

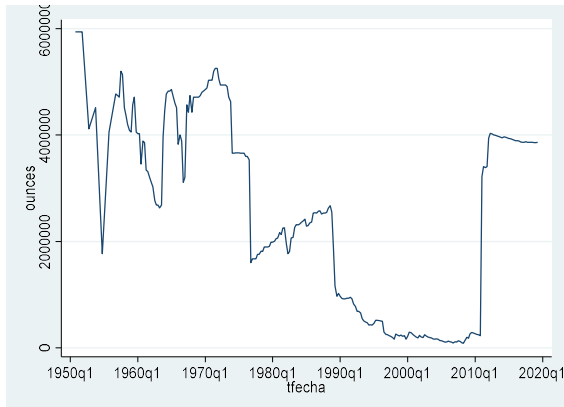
Tarea

1. Realizar un análisis de una serie trimestral empleando SARIMA

Se usará la serie de Kaggle “Country Gold Reserves Annual, quarterly, & monthly” de la pagina <https://www.kaggle.com/datasets/eliasdabbas/gold-reserves-by-country-quarterly> para Mexico.

- a. Análisis previo (hay que diferenciar)

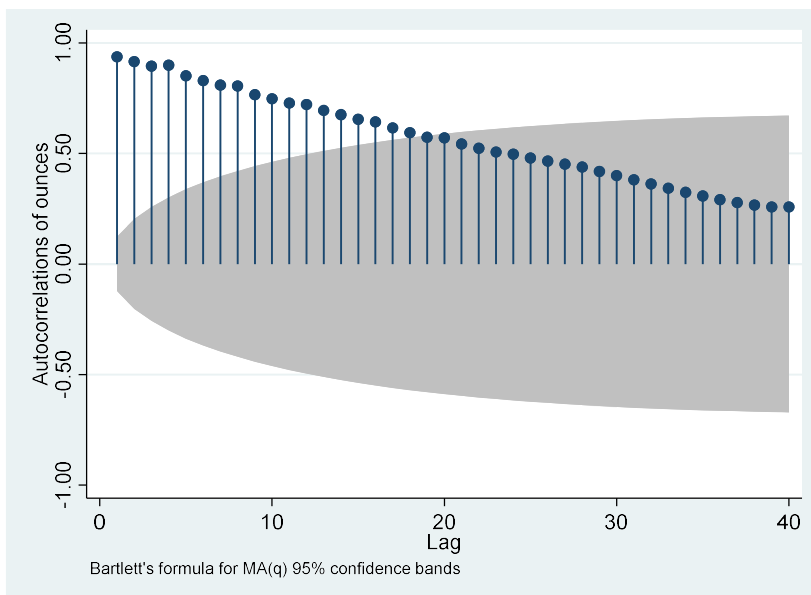
Graficamos la serie



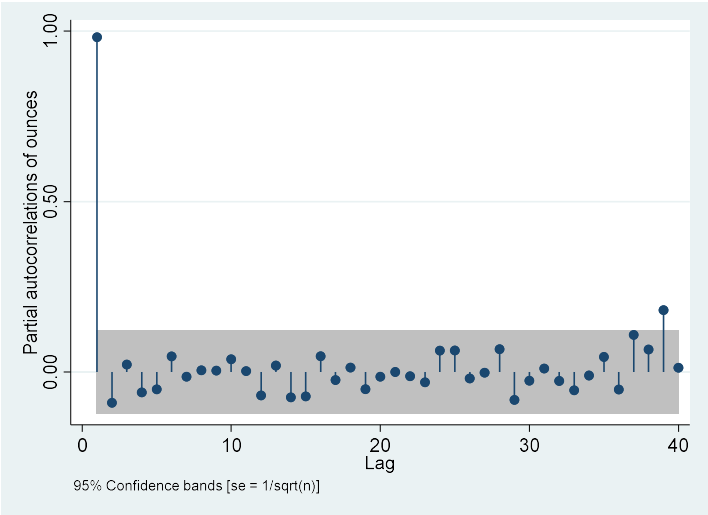
summarize ounces

Variable	Obs	Mean	Std. Dev.	Min	Max
ounces	257	2473928	1758374	81694.98	5943000

Graficando la autocorrelación simple



Graficando la autocorrelación parcial



Verificamos si la serie es estacionaria para diferenciar, como el valor critico es mayor al calculado, el test indica que hay raíz unitaria por lo que no es estacionaria.

