set. The results suggest a smaller number of points of instability, particularly in terms of the events after the 2008 crisis. This suggests that the expanded data set is the most appropriate for explaining the data of several countries. Further, the results may also suggest that the spread of events that followed the 2008 crisis follow a traditional mechanism.

Table 8: Summary of identified points of instability in the estimated models-expanded by misalignment and remaining variables

Points of Instability in the global model extended by misalignment and other variables																							
Country i - Dummy	1973	1974	1975	1976	1977	1979	1980	1981	1983	1984	1985	1986	1993	1994	1999	2000	2001	2002	2003	2005	2008	2009	2010
Austria	X																						
Brazil															X						X	X	
Canada						X																	
Japan															X								X
UK								X															
Sweden						X							X				X	X					
India		X		X																			
Ireland						X										X							
Colombia		X																					
USA		X										X									X		
Uruguay		X	X					X	X														
Spain					X																		
Netherlands											X			X	X								
X - means significance.																							

X - means significance.



6.3. Main findings

The following is a summary of the main conclusions of this study:

- a) There is some evidence of interdependence among the analyzed countries in terms of exchange rate. The dynamics of the US data appear to affect a large number of countries. This is in line with the importance of this country in the global economy;
- b) There is evidence of strong interdependence among European nations;
- c) A shock can only affect the exchange rates of a specific country if it permanently affects its fundamentals;
- d) The mechanism of interdependence is restricted mainly to short-term dynamics;
- e) Since the error correction mechanisms of a particular country are not important in other countries, there is no reason for not employing Gonzalo and Granger's decomposition methodology, which is a usual method for estimating exchange rate misalignment.

7. COMPARISON WITH THE RESULTS OF EXISTING LITERATURE

Dufrenot and Mathieu (2006) do not use non-linear models for analyzing exchange rate misalignment, but the data set is restricted to data from the countries themselves. Pesaran and Smith (2006) and Pesaran and Smith (2005) attempt to model the phenomena using a global model, although they do not apply the methodology to exchange rate misalignment. The authors seek to identify global factors and assess the extent to which such factors help to improve individual models.

Camarero and Ordóñez (2002) estimate exchange rate misalignment using different econometric methodologies:

- a) multivariate cointegration of Johansen (1995);
- b) panel cointegration techniques Pesaran and Shin (1999).

They obtain similar results in the two methods by analyzing data pertaining to the Euro and dollar.

Further, Groen and Kleibergen (2003) conduct a similar analysis to that of Larsson and Lyhagen (2007) and are able to detect interdependence between the dollar and other currencies, as expounded in the monetary approach (Dornbusch (1976) and Mussa (1974) The hypothesis suggested by the monetary approach is not rejected in models in which panel data is used, which is in contrast with models that do not use panel data. In this regard, the results of the abovementioned studies are very close to those of the present study; the only difference lies in the fundamentals used.

Groen and Lombardelli (2004) is also similar to the current study, but their work is limited to model the relationship between the real exchange rate and the Balassa-Samuelson effect without considering the possible effect of the international investment position on the exchange rate.

8. CONCLUSIONS

The objective of this study was to compare different methodologies for determining exchange rate misalignment and test the hypotheses that the exchange rates of different countries are influenced by their own fundamentals as well variables related to the fundamentals of other countries. The second hypothesis is related to the existence of interdependence among countries.

The methodology applied in this study differs from the traditional methodology of calculating exchange rate misalignment using models containing only variables of the particular country under study

in that the methodology used in this study permits the expansion of the set of variables to include information related to other countries.

I employed two different empirical strategies for the purpose of such a comparison. The first one was based on evaluating whether a multivariate time series, which is usually used in the literature, containing exchange data pertaining only to the country in question can be improved by adding variables related to other countries. The second strategy was to utilize a long panel data set that includes variables used to estimate exchange rate misalignment and formally test the hypothesis of no interdependence.

The results suggested that in both strategies there is some evidence of the existence of interdependence. The source of interdependence appears to be most related to short-term factors. However, it must be noted that the generation of long-term effects of the competitive devaluations in one country on the exchange rate of other countries would imply changing these fundamentals.

BIBLIOGRAPHY

- Ahmad, Y. & Craighead, W. D. (2010). Temporal Aggregation and Purchasing Power Parity Persistence. *Journal of International Money and Finance*, 30(5):817–830.
- Alberola, E. & Cervero, S. (1999). Global Equilibrium exchange rate: Euro, Dolar, “Ins”, “Outs” and other major currencies in a Panel Cointegration Framework. Working Paper 99–175, Washington, IMF.
- Beveridge, S. & Nelson, D. B. (1981). A new approach to decomposition of economic time series into permanent and transitory components with particular attention to measurement of the business cycle. *Journal of Monetary Economics*, 7(2):151–174.
- Bilson, J. F. (1979). Recent developments in monetary models of exchange rate determination. *IMF Staff Paper*, 26(2):201–223.
- Camarero, M. & Ordóñez, J. (2002). The euro-dollar exchange rate: is it fundamental? Working paper 798, Venice, CESIFO.
- Cavaliere, G. & Rahbek, A. (2008). Testing for Co-Integration in Vector Autoregressions with Non-Stationary Volatility. *Journal of Econometrics*, 158(10):23.
- Clark, P. B. & MacDonald, R. (1998). Exchange rates and economic fundamentals: a methodological comparison of BEERs and FEERs. International Monetary Fund.
- Cline, W. R. (2008). Estimating consistent fundamental equilibrium exchange rate. Working paper Series 1–26, Washington, Peterson Institute for International Economics.
- Dornbusch, R. (1976). Expectations and Exchange Rate Dynamics. *Journal of Political Economy*, 84(6):1161–1176.
- Dufrenot, G. & Mathieu, L. (2006). Persistent misalignments of the European exchange rates: some evidence from non-linear cointegration. *Applied Economics*, 38:203–229.
- Edwards, S. (1987). Exchange rate misalignment in developing countries. Working paper 442, NBER.
- Eichengreen, B. & Sachs, J. D. (1986). Exchange rates and economic recovery in the 1930s. National Bureau of Economic Research Cambridge, Mass, USA.
- Engle, R. F. & Granger, C. W. J. (1987). Co-integration and Error Correction: Representation, Estimation and Testing. *Econometrica*, 55:251–276.



