



Climate drives among-year variation in natural selection on flowering time

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1 TITLE:

2 Climate drives among-year variation in natural selection on flowering time

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26 ABSTRACT

27 To predict long-term responses to climate change we need to understand how changes in
28 temperature and precipitation elicit both immediate phenotypic responses and changes in
29 natural selection. We used 22 years of data for the perennial herb *Lathyrus vernus* to examine
30 how climate influences flowering phenology and phenotypic selection on phenology. Plants
31 flowered earlier in springs with higher temperatures and higher precipitation. Early flowering
32 was associated with a higher fitness in nearly all years, but selection for early flowering was
33 significantly stronger in springs with higher temperatures and lower precipitation. Climate
34 influenced selection through trait distributions, mean fitness and trait-fitness relationships, the
35 latter accounting for most of the among-year variation in selection. Our results show that
36 climate both induces phenotypic responses and alters natural selection, and that the change in
37 the optimal phenotype might be either weaker, as for spring temperature, or stronger, as for
38 precipitation, than the optimal response.

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INTRODUCTION

Climate change is rapidly altering the environmental conditions experienced by organisms (Diffenbaugh & Field 2013; IPCC 2014). To predict the long-term effects of climate change, we need to assess not only immediate phenotypic responses of individuals, but also how changes in climate affect the optimal phenotypes, natural selection, and evolutionary trajectories of populations. Increased temperatures are expected to influence selection on timing through changes in the relationship between adaptive optima and phenotypic means (e.g. Visser *et al.* 1998, 2006; Gienapp *et al.* 2008; Marrot *et al.* 2018; Valdés *et al.* 2019). The direction of selection induced by increased temperature will depend on how strong plastic responses are compared with changes in the optimal phenotype. If the optimal response is stronger than the response elicited by increased temperatures, then we expect selection for stronger responses at higher temperatures (cf. Conover & Schultz 1995; Conover *et al.* 2009). If the optimal response is weaker, then we expect selection for weaker responses at higher temperatures. Many recent studies have documented phenotypic trait responses to climate, but it is less clear how climate influences the optimal phenotypes, and the direction and intensity of natural selection (Chevin *et al.* 2010; MacColl 2011; Siepielski *et al.* 2017). Moreover, the extent to which variation in climate or other environmental factors influence selection gradients through changes in trait-fitness relationships vs. changes in fitness and trait distributions is still largely unknown (Hunter *et al.* 2018). Long-term studies examining the relationships between climatic variables, phenotypic trait distributions, absolute fitness and trait-fitness relationships, are therefore essential to understand how climate change influences the evolutionary trajectories of plant and animal populations.

Responses to climatic variation often involve changes in the timing of different life cycle events (Visser & Both 2005; Cleland *et al.* 2007; Charmantier & Gienapp 2014). In seasonal environments, timing is crucial for survival, and phenological changes constitute the

64 main way in which organisms can modify their interactions with the abiotic and biotic
65 environment. Short-term plastic responses, such as earlier development with higher spring
66 temperatures or earlier snowmelt, have been documented for many species (Visser *et al.* 1998,
67 2006; Charmantier *et al.* 2008; Anderson *et al.* 2012; CaraDonna *et al.* 2014; Wadgymar *et al.*
68 2018).

69 However, we still know relatively little about how climate influences selection on
70 phenology, or to what extent observed phenotypic changes have a genetic basis (Gienapp *et*
71 *al.* 2008, Franks *et al.* 2014, Merilä & Hendry 2014). A few studies have identified climatic
72 factors as important drivers of temporal variation in selection on timing of reproduction in
73 animals (Réale *et al.* 2003; Visser *et al.* 2015; Marrot *et al.* 2018). Although climate is
74 expected to also influence selection on phenology in plants (Munguía-Rosas *et al.* 2011;
75 Anderson *et al.* 2012; Franks *et al.* 2014), such effects remain largely unexplored.

76 Despite the fundamental importance of temporal variation in selection, few long-term
77 assessments on selection on plant traits exist (Siepielski *et al.* 2017), the longest spanning up
78 to 10-11 years (Campbell & Powers 2015; Thomann *et al.* 2018). In this study, we use 22
79 years of data from permanently marked individuals in a natural population of the spring-
80 flowering understory herb *Lathyrus vernus* to examine the effects of climatic variation on
81 flowering phenology, and on phenotypic selection on phenology. In this species, the
82 development time from onset of shoot growth to flowering strongly depends on temperature.
83 Spring frosts might damage developing shoots or flower buds in some years, and snow might
84 protect shoot buds from extreme low temperatures in early spring. We asked: (1) Does
85 flowering time vary among years in response to climatic conditions during spring? (2) Does
86 the direction and strength of phenotypic selection on flowering time vary among years in
87 response to differences in climatic conditions during spring? and (3) If so, what is the relative

