

# Joint Sparse Estimation with Cardinality Constraint via Mixed-Integer Semidefinite Programming

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## Introduction

- The multiple measurement vectors (MMV) problem aims at jointly estimating multiple signals that share a common sparse support over a known dictionary
- Proposal of a mixed-integer semidefinite programming (MISDP) reformulation for the original formulation of MMV with  $\ell_{2,0}$ -norm constraint

## Mathematical Model and Notations

- Distinct Directions-of-Arrival (DOAs)  $\theta = [\theta_1, \dots, \theta_L]^T$  from  $L$  sources
- Equivalent expression of DOAs in spatial frequencies  $\mu = [\mu_1, \dots, \mu_L]^T$  with  $\mu_l = \pi \cos(\theta_l) \in [-\pi, \pi]$  for  $l = 1, \dots, L$

- Signal Model

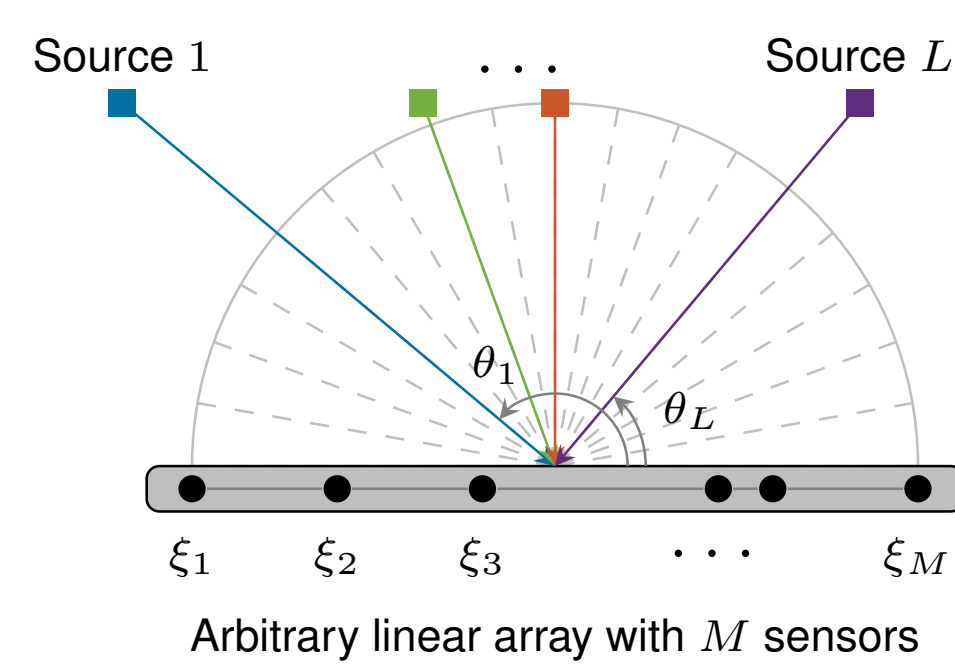
$$\mathbf{Y} = \mathbf{A}(\mu)\Psi + \mathbf{N}$$

$\mathbf{Y} \in \mathbb{C}^{M \times N}$  : Received signal matrix

$\Psi \in \mathbb{C}^{L \times N}$  : Source signal matrix

$\mathbf{N} \in \mathbb{C}^{M \times N}$  : Sensor noise matrix

$N$  : Number of available snapshots



- Steering matrix  $\mathbf{A}(\mu) = [\mathbf{a}(\mu_1), \dots, \mathbf{a}(\mu_L)] \in \mathbb{C}^{M \times L}$  with  $\mathbf{a}(\mu_l) = [e^{j\mu_l \xi_1}, \dots, e^{j\mu_l \xi_M}]^T$  for  $l = 1, \dots, L$