- Faruquee, H. (1995). Long-run determinants of the real exchange rate: A stock Flow Perspective. *IMF Staff Paper*, 42:80–107.
- Froot, K. A. & K. Rogoff, E. (1995). Perspectives on PPP and Long-run Real exchange rates. *Handbook of International Economics*. Amsterdam, North-Holland.
- Glick, R. & Rose, A. K. (1998). Contagion And Trade: Why Are Currency Crises Regional? Working Paper 6806, Washington, NBER.
- Gonzalo, J. & Granger, C. W. J. (1995). Estimation of Common Long-Memory Components in Cointegrated Systems. *Journal of business and Economics Statistics*, 13(1):27–35.
- Gredenhoff, M. & Jacobson, T. (2001). Bootstrap Testing Linear Restrictions on Cointegrating Vectors. *Journal of Business & Economic Statistics*, 19(1):63–72.
- Groen, J. & Lombardelli, C. (2004). Real exchange rates and the relative prices of non-traded and traded goods: an empirical analysis. Working Paper 223, Bank of England.
- Groen, J. J. J. & Kleibergen, F. (2003). Likelihood-Based Cointegration Analysis in Panels of Vector Error-Correction Models. *Journal of Business & Economic Statistics*, 21:295–318.
- Hansen, P. R. (2003). Structural changes in the cointegrated vector autoregressive model. *Journal of Econometrics*, 114:261–295.
- Hendry, D. & Krolzig, H. M. (2005). The Properties of Automatic Gets Modelling. *Economic Journal*, 115:32–61.
- Hendry, D. F. (1995). *Dynamic Econometrics*. Technical report, Oxford, Oxford University Press.
- Johansen, S. (1988). Statistical Analysis of cointegration vectors. *Journal of Economic Dynamics and Control*, 12(2):231–254.
- Johansen, S. (1995). Likelihood-based inference in cointegrated vector autoregressive models. Technical report, Oxford, Oxford University Press.
- Johansen, S. (1999). A Bartlett correction factor for tests on the cointegrating relations. Technical report, Florence, European University Institute.
- Johansen, S. (2000). A small sample correction of the test for cointegrating rank in the vector autoregressive model. Technical report, Florence, European University Institute.
- Juselius, K. (2009). *The Cointegrated VAR Model Methodology and Applications* Oxford. Technical report, Oxford University Press.
- Kubota, M. (2009). Real Exchange Rate Misalignments: Theoretical modelling and empirical evidence. Discussion papers in economics, York, University of York.
- Lane, P. R. & Milesi-Ferretti, G. M. (2007). The external wealth of nations mark II: Revised and extended estimates of foreign assets and liabilities, 1970-2004. *Journal of International Economics*, 73(2):223–250.
- Larsson, R. & Lyhagen, J. (2007). Inference in Panel Cointegration Models with long Panels. *Journal of Business and Economic Statistics*, 25(4):473–483.
- Meese, R. A. & Rogoff, K. (1983). Empirical Exchange models of the seventies: Do they fit out of the sample? *Journal of International Economics*, 14:3–24.

- Mussa, M. (1974). A Monetary Approach to Balance-of-Payments Analysis. *Journal Of Money, Credit And Banking*, 6(3):333–351.
- Mussa, M. (1976). The exchange rate, the balance of payments and monetary policy under a regime of controlled floating. *Scandinavian Journal Of Economics*, 78:228–248.
- Pesaran, M. H. & Shin, Y. (1999). Pooled Mean Group Estimation of Dynamic Heterogenous Panels. *Journal of American Statistical Association*, 94(446):621–634.
- Pesaran, M. H. & Smith, L. V. (2005). What if the UK had Joined the Euro in 1999? An Empirical Evaluation using a Global VAR. Technical report, Cambridge, University of Cambridge Faculty of Economics.
- Pesaran, M. H. & Smith, R. (2006). Macroeconomic modelling with a global perspective. Technical report, Cambridge, Faculty of Economics University of Cambridge.
- Santos, C. & Hendry, D. (2008). Automatic selection of indicators in a fully saturated regression. *Computational Statistics*, 23(2):317–335.
- Shin, Y. (1994). A residual-based test of the null of cointegration against the alternative of no cointegration. *Econometric Theory*, 10(01):91–115.
- Stein, J. (1995). The Fundamental Determinants of the Real Exchange Rate of the U.S. Dollar Relative to Other G-7 Currencies. Working Paper 95–81, IMF.
- Swensen, A. R. (2006). Bootstrap algorithms for testing and determining the cointegration rank in VAR models. *Econometrica*, 74(6):1699–1714.
- Taylor, A. M. (2001). Potential Pitfalls for the Purchasing-Power-Parity Puzzle? Sampling and Specification Biases in Mean-Reversion Tests of the Law of One Price. *Econometrica*, 69(2):473–498.
- Turnovsky, S. J. & Basar, T. (1987). Dynamic Strategic Monetary Policies and Coordination in Interdependent Economies. Working Paper 2467, Washington, NBER.
- Ugai, H. (2007). Effects of the quantitative easing policy: A survey of empirical analyses. *Monetary And Economic Studies-Bank Of Japan*, 25:1–47.
- Williamson, J. (1994). Estimating Equilibrium Exchange Rates. IIE, Washington D.C.



9. APPENDIX A:

In this Appendix, we present the algorithm of the generalized reduced rank regression used to estimate panel VECM parameters and a wild bootstrap procedure to test linear restrictions on panel VECM parameters.

Estimating a panel VECM using generalized reduced rank regression

Defining the following variables and arrays:

$$\begin{aligned} Z_{0t} &= \Delta X_t \\ Z_{1t} &= X_{t-1} \\ Z_{2t} &= [\Delta X_t \dots \Delta X_{t-k} D_t] \end{aligned} \quad (5)$$

$$C = (\Gamma_1, \dots, \Gamma_{k-1}, \Phi) \quad (6)$$

yields

$$Z_{0t} = AB'Z_{1t} + CZ_{2t} + \varepsilon_t \quad (7)$$

Hansen (2003) shows how to estimate the process given in (5). Some definitions are now necessary:

$$\text{vec}(B) = H\varphi + h, \quad (8)$$

where H is known, φ contains free parameters, h is a tool for normalization, and vec is an operator that transforms a matrix in a column vector.¹⁰

$$\text{vec}(A, C) = G\Psi \quad (9)$$

where G is known, and Ψ contains free parameters.

