# BIOS 755: Linear Mixed Models II

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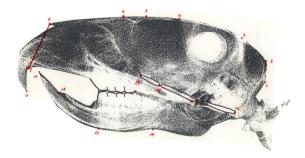
#### Rat Data

- ▶ Randomized experiment in which 50 male Wistar rats are randomized to:
  - ► Control (15 rats)
  - ► Low dose of Decapeptyl (18 rats)
  - ► High dose of Decapeptyl (17 rats)population.
- Question of interest: How does craniofacial growth depend on testosterone production?

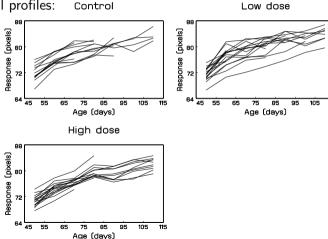
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### Rat Data

- ➤ Treatment starts at 45 days; measurements are taken every 10 days from day 50 on.
- ► The responses are distances (pixels) between well-defined points on x-ray pictures of the skull of each rat:



- ▶ We'll consider only one response: the height of the skull.
- Individual profiles: Control



## Models under consideration

Let's consider the model:

$$Y_{ij} = (\beta_0 + b_{0i}) + (\beta_1 L_i + \beta_2 H_i + \beta_3 C_i + b_{1i}) t_{ij} + \varepsilon_{ij}$$

$$= \begin{cases} (\beta_0 + b_{0i}) + (\beta_1 + b_{1i}) t_{ij} + \varepsilon_{ij}, & \text{if low dose} \\ (\beta_0 + b_{0i}) + (\beta_2 + b_{1i}) t_{ij} + \varepsilon_{ij}, & \text{if high dose} \\ (\beta_0 + b_{0i}) + (\beta_3 + b_{1i}) t_{ij} + \varepsilon_{ij}, & \text{if Control} \end{cases}$$

Where the covariance of the random effects is

$$oldsymbol{D} = cov(oldsymbol{b}) = \left(egin{array}{cc} d_{11} & d_{12} \ d_{12} & d_{22} \end{array}
ight)$$

# Linear Mixed representation

- ▶ What are the X and Z from the linear mixed model that corresponds to this model?
- ▶ What is the implied mean structure?
- ▶ What is the implied variance of  $Y_{ij}$ ?

# Linear Mixed representation

▶ What is the implied marginal Variance of  $Y_{ij}$ ?

$$Var(Y_{ij}) = \left(1 \ t_{ij}\right) D \left(\frac{1}{t_{ij}}\right) + \sigma^2$$
$$= \left(d_{11} + \sigma^2\right) + 2d_{12}t_{ij} + d_{22}t_{ij}^2$$

where

# **Analysis**

► The following model was fitted to the data

$$Y_{ij} = (\beta_0 + b_{1i}) + (\beta_1 L_i + \beta_2 H_i + \beta_3 C_i + b_{2i})t_{ij} + \varepsilon_{ij}$$

▶ The REML estimates obtained from PROC Mixed are:

Effect	Parameter	REMLE (s.e.)
Intercept	$\beta_0$	68.606 (0.325)
Time effects:		, ,
Low dose	$eta_1$	7.503 (0.228)
High dose	$eta_2$	6.877 (0.231)
Control	$\beta_3$	7.319 (0.285)
Covariance of $b_i$ :		
$var(b_{1i})$	$d_{11}$	3.369 (1.123)
$var(b_{2i})$	$d_{22}$	0.000 ( — )
$cov(b_{1i}, b_{2i})$	$d_{12} = d_{21}$	0.090 (0.381)
Residual variance:		
$var(arepsilon_{ij})$	$\sigma^2$	1.445 (0.145)
REML log-likelihood		-466.173

# **Analysis**

- This suggests that the REML likelihood could be further increased by allowing negative estimates for  $d_{22}$ .
- In SAS, this can be done by adding the option nobound to the PROC MIXED statement.

  Parameter restrictions for  $\alpha$

► Results:

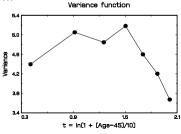
		$d_{ii} \ge 0, \sigma^2 \ge 0$	$d_{ii} \in I\!\!R, \sigma^2 \in I\!\!R$
Effect	Parameter	REMLE (s.e.)	REMLE (s.e.)
Intercept	$\beta_0$	68.606 (0.325)	68.618 (0.313)
Time effects:			
Low dose	$eta_1$	7.503 (0.228)	7.475 (0.198)
High dose	$eta_2$	6.877 (0.231)	6.890 (0.198)
Control	$eta_3$	7.319 (0.285)	7.284 (0.254)
Covariance of $b_i$ :			
$var(b_{1i})$	$d_{11}$	3.369 (1.123)	2.921 (1.019)
$var(b_{2i})$	$d_{22}$	0.000 ( —)	-0.287 (0.169)
$cov(b_{1i}, b_{2i})$	$d_{12} = d_{21}$	0.090 (0.381)	0.462 (0.357)
Residual variance:			
$var(arepsilon_{ij})$	$\sigma^2$	1.445 (0.145)	1.522 (0.165)
REML log-likelihood		-466.173	-465.193