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88 importance of changes in trait means, trait variances, fitness means and trait-fitness
89 relationships for effects of climatic variables on selection gradients?
90 We predicted that higher temperatures during shoot development in March, April and
91 May result in faster development and earlier flowering. We also expected earlier-flowering
92 individuals to benefit more from their earlier start in years when temperatures are high during
93 the early phase of the flowering season. High temperatures might be beneficial because they
94 are associated with increased assimilation, or with earlier emergence and higher activity of
95 pollinators (cf. Kudo & Cooper 2019). On the other hand, low spring temperatures might lead
96 to frost damage of early buds and flowers (J. Ehrlén pers. obs.). We also hypothesized that
97 precipitation during early spring falling as snow might protect overwintering shoot buds from
98 extreme low temperatures (Inouye *et al.* 2002), and lead to an earlier flowering. Because no
99 negative effects of drought during spring have been observed in the study species, we also
100 hypothesized that high precipitation later in spring should be associated with reduced
101 photosynthetic active radiation, later flowering, and weaker selection for early flowering.
102 Taken together, this means that we expected temperature and precipitation to influence both
103 the timing of flowering and mean fitness, but that changes in selection gradients largely
104 would correspond to changes in absolute trait-fitness relationships.

106 MATERIALS AND METHODS

107 **Study species**

108 Individuals of *Lathyrus vernus* grow fairly slow, lack organs for vegetative spread, usually
109 flower only after 10-15 years of vegetative growth and frequently skip flowering in some
110 years (J. Ehrlén pers. obs.). The average conditional life span of flowering individuals has
111 been estimated to 44.3 years (Ehrlén & Lehtilä 2002). Growth of shoots is determinate and
112 the number of flowers is determined in the previous year (Ehrlén & Van Groenendael 2001).

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3 113 One or several erect shoots emerge from the over-wintering rhizome in March-April every
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5 114 year (Ehrlén 1995). At the study site, flowering starts about 4 weeks after shoot emergence,
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7 115 before canopy develops and light availability decreases in early June. Bud development and
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9 116 the onset of flowering therefore depend on climatic conditions during March, April and May.
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11 117 The time from shoot emergence to flowering varies considerably among years (J. Ehrlén
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13 118 unpublished data). Flower number (mean \pm SD = 13.45 ± 17.60) is closely correlated with
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15 119 aboveground biomass and plant resource state (Sola & Ehrlén 2007). Flowering is acropetally
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17 120 sequential within shoots, starting with the basal flower in the basal raceme. The only
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19 121 pollinators of the large flowers are bumblebees (*Bombus* spp.). *Lathyrus vernus* is self-
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21 122 compatible, but lacks mechanisms for autogamy (J. Ehrlén, unpublished data). Timing of
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23 123 flowering is correlated with differences in seed production, and this pattern might be
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25 124 influenced by differences in both resource and pollen availability (Ehrlén 1992). Individuals
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27 125 produce a small number of large seeds (mean \pm SD = 12.0 ± 3.5 mg, N = 200, Ehrlén 2002),
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29 126 and recruitment from seeds is frequent (Ehrlén & Eriksson 1996). Developing seeds are often
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31 127 damaged by the pre-dispersal seed predator *Bruchus atomarius*. Roe deers (*Capreolus*
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33 128 *capreolus*) sometimes consume flowering shoots, and early-flowering individuals are more
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35 129 often grazed than late-flowering (Ehrlén & Münzbergová 2009). There is heritable variation
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37 130 in flowering time among populations of *L. vernus* (Widén & Schiemann 2003), but the
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39 131 amount of genetic variation for flowering time within the study population is not known.
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133 **Data collection**

134 The study was carried out in a population of *L. vernus* in a deciduous forest in the Tullgarn
135 area, SE Sweden (58.9496 N, 17.6097 E), during 1987–1996 and 2006–2017. The study area
136 is quite homogeneous with regard to altitude, slope and canopy cover. A permanent snow
137 cover was lacking in most years, and when present, snow cover was usually less than 20 cm

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138 and disappeared about two months before flowering. In 1987, all flowering individuals in an
139 area of 825 m² were permanently marked and surveyed in each year to 1996. New flowering
140 individuals in the plot were included in the study in each year. No recordings were made
141 1997–2005. In 2006, a new set of individuals in an area of 162 m² within the same population
142 were marked, and surveyed in the same way as the initially marked individuals to 2017. In
143 total, we recorded 2411 flowering events, and followed 606 individuals 1987–1996, and 228
144 individuals 2006–2017.

145 Recordings in each season started when shoots initiated growth in April and continued
146 every fifth day until all plants had finished flowering. At all visits, we recorded the number of
147 buds, open flowers, scars from aborted flowers and fruits. We also recorded if shoots had
148 been grazed. At each visit up to opening of the first flower, we assigned the size of the most
149 advanced bud in all individuals to one of four categories: small, medium, large or very large.

150 We used first flowering day (FFD), i.e. the day of year when the first flower was fully
151 unfolded with the banner petal folded upwards, as an estimate of flowering phenology of
152 individuals. For some of the study years we had data also on median flowering date, and in
153 these years FFD was strongly correlated with mean flowering day (Pearson's $r = 0.89$).
154 Information about the presence of open flowers from each visit allowed us to determine
155 during which 5-day interval an individual had started flowering. To estimate the most likely
156 date within this interval at which the first flower opened, we used information about the size
157 of the most developed bud at the beginning of this 5-day interval, and the number of open
158 flowers at the end of the interval (see Appendix S1 for details).

