

$$y(t) = f[\omega(y(t))] + \varepsilon(t)$$

TC
Tipo de
 cambio

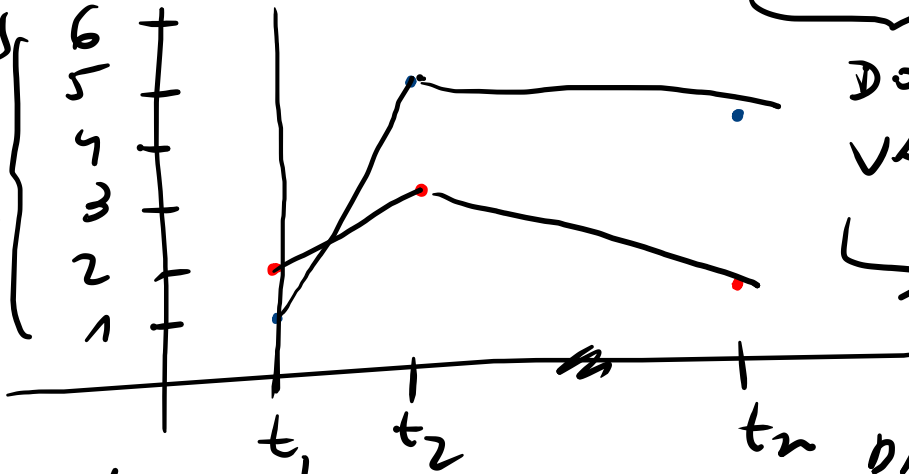
conjunto
 de
 información

ALEATORIO
 variable
 aleatoria

Processos estocásticos

Retornos
Financi
Pass

6
5
4
3
2
1



$\gamma(t, \omega)$ Variável (estados)

Discretas
Dos contínuas
Variáveis

t : tempo

Discretas Contínuas

$\gamma(\omega) \rightarrow VA$

$\gamma(t) \rightarrow$ tempo dio

Process MARKOV \rightarrow AR(1)

✓
Time discrete \rightarrow Chain de MARKOV

$$\begin{matrix} & S & I & R \\ \left\{ \begin{matrix} S \\ I \\ R \end{matrix} \right. & \begin{pmatrix} 0,7 & 0,15 & 0,15 \end{pmatrix} \end{matrix} \rightarrow \underline{\Sigma = 1}$$

$$\omega(y(t)) = \{y(t-j)\}$$

OTRAS
VARIABLES

$$TC_t = f(TC_{t-1}, TC_{t-2}, \dots)$$

Παράδειγμα \rightarrow AR(2)

(no tiene
retardos)

↓
AUTOREGRESIVIDAD

\rightarrow METODO \leftarrow

Addition
 Addition

$$Y(t) = \underbrace{C(t)}_{\text{Cost}} + \underbrace{T(t)}_{\text{Tax}} + \underbrace{S(t)}_{\text{Sales}} + \underbrace{E(t)}_{\text{Error}}$$

