Forecast of Polish Inflation Rate

Erste Group CEE Research Interview Task

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Task Overview

- The task is to analyze the historical dynamics of Polish inflation and forecast the HICP for the first four months of 2025
- My approach begins with a diagnosis of the time series and observation of the key characteristics of time series like stationarity, seasonality, trend and autocorrelation
- I then build and compare 3 models:
 - A basic AR(1) model
 - A univariate Auto ARIMA model selected by AIC minimizing
 - A multivariate ARIMAX model with external macroeconomic variables
- Goals of this task:
 - Create a both technically and economically meaningful forecast
 - Show that my AIC with Erste's CEE Research team is approximately 0:)



Data Overview & Libraries Used

Data description:

- Original <u>data set</u> is downloaded from Eurostat
- Represents percentage change in Polish HICP
- Monthly frequency
- Spans between January 1997 and April 2025
- Last accessed on the 21st of May

For the sake of this task I have split the data in the following 2 data subsets:



During the project I have used the following libraries:

- lubridate
- ggplot2
- tseries
- forecast
- quantmod
- dplyr
- tidyr



Data Overview & Libraries Used







TS Diagnostics (Stationarity)

- Stationarity implies constant mean, variance and autocovariance over time
- Most time series models (like ARIMA) require stationarity to perform well
- Due to the COVID spike in inflation one can't visually decide if the time series is stationary
- Use Augmented Dickey-Fuller Test:
 - At $\alpha = 0.05$
 - H_0 : there is unit root (non-stationary)
 - H_A : there is no unit root (stationary)
- Conclusion: Reject H₀ => the series is stationary at 5% level

HICP Poland (Monthly)



```
##
## Augmented Dickey-Fuller Test
##
## data: train_ts
## Dickey-Fuller = -3.5707, Lag order = 5, p-value = 0.03827
## alternative hypothesis: stationary
```



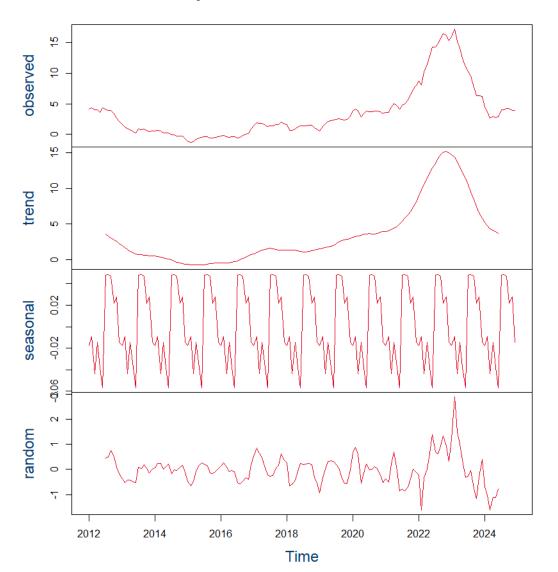
TS Diagnostics (Decomposition)

A time series can be decomposed into **Trend**, **Seasonality** and **Noise**:

- **Trend** not permanent
 - No clear trend before 2021
 - Strong upward movement between 2021 and 2023 due to COVID
 - Fast decline after 2023
- Seasonality regular pattern detected, but variation is small (≈ 0.08)
 - Seasonality is statistically visible but economically minor
 - If seasonality is not impactful, SARIMA model will drop it
- Noise what's left after removing trend and seasonality
 - Stable before 2020 and more volatile around 2021–2023
 - Indicates that external factors likely play a role



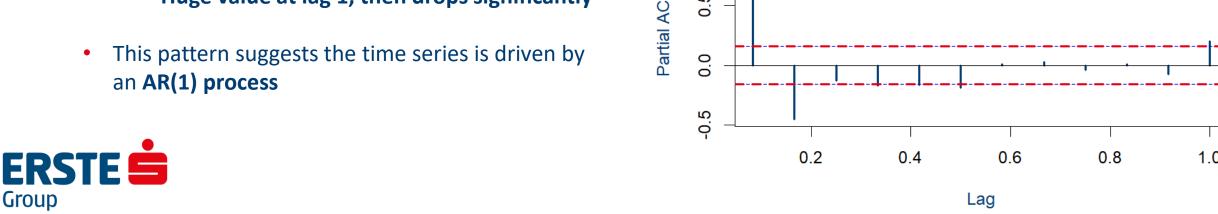
Decomposition of additive time series

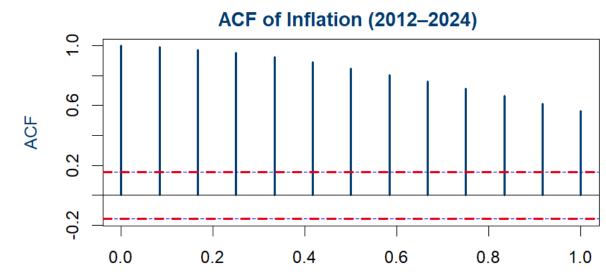


TS Diagnostics (Autocorrelation)