159 The number of open flowers produced by an individual was obtained from counts of
160 the number of buds, flowers, fruits and scars from aborted flowers at each recording. The
161 number of flowers in individuals that were grazed before the first recording of flower number
162 was estimated based on vegetative size (see Appendix S1 for details).

At the time of fruit maturation, we measured vegetative size and recorded the number of mature fruits, intact seeds and seeds damaged by seed predators. To estimate size of individuals, we measured the height and diameter of all shoots, calculated the volume of each shoot as $\pi \times (\text{diameter}/2)^2 \times \text{height}$, and summed the volumes of all shoots (see Appendix S1 for details). Fruits with mature seeds that had not yet opened were examined directly in the field. The number of seeds was counted and pre-dispersal seed predation was identified by the presence of larval entrance holes in a seed. After examination, seeds were dropped to the ground near the parent plant mimicking natural dispersal. Fruits that had opened prior to the recording were brought to the lab where the numbers of intact and preyed seeds were estimated based on the number of placentas and larval entrance holes (see Appendix S1 for details). The number of intact seeds produced by an individual was used as an estimate of fitness.

Weather data for March, April and May 1961–2017 was obtained from the Swedish Meteorological and Hydrological Institute (www.smhi.se). Data from the 26-year period (1961–1986) immediately before the onset of the study, and from 1997–2005 (when no recordings of phenology were made), was used to study climatic variation in the study area over a longer period than the one covered by the study. Daily mean, minimum and maximum temperature values were averaged from two meteorological stations: Oxelösund (58.6777 N, 17.1223 E, 41 km from the study population) and Södertälje (59.2142 N, 17.6289 E, 29 km from the study population). In cases where values for one station were missing, we imputed values using the relationship between values from the two stations (see Appendix S1 for details). Daily precipitation values were obtained from a third station located in Åda (58.9279 N, 17.5358 E, 5 km from the study population).

Statistical analyses

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187 To account for leap years, we converted calendar dates to number of days after the vernal
188 equinox. To describe the yearly position of the flowering season we calculated three
189 measures: start - the date at which 10% of plants had started flowering, mean - the mean first
190 flowering date of all individuals, and end - the date at which 90% of plants had started
191 flowering. We chose the 10 and 90% cutoffs to remove the extremes of the distribution of
192 flowering dates. Other cutoff values (20 and 80%) yielded similar results (not shown). We did
193 not have information about when individuals ceased flowering, and the end of the flowering
194 season thus does not represent the date at which most individuals had ceased flowering. The
195 yearly duration of the flowering season was defined as the number of days between the start
196 and the end of the flowering season.

197 To test our hypotheses regarding the effects of climate on flowering phenology and
198 selection, we calculated 12 variables from weather data: monthly averages of daily minimum,
199 mean and maximum temperatures, and monthly sums of precipitation, for March, April and
200 May in the year of flowering (Table S2). Including also climatic variables from the year
201 preceding flowering did not improve the fit of models (results not shown).

202 To assess how much of variation in FFD of individual plants was explained by year,
203 we calculated the R^2 -value from a linear model of FFD with year as a nominal variable. The
204 effects of climate on individual FFD-values were examined using linear mixed models with
205 plant individual as a random effect (because FFD was sometimes recorded in the same
206 individuals in multiple years) and number of flowers as a covariate. All climatic variables
207 were standardized by subtracting the mean and dividing by the standard deviation. We
208 constructed a candidate model set, using the 12 climatic variables as predictors. To avoid
209 overcomplex models, we did not include climate variables from other months or years, or
210 interactions between climatic predictors. We constrained the set of candidate models by
211 including only models with pairs of predictors having a Pearson correlation coefficient below

0.5. We ranked the resulting candidate models using the corrected second-order Akaike information criterion (AICc). To account for model uncertainty, we performed model averaging of parameter estimates across all models with $\Delta\text{AICc} < 2$. The effects of climate on the start, mean, end and duration of the flowering season were examined using linear models, including climatic predictors that had a significant effect in the averaged model of FFD for individual plants. To explore how long-term trends influenced our estimates of climate effects, we ran the models including also year as a covariate.

To assess total and direct phenotypic selection on FFD, we estimated phenotypic selection differentials and gradients. Fitness, in terms of the number of intact seeds, was relativized within years by dividing individual values by yearly means. Traits were standardized within years by subtracting yearly means and dividing by yearly standard deviations. Selection differentials were calculated for each year using univariate regressions of relative fitness on standardized FFD (Lande & Arnold 1983). Selection gradients were calculated using multiple linear regressions of relative fitness on standardized FFD and number of flowers. We included the number of flowers as a measure of resource state in these analyses (Rausher 1992). We estimated non-linear (quadratic and correlational) selection by examining the effect of quadratic and interaction terms in models including also linear terms (Lande & Arnold 1983; Arnold 1986). Quadratic regression coefficients and their standard errors were doubled (Stinchcombe *et al.* 2008).

