# Linear Discriminant Analysis (part 2)

#### ACM

September 25, 2023

## Reanalysis of Breast Cancer Data Example

In this example we are going to perform a reanalysis of the breast cancer data taking more care towards the (possibility of) non-normality in the results.

In the previous analysis of this data we obtained the following error rates:

LDA: 25.3%Logistic: 26.8%QDA: 28.3%

Logistic won't be re-ran. For LDA and QDA we will do the following:

- LDA: re-run LDA choosing the cutpoing  $(b_0)$  with cross validation (CV) using the same X,
- LDA: re-run LDA after the data have been transformed to have a more normal distribution, and
- QDA: re-run QDA after the data have been transformed to have a more normal distribution.

First, let's read in the data.

```
library(MASS)
wdbct <- read.csv("wpbc.csv")
head(wdbct[, 1:5])</pre>
```

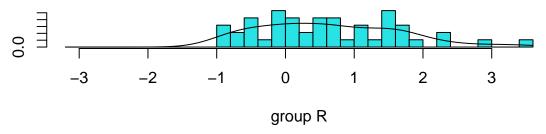
ID	Outcome	Time	$radius\_M$	$texture\_M$
119513	N	31	18.02	27.60
8423	N	61	17.99	10.38
842517	N	116	21.37	17.44
843483	N	123	11.42	20.38
843584	$\mathbf{R}$	27	20.29	14.34
843786	R	77	12.75	15.29

```
X <- matrix(as.numeric(unlist(wdbct[, 4:32])), 198, 29)
Y <- as.factor(wdbct[, 2])
n <- length(Y)</pre>
```

### Plots of the results from LDA/QDA

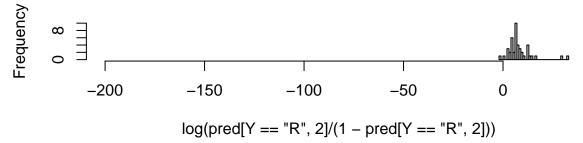
For LDA, here the scores will include the intercept value.

# plot(fit, dimen=1, type="both") -3 -2 -1 0 1 2 3 group N

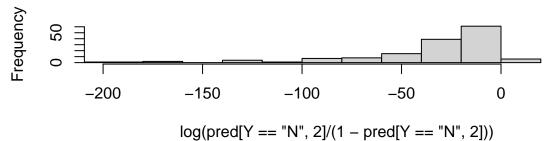


For QDA (a little harder), here the scores will include the intercept value. Notice these are plotted so that they have the same x-axis. This is important towards being able to compare them.

# Histogram of log(pred[Y == "R", 2]/(1 - pred[Y == "R", 2]))



## Histogram of log(pred[Y == "N", 2]/(1 - pred[Y == "N", 2]))



If you want to explore plotting the results a little further, see the klaR package.

#### LDA with $b_0$ choosen via CV

This analysis will choose the intercept using K-fold CV. How this will work:

- 1. Split the data into K sections.
- 2. Get candidate intercept values  $b_{0l}$  for l = 1, 2, ..., L.
- 3. For section k = 1, 2, ..., K being left out we will:
  - a. fit the LDA on the training data,
  - b. predict the outcome for the test data with all  $b_{0l}$  values,
  - c. get the error for all  $b_{0l}$  values, and
  - d. find the optimal intercept  $b_0^k$ .
- 4. Find the mean optimal intercept  $\bar{b}_0 = \frac{1}{K} \sum_{k=1}^{K} b_0^k$ .
- 5. Refit LDA to the whole data, predicted values will be based on  $\bar{b}_0$ .

 $Step\ one:$ 

```
set.seed(123)
K <- 10
CV_ids <- sample(rep(1:K, ceiling(n/K)), n, replace = FALSE)</pre>
```

Step two:

For the candidate values recall that

$$\log \frac{\Pr(G = l | X = x)}{\Pr(G = \ell | X = x)} = \left\{ \log \frac{\pi_k}{\pi_\ell} - \frac{1}{2} (\mu_k + \mu_\ell)' \Sigma^{-1} (\mu_k - \mu_\ell) \right\} + x' \Sigma^{-1} (\mu_k - \mu_\ell)$$

So, we want to center candidate intercept values around  $\log \frac{\pi_k}{\pi_\ell} - \frac{1}{2}(\mu_k + \mu_\ell)'\Sigma^{-1}(\mu_k - \mu_\ell)$ , but we don't know what this is. As a result, we'll have to find it.

