

# Research Positions in Machine Learning and Computer Vision

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## 1 Question 1

Only two equations are supposed to be known:

$$\begin{aligned}\text{Area ( square of side } a) &= a \times a \\ \text{SumIntegers ( } n) &= \frac{n * (n + 1)}{2}\end{aligned}$$

Figure 1 illustrates how a isosceles right triangle of side  $a$  can be approximated by a stack of smaller squares with side  $\frac{a}{2}$ . Each row of the stack contains one more small square than the row below. Figure 2 illustrates how the stack approximates the isosceles right triangle by letting  $n$  tend to infinity. The area covered by the red smaller squares is :

$$\begin{aligned}\text{Area (square of side } \frac{a}{n}) * \text{SumIntegers (} n) &= \frac{a}{n} * \frac{a}{n} * \frac{n * (n + 1)}{2} \\ &= \frac{a^2}{2} (1 + \frac{1}{n}) \\ &\xrightarrow{n \rightarrow \infty} \frac{a^2}{2}\end{aligned}$$

## 2 Question 2

The singular value decomposition theorem states that if  $\mathbf{A}$  is a real  $m \times n$  matrix, then there exist orthogonal matrices  $\mathbf{U} = [\mathbf{u}_1 | \dots | \mathbf{u}_m] \in \mathbb{R}^{m \times m}$  and  $\mathbf{V} = [\mathbf{v}_1 | \dots | \mathbf{v}_n] \in \mathbb{R}^{n \times n}$  such that:

$$\mathbf{U}^T \mathbf{A} \mathbf{V} = \mathbf{\Sigma} = \text{diag}(\sigma_1, \dots, \sigma_p) \in \mathbb{R}^{m \times n}, \quad p = \min\{m, n\},$$

where  $\sigma_1 \geq \sigma_2 \geq \dots \geq \sigma_p \geq 0$ .

By multiplying by blocs,  $\mathbf{A}$  can be written as a sum of rank 1 matrix :

$$\begin{aligned}\mathbf{A} &= \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T \\ &= [\sigma_1 \mathbf{u}_1 | \dots | \sigma_p \mathbf{u}_p] * [\mathbf{v}_1 | \dots | \mathbf{v}_p]^T \\ &= \sum_{i=1}^p \sigma_i \mathbf{u}_i \mathbf{v}_i^T\end{aligned}$$

If  $\mathbf{A} \in \mathbb{R}^{m \times n}$  with  $m < n$ ,  $\text{rank}(\mathbf{A}) = m$ , and  $\mathbf{b} \in \mathbb{R}^m$ , then the linear system  $\mathbf{A} \mathbf{x}_{sol} = \mathbf{b}$  is said to be underdetermined. There are infinitely many solutions. It is the case where the number of observation is not large enough to solve the problem. But we can derive a feasible solution set. The matrix  $\mathbf{\Sigma}$  has this particular shape :

$$\mathbf{\Sigma} = \begin{pmatrix} \sigma_1 & \dots & 0 & 0 & \dots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \dots & \sigma_m & 0 & \dots & 0 \end{pmatrix}_{m \times n}$$