Multiplication

$$Y(t) = C(t) * T(t) * S(t) * E(t)$$

$AR(0) \rightarrow$ SIN METODOS

I I I

INDEPENDENCIA

ESTACIONALIDAD \rightarrow COMPONENTE
SENIOR

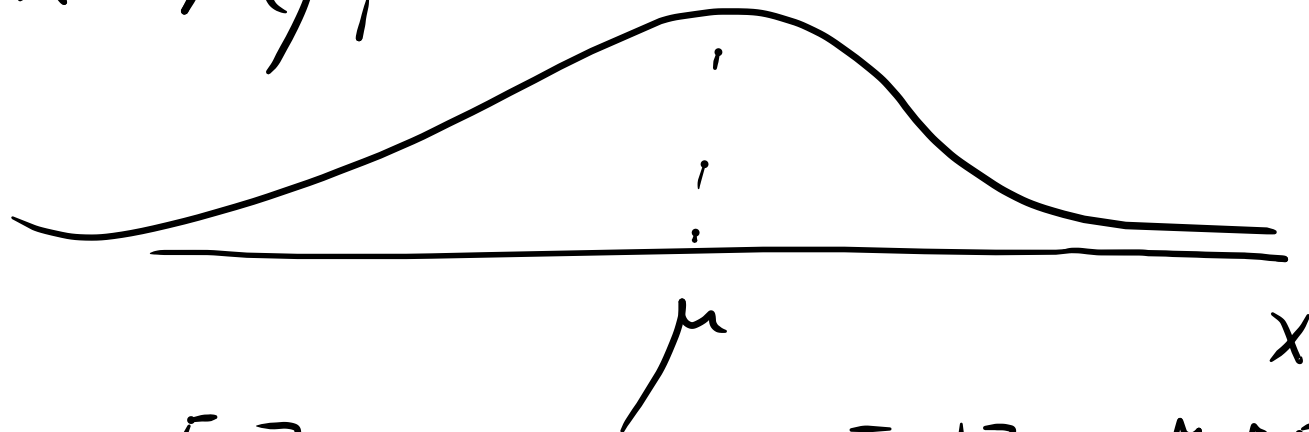
\neq
ESTACIONARIEDAD

ESTRUCTA

DÉBIL

SENIOR

$$X \sim \mathcal{N}(\mu; \sigma)$$



$$E[X] = \mu$$

$$\underbrace{\text{Var}[X] = \sigma^2}$$

$$E[X^k] \rightarrow \text{MOMENTS ABSOLUTE}$$

$$k=1 \rightarrow E[X]$$

$$k=2 \rightarrow E[X^2]$$

$$\text{Var}[X] = \underbrace{E[X^2]}_{\text{OTG}} - \underbrace{E^2[X]}_{\text{OTG}}$$

$$\text{Var}(\bar{x}) = E[(\underbrace{x - \bar{E}(x)}_{\text{}})^2]$$

→ Note the central
DE ORDER 2

$$AS(x) \rightarrow = \frac{mc_3}{\sigma^3}$$

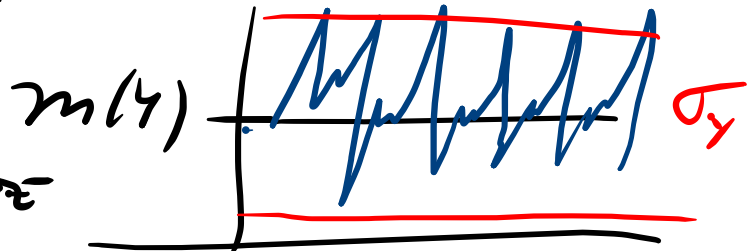
$$K(x) = \frac{mc_4}{\sigma^4} - 3$$

normal

ESTADÍSTICAS DE SEGUNDO
ORDEN

$E[y(t)]$ constante

$\sigma^2[y(t)]$ constante



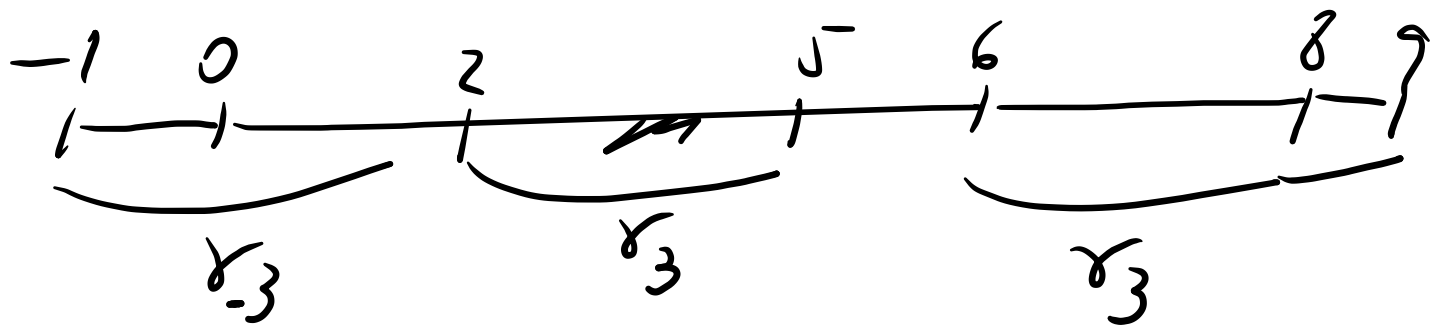
PARADOJA DE SAN PETERSBURGO
 $E[x] = \infty$