Using (6), (7), and (8) to (11), it is possible to estimate the model parameters (see Hansen (2003) for details):

$$\begin{aligned} \text{vec}(\hat{A}, \hat{C}) &= G \left[G' \sum_{t=1}^T \left[\left(\frac{\hat{B}' Z_{1t} Z_{1t}' \hat{B} \hat{B}' Z_{1t} Z_{2t}'}{Z_{2t} Z_{1t}' \hat{B} Z_{2t} Z_{2t}'} \right) \otimes \hat{\Omega}(t)^{-1} \right] G \right]^{-1} x \\ x G' \sum_{t=1}^T \text{vec}(\hat{\Omega}(t)^{-1} Z_{0t} (Z_{1t}' \hat{B}, Z_{2t}')) \end{aligned} \quad (10)$$

$$\begin{aligned} \text{vec}(\hat{B}) &= H \left[H' \sum_{t=1}^T (\hat{A}' \hat{\Omega}(t)^{-1} \hat{A} \otimes Z_{1t} Z_{1t}') H \right]^{-1} x \\ x H' \left[H' \sum_{t=1}^T \text{vec}(Z_{1t} (Z_{0t} - \hat{C} Z_{2t})' \hat{\Omega}(t)^{-1} \hat{A}) - \sum_{t=1}^T \text{vec}(\hat{A}' \hat{\Omega}(t)^{-1} \hat{A} \otimes Z_{1t} Z_{1t}') h \right] + h \end{aligned} \quad (11)$$

¹⁰Further details regarding the properties of such an operator can be found in a more advanced book of econometrics. For example Hendry (1995). Dynamic Econometrics. Oxford, Oxford University Press., or Lutkepohl, H. (2007). New introduction to multiple time series analysis, Springer.

$$\Omega = (T)^{-1} \sum_{t=1}^T \hat{\varepsilon}_t \hat{\varepsilon}_t' \quad (12)$$

$$\hat{\varepsilon}_t = Z_{ot} - \hat{A}\hat{B}'Z_{1t} + \hat{C}Z_{2t}. \quad (13)$$

After convergence is attained according to some established criterion, the likelihood function can be calculated. Based on the values of the restricted and unrestricted general model, it is possible to calculate a likelihood ratio test with an asymptotic chi-squared distribution Larsson and Lyhagen (2007).

Testing linear restrictions on panel VECM parameters

The likelihood ratio procedure described in the previous section suffers from a serious distortion. This encourages analysts to use a bootstrap technique to obtain appropriate critical values. Studies that discuss similar problems are Johansen (1999), Johansen et alii (2000), Gredenhoff and Jacobson (2001), and Swensen (2006). We chose to perform a bootstrap to simulate the appropriate distribution for the test statistic. The results suggest that the size distortion is actually severe.

The procedure used is a wild bootstrap similar to that used in Cavaliere and Rahbek (2008):¹¹

Step 1: Estimate the model given by equation (3) under the null and alternative hypothesis and calculate the statistic likelihood ratio.

Step 2: Save the model parameters estimated under the null hypothesis and check the stability of the system. If the system is stable, proceed to Step 3. If not, abort the procedure.

Step 3: Construct the errors for generating artificial data using equation (12) below from the parameters estimated with the model constraints on the null hypothesis:

$$\hat{u}_t \equiv \hat{\varepsilon}_t w_t, \quad (14)$$

where $\hat{\varepsilon}_t$ is defined as in (11) and $w_t \sim NI(0,1)$.

Step 4: Generate the Q series with the sample size T from equation (1) with the parameter structure of the null hypothesis.

Step 5: Perform the likelihood ratio test described in Step 1 for all artificially generated samples. The values of the test statistics are retained. Calculate the p-value by dividing the total number of times the value of the likelihood ratio statistic obtained with the real sample data is greater than value of the likelihood ratio statistic obtained with artificial data divided by the number of runs.

¹¹The procedure was fully implemented in Matlab software. Further details can be obtained from the author.



10. APPENDIX B: COMPLEMENTARY TABLES

Table 9: Final model selected using the Ox algorithm: Canada.

Canada								
Model 1								
		Cla_spa_1	I:1973	I:1979	I:1986	I:2008	Constant	Cla_canada_1
DLRER_t	Coefficient	-0,0405	-0,0635	0,3249	-0,1633	-0,0680	-6,6033	0,0386
	Standard Error	0,0112	0,0503	0,0492	0,0504	0,0489	1,6940	0,0105
	t	-3,6200	-1,2600	6,6000	-3,2400	-1,3900	-3,9000	3,6700
	p-value	0,10%	21,61%	0,00%	0,28%	17,37%	0,05%	0,09%
DLNFA_t	Coefficient	0,011	0,053	-0,035	0,001	0,069	1,259	0,000
	Standard Error	0,008	0,036	0,036	0,036	0,035	1,224	0,008
	t	1,390	1,470	-0,993	0,036	1,960	1,030	-0,025
	p-value	17,44%	15,13%	32,82%	97,15%	5,86%	31,12%	98,02%
DLBS_t	Coefficient	0,021	-0,041	0,003	0,050	-0,012	3,522	-0,021
	Standard Error	0,006	0,025	0,025	0,025	0,024	0,843	0,005
	t	3,830	-1,630	0,111	2,010	-0,482	4,180	-4,020
	p-value	0,06%	11,31%	91,25%	5,29%	63,27%	0,02%	0,03%
Model 2								
		Cla_spa_1	DLRER_aus_1	DLRER_aut_1	DLRER_uni_1	I:1979	Constant	Cla_canada_1
DLRER_t	Coefficient	-0,0258	0,1455	0,2499	-0,0034	0,3392	-4,5778	0,0312
	Standard Error	0,0105	0,3660	0,3413	0,1256	0,0514	1,6300	0,0106
	t	-2,4500	0,3980	0,7320	-0,0269	6,6000	-2,8100	2,9300
	p-value	1,99%	69,36%	46,93%	97,87%	0,00%	0,84%	0,61%
DLNFA_t	Coefficient	0,005	-0,156	0,238	-0,240	-0,025	0,309	0,004
	Standard Error	0,007	0,234	0,218	0,080	0,033	1,043	0,007
	t	0,685	-0,668	1,090	-2,990	-0,766	0,297	0,569
	p-value	49,85%	50,90%	28,49%	0,54%	44,91%	76,87%	57,35%
DLBS_t	Coefficient	0,022	-0,231	0,053	0,074	0,002	3,582	-0,021
	Standard Error	0,005	0,170	0,158	0,058	0,024	0,756	0,005
	t	4,470	-1,360	0,335	1,270	0,074	4,740	-4,280
	p-value	0,01%	18,28%	74,00%	21,37%	94,15%	0,00%	0,02%

Note: Cla denotes the cointegration vector

Table 10: Final model selected using the Ox algorithm: Austria.