Among-year differences in total and direct phenotypic linear selection on flowering time were evaluated using models including also the interaction between standardized FFD and year. The main effect of year was not included as fitness was relativized within years. Plant individual was included as a random effect. In this analysis, we considered only variation in linear selection because non-linear selection was significant only in very few years (see Results).

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237 To examine if variation in total phenotypic selection among years was related to
238 climatic conditions during spring, we ran models including standardized FFD and the
239 interactions between FFD and the 12 climatic variables. The main effects of climatic variables
240 on fitness were not included as fitness was relativized within years. Plant individual was
241 included as a random effect. For the effects of interactions with climatic variables, we
242 performed model selection and model averaging as explained above. Interaction terms that
243 had a significant effect in the averaged model (based on z , the Wald test statistic of the effect,
244 Table S3) were included in an analysis of deviance using the Wald Chi-square test. We used
245 analogous models to examine among-year variation in both total and direct selection, the
246 latter including also the standardized number of flowers.

247 To assess how much of the among-year variation in total and direct selection was
248 explained by climate, we regressed selection coefficients on the climatic variables whose
249 interactions with FFD were significant in the selection models, accounting for uncertainty in
250 the estimates (i.e. standard errors). Using the R package MCMCglmm (Hadfield 2010), we
251 calculated the posterior distribution of the variance in selection associated with the climatic
252 variables, and applied equation 12 in Hunter *et al.* (2018) to estimate the proportion of the
253 total variation in selection attributed to climate (as the mean of the posterior distribution).

254 To examine how effects of climatic variables on selection gradients for FFD
255 depended on changes in trait means, trait variances, fitness means and trait-fitness
256 relationships, we ran two types of models (cf. Hunter *et al.* 2018). First, we ran four linear
257 mixed models testing the effects of climatic variables and number of flowers on individual
258 yearly absolute FFD values, individual absolute distances from the yearly mean FFD and
259 individual absolute fitness, and the effects of number of flowers, FFD and the interactions of
260 FFD with climatic variables on individual absolute fitness. Each model included the set of
261 standardized climatic predictors that had significant effects on selection gradients for FFD,

and plant individual as a random effect. Second, to assess the relative importance of these four components for among-year variation in selection, we regressed yearly estimates of selection gradients for FFD on yearly estimates of trait means, trait variances, fitness means and FFD-fitness relationships (estimated by the regression coefficient of absolute fitness on FFD and number of flowers).

All statistical analyses were carried out in R version 3.5.1 (R Core Team 2018).

RESULTS

Climate in the study area varied considerably during the period 1961-2017 (Fig. 1A, Figs. S1-S4). Minimum, mean and maximum daily temperature increased significantly from 1961 to 2017 (Table S4, Figs. S1-S3).

Yearly mean FFD ranged from 48.0 to 71.6 (mean = 58.1) days after the vernal equinox (Fig. 1B). Differences among years explained 59% of the total variation in individual FFD-values. Differences in climate explained a large part of the variation in FFD of individuals, flowering occurring significantly earlier in years with higher mean temperatures in April and May, and in years with higher precipitation in March and April (Table 1A, Fig. 2). Climatic predictors explained most of the among-year variation in the start (72%), mean (76%), end (84%), and duration (57%) of the flowering season (Table 1B and C). The start, mean and end of the flowering season were earlier in years with higher mean temperatures in April and May (Table 1B, Fig. 3A-B). The flowering season was longer in years with higher mean April temperatures (Table 1C, Fig. 3C). There were no effects of precipitation on the position and duration of the flowering season. All effects of climatic variables were independent of potential long-term trends (Table S5).

Selection differentials showed that earlier-flowering individuals had significantly higher fitness in 18 of 22 years (mean = 0.44, range = -1.03 - 0.28), while non-linear

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selection was significant in only 2 years (Table 2A and S6). Selection gradients indicated that earlier flowering was favored in 21 of 22 years (mean = 0.28, range = 0.78 - 0.08), but selection differed in strength and was significant in 8 years (Fig. 1D, Table S7). Non-linear direct selection on FFD was significant in one year, and correlational selection on FFD × number of flowers in three years (Table S7)

Differences in total and direct selection among years were related to climatic conditions (Table 2B and S3, Fig. 4). There was stronger selection for early flowering in years with higher minimum April temperatures (Fig. 4A) and lower March precipitation (Fig. 4B). These two climatic variables explained 70 and 68 % of among-year variation in total and direct selection, respectively.

Climatic variables influenced trait means, trait variances, mean fitness, and absolute trait-fitness relationships (Tables S8-11). Variation in trait-fitness relationships explained most of the among-year variation in selection gradients for FFD (57%), while variation in trait means, trait variances and mean fitness among years had little effect (Table S12).

DISCUSSION

Twenty-two years of data for the perennial herb *Lathyrus vernus* showed that variation in climatic conditions during spring was related to both variation in flowering time and differences in phenotypic selection. Plants flowered earlier when springs were warmer and precipitation higher, and early flowering was most advantageous in years with high spring temperatures and low precipitation. Although climatic variation influenced also trait distributions and absolute fitness, among-year variation in phenotypic selection gradients was mostly the result of differences in trait-fitness relationships. Our results imply that climatic conditions are related to not only the expression of phenological traits, but also to natural selection on these traits. They also suggest that climate can elicit plastic responses that are

both co-gradient with the optimal phenotype in a given year, as for spring temperature in our study, and counter-gradient, as for precipitation.

