Misspecification, Sparsity, and Superpopulation Inference with Large-Scale Social Network Data

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February 8, 2016

Overview: Foundations of Applied Statistics

Research agenda: A theory of Applied Statistics.

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"Is this the right method to use to answer my question or make my decision?"

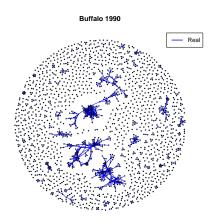
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New desiderata that we can work into modeling decisions.

Generative Network Models



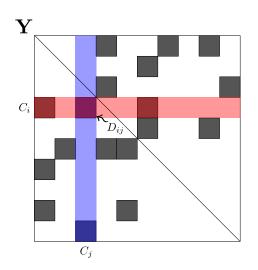
Obtain a set V of n actors.

Explain or predict pairwise outcomes Y_V , potentially using pairwise covariates X_V .

 X_V may be observed or latent.

Running example: inventor collaboration network.

Data Representation



 Y_V entries in **arbitrary sample** space \mathcal{Y} .

Covariates X_V combine observed, latent attributes,

$$X_V^{ij} = f(C_V^i, C_V^j, D_V^{ij}).$$

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Attempt 1: Network Regression

Cox PH regression. (Perry and Wolfe, 2013) Inventor coauthorships in Michigan's motor industry 1982-1988.

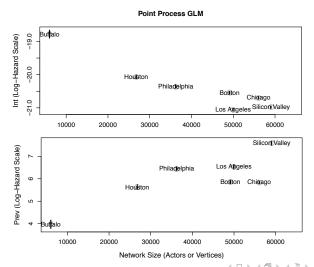
Covariates (Coefs are **log-ratios**):

- post85: After 1985.
- asgnum: Work for same firm.
- prev: Have worked together before.

	lower	est	upper
post85	15.49	15.84	16.20
asgnum	4.65	4.83	5.02
pre	11.36	11.73	12.10
post85:asgnum	-4.77	-4.40	-4.03
post85:prev	-14.57	-14.00	-13.44
asgnum:prev	-5.56	-5.16	-4.76
post85:asgnum:prev	3.91	4.52	5.13

Attempt 2: Regional Comparison Regression

Point process regression. Same time window, different regions.



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What sort of question are you asking?

"Single-Sample"

For a fixed set of actors

- Project forward in time.
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- Compare network samples.
- Predict or pool information across networks.
- Scale local intuition to global network.

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Replications restricted to V.

Replications for any $V \subset \mathbb{V}$.

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Theory so far covers single-sample inference, giving little guidance for superpopulation questions .

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Definition 1 (Random Graph Process).

A random interaction process $Y_{\mathbb{V}}$ is a stochastic process indexed by a countably infinite vertex set V whose finite-dimensional distribution for any finite subset $V \subset \mathbb{V}$ defines an interaction graph Y_V with vertex set V. Denote the law of $Y_{\mathbb{V}}$ as $\mathbb{P}_{\mathbb{V}}$ and the law of a finite-dimensional projection Y_V as \mathbb{P}_V .

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Statistical interpretation: Observed samples are finite subgraphs of population graph. Population graph is of scientific interest.

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Superpopulation: Random graph process

- Induces Kolmogorov consistency on constructed sequences.
- Focus on relationships between finite-dimensional distributions.
- Infinity is useful, but limit is unimportant.

The Method: Parametric MLE

Operational procedure

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Likelihood inference procedure

- ① Propose a model family $\mathcal{P}_{\Theta,\mathbb{V}}$ of models $\mathbb{P}_{\theta,\mathbb{V}}$.
- ② $\mathcal{P}_{\Theta,\mathbb{V}}$ implies a likelihood $\mathbb{P}_{\theta,V}$ on the sampled index set V for each $\theta\in\Theta$. Compute

$$\hat{\theta}_V = \arg\max_{\Theta} \log \mathbb{P}_{\theta, V}(Y_V). \tag{1}$$

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Note: Step 3 is the only difference between single-sample and superpopulation.

Models and Misspecification

Model-Building Tradeoffs

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No parsimonious model can fully represent complex network structure.

Choose one:

- Local structure. Homophily, Heterophily, Transitivity, etc.
- Global structure. Sparsity, Percolation, etc.

Models and Misspecification

Common local approaches

Conditionally independent dyads (regression):

$$P(Y \mid X) = \prod_{i < j < n} P(Y_V^{ij} \mid X_V^{ij}).$$

Infinitely exchangeable dyads (Aldous-Hoover):

$$P(Y \mid X) = \int_{\mathcal{C}} \prod_{i < j < n} P(Y_V^{ij} \mid W_V^{ij}(C_V^i, C_V^j)) dF(C).$$

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Parameter estimates as **measurements** of $\mathbb{P}_{0,V}$.