## DML and MAP Estimators

### Deterministic Maximum Likelihood (DML) Estimator

- Deterministic source signals and spatio-temporal white Gaussian noise assumption  $\mathbf{n}_i \sim \mathcal{CN}(\mathbf{0}, \sigma^2 \mathbf{I}_M) \implies \mathbf{y}_i | \psi_i \sim \mathcal{CN}(\mathbf{A}(\mu)\psi_i, \sigma^2 \mathbf{I}_M)$  for each snapshot  $i$
- Maximization of the likelihood function

$$\max_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \prod_{n=1}^N p(\mathbf{y}_n | \psi_n) \iff \min_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \|\mathbf{A}(\mu)\Psi - \mathbf{Y}\|_F^2$$

### Maximum A Posteriori (MAP) Estimator

- Statistical source signals of uncorrelated Gaussian prior distribution with uniform variances  $\psi_i \sim \mathcal{CN}(\mathbf{0}, \gamma \mathbf{I}_L)$  for each snapshot  $i$
- Maximization of the posterior probability:  $\rho = \sigma^2 / \gamma$

$$\begin{aligned} \max_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \prod_{n=1}^N p(\psi_n | \mathbf{y}_n) &\iff \max_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \prod_{n=1}^N p(\mathbf{y}_n | \psi_n) p(\psi_n) \\ &\iff \min_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \sum_{n=1}^N -\log(p(\mathbf{y}_n | \psi_n)) - \log(p(\psi_n)) \\ &\iff \min_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \|\mathbf{A}(\mu)\Psi - \mathbf{Y}\|_F^2 + \rho \|\Psi\|_F^2 \end{aligned}$$

	DML	MAP
Original	$\min_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \ \mathbf{A}(\mu)\Psi - \mathbf{Y}\ _F^2$	$\min_{\mu \in [-\pi, \pi]^L, \Psi \in \mathbb{C}^{L \times N}} \ \mathbf{A}(\mu)\Psi - \mathbf{Y}\ _F^2 + \rho \ \Psi\ _F^2$
Concentrated	$\min_{\mu \in [-\pi, \pi]^L} \text{tr}(\mathbf{Y}^H \Pi_{\mathbf{A}(\mu)}^\perp \mathbf{Y})$	$\min_{\mu \in [-\pi, \pi]^L} \text{tr}(\mathbf{Y}^H \tilde{\Pi}_{\mathbf{A}(\mu)}^\perp \mathbf{Y})$

- $\Pi_{\mathbf{A}(\mu)}^\perp = \mathbf{I}_M - \mathbf{A}(\mu)(\mathbf{A}(\mu)^H \mathbf{A}(\mu))^{-1} \mathbf{A}(\mu)^H$
- $\tilde{\Pi}_{\mathbf{A}(\mu)}^\perp = \mathbf{I}_M - \mathbf{A}(\mu)(\mathbf{A}(\mu)^H \mathbf{A}(\mu) + \rho \mathbf{I}_L)^{-1} \mathbf{A}(\mu)^H = (\frac{1}{\rho} \mathbf{A}(\mu) \mathbf{A}(\mu)^H + \mathbf{I}_M)^{-1}$

## Sparse Representation

- Sample the field-of-view (FOV) in  $K \gg L$  directions with spatial frequencies  $\nu = [\nu_1, \dots, \nu_K]^T$
- On-grid assumption:  $\{\mu_l\}_{l=1}^L \subset \{\nu_k\}_{k=1}^K$
- Sparse signal model

$$\mathbf{Y} = \mathbf{A}(\nu)\mathbf{X} + \mathbf{N} \quad \begin{aligned} - \mathbf{X} &\in \mathbb{C}^{K \times N}: \text{Row-sparse representation of } \Psi \\ - \mathbf{A}(\nu) &\in \mathbb{C}^{M \times K}: \text{Steering matrix for sampled directions } \nu \end{aligned}$$

- MAP estimator for the sparse model

$$\min_{\mathbf{X} \in \mathbb{C}^{K \times N}, \|\mathbf{X}\|_{2,0} \leq L} \|\mathbf{A}(\nu)\mathbf{X} - \mathbf{Y}\|_F^2 + \rho \|\mathbf{X}\|_F^2$$