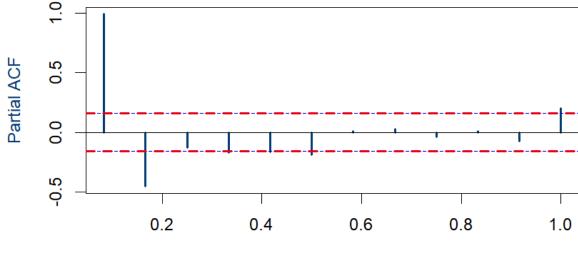
Autocorrelation measures how current values are related to past values

- ACF (Autocorrelation Function)
 - Captures correlation between current and past values (past errors in MA process)
 - Strong correlation at lag 1 (≈ 1)
 - No clear cut-off
- PACF (Partial ACF)
 - shows the pure correlation between a value and its lag (isolates true AR influence)
 - Huge value at lag 1, then drops significantly









Univariate Model - AR(1)

ARIMA(1,0,0) based on diagnostics:

- PACF showed a strong lag-1 spike => use p = 1
- ADF test confirmed stationarity => use d = 0
- ACF decays gradually => no need for MA terms (q = 0)

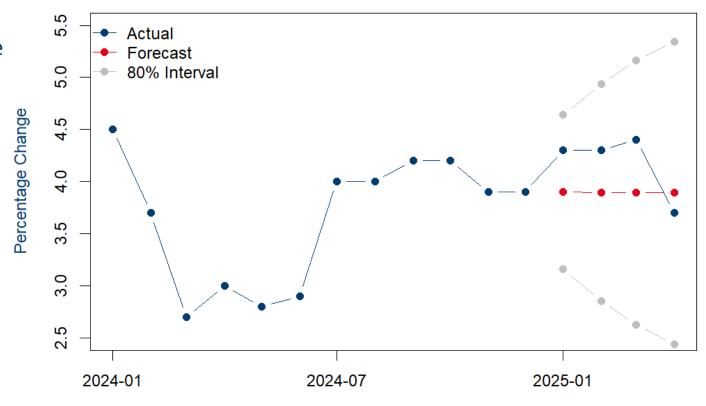
HICP Poland – AR(1) Forecasted vs Observed

Model output:

- AR(1) coefficient ≈ 0.99 => strong persistence
- Residual variance $\sigma^2 \approx 0.34$
- AIC ≈ 281.8

• Forecast:

```
## 2025 Jan 4.3 3.897829
## 2025 Feb 4.3 3.895688
## 2025 Mar 4.4 3.893575
## 2025 Apr 3.7 3.891490
```



Time



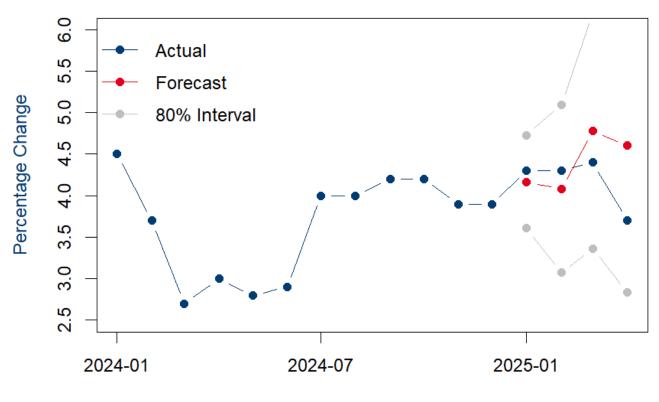
Univariate Model - Auto ARIMA

- auto.arima() to let R choose the best (p,d,q) based on AIC minimization
- Allows seasonality detection (if present uses SARIMA)
- Selected Model ARIMA(1,1,0)(0,0,1)[12]:
 - 1 AR term, 1 order of differencing, 1 seasonal MA term
 - Residual variance $\sigma^2 \approx 0.19$ (better than AR(1))
 - AIC ≈ 198.2 (better than AR(1))
- Forecast:

```
## 2025 Jan 4.3 4.165496
## 2025 Feb 4.3 4.084557
## 2025 Mar 4.4 4.777414
## 2025 Apr 3.7 4.605298
```

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HICP Poland – Auto ARIMA Forecasted vs Observed



Expanding the Model I

Unemployment Rate

- Captures domestic labour market pressure
- According to the Phillips Curve there is inverse relationship between inflation and unemployment
- Lag of 1 month performs best (inflation reacts with a short delay to labour shifts)
- <u>Data</u> from Eurostat, monthly % of population

EUR/PLN Exchange Rate (YoY log diff)

