

# 10 - Continuous MCMC - 01

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Basic Metropolis algorithm

Metropolis-Hastings algorithm

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## Setting

Statistical model with a parameter  $\theta \in \Theta \subset \mathbb{R}^p$ ,  
 $n$  i.i.d.  $d$ -dimensional observations:  $x = (x_1, \dots, x_n)$ .

*Likelihood:*  $f(x_i|\theta)$ ,      *Prior pdf:*  $h(\theta)$ .

*Posterior pdf:*

$$h_x(\theta) = \frac{\prod_{i=1}^n f(x_i|\theta) \cdot h(\theta)}{\underbrace{Z_x}_{\text{yellow box with speech bubble}}} = \frac{f(x, \theta)}{\underbrace{Z_x}_{\text{yellow box with speech bubble}}},$$

$Z_x$  is the *normalization constant*, marginal density of  $x$ ,  
 evaluated at the observed point.

# Metropolis algorithm

We construct a sequence  $\{\theta^{(t)}\}_{t \in \mathbb{N}}$  of points in  $\Theta$ :

A trajectory of a Markov chain whose limit pdf is  $h_x(\theta)$ .

A step in the chain  $\theta^{(t)} \rightarrow \theta^{(t+1)}$  is as follows:

# Metropolis algorithm

- 1 Change proposal (Sample from a candidate generation pdf)
- 2 Acceptance-rejection

# [1] Change proposal (Uniform)

Set a length scale  $\Delta > 0$ . Write  $\theta = \theta^{(t)}$ .

Generate  $\theta' = u$ , a random uniform vector in:

$$\prod_{j=1}^p \text{Unif}(\theta_j - \Delta, \theta_j + \Delta), \quad p = \dim(\theta),$$

that is  $\theta'_j - \theta_j \sim \text{Unif}(-\Delta, \Delta)$ ,  $1 \leq j \leq p$ .

# [1] Change proposal (General pdf)

A *candidate generation* pdf  $g$ , such that:  $g(-\theta) = g(\theta)$ .



Then  $\theta' = \theta + u$ , where  $u \sim g$ .

By analogy with the finite case, write:

$$k(\theta'|\theta) = g(\theta' - \theta). \quad (\text{Sort of "matrix"})$$

Since  $g$  is a pdf,

$$\int_{\theta'} k(\theta'|\theta) d\theta' = 1. \quad (\text{A "stochastic matrix"})$$



## Notation: Transition kernel

The function of the two variables  $\theta$ ,  $\theta'$ :

$$k(\theta'|\theta)$$

is the *transition kernel*.

Here “kernel” means “a function of two variables”, sort of “matrix”.  
Just as in “Kernel Learning”.

## Notation: Remark

Some books prefer:

$$k(\theta, \theta') \quad (\text{equivalent to our}) \quad k(\theta'|\theta).$$

Our motivation is to keep analogy with the matrix and conditional density-oriented notation.

$$k_{ij} = \text{P}(\text{transition to state } j \mid \text{current state is } i).$$

( $j$ -th entry in row  $i$ )

## Metropolis algorithm - [2] Acceptance-rejection step

Generate a random indicator  $I \sim \text{Ber}(p)$  with:

$$p = \min \left\{ 1, \frac{h_x(\theta')}{h_x(\theta)} \right\},$$

- If  $I = 1$ , we accept the update:  $\theta^{(t+1)} = \theta'$ ,
- If  $I = 0$ , we keep:  $\theta^{(t+1)} = \theta$ .

## Intuitive explanation

A jump to  $\theta'$  is proposed. Then:

- If the target distribution density at  $\theta'$  is higher we go there.
- If it is lower we go there only conditionally, with a probability proportional to the decrease in density.

## Intuitive explanation

We wander around the state space  $\Theta$ , and we want to go more often, and spend more time, at regions where the probability density  $h_x(\theta)$  is higher.

Hence, when on a high density area we try to stay, whereas when on a low density area, we move on, perchance we will improve.

