# Assignment

## Option Pricing and Risk - Financial Statistics

## Alexander J Ohrt

## $12~\mathrm{mai},~2022$

## Contents

Problem A - Euler-Maruyama Method	2
Problem B - Option Pricing I	5
Problem C - Option Pricing II	6

## Problem A - Euler-Maruyama Method

First we implement a discretization for the Black-Scholes-Merton (BSM) price dynamics with the parameter set  $(s_0, T, \mu, \sigma) = (12, 1, 0.03, 0.17)$ , with n = 250 steps. We generate  $M_1 = 10$ ,  $M_2 = 100$ ,  $M_3 = 1000$ ,  $M_4 = 10000$  and  $M_5 = 100000$  paths. The paths are plotted in separate figures for the three first cases.

The price dynamics are implemented with two different methods for calculating each path. The R function appended has the parameter version, which is used to select which version one wants to use to calculate each price path; onestep or EM. The first method implemented is the one step solution. Since the Wiener process has the distribution  $W_t \sim N(0,t), \forall t \in (0,t]$ , we can think of such a process as a cumulative summation of normally distributed random variables. Thus, on our uniform discrete grid

$$0 = t_0 < \ldots < t_n = T$$
,

we can simulate the following relation

$$X(t_0) = 0$$
,  $X(t_1) \sim N(0, t_1) = N(0, t_0 + h)$ ,...,  $X(t_n) \sim N(t_0 + nh) = X(t_1) + ... + X(t_{n-1})$ ,

where X represents the Brownian motion, n is the number of steps in the discretization and  $h = \frac{T}{n}$  is the step size. We insert this, replacing  $W_t$ , into the solution of the SDE

$$dS_t = \mu S_t dt + \sigma S_t dW_t,$$

which is

$$S_t = S_0 e^{\mu t} e^{\sigma W_t - \frac{\sigma^2}{2}t},$$

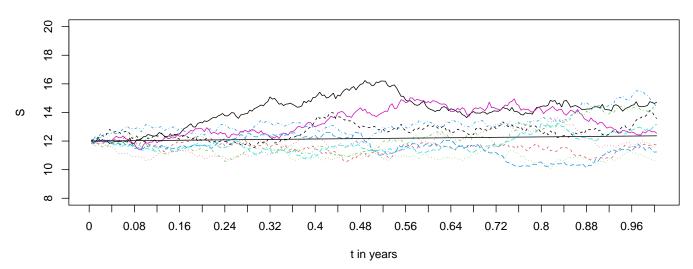
which makes it possible to calculate the one step solution on the grid. The second method implemented is the Euler-Maruyama (EM) scheme

$$\hat{S}_{t_{j+1}} = \hat{S}_{t_j} + \mu h \hat{S}_{t_j} + \sigma \sqrt{h} Z_{t_j} \hat{S}_{t_j},$$

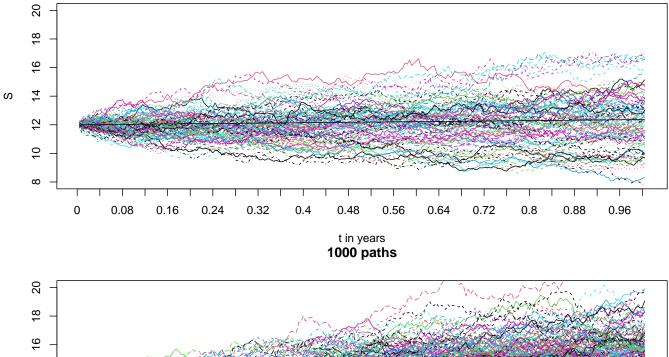
where  $Z_{t_j}$  are standard normally distributed variables.

Notice that the two implementations give slightly different price paths and slightly different estimations. Keep in mind that the results from the onestep method are the ones shown and discussed during the remainder of the problem, even though the results using the EM method are highly comparable.

#### 10 paths



#### 100 paths



ഗ 4 12 9 ω 0 0.08 0.16 0.24 0.32 0.4 0.56 0.64 8.0 0.96 0.48 0.72 0.88 t in years

The Monte Carlo estimator for  $\hat{S}_T$  is calculated separately for each of the values of  $M_i$ ,  $i \in \{1, 2, 3, 4, 5\}$ . The 95% confidence interval (CI) is provided for each estimator. A comparison between these estimators and the analytical solution of  $\mathbb{E}(S_T)$  is done and differences are explained.