Minimal criterion for "usefulness": Stability.

"Similar" inputs Y_V yield "similar" estimates $\hat{\theta}_V$.

Stability: Single sample case

"Similar input" means replications of Y_V from the same finite distribution $\mathbb{P}_{0,V}$.

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Huber 1967 showed MLE is large-sample consistent for a pseudo-true parameter (naming due to Sawa 1978), satisfying

$$\bar{\theta}_{V} = \arg \max_{\theta \in \Theta} \mathbb{E}_{\mathbb{P}_{0}}[\log \mathbb{P}_{\theta, V}(Y_{V})]$$
 (2)

"Similar output" defined by concentration of $\hat{\theta}_V$ in large samples.

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Sparsity Misspecification 2/8/2016

Stability: Superpopulation case

"Similar input" means any sample Y_V drawn from the same superpopulation $\mathbb{P}_{0,\mathbb{V}}$.

Intuitively, outputs $\hat{\theta}_V$ are similar if they **effectively estimate** the same thing.

What does the MLE $\hat{\theta}_V$ effectively estimate when the model is misspecified?

The Effective Estimand of the MLE

Define the effective estimand of the MLE as the finite-sample pseudo-true parameter.

$$\bar{\theta}_V = \arg\max_{\theta \in \Theta} \mathbb{E}_{\mathbb{P}_0}[\log \mathbb{P}_{\theta, V}(Y_V)] \tag{2}$$

The Effective Estimand of the MLE

Define the effective estimand of the MLE as the finite-sample pseudo-true parameter.

$$\bar{\theta}_V = \arg\max_{\theta \in \Theta} \mathbb{E}_{\mathbb{P}_0}[\log \mathbb{P}_{\theta, V}(Y_V)] \tag{2}$$

Justifications:

- Finite-sample concentration (e.g., Spokoiny 2012).
- Fisher-consistency inversion.
- Estimating equation unbiased.
- KL projection plug-in.

Superpopulation Stability Criterion

Criterion 1.

A procedure is superpopulation stable for making inferences about a superopulation process $\mathbb{P}_{0,\mathbb{V}}$ only if, for any finite sample Y_V generated according to $\mathbb{P}_{0,V}$, the effective estimand $\bar{\theta}_V$ of the estimator $\hat{\theta}_V$ is invariant to the indexing set V.

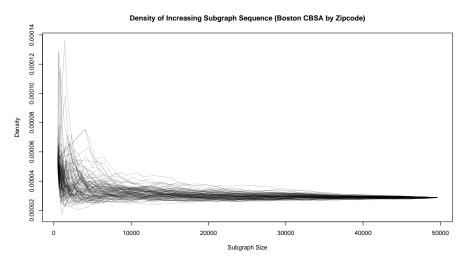
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Test the criterion with top-down specification of superpopulation properties, e.g., sparsity.

Illustration



Formally

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Formally

Define the **density operator**

$$D(Y_V) = \frac{\sum_{ij} \mathbb{I}\{Y_V^{ij} \neq 0\}}{\binom{|V|}{2}}.$$

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Definition 1 (Sparse Graph Process).

Let $Y_{\mathbb{V}}$ be a random graph process on \mathbb{V} . $Y_{\mathbb{V}}$ is *sparse* if and only if for any $\epsilon > 0$ there exists an n such that for any subset of vertices $V \in \mathbb{V}$ with |V| > n the corresponding finite dimensional random graph Y_V has the property $\mathbb{E}(D(Y_V)) < \epsilon$.

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Differences vs single-sample sparsity.

Single-sample

- Defined in terms of random graph sequences.
- Often defined in explicitly non-Kolmogorov-consistent terms.
- Analogy for single sample with very few observed interactions.

Superpopulation

- Property of a random graph process, not a random graph.
- Defines an assumption about the system, not a theoretical object.

Definition

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A model family $\mathcal{P}_{\theta,\mathbb{V}}$ is **sparsity misspecfied** iff for every $\theta \in \Theta$ and every increasing sequence of vertex sets (V_n) ,

$$\frac{\mathbb{E}_{\theta}(D(Y_{V_n}))}{\mathbb{E}_{0}(D(Y_{V_n}))} \to 0 \text{ or } \infty.$$

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For example,

- For CID (under regularity) and exchangeable models, population extension is dense or empty (e.g., Orbanz and Roy, 2013).
- For process models, most lock in a given form for $\epsilon(n)$ (e.g., power law for preferential attachment).

Main Result: Moving Target

Theorem 1 (Moving target theorem).