## Mixed-Integer Semidefinite Programming Reformulation

- Integer programming reformulation for sparse MAP

$$\min_{\mathbf{u} \in \{0,1\}^K, \mathbf{u}^T \mathbf{1} \leq L} \min_{\mathbf{X} \in \mathbb{C}^{K \times N}} \|\mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{X} - \mathbf{Y}\|_F^2 + \rho \|\mathbf{X}\|_F^2 \xrightarrow{\text{concentration}} \min_{\mathbf{u} \in \{0,1\}^K, \mathbf{u}^T \mathbf{1} \leq L} \text{tr}(\mathbf{Y}^H (\frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H + \mathbf{I}_M)^{-1} \mathbf{Y})$$

- Binary vector  $\mathbf{u}$  determines the active directions
- Mixed-Integer semidefinite programming (MISDP) reformulation

$$\min_{\mathbf{u} \in \{0,1\}^K, \mathbf{T} \in \mathbb{S}_+^M} \text{tr}(\mathbf{T}) \quad \text{s.t.} \quad \begin{bmatrix} \frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H + \mathbf{I}_M & \mathbf{Y} \\ \mathbf{Y}^H & \mathbf{T} \end{bmatrix} \succeq 0, \quad \mathbf{u}^T \mathbf{1} \leq L$$

- Since  $\frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H \succ 0$ , by Schur complement formula

$$\begin{bmatrix} \frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H + \mathbf{I}_M & \mathbf{Y} \\ \mathbf{Y}^H & \mathbf{T} \end{bmatrix} \succeq 0 \iff \mathbf{T} - \mathbf{Y}^H (\frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H + \mathbf{I}_M)^{-1} \mathbf{Y} \succeq 0$$

## Interval Relaxation and Randomized Rounding (IRRR) [1]

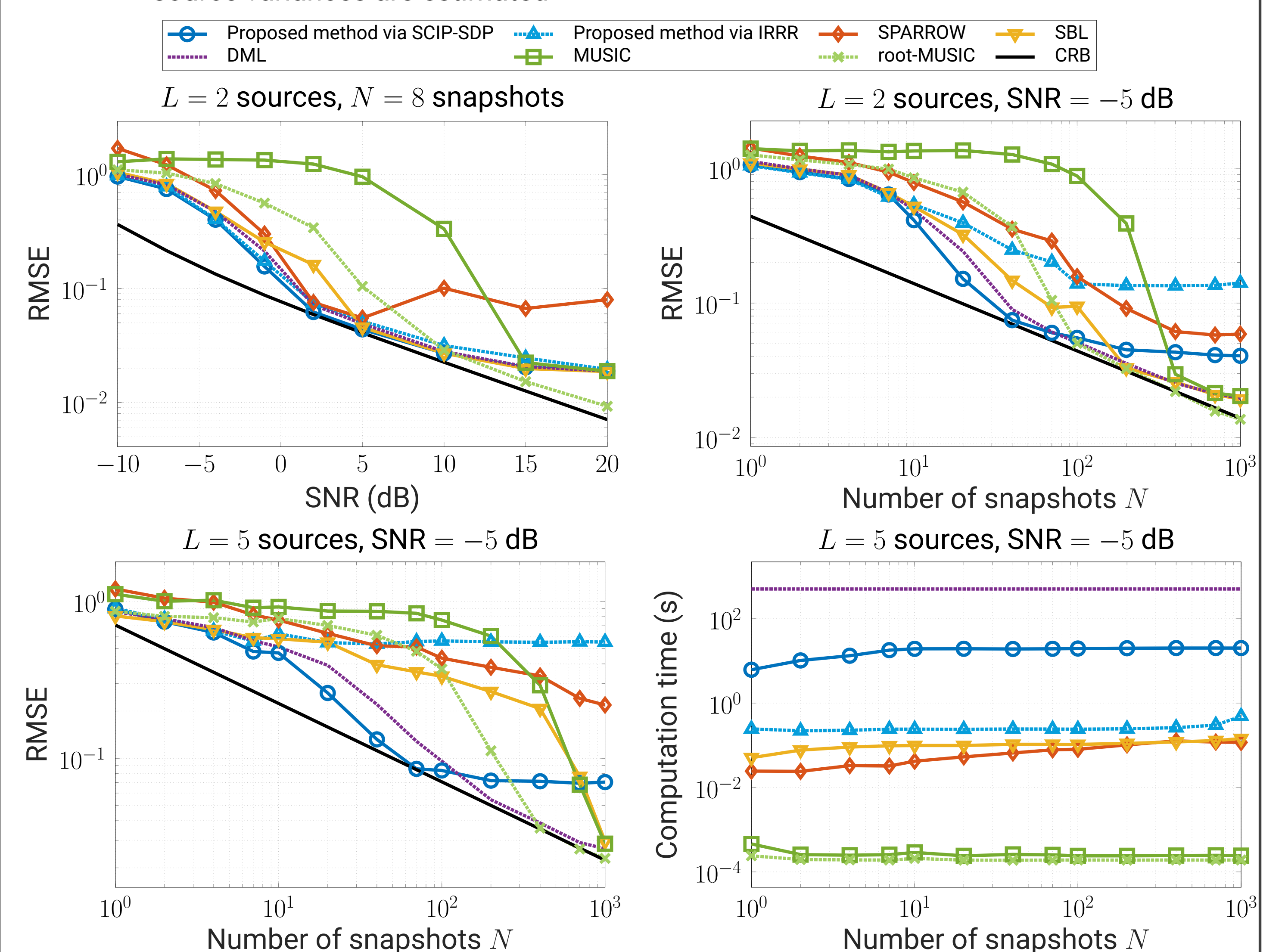
Step 1 **Interval Relaxation**: Solve the convex continuous relaxation