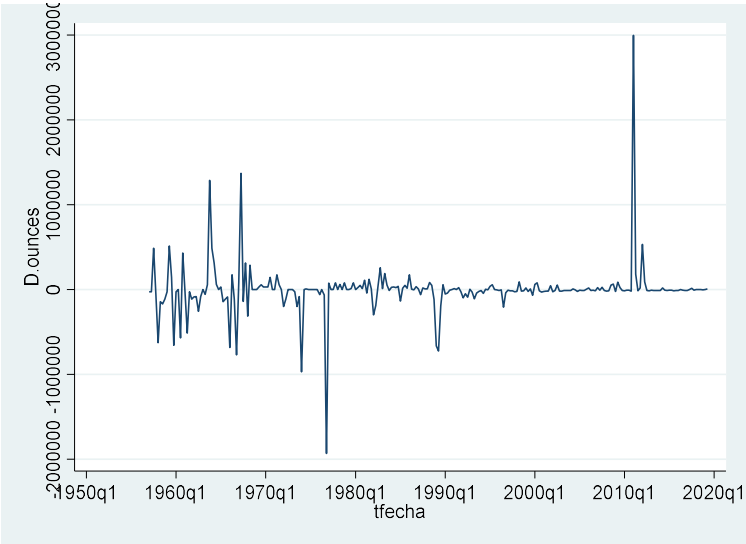
Dickey-Fuller test for unit root

Number of obs = 250

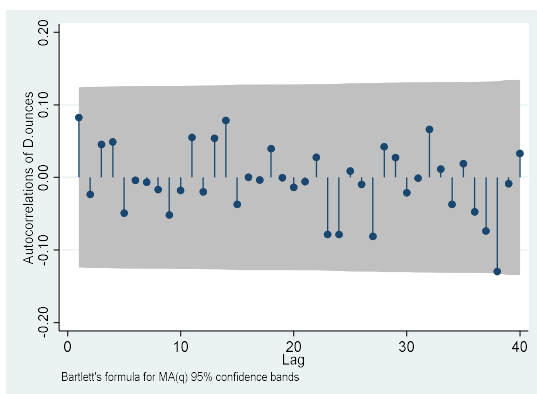
	Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-1.596	-3.460	-2.880	-2.570

MacKinnon approximate p-value for Z(t) = 0.4856

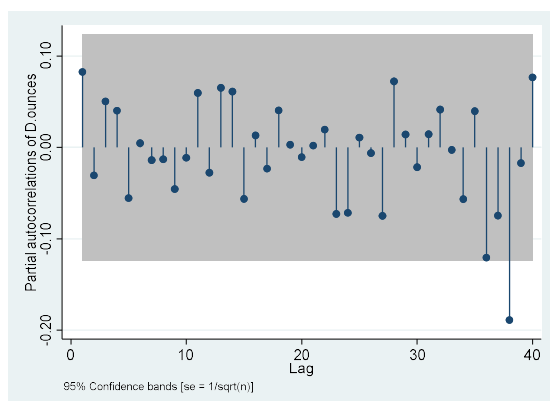
Al diferenciar tenemos:



b. Test de raíz unitaria, test de estacionalidad, correlogramas (AC y PAC)
Graficando la autocorrelación simple



Graficando la autocorrelación parcial



Volvemos a revisar si es estacionaria y comprobamos el valor critico es menor al calculado, la prueba indica que no hay raíz unitaria por lo que es estacionaria.