$$(i) m_1(Y_t) = m_1(Y) = m(Y)$$

$$(ii) E(Y_t^2) < \infty$$

$$(iii) \gamma(Y_{t-j}; Y_t) = \gamma_j(Y) = \gamma_{-j}(Y) \quad (j = 0, \pm 1, \pm 2, \dots)$$

γ
 $\overline{\text{Cov}(Y_{t-j}; Y_t)}$ ~~AUTO~~ COVARIANCES at lag j



Função de Autocorrelação

$$\gamma_j(Y) = E[(y_t - \underbrace{m(Y)}_{\text{média}})(y_{t-j} - \underbrace{m(Y)}_{\text{média}})] \quad (j = 0, 1, 2, \dots)$$

$$\text{cov}(x, y) = E[(\underbrace{x - E(x)}_{\text{desvio}})(\underbrace{y - E(y)}_{\text{desvio}})]$$

Propriedades de ordem 1º e 2º

$$\text{cov}(x, y) = E[xy] - E[x] \cdot E[y]$$

Se $x \perp y$ (independentes) $E[xy] = E[x]E[y]$

$$\Rightarrow \underline{\underline{\text{cov}(x, y) = 0}}$$

INDEPENDENCIA \Rightarrow INCOHERENCIA

$$\rho(x, y) = \frac{\overbrace{\text{Cov}(x, y)}^0}{\sigma_x \cdot \sigma_y}$$

$$\text{Si } \text{Cov}(x, y) = 0 \Rightarrow \rho(x, y) = 0$$

$$\gamma_0(Y) = E[(y_t - m(Y))(y_t - m(Y))] = E[\underbrace{(y_t - m(Y))^2}] = \sigma_Y^2$$