$$\hat{\mathbf{u}} = \arg \min_{\mathbf{u} \in [0,1]^K, \mathbf{T} \in \mathbb{S}_+^M} \text{tr}(\mathbf{T}) \quad \text{s.t.} \quad \begin{bmatrix} \frac{1}{\rho} \mathbf{A}(\nu)\mathbf{D}(\mathbf{u})\mathbf{A}(\nu)^H + \mathbf{I}_M & \mathbf{Y} \\ \mathbf{Y}^H & \mathbf{T} \end{bmatrix} \succeq 0, \quad \mathbf{u}^T \mathbf{1} \leq L$$

Step 2 **Randomized Rounding**: Randomly generate  $T > 0$  binary solutions  $\tilde{\mathbf{u}} \in \{0,1\}^K$  where each entry  $\tilde{u}_k$  independently follows the Bernoulli distribution  $\mathbb{P}[\tilde{u}_k = 1] = \hat{u}_k$  and  $\mathbb{P}[\tilde{u}_k = 0] = 1 - \hat{u}_k$ . Choose the best feasible solution among them.

## Simulation Results

- The MISDP reformulation exactly solved by the SCI-SDP solver [2] based on branch-and-bound
- Continuous SDP problems solved by MOSEK solver
- Uniform linear array with  $M = 8$  sensors; uncorrelated sources;  $K = 100$  grid points
- Comparison sparsity-based DOA estimation methods:
  - SPARROW: Convex method with  $\ell_{2,1}$ -norm regularization
  - Sparse Bayesian Learning (SBL): Similarly uncorrelated Gaussian prior assumption but source variances are estimated



## References

- [1] M. Pilanci, M. J. Wainwright, and L. El Ghaoui, "Sparse learning via Boolean relaxations," *Mathematical Programming*, vol. 151, pp. 63–87, June 2015.
- [2] T. Gally, M. E. Pfetsch, and S. Ulbrich, "A framework for solving mixed-integer semidefinite programs," *Optimization Methods and Software*, vol. 33, pp. 594–632, May 2018.