Austria								
Model 1								
		Cla_india_1	Cla_urug_1	I:1973	I:1985	I:1986	I:2009	Constant
DLRER_t	Coefficient	-0,0268	-0,0081	0,0712	-0,2013	-0,1515	-0,0859	1,8263
	Standard Error	0,0088	0,0115	0,0517	0,0582	0,0601	0,0535	0,8268
	t	-3,0600	-0,7050	1,3800	-3,4600	-2,5200	-1,6100	2,2100
	p-value	0,46%	48,59%	17,84%	0,16%	1,71%	11,85%	3,47%
DLNFA_t	Coefficient	0,007	-0,002	0,005	-0,033	-0,002	0,087	-1,647
	Standard Error	0,006	0,008	0,035	0,039	0,041	0,036	0,558
	t	1,130	-0,281	0,157	-0,843	-0,057	2,400	-2,950
	p-value	26,67%	78,04%	87,64%	40,54%	95,46%	2,25%	0,60%
DLBS_t	Coefficient	-0,003	0,006	-0,071	0,016	-0,015	-0,047	0,537
	Standard Error	0,002	0,002	0,011	0,012	0,012	0,011	0,168
	t	-1,970	2,350	-6,740	1,350	-1,190	-4,330	3,190
	p-value	5,83%	2,52%	0,00%	18,54%	24,18%	0,01%	0,32%
Model 2								
		DLpreco_relativo_ger_1	DLRER_net_1	Dnfa_spa_1	I:1973	Constant U	Cla_austria_1	
DLRER_t	Coefficient	-0,7996	0,8636	-0,4571	0,0417	-0,3130	-0,0042	
	Standard Error	0,6591	0,2869	0,1603	0,0584	0,7003	0,0095	
	t	-1,2100	3,0100	-2,8500	0,7140	-0,4470	-0,4370	
	p-value	23,37%	0,50%	0,75%	48,02%	65,79%	66,46%	
DLNFA_t	Coefficient	0,3231	0,2349	-0,0607	0,0124	-1,1899	-0,0160	
	Standard Error	0,4181	0,1820	0,1017	0,0370	0,4442	0,0060	
	t	0,7730	1,2900	-0,5970	0,3350	-2,6800	-2,6600	
	p-value	44,52%	20,57%	55,45%	74,00%	1,14%	1,19%	
DLBS_t	Coefficient	0,273	0,060	-0,121	-0,059	0,240	0,003	
	Standard Error	0,139	0,061	0,034	0,012	0,148	0,002	
	t	1,960	0,987	-3,570	-4,820	1,620	1,680	
	p-value	5,82%	33,06%	0,11%	0,00%	11,38%	10,27%	

Note: Cla denotes the cointegration vector



Table 11: Final model selected using the Ox algorithm: Colombia.

Colombia									
Model 1									
		Dnfa_col_1	Cla_uni_1	I:1974	I:1977	I:1986	I:2001	Constante	Cla_col_1
DLRER_t	Coefficient	-0,613	0,013	0,089	0,191	-0,189	-0,009	1,321	-0,042
	Standard Error	0,249	0,009	0,058	0,059	0,062	0,061	0,246	0,008
	t	-2,460	1,430	1,540	3,250	-3,070	-0,152	5,370	-5,400
	p-value	0,020	0,163	0,135	0,003	0,004	0,880	0,000	0,000
DLNFA_t	Coefficient	0,778	0,011	-0,013	-0,035	0,026	-0,109	-0,423	0,013
	Standard Error	0,176	0,007	0,041	0,042	0,044	0,043	0,174	0,005
	t	4,410	1,640	-0,313	-0,841	0,606	-2,510	-2,430	2,450
	p-value	0,000	0,111	0,756	0,407	0,549	0,018	0,021	0,020
DLBS_t	Coefficient	0,024	-0,011	0,069	-0,036	0,018	0,039	-0,379	0,012
	Standard Error	0,103	0,004	0,024	0,024	0,026	0,025	0,102	0,003
	t	0,228	-2,880	2,860	-1,480	0,722	1,520	-3,710	3,710
	p-value	82,08%	0,72%	0,75%	14,77%	47,56%	13,86%	0,08%	0,08%
Model 2									
		Dnfa_col_1	DLpreco_relativo_jap_1	Dnfa_ire_1	Dnfa_net_1	Dnfa_spa_1	I:1974	Constante	Cla_col_1
DLRER_t	Coefficient	-0,2543	-0,2278	0,0522	0,1370	-0,3956	0,0687	1,4490	-0,0461
	Standard Error	0,2792	0,2027	0,0682	0,1195	0,2023	0,0702	0,2853	0,0089
	t	-0,9110	-1,1200	0,7650	1,1500	-1,9600	0,9780	5,0800	-5,1800
	p-value	36,95%	26,96%	44,99%	26,04%	5,96%	33,56%	0,00%	0,00%
DLNFA_t	Coefficient	0,595	-0,276	0,133	0,164	-0,131	-0,025	-0,250	0,008
	Standard Error	0,126	0,092	0,031	0,054	0,092	0,032	0,129	0,004
	t	4,710	-3,010	4,290	3,030	-1,440	-0,782	-1,940	1,900
	p-value	0,00%	0,52%	0,02%	0,49%	16,11%	44,02%	6,21%	6,70%
DLBS_t	Coefficient	-0,089	-0,066	-0,057	-0,061	0,266	0,066	-0,420	0,013
	Standard Error	0,084	0,061	0,021	0,036	0,061	0,021	0,086	0,003
	t	-1,060	-1,080	-2,740	-1,690	4,360	3,110	-4,880	4,940
	p-value	29,88%	28,79%	1,00%	10,09%	0,01%	0,40%	0,00%	0,00%

Note: Cla denotes the cointegration vector

Table 12: Final model selected using the Ox algorithm: United Kingdom.

