

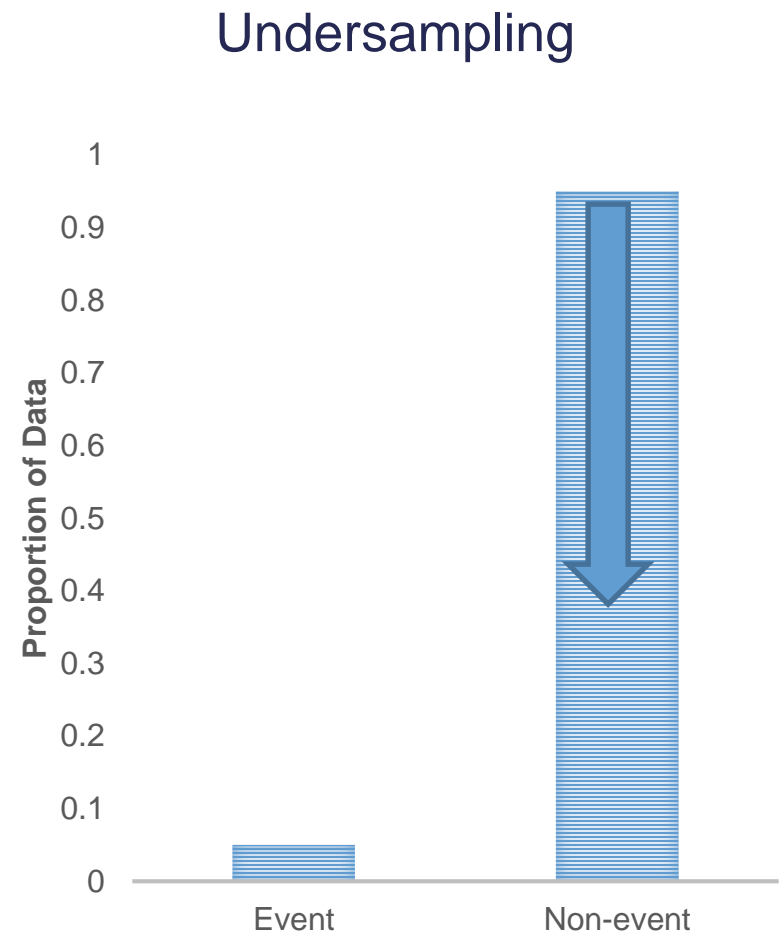
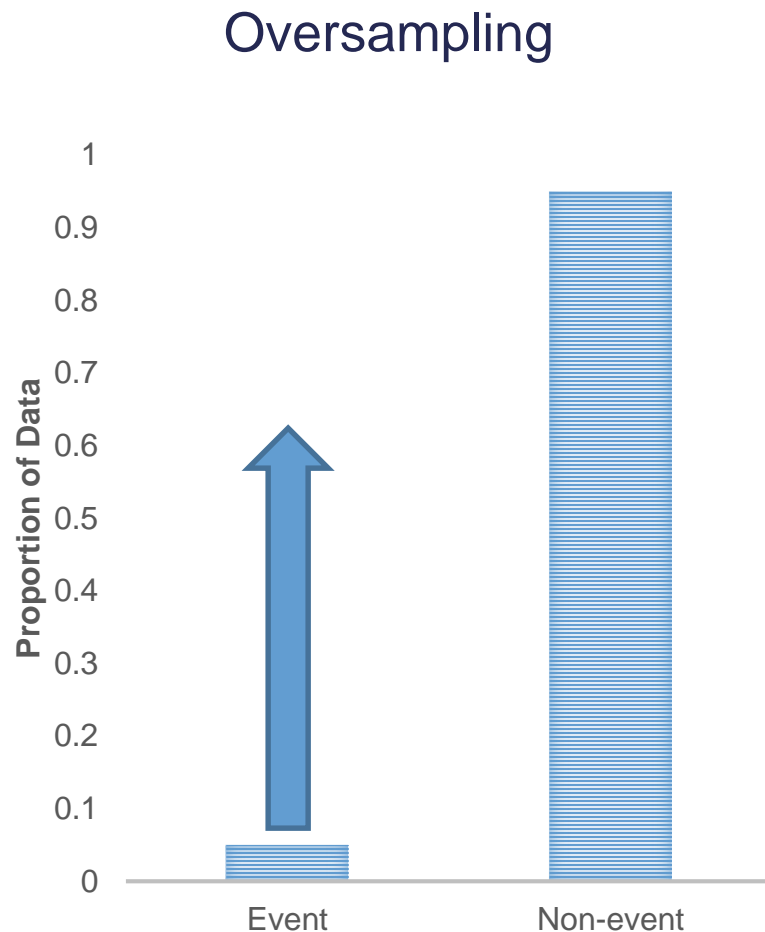
DATA CONSIDERATIONS – EXTRA CONTENT

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RARE EVENT MODELING IN R

Rare Event Sampling Correction



Rare Event Sampling Correction

Oversampling

- Duplicate current event cases in training set to balance better with non-event cases.
- Keep test set as original population proportion.

