

# STA4273 HW 2

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## 6.1

Question 6.1 asks to compute a Monte Carlo estimate of  $\int_0^{\pi/3} \sin(t) dt$

```
m <- 10000
x <- runif(m, 0, pi/3)
theta_hat <- (pi/3)*(mean(sin(x)))
theta <- -cos(pi/3) - (-1)
```

With  $\theta = 0.5$  and  $\hat{\theta} = 0.495228$ , our estimate is 0.95% off from the true value.

## 6.3

Question 6.3 asks to compute Monte Carlo estimate of  $\int_0^{0.5} e^{-x} dx$  by first using the Uniform distribution (where the estimate is denoted  $\hat{\theta}$ ) and then using the exponential distribution ( $\hat{\theta}^*$ ). Afterwards, we compare the variances to see which is a better estimator.

```
theta_hat_sampler <- replicate(1000, expr = {
  x <- runif(m, 0, 0.5)
  theta_hat <- 0.5*mean(exp(-x))
})

theta_star_sampler <- replicate(1000, expr = {
  g <- rexp(m)
  theta_star <- mean(g <= 0.5)
})
var(theta_hat_sampler)
```

```
## [1] 3.346701e-07
```

```
var(theta_star_sampler)
```

```
## [1] 2.349377e-05
```

It appears that  $\hat{\theta}$  has a smaller variance than  $\hat{\theta}^*$ . Verifying this mathematically:

$$\begin{aligned} \text{Var}(\hat{\theta}) &= \frac{1}{4m} (\int_0^{0.5} 2e^{-2u} du - (\int_0^{0.5} 2e^{-u} du)^2) \approx \frac{0.013}{4m} \\ \text{Var}(\hat{\theta}^*) &= \frac{(1-e^{-0.5})(e^{-0.5})}{m} \approx \frac{0.2386}{m} \\ \frac{\text{Var}(\hat{\theta})}{\text{Var}(\hat{\theta}^*)} &= \frac{0.013}{4*0.2386} < 1 \end{aligned}$$

The ratio is less than one, thus  $\hat{\theta}$  is a more efficient estimator.

## 6.4

Question 6.4 asks to write a function to compute a Monte Carlo estimate of the Beta(3, 3) cdf, and use the function to estimate  $F(x)$  for  $x = 0.1, 0.2, \dots, 0.9$ .

```
estimates <- c()
true_prob <- c()
values <- seq(0.1,0.9, 0.1)
cdf_computer <- function(n, p) {
  x <- runif(n, min = 0, max = p)
  ### the function inside mean() is a simplified form of the beta(3,3) pdf
  p*mean((x^2*(1-x)^2*30))
}

for (i in values) {
  estimates <- append(estimates, cdf_computer(100000, i))
  true_prob <- append(true_prob, pbeta(i, 3,3))
}
percent_error <- 100*(estimates - true_prob)/true_prob
round(rbind(values, true_prob, estimates, percent_error), 4)
```

```
##           [,1] [,2] [,3] [,4] [,5] [,6] [,7] [,8]
## values    0.1000 0.2000 0.3000 0.4000 0.5000 0.6000 0.7000 0.8000
## true_prob  0.0086 0.0579 0.1631 0.3174 0.5000 0.6826 0.8369 0.9421
## estimates  0.0086 0.0580 0.1638 0.3177 0.5002 0.6825 0.8352 0.9418
## percent_error 0.0507 0.0810 0.4395 0.0706 0.0393 -0.0058 -0.2035 -0.0338
##           [,9]
## values    0.9000
## true_prob  0.9914
## estimates  0.9905
## percent_error -0.0981
```

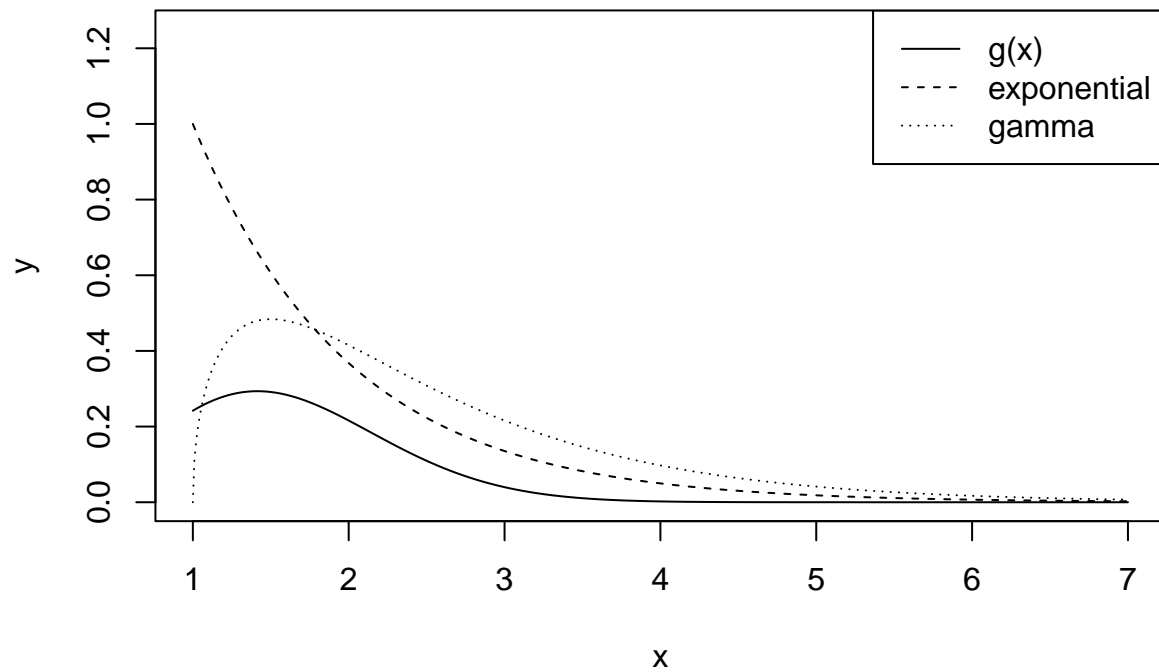
Our estimates seem reasonably close to the true probability.