Because survival in *L. vernus* is high, and time from seed germination to flowering is very long, we do not expect plants recruited during the study to have flowered. We thus assume that the observed variation in FFD-values among years was mainly due to plastic responses. Moreover, the fact that the rank order of FFD-values was weakly correlated among years (Pearson r , mean \pm SD = 0.16 ± 0.17 , $N = 22$ pairs of adjacent years), indicates that there are no consistently early- or late-flowering individuals.

Phenotypic responses to among-year climatic variation were strong, individual plants starting flowering on average 1.8 and 2.5 days earlier with a one-degree increase in mean temperature in April and May, respectively. This is a slightly weaker response to temperatures than the average shifts recorded in other studies (Fitter *et al.* 1995; Miller-Rushing & Primack 2008; Ibáñez *et al.* 2010; Iler *et al.* 2013; Tansey *et al.* 2017). Given the effects of spring temperatures on flowering phenology observed in our system, the recorded increase in spring temperatures over the last 57 years is expected to result in an earlier flowering of *L. vernus*. We did indeed observe such an advance in our data, but this trend was stronger than what would be expected based only on changes in temperature (results not shown). This suggests that other factors, such as increases in individual size during the study contributed to the trend.

For early-flowering species in seasonal environments, higher temperatures up to flowering start can extend the flowering season, while higher temperatures during flowering might shorten it (Theobald *et al.* 2017). In agreement, we found that for *L. vernus* the duration of the flowering season increased with increasing temperatures in April, the effect on the start of the flowering season being much stronger than on the end. In contrast, May temperature

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336 had roughly equal effects on the start and end of the flowering season, and did thus not affect
337 the duration.

338 In temperate environments with little drought stress during spring, the expected
339 phenological responses to precipitation are less clear than responses to temperature.
340 Experimental studies examining responses of phenology to precipitation have either found no
341 effects (Phoenix *et al.* 2001; Cleland *et al.* 2006), or variable responses (Peñuelas *et al.* 2004).
342 In *L. vernus*, higher precipitation during early spring was associated with an earlier flowering.
343 This might be because early-spring precipitation is associated with lower risk of frost damage
344 if falling as snow, and with relatively mild temperatures if falling as rain. However, the
345 effects of precipitation were weaker than those of temperature (plants flowered on average 0.5
346 and 0.2 days earlier with a 10 mm increase in precipitation in March and April, respectively).

347 In our study, the strength of selection on flowering time varied significantly over the
348 22 study years. Still, there was a tendency towards early flowering being favored in almost all
349 years, which agrees with many previous studies (Harder & Johnson 2009; Munguía-Rosas *et*
350 *al.* 2011; Austen *et al.* 2017). Several explanations for such persistent phenotypic selection for
351 earlier flowering have been suggested (Forrest 2014; Ehrlén 2015; Austen *et al.* 2017). For
352 example, selection estimates are often based on a single fitness component, such as seed
353 production in one year, and trade-offs between current and future reproduction can potentially
354 counteract observed effects on seed production. In *L. vernus*, such costs of reproduction are
355 small (Ehrlén & van Groenendael 2001), and therefore unlikely to eliminate the observed
356 selection for earlier flowering, but might potentially weaken it. Another suggested
357 explanation for observed persistent phenotypic selection for earlier flowering is that
358 individuals in a better condition both flower earlier and produce more seeds. In our study, we
359 did not experimentally examine how condition influenced the relationship between phenology
360 and fitness. However, the fact that the advantage of early-flowering differed significantly

among years with climatic conditions in spring, suggests that there is a causal relationship between flowering time and fitness, which is influenced by climate.

Identifying the environmental drivers of natural selection is key to understand past evolution and to predict the evolutionary implications of ongoing environmental changes. Yet, the factors underlying temporal variation in natural selection on many traits, like plant flowering phenology, remain poorly known (but see Siepielski *et al.* 2017). One suggested driver of variation in selection is climatic fluctuations (Réale *et al.* 2003; Siepielski *et al.* 2009, 2017; Visser *et al.* 2015; Siepielski *et al.* 2017; Marrot *et al.* 2018). In our study, which included one of the longest records of natural selection in plants to date, we found that most of observed among-year variation in selection could be explained by climatic variation. The strength of phenotypic selection for early flowering was associated with climatic variables, increasing with minimum temperatures in April and decreasing with precipitation in March. High temperatures early during the flowering period are likely to be associated with beneficial conditions, in terms of more rapid development, higher light availability (Kudo *et al.* 2008), decreased frost damage risk (Inouye 2008; Pardee *et al.* 2018), or higher pollinator availability (Kudo & Ida 2013), and early-flowering individuals might benefit most from this. Our results agree with the results of previous studies on birds (Visser *et al.* 2015; Marrot *et al.* 2018) and mammals (Réale *et al.* 2003), suggesting that stronger selection for earlier reproduction in warmer springs might occur in many groups of organisms in temperate seasonal environments.