Dickey-Fuller test for unit root Number of obs = 249

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-14.466	-3.460	-2.880	-2.570

MacKinnon approximate p-value for Z(t) = 0.0000

Phillips-Perron test for unit root Number of obs = 249
Newey-West lags = 4

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(rho)	-230.649	-20.297	-13.998	-11.199
Z(t)	-14.478	-3.460	-2.880	-2.570

MacKinnon approximate p-value for Z(t) = 0.0000

Con kpss vemos que en todos el calculado es menor al crítico por lo que aceptamos H_0 es No estacionario

KPSS test for D.ounces

Maxlag = 5 chosen by Schwert criterion
Autocovariances weighted by Bartlett kernel

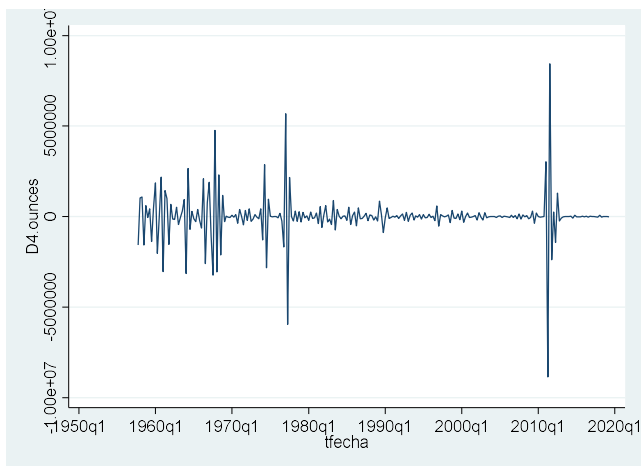
Critical values for H_0 : D.ounces is trend stationary

10%: 0.119 5% : 0.146 2.5%: 0.176 1% : 0.216

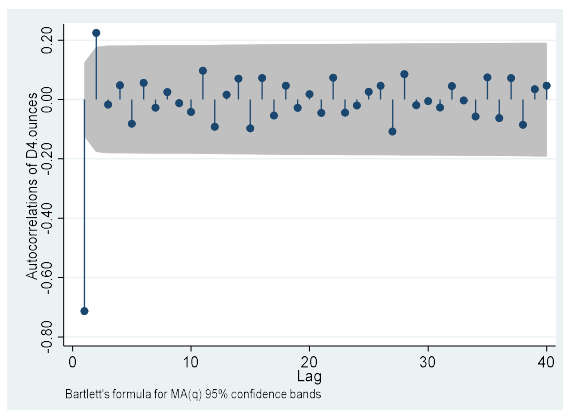
Lag order	Test statistic
0	.0768
1	.0714
2	.0711
3	.0698
4	.068
5	.068

c. Estimación de modelo SARIMA

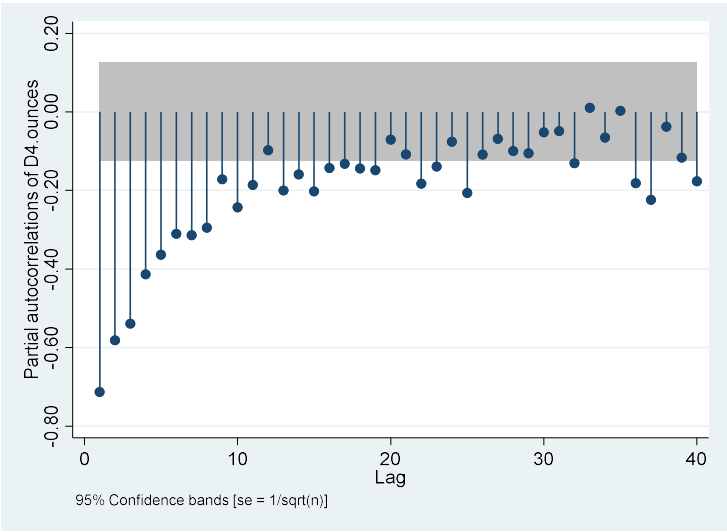
Al ser una serie trimestral graficamos con 4 rezagos



Graficando la autocorrelación simple



Graficando la autocorrelación parcial



Estimando el modelo tenemos las siguientes opciones

Con un sarima (1,1,1,4)

```
ARIMA regression

Sample: 1951q4 - 2019q2, but with gaps      Number of obs   =      253
                                           Wald chi2(2)     =    187.00
Log likelihood = -3759.474                  Prob > chi2      =    0.0000
```

