

06 - Conjugate models - 01

Master in Foundations of Data Science
Bayesian Statistics and Probabilistic Programming
Fall 2018-2019

Josep Fortiana

Wednesday, October 24, 2018

Universitat de Barcelona

06 - Conjugate models - 01

The negative binomial distribution

The Poisson-Gamma model

06 - Conjugate models - 01

The negative binomial distribution

The Poisson-Gamma model

First definition

A sequence of independent binary 0/1 experiments, whose indicators are equally distributed as a $\text{Ber}(p)$, $p \in (0, 1)$.

The number X of realizations needed to obtain a number $r \in \mathbb{N}$ of successes (1's), is a r.v. following the *negative binomial distribution* with *size* r and probability p .

Alternative: $Y = X - r =$ number of failures (0's) before obtaining a number r of successes.

Probability mass function

$$P(x) = \binom{x-1}{r-1} \cdot (1-p)^{x-r} \cdot p^r, \quad x = 1, 2, \dots,$$

$$P(y) = \binom{y+r-1}{r-1} \cdot (1-p)^y \cdot p^r, \quad y = 0, 1, 2, \dots,$$

The second one is more usual (see e.g. the `dnbinom` function in R).

Alternative (and the reason for the name)

$$P(y) = \binom{-r}{y} \cdot p^r \cdot (-q)^y,$$

where $q = 1 - p$. Indeed:

$$\begin{aligned} \binom{-r}{y} &= \frac{(-r) \cdot (-r-1) \cdot \dots \cdot (-r-y+1)}{y!} \\ &= (-1)^y \cdot \binom{y+r-1}{r-1}. \end{aligned}$$

General definition

For an integer r ,

$$\binom{y+r-1}{r-1} = \frac{(y+r-1)!}{(r-1)! \cdot y!} = \frac{\Gamma(y+r)}{\Gamma(r) \cdot y!}.$$

The right hand is valid for real $r > 0$. Thus the pmf:

$$P(y) = \frac{\Gamma(y+r)}{\Gamma(r) \cdot y!} \cdot (1-p)^y \cdot p^r, \quad y = 0, 1, 2, \dots,$$

defines the $\text{NegBin}(r, p)$, for $r > 0$ and probability p .

Relation to the geometric distribution

For $r = 1$, the $\text{NegBin}(1, p)$ is the $\text{Geom}(p)$.

For integer r , a r.v. distributed as $\text{NegBin}(r, p)$ can be considered as the sum of r i.i.d. copies of a $\text{Geom}(p)$.

Expectation, variance of a negative binomial

For $Y \sim \text{NegBin}(r, p)$,

$$E(Y) = \mu \equiv r \cdot \frac{1-p}{p}$$

$$\text{var}(Y) = \sigma^2 \equiv r \cdot \frac{1-p}{p^2} = \frac{\mu}{p} = \mu + \frac{\mu^2}{r}.$$

06 - Conjugate models - 01

The negative binomial distribution

The Poisson-Gamma model

Likelihood

The likelihood is:

$$(y|\lambda) \sim \text{Poisson}(\lambda),$$

with pmf:

$$f(y|\lambda) = e^{-\lambda} \frac{\lambda^y}{y!}, \quad y = 0, 1, 2, \dots, \quad \lambda > 0,$$

Conjugate prior pdf

λ 's prior is:

$$\text{Gamma}(\alpha, \beta), \quad \alpha, \beta > 0,$$

with pdf:

$$h(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}, \quad \alpha, \beta, \lambda > 0.$$

Joint pdf

The joint “density” of (y, λ) is:

$$f(y, \lambda) = f(y|\lambda) \cdot h(\lambda) = \frac{\beta^\alpha}{\Gamma(\alpha)} \cdot \frac{\lambda^{\alpha+y-1}}{y!} e^{-(\beta+1)\lambda},$$

for $\alpha, \beta, \lambda > 0$.

Marginal pmf of y (Prior predictive pmf)

To integrate with respect to λ we split $f(y, \lambda)$:

$$\frac{\beta^\alpha}{\Gamma(\alpha)} \cdot \frac{1}{y!} \cdot \frac{\Gamma(\alpha + y)}{(\beta + 1)^{(\alpha + y)}} \times \frac{(\beta + 1)^{(\alpha + y)}}{\Gamma(\alpha + y)} \cdot \lambda^{\alpha + y - 1} e^{-(\beta + 1)\lambda},$$

The second factor is a $\text{Gamma}(\alpha + y, \beta + 1)$ pdf, which integrates to 1.

Marginal pmf of x

The result of integrating out λ is the marginal pmf of y :

$$f(y) = \frac{\Gamma(\alpha + y)}{\Gamma(\alpha) y!} \frac{\beta^\alpha}{(\beta + 1)^{\alpha+y}}, \quad y = 0, 1, \dots, \quad \alpha, \beta > 0$$

Noting that:

$$\frac{\beta^\alpha}{(\beta + 1)^{\alpha+y}} = \left(\frac{1}{\beta + 1} \right)^y \cdot \left(\frac{\beta}{\beta + 1} \right)^\alpha,$$

Marginal pmf of y

We identify $f(y)$ as the pdf of a $\text{NegBin}(r, p)$, a negative binomial r.v. , with parameters

$$r = \alpha, \quad \text{and} \quad p = \frac{\beta}{\beta + 1}.$$

As always, this is the *prior predictive* distribution for the observed y .

Posterior pdf

We apply Bayes' formula, dividing $f(y, \lambda)$ by the marginal pmf, and we obtain the posterior pdf of λ , given y , the first factor above:

$$h(\lambda|y) = \frac{(\beta + 1)^{\alpha+y}}{\Gamma(\alpha + y)} \lambda^{\alpha+y-1} e^{-(\beta+1)\lambda}, \quad \lambda > 0,$$

which is a $\text{Gamma}(\alpha + y, \beta + 1)$.

Case of an n -sample: posterior pdf

For y_1, \dots, y_n i.i.d. $\sim \text{Poisson}(\lambda)$, the sum:

$$y = \sum_{i=1}^n y_i \sim \text{Poisson}(n \lambda)$$

Thus, for a prior $\lambda \sim \text{Gamma}(\alpha, \beta)$ and n observed $\text{Poisson}(\lambda)$ data, the posterior pdf is $\text{Gamma}(\alpha + y, \beta + n)$.

Case of an n -sample: prior predictive pmf

Similarly, for n observed $\text{Poisson}(\lambda)$ data, the prior predictive (marginal) distribution of the total count number $y = \sum_{i=1}^n y_i$ is a $\text{NegBin}(r, p)$, a negative binomial r.v. , with parameters:

$$r = \alpha, \quad \text{and} \quad p = \frac{\beta}{\beta + n}.$$

An n -sample with different exposures

In many applications we find data of the form:

$$y_i \sim \text{Poisson}(\lambda_i), \quad \text{where } \lambda_i = x_i \cdot \theta, \quad 1 \leq i \leq n.$$

The values x_i are known positive values of an explanatory variable, usually called *exposure*, and θ is the common *rate* parameter.

Posterior pdf for an n -sample with different exposures

For a prior $\theta \sim \text{Gamma}(\alpha, \beta)$,
observations $\mathbf{y} = (y_1, \dots, y_n)$, corresponding to
known exposures $\mathbf{x} = (x_1, \dots, x_n)$,
the posterior pdf is:

$$\theta|\mathbf{y} \sim \text{Gamma}(\alpha + y, \beta + x),$$

where $y = \sum_{i=1}^n y_i$ and $x = \sum_{i=1}^n x_i$.