Let (V_n) be an increasing sequence of vertex sets from \mathbb{V} . Suppose the following hold:

- ① For some finite n, $\mathbb{E}_0(D(Y_{V_n})) > 0$.
- ② The inferential family $\mathcal{P}_{\Theta,\mathbb{V}}$ is sparsity misspecified for the true population process $\mathbb{P}_{0,\mathbb{V}}$.
- 3 The inferential model is responsive to the sample density $D(Y_{V_n})$ under the true population process and

$$|\mathbb{E}_{\bar{\theta}}(D(Y_{V_n})) - \mathbb{E}_0(D(Y_{V_n}))| = O(\epsilon_{\ell}(n)). \tag{3}$$

Then, $\bar{\theta}_{V_n}$ varies with n in the sense that for any n, there exists an n' > n such that $\bar{\theta}_{V_n} \neq \bar{\theta}_{V_{n'}}$, and the MLE of the model violates Criterion 1.

Example: Poisson Regression

Setup

Question: How do firms influence collaboration dynamics?

 Y_V collaboration counts; X_V^{ij} indicates shared firm.

Assumptions:

- ullet The true collaboration-generating process $Y_{0,\mathbb{V}}$ is sparse .
- All firms have finite size.
- A non-vanishing fraction of firms have a positive number of expected within-firm interactions.

Example: Poisson Regression

Model and effective estimand

Model:

$$Y_{V_n}^{ij} \stackrel{\perp}{\sim} \operatorname{Pois}(\exp(\theta_1 + X_{V_n}^{ij}\theta_2)),$$
 (4)

Effective Estimands:

$$\bar{\theta}_{1V} = \log \left(\frac{\sum_{ij} \mathbb{E}_0(Y_V^y \mid X_V^y = 0)(1 - X_V^y)}{\sum_{ij} (1 - X_V^{ij})} \right)$$
(5)

$$\bar{\theta}_{2V} = \log \left(\frac{\sum_{ij} \mathbb{E}_{0}(Y_{V}^{ij} \mid X_{V}^{ij} = 1) X_{V}^{ij}}{\sum_{ij} X_{V}^{ij}} \middle/ \frac{\sum_{ij} \mathbb{E}_{0}(Y_{V}^{ij} \mid X_{V}^{ij} = 0) (1 - X_{V}^{ij})}{\sum_{ij} (1 - X_{V}^{ij})} \right).$$
(6)

What's wrong with this picture?

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Salvaging conditional independence

What can we estimate with local models?

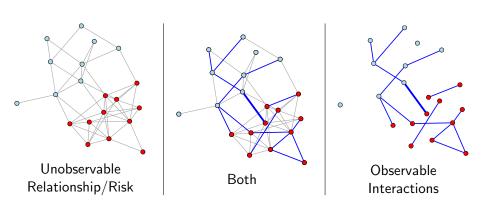
Two-stage process. Complex **relationship** structure R. Y simple conditional on relationships.

$$\mathbb{P}_{0,V} = \mathbb{P}_0(R) \prod_{i < i < n} \mathbb{P}_0(Y_V^{ij} \mid R_V, X_V).$$

Changes question:

$$\mathbb{P}_{\theta,V}(Y_V \mid X_V) \to \mathbb{P}_{\beta,V}(Y_V \mid R_V, X_V).$$

Conditionally Independent Relationship model



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Partial likelihood inference

Let
$$A_V = \mathbb{I}\{Y_V \neq 0\}$$
.

Exploit conditional distribution

$$\mathbb{P}_{\beta}(Y_V \mid A_V, X_V) = \frac{\mathbb{P}_{\beta}(Y_V \mid R_V, X_V)}{1 - \mathbb{P}_{\beta}(Y_V = 0 \mid R_V X_V)},$$

or the zero-truncated likelihood.

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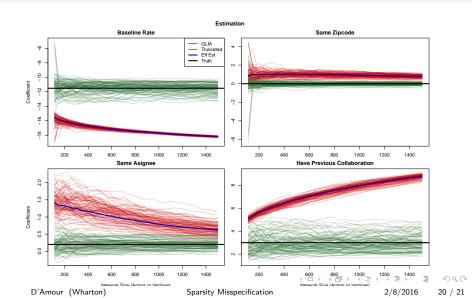
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Bonus: Computation is $O(\sum_{ii} A_V)$.

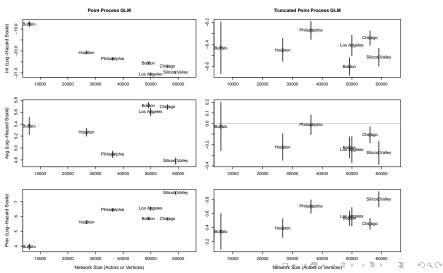
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Simulated success



Real success



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Modeling *less* of the system by seeking invariances make a model *more* scientifically relevant. **Invariance** can be better than a bad explanation.

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