S4.ounces	OPG		z	P> z	[95% Conf. Interval]	
	Coef.	Std. Err.				
ounces						
_cons	-29006.1	57807.74	-0.50	0.616	-142307.2	84294.99
ARMA4						
ar						
L1.	.9418268	.0794862	11.85	0.000	.7860368	1.097617
ma						
L1.	-.9999889	155.7668	-0.01	0.995	-306.2973	304.2974
/sigma	681693	5.31e+07	0.01	0.495	0	1.05e+08

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
.	253	.	-3759.474	4	7526.947	7541.081

Note: BIC uses N = number of observations. See [\[R\] BIC note](#).

Con un sarima(1,1,2,4)

ARIMA regression

Sample: 1951q4 - 2019q2, but with gaps Number of obs = 253
Wald chi2(3) = 0.10
Log likelihood = -3761.175 Prob > chi2 = 0.9923

S4.ounces	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
ounces						
_cons	-20065.92	48905.37	-0.41	0.682	-115918.7	75786.85
ARMA4						
ar						
L1.	.0254688	3.437931	0.01	0.994	-6.712753	6.763691
ma						
L1.	-.0339756	3.415305	-0.01	0.992	-6.727851	6.6599
L2.	-.0205673	.0781546	-0.26	0.792	-.1737474	.1326129
/sigma	692059.4	15978.61	43.31	0.000	660741.9	723376.9

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
.	253	.	-3761.175	5	7532.351	7550.017

Note: BIC uses N = number of observations. See [\[R\] BIC note](#).

Con un sarima (2,1,1,4)

ARIMA regression

Sample: 1951q4 - 2019q2, but with gaps Number of obs = 253
Wald chi2(4) = 25986.70
Log likelihood = -3752.333 Prob > chi2 = 0.0000

S4.ounces	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
ounces						
_cons	-19916.52	45163.51	-0.44	0.659	-108435.4	68602.33
ARMA4						
ar						
L1.	-.031502	.0416416	-0.76	0.449	-.113118	.0501141
L2.	-.9873568	.0290745	-33.96	0.000	-1.044342	-.9303719
ma						
L1.	-.0232064	1.49115	-0.02	0.988	-2.945808	2.899395
L2.	.999974	127.7715	0.01	0.994	-249.4275	251.4275
/sigma	651640.7	4.16e+07	0.02	0.494	0	8.22e+07

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

Akaike's information criterion and Bayesian information criterion

Model	N	ll(null)	ll(model)	df	AIC	BIC
.	253	.	-3752.333	6	7516.666	7537.867

Note: BIC uses N = number of observations. See [\[R\] BIC note](#).

De estos tres elegimos el sarima (1,1,2,4)

d. Verificación de supuestos (estacionariedad principalmente)

Probamos estacionariedad

```
Number of observations : 246
Deterministic variables : Seasonal dummies
Optimal lag selection method: Modified AIC
Lags tested: 5
Augmented by lags : 0
```

	Stat	1% critical	5% critical	10% critical
t[0]	-7.110	-3.424	-2.845	-2.548
t[Pi]	-8.007	-3.423	-2.845	-2.548
F[Pi/2]	54.224	8.861	6.628	5.592
F[All seas]	58.393	7.666	5.944	5.132
F[All]	56.065	7.001	5.556	4.869

Testing regression:

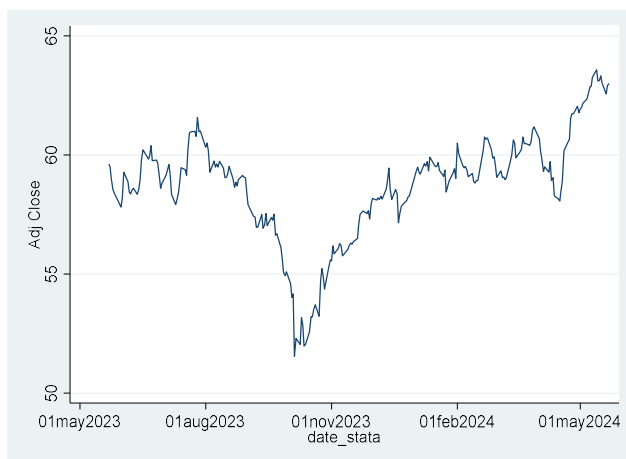
Revisamos t[Pi] y F[all], vemos que son mayores en valor absoluto por lo que sugiere no estacionariedad en los datos.

