17.5 Correlated Process and Observation Noise

We shall follow a development from Brown & Hwang (1997), pages 296–298 and start by explicitly stating a somewhat changed version of the usual covariance conditions

$$E\{\boldsymbol{\epsilon}_{k}\boldsymbol{\epsilon}_{j}^{\mathrm{T}}\} = \begin{cases} \Sigma_{\epsilon,k} & \text{if } j = k \\ 0 & \text{if } j \neq k \end{cases}$$

$$E\{\boldsymbol{e}_{k}\boldsymbol{e}_{j}^{\mathrm{T}}\} = \begin{cases} \Sigma_{e,k} & \text{if } j = k \\ 0 & \text{if } j \neq k \end{cases}$$

$$E\{\boldsymbol{\epsilon}_{k}\boldsymbol{e}_{j}^{\mathrm{T}}\} = C_{k} & \text{for all } j \text{ and } k.$$

The last equation states that process noise ϵ_k and observation noise e_k are correlated. We want to study the modifications introduced by this change to the usual filter equations.

Filtering of the state vector:

$$\hat{x}_{k|k} = \hat{x}_{k|k-1} + K_k(b_k - A_k \hat{x}_{k|k-1}).$$

The system equation is

$$\boldsymbol{x}_k = F_{k-1}\boldsymbol{x}_{k-1} + \boldsymbol{\epsilon}_k$$

and hence the estimation error

$$v_{k|k} = x_k - \hat{x}_{k|k}$$

$$= x_k - (\hat{x}_{k|k-1} + K_k(b_k - A_k \hat{x}_{k|k-1}))$$

$$= (I - K_k A_k)(x_k - \hat{x}_{k|k-1}) - K_k e_k$$

$$= (I - K_k A_k)v_{k|k-1} - K_k e_k$$

with covariance matrix $P_{k|k} = E\{v_{k|k}v_{k|k}^{\mathrm{T}}\}$.

Now we compute the covariance between $v_{k|k-1}$ and e_k :

$$E\{v_{k|k-1}e_k^{\mathrm{T}}\} = E\{(x_k - \hat{x}_{k|k-1})e_k^{\mathrm{T}}\}\$$

= $E\{(F_{k-1}x_{k-1} + \epsilon_k - F_{k-1}\hat{x}_{k-1|k-1})e_k^{\mathrm{T}}\}.$

 e_k is neither correlated with x_{k-1} nor $\hat{x}_{k-1|k-1}$ due to their whiteness. Therefore we get

$$E\{\boldsymbol{v}_{k|k-1}\boldsymbol{e}_{k}^{\mathrm{T}}\}=E\{\boldsymbol{\epsilon}_{k}\boldsymbol{e}_{k}^{\mathrm{T}}\}=C_{k}.$$

We return to the main development. The covariance matrix for the estimation error is

$$P_{k|k} = E\{\mathbf{v}_{k|k}\mathbf{v}_{k|k}^{T}\}$$

$$= E\{(I - K_{k}A_{k})\mathbf{v}_{k|k-1} - K_{k}\mathbf{e}_{k})(I - K_{k}A_{k})\mathbf{v}_{k|k-1} - K_{k}\mathbf{e}_{k})^{T}\}$$

$$= (I - K_{k}A_{k})P_{k|k-1}(I - K_{k}A_{k})^{T} + K_{k}\Sigma_{e,k}K_{k}^{T}$$

$$- (I - K_{k}A_{k})C_{k}K_{k}^{T} - K_{k}C_{k}^{T}(I - K_{k}A_{k})^{T}.$$
(17.1)

This is a general expression for the error covariance matrix and valid for any gain matrix K_k . The two last terms are introduced due to the cross-correlation C_k . For $C_k = 0$ we recover the earlier Joseph form of the variance.

We want to rewrite (17.1), and collect a factor from the first and fourth terms and manipulate the other terms:

$$P_{k|k} = P_{k|k-1} - K_k (A_k P_{k|k-1} + C_k^{\mathrm{T}}) + K_k (A_k P_{k|k-1} A_k^{\mathrm{T}} + \Sigma_{e,k} + A_k C_k + C_k^{\mathrm{T}} A_k^{\mathrm{T}}) K_k^{\mathrm{T}} - (P_{k|k-1} A_k^{\mathrm{T}} + C_k)^{\mathrm{T}} K_k^{\mathrm{T}}$$
(17.2)

The Kalman matrix associated with a minimum estimation error is determined by setting $\frac{\partial \operatorname{tr}(P_{k|k})}{\partial K_i} = 0$ or

$$-2(A_k P_{k|k-1})^{\mathrm{T}} + 2K_k A_k P_{k|k-1} A_k^{\mathrm{T}} + 2K_k \Sigma_{e,k} - 2C_k + 2K_k A_k C_k + 2K_k C_k^{\mathrm{T}} A_k^{\mathrm{T}} = 0.$$

This yields

$$K_k = (P_{k|k-1}A_k^{\mathrm{T}} + C_k)(A_k P_{k|k-1}A_k^{\mathrm{T}} + \Sigma_{e,k} + A_k C_k + C_k^{\mathrm{T}} A_k^{\mathrm{T}})^{-1}.$$
 (17.3)

Let $C_k \to 0$ and this equation reduces to the gain in the zero cross-correlation model

$$K_k = P_{k|k-1} A_k^{\mathrm{T}} (A_k P_{k|k-1} A_k^{\mathrm{T}} + \Sigma_{e,k})^{-1}$$
(17.4)

as it should be. Finally we insert (17.3) into (17.2) and get

$$P_{k|k} = (I - K_k A_k) P_{k|k-1} - K_k C_k^{\mathrm{T}}.$$
(17.5)

The prediction equations are not affected by the cross-correlation between ϵ_k and e_k because of the whiteness property of each.

A possible correlation between the process noise ϵ_k and the observation noise e_k leads to modifications of the Kalman gain matrix K_k and the expression for the filter covariance matrix $P_{k|k}$. The modifications involving terms with C_k are in (17.3) and (17.5).

Useful partial derivatives of traces and the determinant:

$$\frac{\partial \operatorname{tr}(A)}{\partial A} = I$$

$$\frac{\partial \operatorname{tr}(AB)}{\partial A} = \begin{cases} B^{\mathrm{T}} & \text{if } A \text{ is asymmetric} \\ B + B^{\mathrm{T}} - \operatorname{diag}(B), & \text{if } A \text{ is symmetric} \end{cases}$$

$$\frac{\partial \operatorname{tr}(BAC)}{\partial A} = B^{\mathrm{T}}C^{\mathrm{T}}$$

$$\frac{\partial \operatorname{tr}(ABA^{\mathrm{T}})}{\partial A} = A(B + B^{\mathrm{T}})$$

$$\frac{\partial \operatorname{tr}(e^{A})}{\partial A} = e^{A^{\mathrm{T}}}$$

$$\frac{\partial |BAC|}{\partial A} = |BAC|(A^{-1})^{\mathrm{T}}$$

REFERENCES

Brown, Robert Grover & Hwang, Patrick Y. C. (1997). *Introduction to Random Signals and Applied Kalman Filtering*. John Wiley & Sons, Inc., New York, 3rd edition.