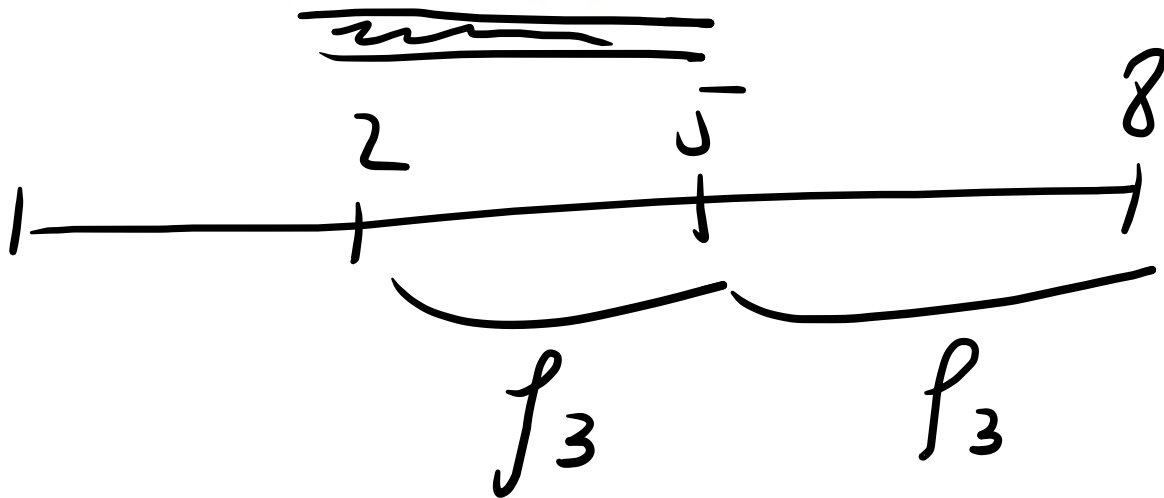
✓ AUTOCOVARIANZA DE
ERRORES CON

$$E[(x - \bar{e}(x))^2] = \sigma_x^2$$

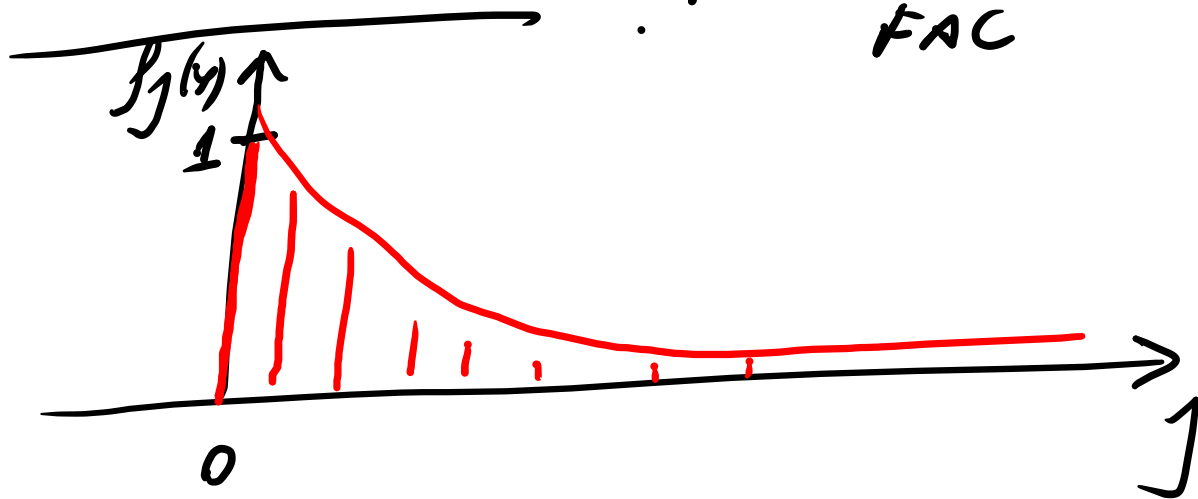
$$\rho_j(Y) = \frac{\gamma_j(Y)}{\gamma_0(Y)} \quad (j = 0, 1, 2, \dots)$$

Donde se cumple que $|\rho_j(Y)| \leq 1$

Se verifica que si $j = 0 \Rightarrow \rho_0(Y) = 1$



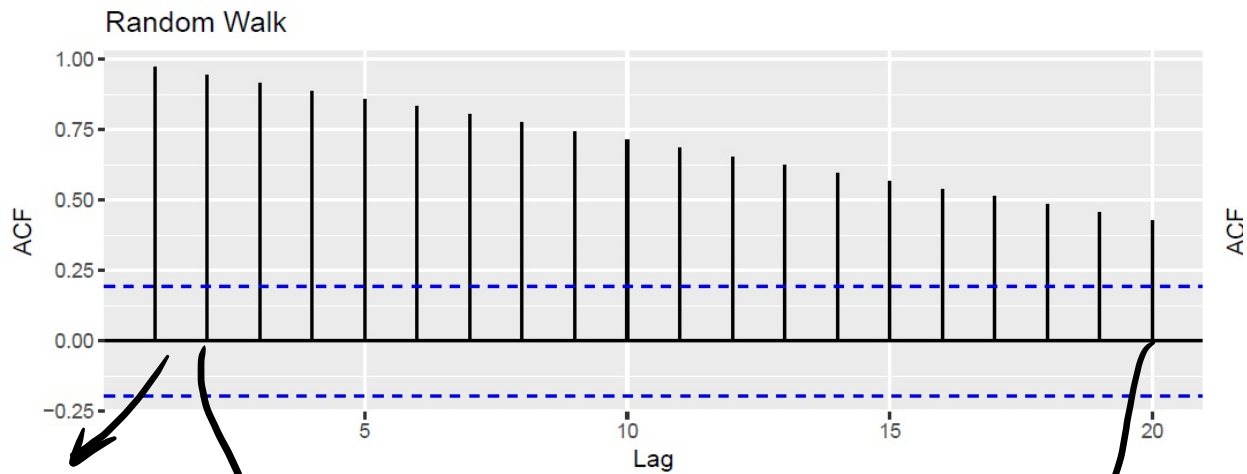
condensados \rightarrow gráfica de la
FAC