- Realizar un análisis de un precio de una acción de forma diaria empleando ARCH, GARCH, etc.

Analizamos a coca cola (KO)

a. Análisis previo (hay que diferenciar)

Graficamos la serie



Vemos un resumen de los datos

```
. summarize adjclose
```

Variable	Obs	Mean	Std. Dev.	Min	Max
adjclose	253	58.6689	2.28728	51.55385	63.58

Hacemos la prueba de raíz unitaria y vemos que el valor critico es mayor al calculado, el test indica que hay raíz unitaria por lo que no es estacionaria.

Dickey-Fuller test for unit root Number of obs = 198

Test Statistic	Interpolated Dickey-Fuller			
	1% Critical Value	5% Critical Value	10% Critical Value	
Z(t)	-0.772	-3.477	-2.883	-2.573

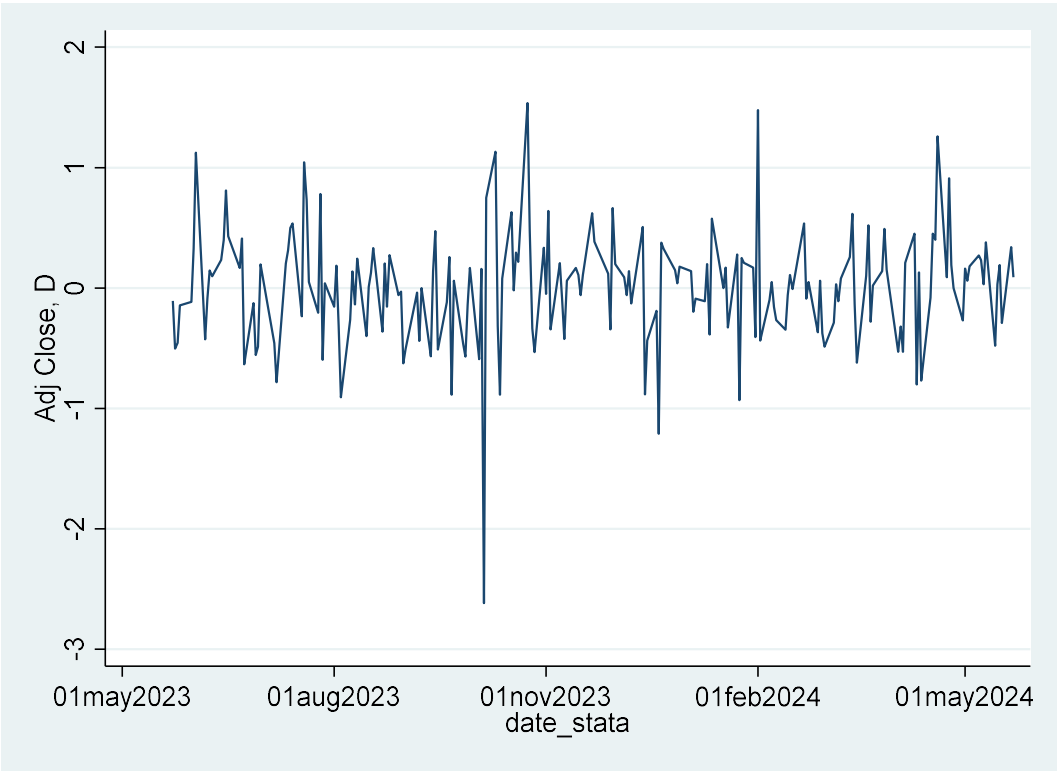
MacKinnon approximate p-value for Z(t) = 0.8273

Phillips-Perron test for unit root Number of obs = 198
Newey-West lags = 4

		Interpolated Dickey-Fuller		
	Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
Z(rho)	-2.274	-20.127	-13.896	-11.131
Z(t)	-0.768	-3.477	-2.883	-2.573

MacKinnon approximate p-value for Z(t) = 0.8283

Diferenciamos para hacerla estacionaria



Corroboramos nuevamente que sea estacionaria y vemos que si es estacionaria. Al ver que el valor critico es menor al valor calculado en los test.