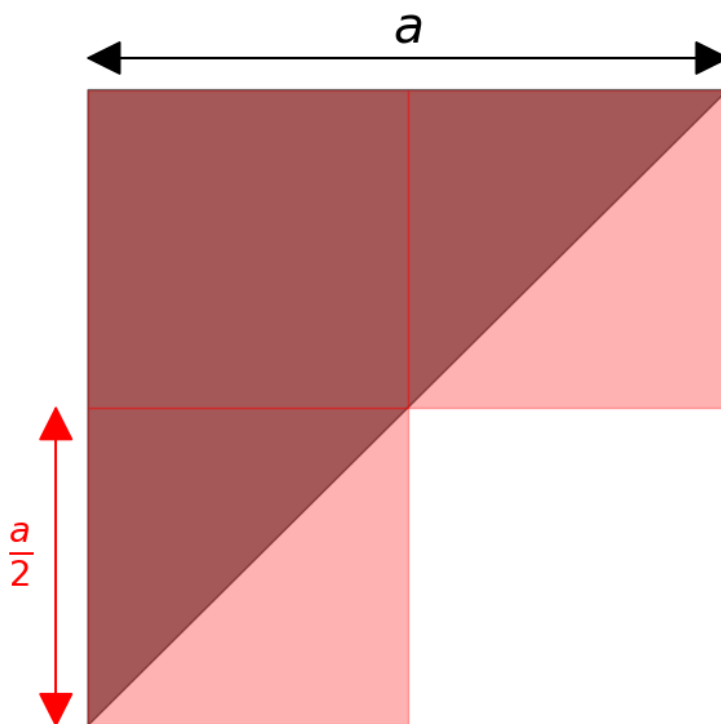


Figure 1: Squares of side  $\frac{a}{2}$  realises a cover of the isosceles right triangle

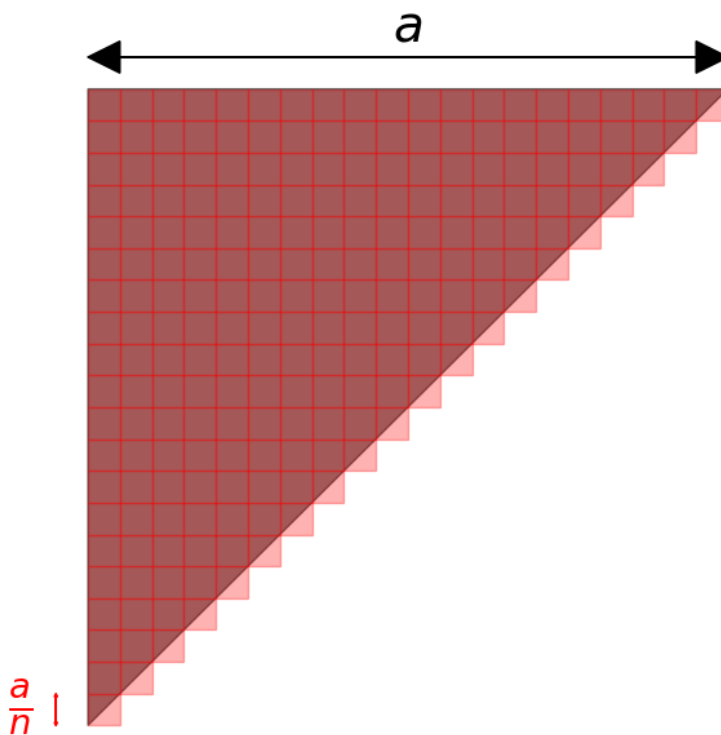


Figure 2: Stacking smaller squares gives a better covering of the isosceles right triangle

where  $\forall i = 1, \dots, m, \quad \sigma_i > 0$ .

To find a particular solution, we compute the squared error which is a function of  $\mathbf{x} \in \mathbb{R}^m$ :

$$\begin{aligned} e(\mathbf{x}) &= \|\mathbf{Ax} - \mathbf{b}\|^2 \\ &= (\mathbf{Ax} - \mathbf{b})^T (\mathbf{Ax} - \mathbf{b}) \\ &= \mathbf{x}^T \mathbf{A}^T \mathbf{Ax} - \mathbf{b}^T \mathbf{Ax} - \mathbf{x}^T \mathbf{A}^T \mathbf{b} + \mathbf{b}^T \mathbf{b} \\ &= \mathbf{x}^T \mathbf{A}^T \mathbf{Ax} - 2\mathbf{b}^T \mathbf{Ax} + \mathbf{b}^T \mathbf{b} \end{aligned}$$

We compute the gradient and set it to zero to get the minimum of the squared error:

$$\nabla e(\mathbf{x}_{\text{part}}) = 2\mathbf{A}^T \mathbf{Ax}_{\text{part}} - 2\mathbf{A}^T \mathbf{b} = \mathbf{0}_{\mathbb{R}^n}$$