Undersampling

- Randomly sample current non-event cases to keep in the training set to balance with event cases.
- Keep test set as original population proportion.

Rare Event Sampling – R

```
train_id <- sample(seq_len(nrow(churn)),  
                  size = floor(0.7*nrow(churn)))
```

```
train <- churn[train_id,]  
valid <- churn[-train_id,]
```

```
table(train$churn)
```

```
##
```

```
## FALSE  TRUE
```

```
##  1995   107
```

```
table(valid$churn)
```

```
##
```

```
## FALSE  TRUE
```

```
##   855    47
```

Rare Event Sampling – R

```
prop.table(table(train$churn))

##
##      FALSE      TRUE
## 0.9490961 0.0509039

inputs <- train[,1:18]
target <- train[,19]
over_sam <- ubOver(X = inputs, Y = target)
train_o <- cbind(over_sam$X, over_sam$Y)
train_o$churn <- train_o$`over_sam$Y`
train_o$`over_sam$Y` <- NULL

table(train_o$churn)

##
## FALSE  TRUE
## 1995   1995
```

Rare Event Sampling – R

```
inputs <- train[,1:18]
target <- train[,19]
under_sam <- ubUnder(X = inputs, Y = target)
train_u <- cbind(under_sam$X, under_sam$Y)
train_u$churn <- train_u$`under_sam$Y`
train_u$`under_sam$Y` <- NULL
```

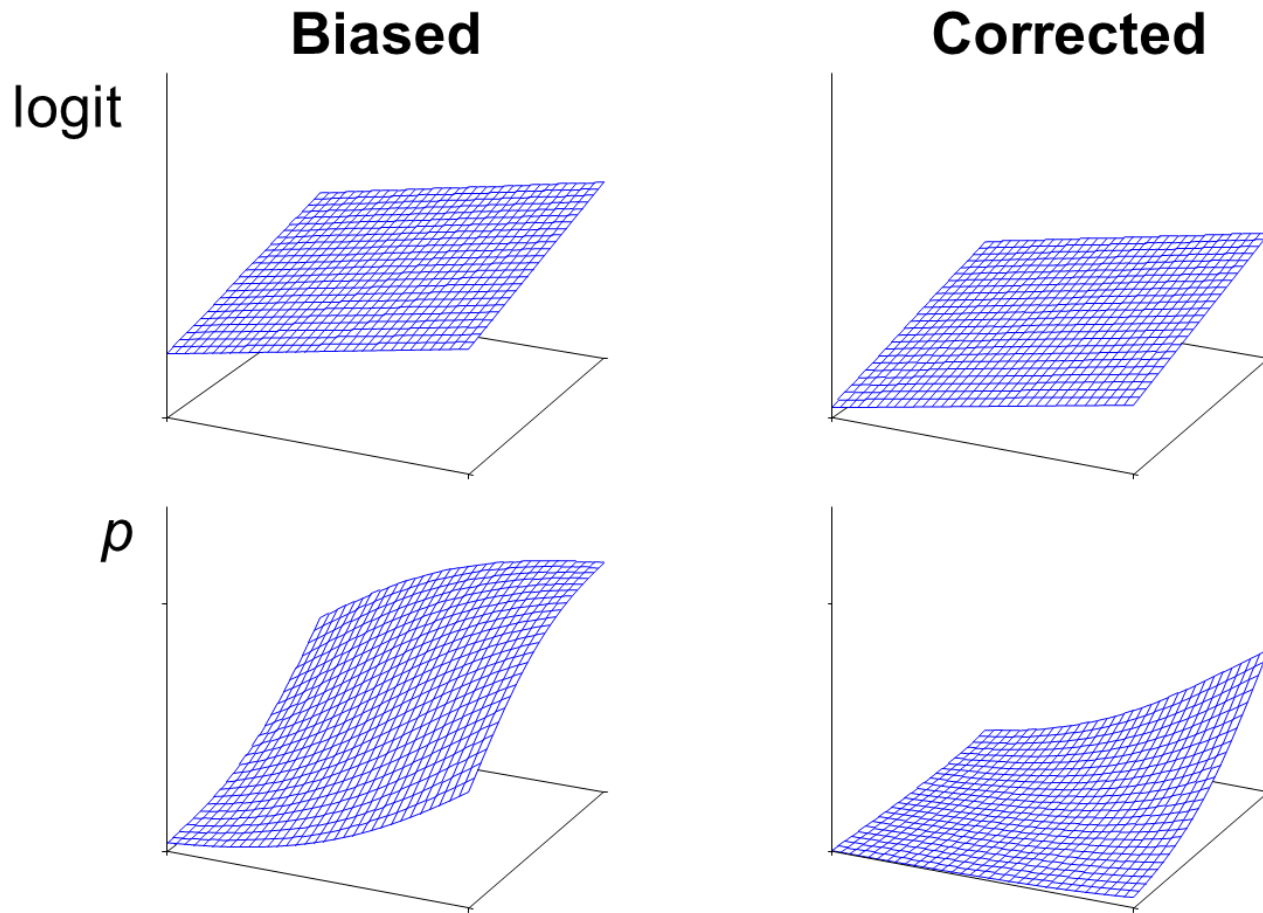
```
table(train_u$churn)
```

```
##
```

```
## FALSE TRUE
```

```
## 107 107
```