The Monte Carlo estimator for  $\hat{S}_T$  is simply calculated by averaging the values of all the different BSM price paths plotted above at time T=1. Thus, in practice, we only need the last value of the price paths to calculate this estimator (and its standard error). The  $(1-\alpha)\% = (1-0.05)\% = 95\%$  CIs are calculated by finding the standard error of the values of all the different BSM price paths at time T and using the approximation given by

$$CI_{\alpha} = \left(\hat{S}_T - z_{\alpha/2} \frac{se_{\hat{S}_T}}{\sqrt{M}}, \hat{S}_T + z_{\alpha/2} \frac{se_{\hat{S}_T}}{\sqrt{M}}\right),$$

where  $se_{\hat{S}_T}$  is the aforementioned standard error,  $z_{\alpha/2}$  is the  $1-\alpha$  quantile of the standard normal distribution and M is the number of price paths simulated. This is an asymptotically valid  $1-\alpha$  CI, which means it becomes closer to the exact analytic value when M is increased.

Table 1: Monte Carlo Estimation for S T, varying M

	M1 = 10	M2 = 100	M3 = 1000	M4 = 10000	M5 = 100000
Est.	12.79035	12.27904	12.41963	12.32897	12.36536
Lower CI	11.86917	11.90612	12.28574	12.28727	12.35226
Upper CI	13.71154	12.65195	12.55352	12.37067	12.37845

Now, what is the analytical solution for  $\mathbb{E}(S_T)$ ? We know that the process of the risky asset in  $t \in [0, T]$  is distributed as

$$S_t \sim \mathcal{L}N(\mu^*, \sigma^{*2}),$$

where  $\mathbb{E}(S_T) = \mu^*$ . In addition, we know that, when  $W_t \sim N(0,t)$ , it follows that

$$X_t \sim N(\ln S_0 - \left(\frac{\sigma^2}{2} - \mu\right)t, \sigma^2 t) = N(\mu_{X_t}, \sigma_{X_t}^2),$$

where  $S_t = e^{X_t}$ . Finally, we know that the expected value of the log-normally distributed variable  $S_t$  is  $\exp\left(\mu_{X_t} + \frac{\sigma_{X_t}^2}{2}\right)$ . This means that the analytical solution for  $\mathbb{E}(S_T)$  is

$$\mu^* = \exp\left(\mu_{X_t} + \frac{\sigma_{X_t}^2}{2}\right) = \exp\left(\ln S_0 - \left(\frac{\sigma^2}{2} - \mu\right)t + \frac{\sigma^2 t}{2}\right) = S_0 e^{\mu t},$$

which in this case has the numerical value

#### **#>** [1] 12.36545

From the results above we can clearly see that the MC estimations move closer and closer to the analytical solution when the number of paths M is increased. For  $M_5$  the estimation is precise to the first three decimals (depending on which version one uses; this is true for onestep), which I would regard as a very good estimation. We also see that the CI's clearly change with the increase of M; the lower value of the CI's increase with M and the upper value of the CI's decrease with M, meaning that the 95% CI's become narrower when the number of paths increase. This is what we expect from the MC theory, since the estimators are unbiased and, as stated by the strong Law of Large Numbers, the sample average converges a.s. to the true expected value.

Now we fix the number of paths  $M^* = 1000$  and vary the values of n, i.e. the number of steps, while repeating the discussion done above.

Table 2: Monte Carlo Estimation for S\_T, varying n

	n1 = 12	n2 = 24	n3 = 250	n4 = 1000
Est.	12.38563	12.27697	12.43257	12.40517
Lower CI	12.24430	12.15002	12.30242	12.27156
Upper CI	12.52696	12.40391	12.56272	12.53879

We can make several similar observations in this case;

Notice however that is is the number of paths M, and not the number of discretization points n, that yields dramatic differences when the value is increased. In other words, the estimates move closer to the true value when n is increased while M is fixed, but the changes seem to be more dramatic when n is fixed and M is increased. It is however important to notice that, in our experiment, n is only changed across three orders of magnitude, while M is changed across five orders of magnitude, which might lead to a somewhat biased observation or discussion.

### Problem B - Option Pricing I

We calculate the price of a Call option with parameter set  $(s_0, T, r, \sigma, K) = (24, 1.5, 0.02, 0.22, 23.5)$ , using the Black-Scholes-Merton (BSM) formula.

#### #> [1] 3.149899

The price is approximately equal to 3.15, when rounded to 2 decimals.

We implement a Monte Carlo estimator for the price of this option by simulating paths with the Euler-Maruyama method for steps n = 10, 100, 1000 and paths M = 10, 100, 1000, 10000.