## Meaning of a negative variance component

► Fitted variance function

$$Var(Y_{ij}) = (\hat{d}_{11} + \hat{\sigma}^2) + 2\hat{d}_{12}t_{ij} + \hat{d}_{22}t_{ij}^2$$
  
= 4.443 + 0.924t - 0.287t<sub>ij</sub><sup>2</sup>

► The suggested negative curvature in the variance function is supported by the sample variance function:



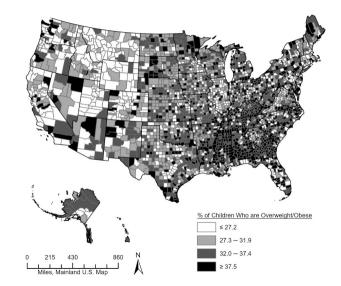
# Inference for Linear Mixed Models

## Introduction

- lacktriangle In most applications, inference is focused on the fixed effects, eta .
- However, in some studies we may want to predict (or "estimate") subject-specific response profiles.
- ightharpoonup Technically, because the  $b_i$ 's are random, we customarily talk of "predicting" the random effects rather than "estimating" them.
- ► The predicted random effects can be used to predict values for all levels in the data.

# Example: County level disease rates/summaries.

- ► The National Study of Children's Health (NSCH) gathers data from roughly 50K 2–17 year-old children at each survey.
- The NSCH variables include indicators of ADHD, ASD, and many other conditions.
- ▶ It also includes BMI percentile, and other continuous variables.
- ► A regression model with a random county level intercept was fitted.
- ► To predict (estimate) county level values, we needed to predict the value of the random effect.



# Conditional Expectation

For a RE ANOVA model on  $Y_{i1}, Y_{i2}, \ldots, Y_{in_i}$  we have

$$oldsymbol{Y}_{ij} = \mu + b_i + e_{ij}$$

with  $b_i \sim N(0, G)$  and  $e_i \sim N(0, \sigma^2)$ .

Under this model, the predicted value of the random intercept is

$$E(b_i|Y_i) = \frac{n_i G}{n_i G + \sigma^2} (\bar{Y}_i - \mu)$$

where  $Var(b_i) = G$ 

# Best Linear Unbiased Predictor (BLUP)

When  $\Sigma_i$  is known, the estimator for  $\boldsymbol{\beta}$  can be obtained by using ML weighted least square. Then the prediction of  $\boldsymbol{b}_i$  is given by

$$GZ_i'\Sigma_i^{-1}(Y_i-X_i\hat{\boldsymbol{\beta}}),$$

where

$$\Sigma_i = \mathsf{var}(oldsymbol{Y}_i) = oldsymbol{Z}_i oldsymbol{GZ}_i' + oldsymbol{R}_i$$

▶ This is known as the Best Linear Unbiased Predictor (or BLUP).

## The BLUP Estimation of Individual Mean

Finally, the ith subject's predicted response profile is,

$$\hat{\mathbf{Y}}_{i} = \mathbf{X}_{i}\hat{\boldsymbol{\beta}} + \mathbf{Z}_{i}\hat{\mathbf{b}}_{i} 
= \mathbf{X}_{i}\hat{\boldsymbol{\beta}} + \mathbf{Z}_{i}\hat{\mathbf{G}}\mathbf{Z}_{i}'\hat{\boldsymbol{\Sigma}}_{i}^{-1}(\mathbf{Y}_{i} - \mathbf{X}_{i}\hat{\boldsymbol{\beta}}) 
= (\hat{\mathbf{R}}\hat{\boldsymbol{\Sigma}}^{-1})\mathbf{X}_{i}\hat{\boldsymbol{\beta}} + (\mathbf{I} - \hat{\mathbf{R}}_{i}\hat{\boldsymbol{\Sigma}}_{i}^{-1})\mathbf{Y}_{i}$$

That is, the *i*th subject's predicted response profile is a weighted combination of the population-averaged mean response profile,  $\mathbf{X}_i\hat{\boldsymbol{\beta}}$ , and the *i*th subject's observed response profile  $\mathbf{Y}_i$ .

## Example: Country level rates of malnutrition.

- ► Ending malnutrition is a key outcome in the Sustainable Development Goal (SDG)
- ► The 2nd SDG calls for achieving, by 2025, a reduction of stunting and wasting and halt the rise in overweight in children under 5 years of age.
- Monitoring countries' progress toward the achievement of their SDG targets is an important task, but data sparsity makes monitoring trends challenging.
- The model fitted for this analysis was

$$Y_{ijk} = \beta' \mathbf{X}_{ijk} + \mathbf{b}_i' \mathbf{B}_{ij} + \mathbf{b}_{ij}' \mathbf{B}_{ijk} + \epsilon_{ij}, \tag{1}$$

for region i, country j, and observation k.