Effect of Oversampling



Adjustments to Oversampling

- When the sample proportion is out of line with the population proportion, adjustments need to be made to correct the bias.
- 2 Methods:
 1. Adjusting the intercept
 2. Weighting observations

Adjust Intercept – R

```
logit.model <- glm(churn ~ factor(international.plan) +  
                    factor(voice.mail.plan) +  
                    total.day.charge +  
                    customer.service.calls,  
                    data = train_u,  
                    family = binomial(link = "logit"))  
summary(logit.model)
```

Adjust Intercept – R

```
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)    -4.71202     0.77589  -6.073 1.25e-09 ***
## factor(international.plan)yes  2.91300     0.58964   4.940 7.80e-07 ***
## factor(voice.mail.plan)yes    -0.30174     0.43242  -0.698   0.485
## total.day.charge      0.09769     0.01829   5.341 9.26e-08 ***
## customer.service.calls  0.63437     0.12268   5.171 2.33e-07 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 296.67  on 213  degrees of freedom
## Residual deviance: 217.88  on 209  degrees of freedom
## AIC: 227.88
```

Adjust Intercept – R

```
valid_p_bias <- predict(logit.model, newdata = valid,  
                        type = "response")  
valid_p <- (valid_p_bias*0.5*(154/3004))/  
            ((1-valid_p_bias)*0.5*(2850/3004) +  
             valid_p_bias*0.5*(154/3004))
```

Weighting Adjustment – R

```
train_u$weight <- ifelse(train_u$churn == 'TRUE', 1, 18.49)

logit.model.w <- glm(churn ~ factor(international.plan) +
                     factor(voice.mail.plan) +
                     total.day.charge +
                     customer.service.calls,
                     data = train_u,
                     family = binomial(link = "logit"),
                     weights = weight)

summary(logit.model.w)
```

Weighting Adjustment – R

```
## Coefficients:
```

##	Estimate	Std. Error	z value	Pr(> z)	
## (Intercept)	-7.29928	0.52017	-14.033	< 2e-16	***
## factor(international.plan)yes	2.57544	0.28830	8.933	< 2e-16	***
## factor(voice.mail.plan)yes	-1.08471	0.33129	-3.274	0.00106	**
## total.day.charge	0.10566	0.01363	7.754	8.90e-15	***
## customer.service.calls	0.42969	0.07811	5.501	3.78e-08	***

```
## ---
```

```
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
##
```

```
## (Dispersion parameter for binomial family taken to be 1)
```

```
##
```

```
##      Null deviance: 843.97  on 213  degrees of freedom
```

```
## Residual deviance: 681.96  on 209  degrees of freedom
```

```
## AIC: 686.82
```

CONTRASTS

Testing Individual Contrasts

- Instead of testing all possible combinations of odds ratios, you may only be interested in certain comparison, or a linear combination of comparisons.
- These are called **contrasts**.
- For example:
 - Group A vs. Group B
 - Group A vs. the average of Group B and Group C
 - Etc.

Testing Individual Contrasts – SAS

```
proc logistic data=churn_t;  
  class international_plan(ref='no')  
    voice_mail_plan(ref='no')  
    customer_service_calls(ref='0') / param=ref;  
  model churn(event='TRUE') = international_plan  
    voice_mail_plan  
    total_day_charge  
    customer_service_calls;  
  
  weight weights;  
  oddsratio customer_service_calls_c / diff=all;  
  
run;  
quit;
```