Notice that for the path-independent options, like standard European call and put options, it is not needed to save the price path as is done in my implementation. This is done to keep the function as general as possible, in order to re-use it for the rest of the assignment. To increase computational efficiency and lessen the use of memory one could simply iteratively overwrite one variable, instead of saving the entire price path history.

Assuming that  $\mu = r$ , we use the MC estimator to calculate the price of the option.

Table 3: Relative Errors

	n = 10	n = 100	n = 1000
M = 10	0.5821167	0.5566707	0.6125545
M = 100	0.0255317	0.0021910	0.3364722
M = 1000	0.0522963	0.0273627	0.0263294
M = 10000	0.0066211	0.0508549	0.0291815
M = 100000	0.0035652	0.0019417	0.0029947

Table 4: Absolute Errors

	n = 10	n = 100	n = 1000
M = 10	1.8336091	1.7534568	1.9294852
M = 100	0.0804222	0.0069013	1.0598536
M = 1000	0.1647280	0.0861897	0.0829349
M = 10000	0.0208557	0.1601878	0.0919187
M=100000	0.0112299	0.0061162	0.0094331

How can these results be interpreted in view of n and M?

First of all, we can clearly see that, in the majority of cases, the absolute values of the errors decreases when the number of paths M increases. This is coherent with what we expect, as noted earlier, based on the MC estimator. The best value obtained for the relative error is  $\approx -0.0019$ , meaning that we are able to get within 0.2% of the closed form solution, which is very accurate.

Secondly, we notice that it is the change in M that leads to dramatic changes (often of an order of magnitude) in the errors, as already noted in Problem A, where a change in n does not lead to the same dramatic changes BE SURE THAT THIS MAKES SENSE!

Additionally, we notice that the lowest errors tend to be found in the lower right corner of the tables, which is as expected, since this is the area of the table with the finest discretized grids (large n) and more paths (large M).

#### WHAT ELSE CAN I COMMENT ON HERE?!

Moreover, notice that all values are smaller in Table 3 (relative errors) compared to their respective value in Table 4 (absolute errors), since the relative errors are calculated in almost the same way as the absolute errors, where the only difference is not taking absolute values and dividing by the BSM price ( $\approx 3.15$ ). Notice that if the absolute value is added in the calculation of the relative errors, the values are all the same, with the five negative signs disappearing.

I cannot find a clear pattern in the negative signs that appear in the relative errors, which means that it is not clear that some combinations of n and M overestimate the BSM price.

The results that are yielded for values of M smaller than 1000 are useless, since they are very bad estimations far from the BSM price. When setting M = 1000 we are getting into the 5% domain of relative errors, which is further improved when M is further increased. Thus, we see that we need a (relatively) substantial amount of paths in order for the MC estimator to gain useful estimations - it is not enough with 10 or 100 paths.

### Problem C - Option Pricing II

A Monte Carlo estimator is implemented for an Asian Call Option. This Asian Call Option averages prices every Friday. We set n=252 and assume that n=1 is Monday. This means that we average the prices  $t_5, t_{10}, t_{15}, \ldots$  The payoff profile for this option is thus  $(\bar{S}-K)^+$ , where  $\bar{S}$  is the arithmetic average over all the prices  $t_i, i \in \{5, 10, 15, \ldots\}$ . We set M=10000 and use the parameter set  $(s_0, T, r, \sigma, K) = (20, 1, 0.02, 0.24, 20)$ .

#### #> [1] 1.174624

As we can see from the result above, the price estimated via the MC estimator is 1.1746.

A Monte Carlo estimator is implemented for a Lookback option with payoff profile  $(S_{max} - K)^+$ , where  $S_{max}$  refers to the maximum price during the time to maturity of the option. We use the parameter set  $(s_0, T, r, \sigma, K) = (22, 2, 0.02, 0.29, 21)$ .

We calculate estimations and 95% confidence intervals of this option price for  $M_1 = 1000$ ,  $M_2 = 10000$  and  $M_3 = 100000$  paths.

Table 5: Monte Carlo Estimation for Lookback Option, Varying M

	M1 = 1000	M2 = 10000	M3 = 100000
Est.	9.331903	9.066059	9.187286
Lower CI	8.839099	8.913349	9.137597
Upper CI	9.824706	9.218769	9.236975

Similar observations as in earlier problems can be made concerning how these CI's change with increasing M; they become narrower with increasing number of paths, which is the behavior we expect.