Although precipitation has recently been suggested to be an important driver of selection (Siepielski *et al.* 2017), we still lack studies relating long-term variation in precipitation to variation in phenotypic selection. Our results show that higher March precipitation was associated with a weaker selection for early flowering. A possible explanation for this pattern is that early-spring precipitation is associated with smaller fitness

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386 advantages of earlier-flowering individuals because growth conditions and light availability
387 are poorer, and pollinator abundance lower, early in the season. Higher precipitation in early
388 spring might also increase the intensity of antagonistic interactions, such as mollusk herbivory
389 (Crawford-Sidebotham 1972), which have strong negative effects on fitness in *L. vernus*
390 (Ehrlén 2002).

391 To fully understand the causes of variation in natural selection, we need not only to
392 establish relationships between environmental factors and selection estimates, but also to
393 examine the ecological mechanisms underlying such relationships (Chevin *et al.* 2010, Hunter
394 *et al.* 2018). The pathways through which climate influences selection are still largely
395 unexplored, and studies examining how effects act through changes in trait distributions,
396 mean fitness and trait-fitness relationships are important to further our understanding (Hunter
397 *et al.* 2018). Our results show that even if climatic variation influenced both trait distributions
398 and mean fitness, most of the among-year variation in selection gradients for flowering time
399 was explained by effects of climatic factors on trait-fitness relationships. This means that
400 differences in selection among years were mostly a consequence of differences in trait-
401 absolute fitness relationships, and that climatic conditions during spring influenced how much
402 fitness increased with an earlier flowering.

403 Our 22-year assessment of among-year variation in selection on timing of flowering
404 shows that climatic variation not only influences the expression of phenological traits, but
405 also natural selection on these traits. They also show that effects of climate on selection are
406 largely the result of changes in trait-fitness relationships. Lastly, our results illustrate that
407 climatic variables can elicit plastic responses that are either co- or counter-gradient with the
408 optimal response (cf. Conover & Schultz 1995; Conover *et al.* 2009). In our study, higher
409 spring temperatures were associated with both an earlier flowering and an increased fitness
410 advantage for plants flowering early, suggesting that the plastic response was weaker than the

optimal response. In contrast, higher precipitation before the flowering season was associated with an earlier phenology, but also with a decreased fitness advantage for plants flowering early, suggesting that the observed response was stronger than the optimal response. An important broader implication of this study is thus that in order to predict long-term responses of life-history traits to global climate change, we need to consider both the immediate effects on trait expression and the potential evolutionary responses to climate-induced changes in selection.

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3 425 Tables
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5 426 Table 1: Results of models testing for effects of climatic factors on (A) first flowering date
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7 427 (FFD) of individual plants, and on three measures of the yearly position (B) and duration (C)
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9 428 of the flowering season of *Lathyrus vernus* over 22 years. The position of the flowering
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11 429 season was defined by the start, i.e. the date when 10% of the plants had started flowering, the
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13 430 mean, i.e. the mean first flowering date, and the end, i.e. the date when 90% of the plants had
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15 431 started flowering. The duration of the flowering season was defined by the number of days
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17 432 between the start and the end. In (A), linear mixed models with plant individual as a random
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19 433 effect and number of flowers as a covariate were used. Also shown are averaged estimates
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21 434 and z-values across all candidate models with $\Delta AICc < 2$, and values of marginal and
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23 435 conditional pseudo- R^2 (representing, respectively, the variance explained by the fixed effects
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25 436 and by the entire model, Nakagawa & Schielzeth 2013) and variance and standard deviation
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27 437 of the random effect for the best model resulting from model selection. Estimates are on
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29 438 standardized scales.
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| Response variable | Predictor variable | Estimate | Std. error | z | P |
|--|---------------------------|----------|------------|----------------|--------|
| A) FFD N = 2411 R ² _{marg} = 0.577 R ² _{cond} = 0.628 | Precipitation March | -0.716 | 0.105 | 6.85 | <0.001 |
| | Precipitation April | -0.345 | 0.125 | 2.77 | 0.006 |
| | Mean temperature March | -0.074 | 0.136 | 0.54 | 0.587 |
| | Max temperature March | -0.236 | 0.185 | 1.28 | 0.201 |
| | Mean temperature April | -2.188 | 0.134 | 16.29 | <0.001 |
| | Mean temperature May | -3.754 | 0.113 | 33.17 | <0.001 |
| | Number of flowers | -2.407 | 0.101 | 23.86 | <0.001 |
| | | Variance | | Std. deviation | |
| | Plant individual (random) | 2.708 | | 1.65 | |
| | | Estimate | Std. error | t | P |
| B) Position of the fl. season | | | | | |
| Start N = 22 R ² = 0.719 | Precipitation March | -0.993 | 0.888 | -1.12 | 0.279 |
| | Precipitation April | -0.353 | 0.938 | -0.38 | 0.712 |