Let's take a look at what we get for the non-intercept portion  $b'_1X$  for the whole data. This is okay to do since this is just giving us candidate  $b_0$  values. We're not keeping any parameter estimates.

Here's  $b_1'X$  for the full data

```
fit <- lda(Y ~ X)
summary(c(X %*%fit$scaling))</pre>
```

Min.	1st Qu.	Median	Mean	3rd Qu.	Max.
-14.28188	-12.18136	-11.54673	-11.45418	-10.78047	-7.705099

Giving us our candidate values:

```
b0_seq <- seq(-16,-5,0.01)
```

Step three:

part a:

Fit to the data

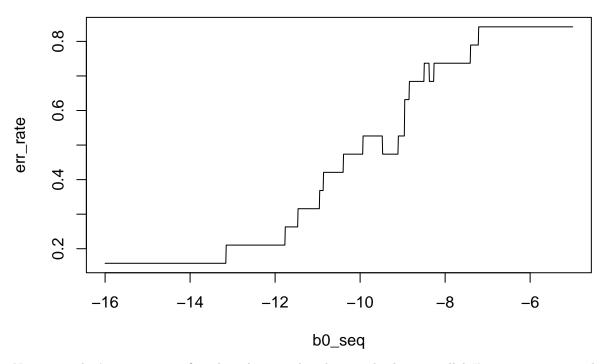
```
k = 1
X_tr <- X[CV_ids != k,]
Y_tr <- Y[CV_ids != k]

X_tst<- X[CV_ids == k,]
Y_tst<- Y[CV_ids == k]

fit_tr <- lda(Y_tr ~ X_tr)</pre>
```

part b and c: Predict for the test data with all  $b_{0l}$  values get error rates

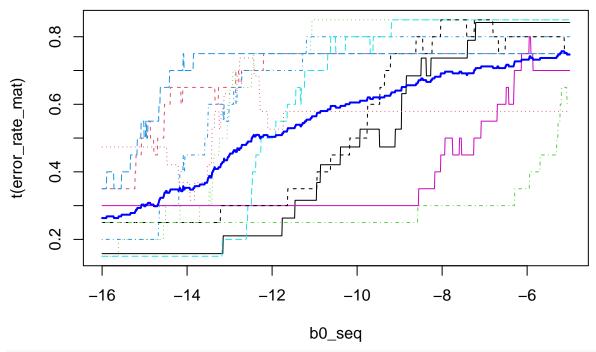
```
err_rate <- NULL
for(1 in 1:length(b0_seq)){
    Y_pred_fct = rep("R",length(Y_tst))
    ## c(X_tst %*%fit_tr$scaling) > b0_seq[1]
    ## TRUE = "N" and FALSE = "R"
    Y_pred_fct[c(X_tst %*%fit_tr$scaling) > b0_seq[1]] <- "N"
    ## Error rate for this b0
    t_err_rate <- mean(Y_pred_fct != Y_tst)
    ## Keep the results
    err_rate <- c(err_rate, t_err_rate)
}
plot(b0_seq, err_rate, type = "l")</pre>
```



Notice, we don't get one specific value. As a result, when we do this over all k I'm going to retain the error rates for all candidate  $b_0$  values and not just the optimal value.

Now, we'll put this all together.

```
# Full error rate matrix
error_rate_mat <- NULL</pre>
for(k in 1:K){
X_tr <- X[CV_ids != k,]</pre>
Y_tr <- Y[CV_ids != k]</pre>
X_{tst} - X[CV_{ids} == k,]
Y_tst<- Y[CV_ids == k]
fit_tr <- lda(Y_tr ~ X_tr)</pre>
err_rate <- NULL</pre>
for(l in 1:length(b0_seq)){
  Y_pred_fct = rep("R",length(Y_tst))
  ## TRUE = "N" and FALSE = "R"
  Y_pred_fct[c(X_tst %*%fit_tr$scaling) > b0_seq[1]] <- "N"
  ## Error rate for this b0
  t_err_rate <- mean(Y_pred_fct != Y_tst)</pre>
  ## Keep the results
  err_rate <- c(err_rate, t_err_rate)</pre>
## Append the results for this fold onto the rest of the results.
error_rate_mat <- rbind(error_rate_mat,</pre>
                          err_rate)
}
matplot(b0_seq, t(error_rate_mat), type = "1")
avg_err_rate <- apply(error_rate_mat, 2, mean)</pre>
lines(b0_seq, avg_err_rate, lwd=2, col="blue")
```



b0\_seq[which(avg\_err\_rate==min(avg\_err\_rate))]

```
## [1] -16.00 -15.99 -15.98 -15.97 -15.96 -15.95 -15.94 -15.93 -15.92 -15.91
## [11] -15.90 -15.72 -15.71 -15.70 -15.69 -15.68 -15.67 -15.66 -15.65 -15.64
## [21] -15.63 -15.62
```

However, we know these values are smaller than the smallest  $b'_1X$  when ran with the full data.