## 6.13

6.13 asks to find two functions to estimate  $\int_1^\infty \frac{x^2}{\sqrt{2\pi}} e^{-x^2} dx$ , then to compare the functions to find the better estimator.

The two functions we will look at are the exponential distribution, and the gamma distribution (both shifted by 1 in order to match the support). Comparing those to the given function:

```
x <- seq(1,7,0.01)
y <- x^2 * exp(-x^2/2)/sqrt(2*pi)
yexp <- dexp(x-1)
ygam <- dgamma(x-1, 3/2, 1)
plot(x,y, type = "l", ylim = c(0,1.25))
lines(x, yexp, lty = 2)
lines(x, ygam, lty = 3)
legend("topright", legend = c("g(x)", "exponential", "gamma"), lty = 1:3)
```

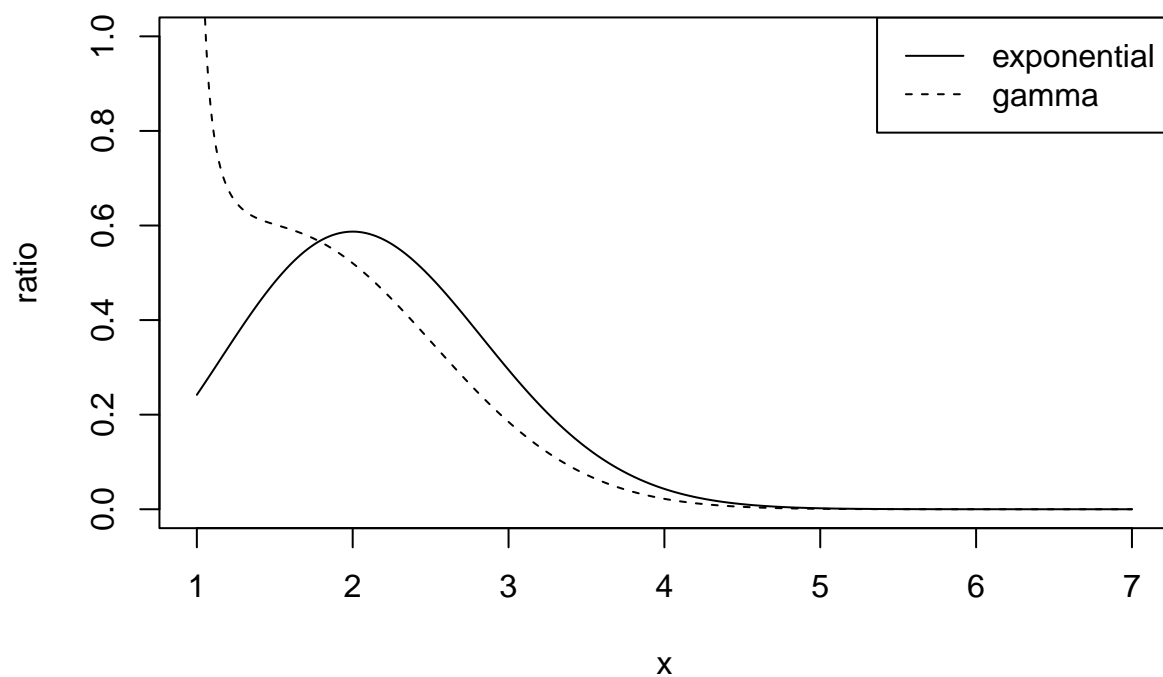


Then creating functions to compare the ratios to minimize errors:

```

expratio <- y/yexp
gamratio <- y/ygam
plot(x, expratio, type = "l", ylim = c(0,1), ylab = "ratio")
lines(x, gamratio, lty = 2)
legend("topright", legend = c("exponential", "gamma"), lty = c(1,2))

```



It seems the modified exponential function would be the better fit, since the importance function seems more “constant”.

## 6.14

6.14 asks to find the estimate of the integral referenced in 6.13

Using the modified exponential distribution to sample:

```
x <- rexp(10000) + 1
est <- mean((x^2 * exp(-x^2/2))/sqrt(2*pi)/dexp(x-1))
```

Our estimate for the integral is 0.4024391