- Poland's largest trade partners (Germany, France) operate in EUR
- EUR/PLN affects prices of cars, industrial equipment and other euro-denominated imports
- Used log return over 12 months + 1-month lag to reflect pricing contract delays
- Data from Yahoo Finance (quantmod)

Oil Price (in PLN)

- Oil is a key input for transportation, heating and production costs
- Oil shocks are historically linked to inflation spikes
- Measured in PLN to reflect true domestic cost (USD price * USD/PLN rate)
- Lag of 9 months best captures delayed cost pass-through
- <u>Data</u> from FRED and Yahoo Finance (quantmod)

USD/PLN Exchange Rate (YoY log diff)

- Many key commodities (especially oil and food) are priced in USD
- Depreciation of the PLN makes imports more expensive (inflationary pressure)
- Used 12-month log return to smooth volatility and capture trend
- Data from Yahoo Finance (quantmod)



Expanding the Model II

Global Food Price Index

- Food makes up a large share of Poland's consumer basket
- Index reflects global agricultural prices (wheat, oils, dairy, meat)
- Captures import-driven food inflation (e.g. droughts, trade restrictions, supply chain distortions, wars)
- Data from FAO

Average Nominal Gross Wage

- Wages reflect domestic demand strength and costpush inflation
- As wages increase, consumption increases and firms pass labour costs to prices
- Lag of 2 months chosen based on test performance
- <u>Data</u> from Statistics Poland (GUS)

10-Year Government Bond Yield

- Long-term yield reflects market expectations of future inflation and monetary policy
- Yields increase when a higher inflation or tighter monetary stance is expected
- Used 6-month lag based on test performance
- <u>Data</u> from FRED



Multivariate Model - ARIMAX

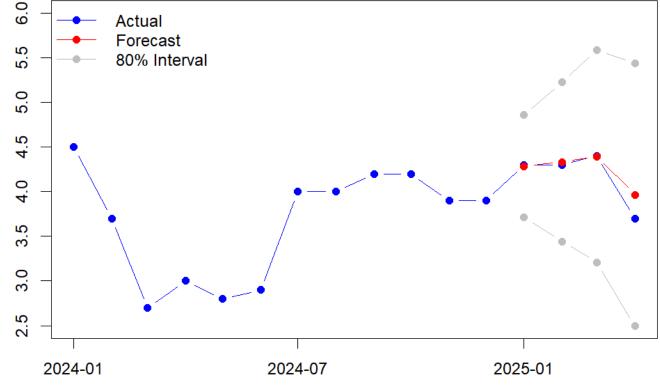
- Trained using auto.arima() with exogenous variables
- Selected Model ARIMA(2,0,1)(1,0,0)[12] + 11 regressors:
 - 2 AR terms, 1 MA term and 1 seasonal AR term (12 months)
 - Residual variance $\sigma^2 \approx 0.199$
 - AIC ≈ 211.1 (slightly worse than the univariate model)
- Forecast:

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```
## 2025 Jan 4.3 4.284120
## 2025 Feb 4.3 4.332600
## 2025 Mar 4.4 4.394824
## 2025 Apr 3.7 3.965498
```

Percentage Change

HICP Poland - ARIMAX Forecasted vs Observed



Time

Multivariate Model - ARIMAX

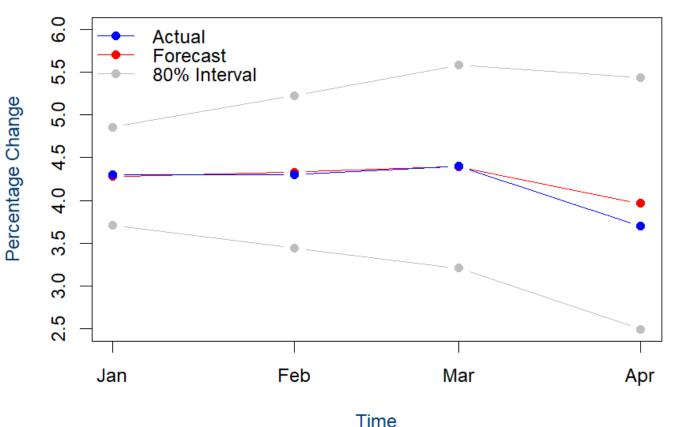
• Forecast:

```
## 2025 Jan 4.3 4.284120
## 2025 Feb 4.3 4.332600
## 2025 Mar 4.4 4.394824
## 2025 Apr 3.7 3.965498
```

- AIC measures in-sample fit and penalizes model complexity => Slightly higher AIC in ARIMAX is due to more parameters (external variables)
- ARIMAX achieved better forecast accuracy for the forecasting period (2025/01 – 2025/04)
- External variables helped capture real-world shocks
- This improved out-of-sample generalization, even with a slight AIC trade-off

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HICP Poland 2025 - ARIMAX Forecasted vs Observed



Thank you!