What does this tell us about the best predictions? What does this tell us about the X data?

```
mean(Y == "R")
```

#### ## [1] 0.2373737

In the previous analysis of this data we obtained the following error rates:

LDA: 25.3%Logistic: 26.8%QDA: 28.3%

Note that instead of minimizing the error rate we could minimize the "total cost" where we may have

```
t_cost <- 3*I(Y_pred_fct == "N" & Y_tst == "R") +
1*I(Y_pred_fct == "R" & Y_tst == "N")</pre>
```

## Finding "more Normal" predictors

For this re-analysis, we'll focus on getting the data to look more like a normal distribution.

While there are more sophisticated methods, here, I'm just going to identify predictors that could use a log transformation and transform them.

I identified right-skewed predictors via histograms (not shown):

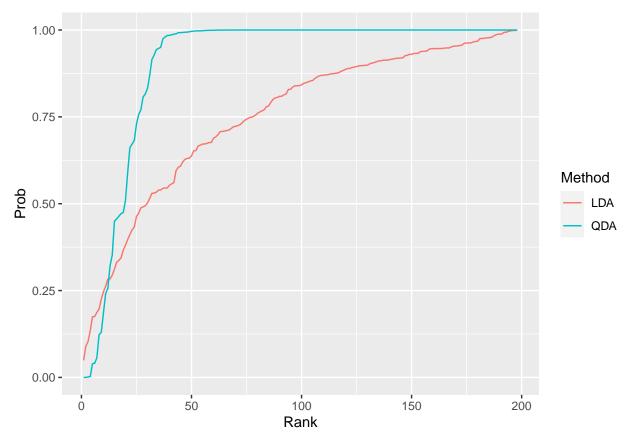
Based on these, I transformed the following variables

```
trans_ids <- c(10:20, 24, 26, 29)
X_trans <- X
X_trans[,trans_ids] <- log(X[,trans_ids])</pre>
```

Then I rechecked the data, and was happy.

```
par(mfrow = c(5,6))
for(k in 1:29){
  hist(X_trans[,k],
      main = paste0("Histogram of X",k),
      breaks = 30)
}
```

Now, we'll refit LDA and QDA to the transformed data and see if we do better.



Assess the accuracy of the prediction percent correct for each category of K.

```
ct <- table(WDB_DF$Y, rep("N",n))
ct</pre>
```

First for LDA with new  $b_0$ :

$$\begin{array}{c|c} \underline{/} & N \\ \hline N & 151 \\ R & 47 \end{array}$$

```
prop.table(ct)
```

/	N
N	0.7626263
R	0.2373737

```
tct <- table(WDB_DF$Y, fit$class)
tct</pre>
```

Second for LDA with transformed data:

/	N	R
N	135	16
R	34	13

#### prop.table(tct)

/	N	R
N R	0.6818182 0.1717172	0.0808081 $0.0656566$

```
qt <- table(WDB_DF$Y, q_fit$class)
qt</pre>
```

#### Third for QDA:

```
/ N R
N 137 14
R 42 5
```

#### prop.table(qt)

/	N	R
N	0.6919192	0.0707071
R	0.2121212	0.0252525

#### Total error rate for all three methods:

```
LDA_acc <- 1-sum(diag(prop.table(ct)))
LDAtr_acc <- 1-sum(diag(prop.table(tct)))
QDA_acc <- 1-sum(diag(prop.table(qt)))

res <- matrix(c(LDA_acc, LDAtr_acc, QDA_acc))
rownames(res) <- c("LDA new b0", "LDA trans", "QDA trans")
colnames(res) <- "N-fold Error Rate"
res</pre>
```

	N-fold Error Rate
LDA new b0	0.2373737
LDA trans	0.2525253
QDA trans	0.2828283

In the previous analysis of this data we obtained the following error rates:

LDA: 25.3%Logistic: 26.8%QDA: 28.3%