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|--|--|------------------------|--------|-------|-------|--------|
| | | Mean temperature April | -4.465 | 1.001 | -4.46 | <0.001 |
| | | Mean temperature May | -3.531 | 0.823 | -4.29 | <0.001 |
| <hr/> | | | | | | |
| Mean N = 22 R ² = 0.764 | | Precipitation March | -0.803 | 0.724 | -1.11 | 0.283 |
| | | Precipitation April | -0.627 | 0.765 | -0.82 | 0.423 |
| | | Mean temperature April | -3.429 | 0.816 | -4.20 | <0.001 |
| | | Mean temperature May | -4.004 | 0.671 | -5.97 | <0.001 |
| <hr/> | | | | | | |
| End N = 22 R ² = 0.844 | | Precipitation March | -0.461 | 0.549 | -0.84 | 0.412 |
| | | Precipitation April | -0.810 | 0.580 | -1.40 | 0.180 |
| | | Mean temperature April | -2.324 | 0.618 | -3.76 | 0.002 |
| | | Mean temperature May | -4.555 | 0.508 | -8.96 | <0.001 |
| <hr/> | | | | | | |
| C) Duration or the fl. season | | | | | | |
| N = 22 R ² = 0.567 | | Precipitation March | 0.532 | 0.604 | 0.88 | 0.390 |
| | | Precipitation April | -0.458 | 0.638 | -0.72 | 0.483 |
| | | Mean temperature April | 2.141 | 0.680 | 3.15 | 0.006 |
| | | Mean temperature May | -1.024 | 0.560 | -1.83 | 0.085 |

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Table 2: Phenotypic selection models testing for (A) among-year differences and (B) effects of climatic factors on total (direct and indirect) and direct phenotypic selection on flowering time in *Lathyrus vernus* in 22 study years. Results are from (A) linear mixed models including the interaction first flowering date \times year, testing for among-year differences in selection differentials and selection gradients for first flowering date, and (B) linear mixed models including the interactions of first flowering date with climatic variables that were significant according to model selection (see Table S2). For direct phenotypic selection, number of flowers is included as a condition trait. Fitness was estimated by the number of intact seeds. Traits were standardized and fitness relativized within years before analyses. Plant individual was included as a random effect.

| Predictor variable | Total selection | | | Direct selection | | |
|--|-----------------|--------------------|--------|------------------|--------------------|--------|
| | χ^2 | Degrees of freedom | P | χ^2 | Degrees of freedom | P |
| A) First flowering date | 110.18 | 1 | <0.001 | 33.89 | 1 | <0.001 |
| Number of flowers | - | - | | 64.79 | 1 | <0.001 |
| First flowering date : year | 36.46 | 21 | 0.019 | 37.87 | 21 | 0.013 |
| B) First flowering date | 110.44 | 1 | <0.001 | 33.93 | 1 | <0.001 |
| Number of flowers | - | - | | 65.55 | 1 | <0.001 |
| First flowering date : Precipitation March | 12.48 | 1 | <0.001 | 14.80 | 1 | <0.001 |
| First flowering date : Precipitation April | 3.48 | 1 | 0.062 | 3.83 | 1 | 0.050 |
| First flowering date : Min temperature April | 10.73 | 1 | 0.001 | 9.90 | 1 | 0.002 |

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451 Figures

452 Figure 1: Variation in temperature, flowering time, fitness and selection on flowering time
453 during the study period. Symbols represent means (dots) and standard errors (whiskers) for
454 yearly values of (A) mean daily temperature during March (black circles), April (white
455 circles) and May (grey triangles), (B) first flowering date, calculated as the number of days
456 from the vernal equinox, (C) mean fitness, calculated as the number of intact, non-predated
457 seeds, and (D) linear selection gradient for first flowering date. In (A), data points are slightly
458 jittered along the x-axis in order to prevent overplotting. In (D), black circles indicate
459 significant selection gradients at $\alpha = 0.05$ and the grey dashed line indicates a selection
460 gradient not significantly different from 0. Variation in other climatic variables is presented in
461 Figures S1-S4.

462 Figure 2: Relationship between first flowering date of individual plants in different years and
463 (A) mean daily temperature in April, (B) mean daily temperature in May, (C) sum of
464 precipitation in March, and (D) sum of precipitation in April. Lines represent linear model
465 fits.