United Kindgom							
		I:1973	I:1981	I:1984	I:2009	Constant	Cla_uni_1
DLRER_t	Coefficient	-0,105	0,011	0,197	-0,012	0,006	0,005
	Standard Error	0,058	0,058	0,059	0,058	0,010	0,009
	t	-1,810	0,187	3,310	-0,205	0,640	0,597
	p-value	7,96%	85,29%	0,22%	83,85%	52,69%	55,44%
DLNFA_t	Coefficient	-0,008	0,061	0,067	-0,177	-0,004	0,020
	Standard Error	0,060	0,060	0,061	0,059	0,010	0,009
	t	-0,142	1,030	1,100	-2,980	-0,415	2,120
	p-value	88,82%	31,04%	28,10%	0,54%	68,11%	4,16%
DLBS_t	Coefficient	-0,078	0,147	0,013	0,070	0,002	-0,008
	Standard Error	0,030	0,030	0,030	0,030	0,005	0,005
	t	-2,610	4,950	0,412	2,370	0,419	-1,680
	p-value	1,36%	0,00%	68,27%	2,40%	67,81%	10,16%
Model 2							
		DLRER_por_1	Dnfa_bra_1	Dnfa_den_1	I:1981	Constant	Cla_uni_1
DLRER_t	Coefficient	-0,5112	-0,2089	-0,1362	0,0032	0,0097	0,0025
	Standard Error	0,2184	0,1514	0,2044	0,0624	0,0100	0,0099
	t	-2,3400	-1,3800	-0,6660	0,0506	0,9750	0,2540
	p-value	2,54%	17,68%	50,97%	95,99%	33,67%	80,14%
DLNFA_t	Coefficient	0,093	-0,645	-0,548	0,025	-0,007	0,015
	Standard Error	0,154	0,106	0,144	0,044	0,007	0,007
	t	0,608	-6,060	-3,810	0,573	-1,000	2,130
	p-value	54,77%	0,00%	0,06%	57,05%	32,38%	4,08%
DLBS_t	Coefficient	0,361	0,044	-0,158	0,143	0,002	-0,006
	Standard Error	0,101	0,070	0,095	0,029	0,005	0,005
	t	3,570	0,626	-1,670	4,940	0,485	-1,230
	p-value	0,11%	53,54%	10,45%	0,00%	63,07%	22,66%

Note: Cla denotes the cointegration vector



Table 13: Final model selected using the Ox algorithm: India

India								
Model 1								
		Dnfa_ind_1	I:1974	I:1976	I:1993	I:2009	Constant U	Cla_india_1 U
DLRER_t	Coefficient	0,726	0,135	-0,117	0,124	-0,074	0,539	-0,0284
	Standard Error	0,435	0,057	0,056	0,061	0,059	0,174	0,0087
	t	1,670	2,380	-2,080	2,030	-1,250	3,100	-3,2700
	p-value	10,49%	2,35%	4,60%	5,11%	22,18%	0,40%	0,26%
DLNFA_t	Coefficient	0,720	-0,006	0,027	0,063	-0,009	-0,011	0,000
	Standard Error	0,132	0,017	0,017	0,019	0,018	0,053	0,003
	t	5,460	-0,367	1,580	3,380	-0,473	-0,216	0,186
	p-value	0,00%	71,58%	12,47%	0,19%	63,97%	83,07%	85,34%
DLBS_t	Coefficient	-0,041	-0,074	-0,020	-0,021	0,056	-0,162	0,008
	Standard Error	0,109	0,014	0,014	0,015	0,015	0,043	0,002
	t	-0,374	-5,240	-1,410	-1,360	3,790	-3,720	3,880
	p-value	71,08%	0,00%	16,88%	18,31%	0,06%	0,08%	0,05%
Model 2								
		Dnfa_ind_1	DLpreco_relativo_usa	Dnfa_usa_1	I:1974	I:1976	Constant	Cla_india_1
DLRER_t	Coefficient	0,6724	-0,2451	0,4215	0,1345	-0,1429	0,3377	-0,0181
	Standard Error	0,4217	0,1161	0,2970	0,0560	0,0559	0,1555	0,0078
	t	1,5900	-2,1100	1,4200	2,4000	-2,5600	2,1700	-2,3300
	p-value	12,06%	4,27%	16,56%	2,23%	1,55%	3,74%	2,61%
DLNFA_t	Coefficient	0,673	-0,049	-0,031	-0,010	0,022	-0,070	0,004
	Standard Error	0,149	0,041	0,105	0,020	0,020	0,055	0,003
	t	4,530	-1,210	-0,297	-0,492	1,120	-1,290	1,290
	p-value	0,01%	23,63%	76,86%	62,59%	26,91%	20,78%	20,78%
DLBS_t	Coefficient	-0,040	0,071	-0,355	-0,078	-0,008	-0,074	0,004
	Standard Error	0,082	0,022	0,057	0,011	0,011	0,030	0,002
	t	-0,486	3,160	-6,180	-7,250	-0,710	-2,460	2,640
	p-value	63,03%	0,34%	0,00%	0,00%	48,27%	1,96%	1,27%

Note: Cla denotes the cointegration vector

Table 14: Final model selected using the Ox algorithm: Ireland.