The issue compared to usual machine learning problem is that we cannot invert  $\mathbf{A}^T \mathbf{A}$  to derive the normal equation as  $\text{rank}(\mathbf{A}) < n$ . The SVD decomposition will help us to find an explicit particular solution. Let's define  $\boldsymbol{\alpha} = \mathbf{V}^T \mathbf{x}_{\text{part}}$ :

$$\begin{aligned} \mathbf{A}^T \mathbf{Ax}_{\text{part}} &= \mathbf{A}^T \mathbf{b} \\ (\mathbf{U}\boldsymbol{\Sigma}\mathbf{V}^T)^T \mathbf{U}\boldsymbol{\Sigma}\mathbf{V}^T \mathbf{x}_{\text{part}} &= (\mathbf{U}\boldsymbol{\Sigma}\mathbf{V}^T)^T \mathbf{b} \\ \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T \mathbf{U}\boldsymbol{\Sigma}\mathbf{V}^T \mathbf{x}_{\text{part}} &= \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T \mathbf{b} \\ \boldsymbol{\Sigma}^T \boldsymbol{\Sigma} \boldsymbol{\alpha} &= \boldsymbol{\Sigma}^T \mathbf{U}^T \mathbf{b} \\ \forall i = 1, \dots, m, \quad \sigma_i^2 \alpha_i &= \sigma_i \mathbf{u}_i^T \mathbf{b} \\ \forall i = 1, \dots, m, \quad \alpha_i &= \frac{\mathbf{u}_i^T \mathbf{b}}{\sigma_i} \end{aligned}$$

We derive  $\mathbf{x}$  from the  $\alpha_i$ , by taking  $\alpha_i = 0$  if  $i = m+1, \dots, n$ :

$$\mathbf{x}_{\text{part}} = \mathbf{V}\boldsymbol{\alpha} = \sum_{i=1}^m \alpha_i \mathbf{v}_i = \sum_{i=1}^m \frac{\mathbf{u}_i^T \mathbf{b}}{\sigma_i} \mathbf{v}_i = \sum_{i=1}^m \mathbf{v}_i \frac{\mathbf{u}_i^T \mathbf{b}}{\sigma_i} = \left( \sum_{i=1}^m \frac{1}{\sigma_i} \mathbf{v}_i \mathbf{u}_i^T \right) \mathbf{b}$$

The matrix in brackets is named the pseudoinverse, and it is noted  $\mathbf{A}^\dagger$ . We recognise the rank 1 matrix decomposition with :

$$\begin{aligned} \mathbf{A}^\dagger &= \sum_{i=1}^m \frac{1}{\sigma_i} \mathbf{v}_i \mathbf{u}_i^T \\ &= \left[ \frac{1}{\sigma_1} \mathbf{v}_1 \mid \dots \mid \frac{1}{\sigma_m} \mathbf{v}_m \right] * [\mathbf{u}_1 \mid \dots \mid \mathbf{u}_m]^T \\ &= \mathbf{V}\boldsymbol{\Sigma}^\dagger \mathbf{U}^T \end{aligned}$$

where  $\boldsymbol{\Sigma}^\dagger = \text{diag}(\frac{1}{\sigma_1}, \dots, \frac{1}{\sigma_m}) \in \mathbb{R}^{n \times m}$ . It can also be computed by the formula:

$$\begin{aligned} \mathbf{A}^T (\mathbf{AA}^T)^{-1} &= \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T (\mathbf{U}\boldsymbol{\Sigma}\mathbf{V}^T \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T)^{-1} \\ &= \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T (\mathbf{U} \text{diag}(\sigma_1^2, \dots, \sigma_p^2) \mathbf{U}^T)^{-1} \\ &= \mathbf{V}\boldsymbol{\Sigma}^T \mathbf{U}^T \mathbf{U} \text{diag}(\frac{1}{\sigma_1^2}, \dots, \frac{1}{\sigma_p^2}) \mathbf{U}^T \\ &= \mathbf{V}\boldsymbol{\Sigma}^\dagger \mathbf{U}^T \\ &= \mathbf{A}^\dagger \end{aligned}$$