Testing Individual Contrasts – SAS

Odds Ratio Estimates and Wald Confidence Intervals			
Odds Ratio	Estimate	95% Confidence Limits	
customer_service_calls 1 vs 0	1.082	0.125	9.397
customer_service_calls 2 vs 0	0.950	0.096	9.437
customer_service_calls 3 vs 0	1.246	0.088	17.694
customer_service_calls 4 vs 0	26.009	1.517	445.849
customer_service_calls 5 vs 0	13.653	0.634	293.830
customer_service_calls 6 vs 0	19.742	0.218	>999.999
customer_service_calls 7 vs 0	>999.999	<0.001	>999.999
customer_service_calls 1 vs 2	1.140	0.136	9.579
customer_service_calls 1 vs 3	0.869	0.071	10.572
customer_service_calls 1 vs 4	0.042	0.003	0.610
customer_service_calls 1 vs 5	0.079	0.004	1.482
customer_service_calls 1 vs 6	0.055	<0.001	4.643
customer_service_calls 1 vs 7	<0.001	<0.001	>999.999
customer_service_calls 2 vs 3	0.762	0.058	9.965
customer_service_calls 2 vs 4	0.037	0.002	0.556
customer_service_calls 2 vs 5	0.070	0.004	1.339
customer_service_calls 2 vs 6	0.048	<0.001	4.602
customer_service_calls 2 vs 7	<0.001	<0.001	>999.999
customer_service_calls 3 vs 4	0.048	0.002	1.029
customer_service_calls 3 vs 5	0.091	0.003	2.451
customer_service_calls 3 vs 6	0.063	<0.001	6.951
customer_service_calls 3 vs 7	<0.001	<0.001	>999.999
customer_service_calls 4 vs 5	1.905	0.070	51.556
customer_service_calls 4 vs 6	1.317	0.010	170.540
customer_service_calls 4 vs 7	<0.001	<0.001	>999.999
customer_service_calls 5 vs 6	0.692	0.005	106.159
customer_service_calls 5 vs 7	<0.001	<0.001	>999.999
customer_service_calls 6 vs 7	<0.001	<0.001	>999.999

Testing Individual Contrasts – SAS

```
proc logistic data=churn_t;  
  class international_plan(ref='no')  
    voice_mail_plan(ref='no')  
    customer_service_calls(ref='0') / param=ref;  
  model churn(event='TRUE') = international_plan  
    voice_mail_plan  
    total_day_charge  
    customer_service_calls / clodds=pl;  
  
  weight weights;  
  test customer_service_cal1 = customer_service_cal2;  
  test customer_service_cal1 = 0.25*customer_service_cal4 +  
    0.25*customer_service_cal5 +  
    0.25*customer_service_cal6 +  
    0.25*customer_service_cal7;  
  
run;  
quit;
```

Testing Individual Contrasts – SAS

Linear Hypotheses Testing Results			
Label	Wald Chi-Square	DF	Pr > ChiSq
Test 1	0.0145	1	0.9043
Test 2	0.0000	1	0.9951

Testing Individual Contrasts – R

```
train_u$fcsc <- factor(train_u$customer.service.calls)
```

```
logit.model.w.2 <- glm(churn ~ factor(international.plan) +  
                        factor(voice.mail.plan) +  
                        total.day.charge +  
                        fcsc,  
                        data = train_u,  
                        family = binomial(link = "logit"),  
                        weights = weight)
```

```
summary(glht(logit.model.w.2, linfct = mcp(fcsc = "Tukey")))
```

Testing Individual Contrasts – R

Linear Hypotheses:

##		Estimate	Std. Error	z value	Pr(> z)
## 1 - 0 == 0	-0.4461	1.0427	-0.428	0.999	
## 2 - 0 == 0	-0.2486	1.0901	-0.228	1.000	
## 3 - 0 == 0	0.1342	1.2275	0.109	1.000	
## 4 - 0 == 0	0.8636	1.1639	0.742	0.987	
## 5 - 0 == 0	2.3143	1.5825	1.462	0.727	
## 6 - 0 == 0	20.5210	1865.2667	0.011	1.000	
## 2 - 1 == 0	0.1975	1.0625	0.186	1.000	
## 3 - 1 == 0	0.5803	1.1901	0.488	0.999	
## 4 - 1 == 0	1.3097	1.1705	1.119	0.904	
## 5 - 1 == 0	2.7604	1.5486	1.782	0.508	
## 6 - 1 == 0	20.9671	1865.2667	0.011	1.000	
## 3 - 2 == 0	0.3828	1.2285	0.312	1.000	
## 4 - 2 == 0	1.1122	1.2322	0.903	0.965	
## 5 - 2 == 0	2.5629	1.5755	1.627	0.617	
## 6 - 2 == 0	20.7696	1865.2668	0.011	1.000	
## 4 - 3 == 0	0.7294	1.3468	0.542	0.998	
## 5 - 3 == 0	2.1801	1.6573	1.315	0.814	
## 6 - 3 == 0	20.3868	1865.2668	0.011	1.000	
## 5 - 4 == 0	1.4507	1.6966	0.855	0.973	
## 6 - 4 == 0	19.6573	1865.2668	0.011	1.000	
## 6 - 5 == 0	18.2067	1865.2670	0.010	1.000	

(Adjusted p values reported -- single-step method)