Ireland									
Model 1									
		Cla_india_1	I:1974	I:1979	I:2000	I:2003	I:2009	Constant	Cla_ire_1
DLRER_t	Coefficient	-0,011	-0,006	0,170	-0,053	0,084	-0,067	-0,0160	-0,00377426
	Standard Error	0,006	0,038	0,038	0,038	0,038	0,041	0,4208	0,007132
	t	-1,860	-0,145	4,450	-1,400	2,200	-1,640	-0,0381	-0,529
	p-value	7,31%	88,55%	0,01%	17,17%	3,57%	11,17%	96,99%	0,6004
DLNFA_t	Coefficient	0,013	-0,022	-0,019	-0,584	-0,007	-0,197	-3,921	-0,056
	Standard Error	0,021	0,131	0,131	0,130	0,130	0,139	1,436	0,024
	t	0,613	-0,169	-0,143	-4,500	-0,050	-1,420	-2,730	-2,290
	p-value	54,44%	86,70%	88,74%	0,01%	96,02%	16,57%	1,03%	2,92%
DLBS_t	Coefficient	0,019	0,027	0,261	-0,019	-0,004	0,015	0,761	0,017
	Standard Error	0,006	0,137	0,082	0,105	0,035	0,044	0,364	0,006
	t	3,370	0,199	3,170	-0,180	-0,110	0,337	2,090	2,790
	p-value	0,20%	84,39%	0,34%	85,85%	91,34%	73,87%	4,49%	0,90%
Model 2									
		Cla_india_1	Dnfa_den_1	Dnfa_gre_1	Dnfa_uni_1	I:1979	I:2000	Constant	Cla_ire_1
DLRER_t	Coefficient	-0,010	-0,355	-0,154	0,117	0,168	-0,020	-0,3820	-0,00872603
	Standard Error	0,006	0,136	0,081	0,104	0,034	0,043	0,3596	0,006191
	t	-1,690	-2,620	-1,890	1,130	4,880	-0,456	-1,0600	-1,41
	p-value	10,09%	1,36%	6,78%	26,66%	0,00%	65,15%	29,63%	0,1686
DLNFA_t	Coefficient	0,022	0,471	-0,042	-0,861	0,034	-0,566	-4,107	-0,056
	Standard Error	0,019	0,450	0,269	0,344	0,114	0,143	1,193	0,021
	t	1,160	1,050	-0,157	-2,500	0,297	-3,950	-3,440	-2,710
	p-value	25,69%	30,28%	87,66%	1,80%	76,82%	0,04%	0,17%	1,10%
DLBS_t	Coefficient	0,028	-0,109	0,007	0,007	-0,085	0,131	0,440	0,015
	Standard Error	0,004	0,023	0,023	0,023	0,023	0,025	0,256	0,004
	t	7,390	-4,650	0,290	0,318	-3,650	5,290	1,720	3,520
	p-value	0,00%	0,01%	77,41%	75,26%	0,10%	0,00%	9,61%	0,14%

Note: Cla denotes the cointegration vector



Table 15: Final model selected using the Ox algorithm: Japan.

Japan								
Model 1								
		Cla_swe_1	I:1975	I:1999	I:2005	I:2010	Constant	Cla_jap_1
DLRER_t	Coefficient	0,0366	-0,0370	0,1298	-0,0916	-0,0024	2,7562	-0,0093
	Standard Error	0,0151	0,0938	0,0933	0,0946	0,1027	1,1780	0,0153
	t	2,4200	-0,3940	1,3900	-0,9690	-0,0237	2,3400	-0,6050
	p-value	2,14%	69,63%	17,38%	33,98%	98,13%	2,57%	54,92%
DLNFA_t	Coefficient	-0,014	0,045	-0,153	-0,088	-0,083	-0,806	-0,012
	Standard Error	0,004	0,024	0,024	0,024	0,027	0,305	0,004
	t	-3,630	1,860	-6,350	-3,620	-3,120	-2,650	-3,020
	p-value	0,10%	7,28%	0,00%	0,10%	0,38%	1,25%	0,49%
DLBS_t	Coefficient	-0,021	-0,119	-0,087	0,056	0,253	-1,707	0,013
	Standard Error	0,008	0,052	0,051	0,052	0,057	0,649	0,008
	t	-2,510	-2,310	-1,700	1,080	4,470	-2,630	1,520
	p-value	1,74%	2,77%	9,95%	28,72%	0,01%	1,30%	13,78%
Model 2								
		DLpreco_relativo_1	I:1999	I:2010	Constant	Cla_jap_1		
DLRER_t	Coefficient	-0,1842	0,0926	0,0111	-0,0954	0,0073		
	Standard Error	0,2031	0,0966	0,1069	0,2191	0,0143		
	t	-0,9070	0,9580	0,1030	-0,4350	0,5070		
	p-value	37,07%	34,48%	91,82%	66,60%	61,57%		
DLNFA_t	Coefficient	0,033	-0,133	-0,068	0,207	-0,012		
	Standard Error	0,068	0,032	0,036	0,073	0,005		
	t	0,480	-4,130	-1,920	2,830	-2,610		
	p-value	63,40%	0,02%	6,34%	0,77%	1,32%		
DLBS_t	Coefficient	-0,158	-0,067	0,227	-0,043	0,001		
	Standard Error	0,118	0,056	0,062	0,128	0,008		
	t	-1,330	-1,190	3,640	-0,333	0,103		
	p-value	19,19%	24,20%	0,09%	74,13%	91,82%		

Note: Cla denotes the cointegration vector

Table 16: Final model selected using the Ox algorithm: Netherlands.

Netherlands									
Model 1									
		DLpreco_relativo_net_1	Cla_swe_1	I:1984	I:1985	I:1999	I:2003	Constant U	Cla_net_1
DLRER_t	Coefficient	0,2716	0,0125	-0,0353	-0,0017	-0,0019	0,0623	0,9294	0,0102378
	Standard Error	0,1346	0,0050	0,0263	0,0314	0,0271	0,0268	0,3468	0,004584
	t	2,0200	2,4900	-1,3400	-0,0529	-0,0714	2,3200	2,6800	2,23
	p-value	5,23%	1,83%	18,93%	95,82%	94,35%	2,68%	1,17%	0,0329
DLNFA_t	Coefficient	-0,009	0,016	0,033	0,040	0,338	0,278	1,041	-0,020
	Standard Error	0,403	0,015	0,079	0,094	0,081	0,080	1,039	0,014
	t	-0,022	1,070	0,413	0,429	4,170	3,460	1,000	-1,480
	p-value	98,30%	29,40%	68,21%	67,06%	0,02%	0,16%	32,43%	14,78%
DLBS_t	Coefficient	0,236	-0,011	-0,108	0,167	-0,001	-0,003	-0,763	-0,002
	Standard Error	0,110	0,004	0,022	0,026	0,022	0,022	0,284	0,004
	t	2,140	-2,590	-5,020	6,510	-0,026	-0,124	-2,680	-0,525
	p-value	4,02%	1,44%	0,00%	0,00%	97,98%	90,23%	1,16%	60,30%
Model 2									
		DLRER_por_1	Dnfa_au_1	Dnfa_swe_1	I:1985	I:1994	I:1999	Constant U	Cla_net_1
DLRER_t	Coefficient	-0,0537	-0,1940	-0,0863	-0,0546	-0,0184	-0,0124	0,0735	0,0169529
	Standard Error	0,1103	0,1439	0,0497	0,0315	0,0311	0,0309	0,0204	0,004875
	t	-0,4870	-1,3500	-1,7400	-1,7300	-0,5920	-0,4020	3,6100	3,48
	p-value	62,97%	18,72%	9,23%	9,33%	55,85%	69,04%	0,11%	0,0015
DLNFA_t	Coefficient	0,152	1,029	-0,373	0,097	-0,310	0,418	-0,046	-0,0103384
	Standard Error	0,172	0,225	0,078	0,049	0,049	0,048	0,032	0,007611
	t	0,884	4,580	-4,810	1,970	-6,380	8,650	-1,440	-1,36
	p-value	38,35%	0,01%	0,00%	5,78%	0,00%	0,00%	15,91%	0,1841
DLBS_t	Coefficient	0,368	-0,391	-0,035	0,123	-0,006	0,005	-0,053	-0,0129943
	Standard Error	0,093	0,121	0,042	0,027	0,026	0,026	0,017	0,004114
	t	3,960	-3,220	-0,841	4,610	-0,245	0,206	-3,100	-3,16
	p-value	0,04%	0,30%	40,70%	0,01%	80,77%	83,78%	0,41%	0,35%