Now let's compute the general solution of the equation, so  $\mathbf{Ax}_{\text{gene}} = \mathbf{0}_{\mathbb{R}^m}$ . It is equivalent to derive the null space of  $\mathbf{A}$  and, as  $\mathbf{U}$  is invertible, equivalent to derive the null space of  $\boldsymbol{\Sigma}\mathbf{V}^T$ . Let's define  $\mathbf{x}' = \mathbf{V}^T \mathbf{x}_{\text{gene}}$ . As  $\boldsymbol{\Sigma}$  is diagonal in  $\mathbb{R}^{m \times n}$ , a vector in the null space has the shape  $(0, \dots, 0, x'_{m+1}, \dots, x'_n)^T$ . As  $\mathbf{V}^T$  is orthogonal, the null space of  $\mathbf{A}$  is  $\text{span}(\mathbf{v}_{m+1}, \dots, \mathbf{v}_n)$ .

We can also derive a formula based on  $\mathbf{A}$ . Let's define:

$$\mathbf{P} = \mathbf{I}_n - \mathbf{A}^\dagger \mathbf{A} = \mathbf{I}_n - \mathbf{A}^T (\mathbf{A} \mathbf{A}^T)^{-1} \mathbf{A}$$

By substituting the respective SVD:

$$\begin{aligned} \mathbf{P} &= \mathbf{I}_n - \mathbf{V} \mathbf{\Sigma}^\dagger \mathbf{U}^T \mathbf{U} \mathbf{\Sigma} \mathbf{V}^T \\ &= \mathbf{I}_n - [\mathbf{v}_1 | \dots | \mathbf{v}_n] \begin{pmatrix} \mathbf{I}_m & \mathbf{0} \\ \mathbf{0} & \mathbf{0} \end{pmatrix} [\mathbf{v}_1 | \dots | \mathbf{v}_n]^T \\ &= \mathbf{I}_n - \sum_{i=1}^m \mathbf{v}_i \mathbf{v}_i^T \end{aligned}$$

We decompose any vector  $\mathbf{x}$  on the orthogonal basis formed by the  $\mathbf{v}_i$  :  $\mathbf{x} = \sum_{j=1}^n \nu_j \mathbf{v}_j$ . We can compute the image of  $\mathbf{P}$ :

$$\begin{aligned} \mathbf{P} \mathbf{x} &= \sum_{j=1}^n \nu_j \mathbf{v}_j - \sum_{i=1}^m \left( \mathbf{v}_i \mathbf{v}_i^T \sum_{j=1}^n \nu_j \mathbf{v}_j \right) \\ &= \sum_{j=1}^n \nu_j \mathbf{v}_j - \sum_{i=1}^m \nu_i \mathbf{v}_i \\ &= \sum_{j=m+1}^n \nu_j \mathbf{v}_j \end{aligned}$$

We have found that the image of  $\mathbf{P}$  is the null space of  $\mathbf{A}$ .

Finally  $\mathbf{x}_{sol} = \mathbf{x}_{part} + \mathbf{x}_{gene}$  :

$$\boxed{\mathbf{x}_{sol} = \mathbf{V} \mathbf{\Sigma}^\dagger \mathbf{U}^T \mathbf{b} + \text{span}(\mathbf{v}_{m+1}, \dots, \mathbf{v}_n)} \quad (1)$$

$$\boxed{\mathbf{x}_{sol} = \mathbf{A}^T (\mathbf{A} \mathbf{A}^T)^{-1} \mathbf{b} + (\mathbf{I}_n - \mathbf{A}^T (\mathbf{A} \mathbf{A}^T)^{-1} \mathbf{A}) \mathbf{x} \quad \forall \mathbf{x}} \quad (2)$$

### 3 Question 3

To compute the gradient of  $E(u) = \sum_{i=1}^2 \sum_{j=1}^2 \arctan(u_{i,j} - u_{i-1,j})$  from sums starting from  $\mathbf{1}$  with the assumption that  $u_{0,1} = u_{0,2} = 0$ , we need to compute the partial derivative with respect to  $u_{1,1}$ ,  $u_{1,2}$ ,  $u_{2,1}$  and  $u_{2,2}$ .