Dickey-Fuller test for unit root

Number of obs = 145

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-12.023	-3.495	-2.887	-2.577

MacKinnon approximate p-value for Z(t) = 0.0000

Phillips-Perron test for unit root

Number of obs = 145

Newey-West lags = 4

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(rho)	-145.225	-19.950	-13.790	-11.060
Z(t)	-12.021	-3.495	-2.887	-2.577

MacKinnon approximate p-value for Z(t) = 0.0000

b. Test ARCH con n rezagos

. reg adjclose

Source	SS	df	MS	Number of obs	=	253
Model	0	0	.	F(0, 252)	=	0.00
Residual	1318.37604	252	5.23165096	Prob > F	=	.
Total	1318.37604	252	5.23165096	R-squared	=	0.0000
				Adj R-squared	=	0.0000
				Root MSE	=	2.2873

adjclose	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
_cons	58.6689	.1438001	407.99	0.000	58.38569 58.9521

Number of gaps in sample: 54

LM test for autoregressive conditional heteroskedasticity (ARCH)

lags(p)	chi2	df	Prob > chi2
1	173.980	1	0.0000

H0: no ARCH effects vs. H1: ARCH(p) disturbance

c. Estimación de modelo

Estimado un arch

ARCH family regression

Sample: 22may2023 - 22may2024, but with gaps Number of obs = 253
Distribution: Gaussian Wald chi2(.) = .
Log likelihood = -441.8821 Prob > chi2 = .

adjclose	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
adjclose _cons	59.22553	.0473809	1249.99	0.000	59.13266	59.31839
ARCH						
arch L1.	.9307497	.1212165	7.68	0.000	.6931698	1.16833
_cons	.1788525	.0401575	4.45	0.000	.1001453	.2575598

Estimando un garch

ARCH family regression

Sample: 22may2023 - 22may2024, but with gaps Number of obs = 253
Distribution: Gaussian Wald chi2(.) = .
Log likelihood = -439.6973 Prob > chi2 = .

adjclose	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
adjclose _cons	59.24767	.0480749	1232.40	0.000	59.15344	59.34189
ARCH						
arch L1.	.954494	.1185837	8.05	0.000	.7220742	1.186914
garch L1.	-.0285958	.0151416	-1.89	0.059	-.0582727	.0010812
_cons	.2219635	.0634891	3.50	0.000	.0975271	.3463999

El modelo GARCH es el que mejor modela la serie.

d. Verificación de supuestos

Verificamos autocorrelación

Dado que el valor p es menor que el nivel de significancia, se rechaza la hipótesis nula de que los residuos son ruido blanco.

Portmanteau test for white noise

Portmanteau (Q) statistic = 2260.0228
Prob > chi2(40) = 0.0000

Verificación de la No Heterocedasticidad en los Residuos

Como el p-valor es menor al 5% no hay heterocedastsidad, los residuos son homocedasticos.

Time-series regression

Sample: 22may2023 - 22may2024, but with gaps Number of obs = 253
Distribution: Gaussian Wald chi2(.) = .
Log likelihood = -567.8134 Prob > chi2 = .

resid	OPG		z	P> z	[95% Conf. Interval]	
	Coef.	Std. Err.				
_cons	-.5787732	.1594846	-3.63	0.000	-.8913573	-.2661891
/SIGMA2	5.210973	.4229635	12.32	0.000	4.381979	6.039966

Verificación de Normalidad de los Residuos

Como el p-valor es menor al 5% aceptamos que no los residuos siguen una distribución normal

. sktest resid

Skewness/Kurtosis tests for Normality					
Variable	Obs	Pr(Skewness)	Pr(Kurtosis)	joint	
				adj chi2(2)	Prob>chi2
resid	253	0.0000	0.0122	21.67	0.0000