Note: Cla denotes the cointegration vector



Table 17: Final model selected using the Ox algorithm: Spain.

Spain									
Model 1									
		Cla_eua_1	I:1977	I:1979	I:1983	I:1993	I:2009	Constant U	Cla_spa_1
DLRER_t	Coefficient	-0,0139	-0,0089	0,1505	-0,1610	-0,1093	0,0086	0,1118	-0,00113962
	Standard Error	0,0043	0,0272	0,0282	0,0276	0,0277	0,0281	0,5760	0,004859
	t	-3,2400	-0,3270	5,3400	-5,8300	-3,9500	0,3070	0,1940	-0,235
	p-value	0,29%	74,58%	0,00%	0,00%	0,04%	76,10%	84,73%	0,8161
DLNFA_t	Coefficient	0,003	0,000	0,024	-0,024	-0,011	-0,190	-2,280	-0,020
	Standard Error	0,008	0,051	0,053	0,052	0,052	0,053	1,086	0,009
	t	0,330	0,005	0,460	-0,454	-0,215	-3,600	-2,100	-2,180
	p-value	74,36%	99,61%	64,87%	65,28%	83,13%	0,11%	4,40%	3,68%
DLBS_t	Coefficient	-0,002	0,051	-0,008	0,030	-0,003	0,033	0,813	0,007
	Standard Error	0,002	0,015	0,015	0,015	0,015	0,015	0,307	0,003
	t	-0,782	3,530	-0,499	2,040	-0,180	2,170	2,640	2,750
	p-value	44,01%	0,13%	62,12%	4,96%	85,86%	3,75%	1,27%	0,98%
Model 2									
		Cla_india_1	DLRER_usa_1	DLRERtur_1	Dnfa_bra_1	Dnfa_ger_1	I:1977	Constant U	Cla_spa_1
DLRER_t	Coefficient	-0,0015	-0,3205	0,0324	0,1248	0,2015	-0,0026	-0,5473	-0,00524426
	Standard Error	0,0069	0,1376	0,0738	0,1187	0,2263	0,0461	0,8640	0,007551
	t	-0,2150	-2,3300	0,4380	1,0500	0,8900	-0,0561	-0,6330	-0,695
	p-value	83,11%	2,65%	66,41%	30,12%	38,01%	95,57%	53,11%	0,4925
DLNFA_t	Coefficient	0,013	-0,015	0,053	-0,511	-0,415	0,015	-2,754	-0,0222619
	Standard Error	0,006	0,128	0,069	0,111	0,211	0,043	0,805	0,007031
	t	2,010	-0,117	0,768	-4,620	-1,970	0,355	-3,420	-3,17
	p-value	5,28%	90,74%	44,82%	0,01%	5,79%	72,51%	0,18%	0,0035
DLBS_t	Coefficient	-0,005	0,157	-0,074	0,048	-0,137	0,048	0,670	0,00516866
	Standard Error	0,001	0,028	0,015	0,025	0,047	0,010	0,179	0,00156
	t	-3,540	5,510	-4,850	1,960	-2,930	5,060	3,750	3,31
	p-value	0,13%	0,00%	0,00%	5,87%	0,63%	0,00%	0,07%	0,24%

Note: Cla denotes the cointegration vector

Table 18: Final model selected using the Ox algorithm: Sweden.

Sweden									
Model 1									
		Cla_austria_1	Dnfa_fra_1	Dnfa_uni_1	I:1979	I:1993	I:2002	Constant U	Cla_swe_1 U
DLRER_t	Coefficient	0,0188	-0,2298	-0,3052	0,2703	-0,1952	0,0715	1,1336	-0,0033922
	Standard Error	0,0068	0,1817	0,1138	0,0397	0,0409	0,0424	0,6502	0,006043
	t	2,7400	-1,2600	-2,6800	6,8100	-4,7800	1,6900	1,7400	-0,561
	p-value	1,01%	21,55%	1,16%	0,00%	0,00%	10,17%	9,12%	0,5786
DLNFA_t	Coefficient	-0,016	1,100	-0,055	-0,037	-0,188	-0,592	-1,902	-0,011
	Standard Error	0,010	0,274	0,171	0,060	0,062	0,064	0,979	0,009
	t	-1,520	4,020	-0,322	-0,620	-3,060	-9,280	-1,940	-1,180
	p-value	13,77%	0,03%	74,93%	53,95%	0,45%	0,00%	6,11%	24,69%
DLBS_t	Coefficient	0,001	0,114	0,164	0,001	0,044	-0,013	1,144	0,015
	Standard Error	0,007	0,183	0,115	0,040	0,041	0,043	0,656	0,006
	t	0,169	0,620	1,430	0,020	1,060	-0,295	1,740	2,420
	p-value	86,67%	53,99%	16,23%	98,40%	29,78%	76,97%	9,13%	2,15%
Model 2									
		DLpreco_relativo_swe_1	Cla_bra_1	I:1979	I:1993	I:2001	I:2002	Constant U	Cla_swe_1 U
DLRER_t	Coefficient	0,3016	-0,0109	0,2715	-0,1913	-0,0956	0,0172	-0,3214	-0,00443505
	Standard Error	0,1768	0,0075	0,0463	0,0449	0,0451	0,0453	0,5119	0,007206
	t	1,7100	-1,4600	5,8700	-4,2600	-2,1200	0,3800	-0,6280	-0,615
	p-value	9,80%	15,32%	0,00%	0,02%	4,21%	70,64%	53,48%	0,5427
DLNFA_t	Coefficient	-0,202	0,008	-0,031	-0,170	0,211	-0,477	-0,725	-0,0103196
	Standard Error	0,263	0,011	0,069	0,067	0,067	0,067	0,762	0,01073
	t	-0,769	0,679	-0,446	-2,540	3,140	-7,070	-0,951	-0,962
	p-value	44,75%	50,24%	65,87%	1,63%	0,37%	0,00%	34,89%	0,3435
DLBS_t	Coefficient	0,281	-0,008	0,022	0,049	0,010	-0,015	1,365	0,0191575
	Standard Error	0,151	0,006	0,040	0,038	0,039	0,039	0,438	0,006161
	t	1,860	-1,310	0,547	1,290	0,267	-0,383	3,120	3,11
	p-value	7,27%	20,13%	58,80%	20,81%	79,13%	70,43%	0,39%	0,40%