For  $u_{1,1}$ :

$$\begin{aligned} \frac{\partial E}{\partial u_{1,1}} &= \frac{\partial}{\partial u_{1,1}} [\arctan(u_{1,1} - 0) + \arctan(u_{2,1} - u_{1,1})] \\ &= \frac{1}{1 + u_{1,1}^2} + \frac{-1}{1 + (u_{2,1} - u_{1,1})^2} \end{aligned}$$

For  $u_{1,2}$ :

$$\begin{aligned} \frac{\partial E}{\partial u_{1,2}} &= \frac{\partial}{\partial u_{1,2}} [\arctan(u_{1,2} - 0) + \arctan(u_{2,2} - u_{1,2})] \\ &= \frac{1}{1 + u_{1,2}^2} + \frac{-1}{1 + (u_{2,2} - u_{1,2})^2} \end{aligned}$$

For  $u_{2,1}$ :

$$\begin{aligned}\frac{\partial E}{\partial u_{2,1}} &= \frac{\partial}{\partial u_{2,1}} [\arctan(u_{2,1} - u_{1,1})] \\ &= \frac{1}{1 + (u_{2,1} - u_{1,1})^2}\end{aligned}$$

For  $u_{2,2}$ :

$$\begin{aligned}\frac{\partial E}{\partial u_{2,2}} &= \frac{\partial}{\partial u_{2,2}} [\arctan(u_{2,2} - u_{1,2})] \\ &= \frac{1}{1 + (u_{2,2} - u_{1,2})^2}\end{aligned}$$

Finally :

$$\begin{aligned}\nabla E(u) &= \begin{pmatrix} \frac{\partial E}{\partial u_{1,1}} & \frac{\partial E}{\partial u_{1,2}} \\ \frac{\partial E}{\partial u_{2,1}} & \frac{\partial E}{\partial u_{2,2}} \end{pmatrix} \\ \nabla E(u) &= \begin{pmatrix} \frac{1}{1+u_{1,1}^2} + \frac{-1}{1+(u_{2,1}-u_{1,1})^2} & \frac{1}{1+u_{1,2}^2} + \frac{-1}{1+(u_{2,2}-u_{1,2})^2} \\ \frac{1}{1+(u_{2,1}-u_{1,1})^2} & \frac{1}{1+(u_{2,2}-u_{1,2})^2} \end{pmatrix}\end{aligned}$$

## 4 Question 4

Given the probability density function :

$$p(x; \alpha, \epsilon) = \frac{1}{Z} e^{-\alpha x^2 - \epsilon}$$

Knowing that the  $x(i)$  are independent and identically distributed samples, the likelihood function  $\mathcal{L}$  is :

$$\begin{aligned}\mathcal{L}(\alpha; x(1), \dots, x(m), \epsilon) &= \prod_{i=1}^m p(x(i); \alpha, \epsilon). \\ &= \frac{1}{Z^m} e^{-\sum_{i=1}^m (\alpha x(i)^2) - m\epsilon}\end{aligned}$$

The likelihood function is maximum when  $\alpha$  is as small as possible :

$$\hat{\alpha}_{MLE} \rightarrow 0$$

It is clearly not a good estimator as it does not depend on  $x(i)$ .

We could try the first moment method, but :

$$E[X] = \int_{-\infty}^{\infty} x \frac{1}{Z} e^{-\alpha x^2 - \epsilon} dx = 0$$

as the integrand is an odd function.

For the second moment :

$$E[X^2] = \int_{-\infty}^{\infty} x^2 \frac{1}{Z} e^{-\alpha x^2 - \epsilon} dx$$

Let's do an integration by parts:

$$\begin{aligned} u &= \frac{x}{Z} & u' &= \frac{1}{Z} \\ v &= \frac{1}{-2\alpha} e^{-\alpha x^2 - \epsilon} & v' &= x e^{-\alpha x^2 - \epsilon} \end{aligned}$$

We get the primitive part equals to zero and:

$$E[X^2] = \frac{Z}{Z2\alpha} = \frac{1}{2\alpha}$$

Applying the method of the moments:

$$\hat{\alpha}_{MM} = \frac{1}{2} \frac{n}{\sum_{i=1}^n x(i)^2}$$