## No denominators

Since the target pdf appears only in the quotient:

$$\frac{h_x(\theta')}{h_x(\theta)},$$

$Z_x$  is NOT required. Only the joint pdf  $f(x, \theta)$ :

$$p = \min \left\{ 1, \frac{f(x, \theta')}{f(x, \theta)} \right\}.$$

## Scale in candidate generation

$\Delta$  in the uniform case, in general dispersion parameter(s) in the candidate generation pdf  $g$ .

Tradeoff between:

- Small  $\Delta$ , high acceptance probability, slow displacement in  $\Theta$ ,
- Large  $\Delta$ , a swift displacement, small acceptance probability.

## Resulting Markov chain

In the continuous state space  $\Theta$ .

Transition kernel (“matrix”):

$$P(\theta' | \theta) = k(\theta' | \theta) \cdot \min \left\{ 1, \frac{h_x(\theta')}{h_x(\theta)} \right\}$$



## Detailed balance condition

Multiplying by  $h_x(\theta)$ ,

$$h_x(\theta) \cdot P(\theta' | \theta) = k(\theta' | \theta) \cdot \min \{h_x(\theta), h_x(\theta')\},$$



By the symmetry of  $k(\theta' | \theta)$  this is equal to:

$$k(\theta | \theta') \cdot \min \{h_x(\theta'), h_x(\theta)\} = h_x(\theta') \cdot P(\theta | \theta').$$

Hence it is a *time-reversible* Markov chain.

## The target pdf $h_x(\theta)$ is the limit probability

Indeed:

$$\begin{aligned}\int_{\theta \in \Theta} h_x(\theta) \cdot P(\theta' | \theta) d\theta &= \int_{\theta \in \Theta} h_x(\theta') \cdot P(\theta | \theta') d\theta \\ &= h_x(\theta') \cdot \int_{\theta \in \Theta} P(\theta | \theta') d\theta = h_x(\theta').\end{aligned}$$

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## References

Nicholas Metropolis, Arianna W. Rosenbluth, Marshall N. Rosenbluth, Augusta H. Teller, and Edward Teller (1953), *Equation of State Calculations by Fast Computing Machines*, J. Chemical Physics, Vol. 21, pp. 1087–1092.

Wilfred Keith Hastings (1970), *Monte Carlo sampling methods using Markov chains and their applications*, Biometrika 57, 97-109.

## Description

A generalization, with a non-symmetric candidate proposal kernel  $k(\cdot | \cdot)$ .

The acceptance rule is modified to compensate.

Now we accept  $\theta^{(m+1)} = \theta'$  with probability:

$$\min \left\{ 1, \frac{h_x(\theta') \cdot k(\theta | \theta')}{h_x(\theta) \cdot k(\theta' | \theta)} \right\},$$

## Description

The transition kernel is:

$$P(\theta' | \theta) = k(\theta' | \theta) \cdot \min \left\{ 1, \frac{h_x(\theta') \cdot k(\theta | \theta')}{h_x(\theta) \cdot k(\theta' | \theta)} \right\}.$$

## Detailed balance condition

Multiplying by  $h_x(\theta)$ ,

$$\begin{aligned}h_x(\theta) \cdot P(\theta' | \theta) &= k(\theta' | \theta) \cdot \min \left\{ h_x(\theta), \frac{h_x(\theta') \cdot k(\theta | \theta')}{k(\theta' | \theta)} \right\} \\&= \min \{ h_x(\theta) \cdot k(\theta' | \theta), h_x(\theta') \cdot k(\theta | \theta') \} \\&= h_x(\theta') \cdot P(\theta | \theta'),\end{aligned}$$

hence the chain is reversible with respect to  $h_x(\theta)$ ,  
the stationary distribution.

## Construction of $k(\cdot, \cdot)$

Based on a random walk.

From  $x$ , the *proposed*  $y$  is equal to  $x$  plus a random

$$z = y - x,$$

generated following a pdf  $g$ .

$$k(x, y) = g(z) = g(y - x).$$

When  $g$  is a symmetric pdf, we recover the Metropolis algorithm, where  $k(\cdot, \cdot)$  is a symmetric kernel.