Note: Cla denotes the cointegration vector



Table 19: Final model selected using the Ox algorithm: Uruguay.

Uruguay									
Model 1									
		Cla_eua_1	I:1974	I:1975	I:1980	I:1981	I:1983	Constant U	Cla_urru_1 U
DLRER_t	Coefficient	0,0514	0,7462	-0,6785	0,1710	0,0575	-0,4282	-0,8644	-0,0155401
	Standard Error	0,0120	0,0793	0,0756	0,0775	0,0767	0,0766	0,2130	0,01362
	t	4,2900	9,4100	-8,9700	2,2100	0,7500	-5,5900	-4,0600	-1,14
	p-value	0,02%	0,00%	0,00%	3,48%	45,86%	0,00%	0,03%	0,2625
DLNFA_t	Coefficient	-0,004	-0,096	-0,055	-0,012	-0,040	-0,263	0,051	-0,012
	Standard Error	0,007	0,044	0,042	0,043	0,042	0,042	0,117	0,007
	t	-0,638	-2,200	-1,330	-0,293	-0,937	-6,250	0,433	-1,640
	p-value	52,80%	3,50%	19,38%	77,18%	35,59%	0,00%	66,82%	11,15%
DLBS_t	Coefficient	0,004	-0,020	0,006	0,035	0,096	-0,020	-0,079	-0,006
	Standard Error	0,003	0,018	0,017	0,017	0,017	0,017	0,047	0,003
	t	1,600	-1,130	0,337	2,040	5,650	-1,150	-1,660	-2,070
	p-value	11,96%	26,56%	73,81%	5,03%	0,00%	26,04%	10,76%	4,64%
Model 2									
		Cla_eua_1	Dnfa_usa_1	I:1974	I:1975	I:1981	I:1983	Constant U	Cla_urru_1 U
DLRER_t	Coefficient	0,0566	0,5488	0,7580	-0,6762	0,0396	-0,4295	-0,9442	-0,0164788
	Standard Error	0,0123	0,4159	0,0833	0,0794	0,0802	0,0802	0,2194	0,01426
	t	4,6100	1,3200	9,1000	-8,5200	0,4930	-5,3600	-4,3000	-1,16
	p-value	0,01%	19,66%	0,00%	0,00%	62,53%	0,00%	0,02%	0,2566
DLNFA_t	Coefficient	-0,005	0,325	-0,090	-0,049	-0,043	-0,265	0,059	-0,012
	Standard Error	0,006	0,211	0,042	0,040	0,041	0,041	0,111	0,007
	t	-0,752	1,540	-2,130	-1,230	-1,050	-6,520	0,533	-1,730
	p-value	45,78%	13,35%	4,10%	22,96%	30,22%	0,00%	59,76%	9,37%
DLBS_t	Coefficient	0,005	-0,293	-0,025	-0,001	0,098	-0,018	-0,098	-0,006
	Standard Error	0,002	0,078	0,016	0,015	0,015	0,015	0,041	0,003
	t	2,360	-3,740	-1,610	-0,040	6,480	-1,170	-2,380	-2,290
	p-value	2,46%	0,07%	11,80%	96,83%	0,00%	25,08%	2,37%	2,88%

Note: Cla denotes the cointegration vector

Table 20: Summary of the test results for cointegration using univariate and multivariate methodologies.

Test	Johansen		Shin		Modified Shin test		Engle and Granger	
	Number of cointegrated relationship - 5% level - Trace	Number of cointegrated relationship - 5% level - Maximum Eigenvalue	Is the null rejected at 5% level?	Is the null rejected at 1% level?	Is the null rejected at 5% level?	Is the null rejected at 1% level?	Is the null rejected at 5% level?	Is the null rejected at 1% level?
South Africa	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Germany	1	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Australia	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Austria	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Canada	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
South Korea	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Cingapura	0	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected
Colombia	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Denmark	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected
Spain	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
United States	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Finland	1	1	The null is rejected	The null is rejected	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
France	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected
Greece	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected
Netherlands	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected

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Table 20 – Continued

Test	Johansen		Shin		Modified Shin test		Engle and Granger	
	Number of cointegrated relationship - 5% level - Trace	Number of cointegrated relationship - 5% level - Maximum Eigenvalue	Is the null rejected at 5% level?	Is the null rejected at 1% level?	Is the null rejected at 5% level?	Is the null rejected at 1% level?	Is the null rejected at 5% level?	Is the null rejected at 1% level?
India	1	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Ireland	0	1	The null is rejected	The null is rejected	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted
Italy	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null is rejected	The null is rejected
Japan	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Mexico	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Portugal	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
United Kindgon	0	0	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Sweedden	0	0	The null is rejected	The null is rejected	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Uruguay	1	1	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted
Turkey	0	0	The null is rejected	The null is rejected	The null hypothesis of cointegration is accepted	The null hypothesis of cointegration is accepted	The null of no cointegration is accepted	The null of no cointegration is accepted

Bold values suggests evidence in favor of cointegration.