$$f_0(y) = \frac{\gamma_0(y)}{\gamma_0(y)} = 1$$

$$|f_j(y)| \leq 1$$

$$\underline{-1 \leq f_j(y) \leq 1}$$



$$\rho_0 = \frac{\sigma_0}{\sigma_0}$$

$$\rho_1 = \frac{\sigma_1}{\sigma_0}$$

$$\rho_0 = \frac{\sigma_0}{\sigma_0}$$

$$Y_t = C_t + T_t + S_t + \epsilon_t$$

NO ESTÁ CONNORMADA

ARMA(p, q)

$$\hat{Y}_t = Y_t - \hat{T}_t$$

SARIMA(p, d, q)

$$\hat{T}_t = \hat{\alpha} + \hat{\beta} t \quad \underline{\underline{MC}}$$

$$Y_t = Y_{t-1} + \epsilon_t \quad \text{Random walk (AR(1))}$$

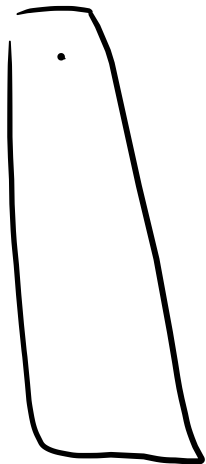
$$Y_t - Y_{t-1} = \epsilon_t$$

Widened
Blank

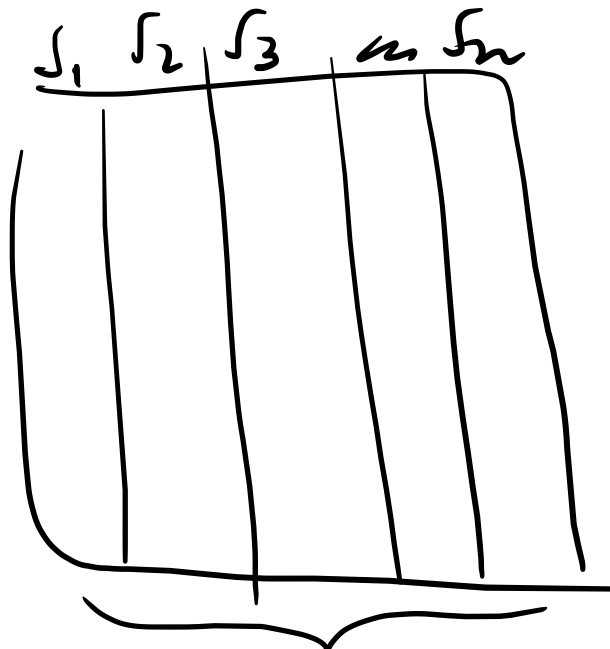
$$\nabla Y_t = \epsilon_t \quad \downarrow \text{Widened Blank}$$

$$E(\nabla Y_t) = E(\epsilon_t) = 0$$

$$\text{Var}(\nabla Y_t) = \text{Var}(\epsilon_t) = \sigma_\epsilon^2$$



fechas



ARRAYS DE SEMEIO

AS . XTS (ARRAY/VECTOR,
ORDER BY = fechas)

$$Y_t - Y_{t-12} = \underbrace{\nabla_{12}}_{\text{lag 12}} Y_t$$

$$\nabla \leftarrow \text{lag 1}$$

$$\nabla Y_t = \underline{Y_t} - Y_{t-1}$$

$$\xrightarrow{\text{diff}} \Delta Y_t = \underline{Y_{t+1}} - Y_t$$

$$\nabla \nabla_{12} Y_t =$$

$$\nabla_{12} Y_t - \nabla_{12} Y_{t-1}$$

12/216	100
08/224	410
03/221	430

Infusion

$$\frac{430}{410} - 1 = \%L$$