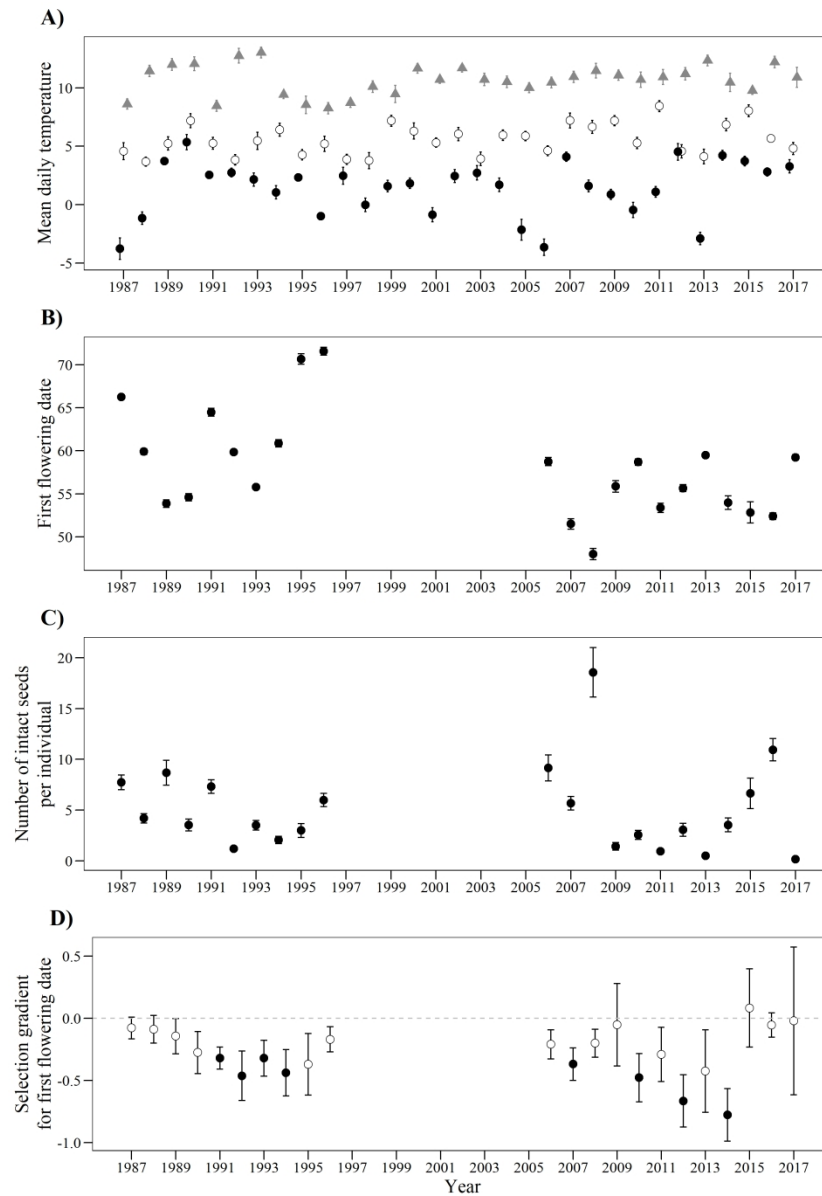
466 Figure 3: Relationship between the start, mean and end of the flowering season, and (A) mean
467 daily temperature in April and (B) mean daily temperature in May, and (C) between the
468 duration of the flowering season and mean daily temperature in April. The symbols in (A) and
469 (B) denote the start, i.e. the date when 10% of the plants had started flowering (triangles and
470 dashed line), the mean, i.e. the mean first flowering date (points and solid line) and the end,
471 i.e. the date when 90% of the plants had started flowering (stars and dotted line), of the
472 flowering season. The duration of the flowering season was measured as the number of days
473 between the start and the end of the season. Lines represent linear model fits.

474 Figure 4: Partial regression plots (i.e. added variable plots) showing the independent effects of
475 climatic variables on selection gradients for first flowering date (FFD) in *Lathyrus vernus* in

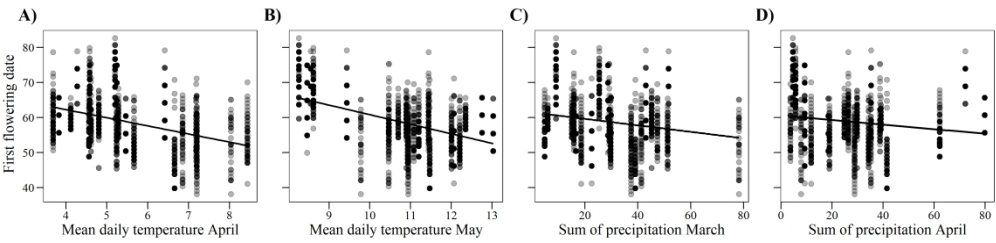
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22 study years. The effect of standardized first flowering date on relative fitness, measured as
the relative number of intact seeds, is shown for (A) different minimum daily temperatures in
April and (B) different sums of precipitation in March. Lines represent linear model fits.
Values were calculated from models in Table 2.

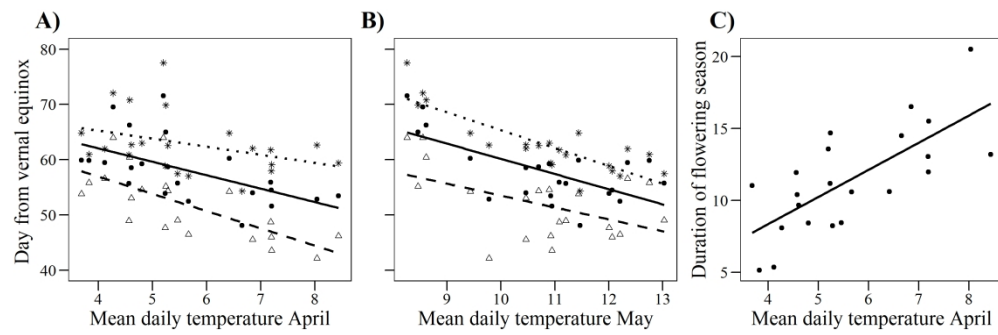
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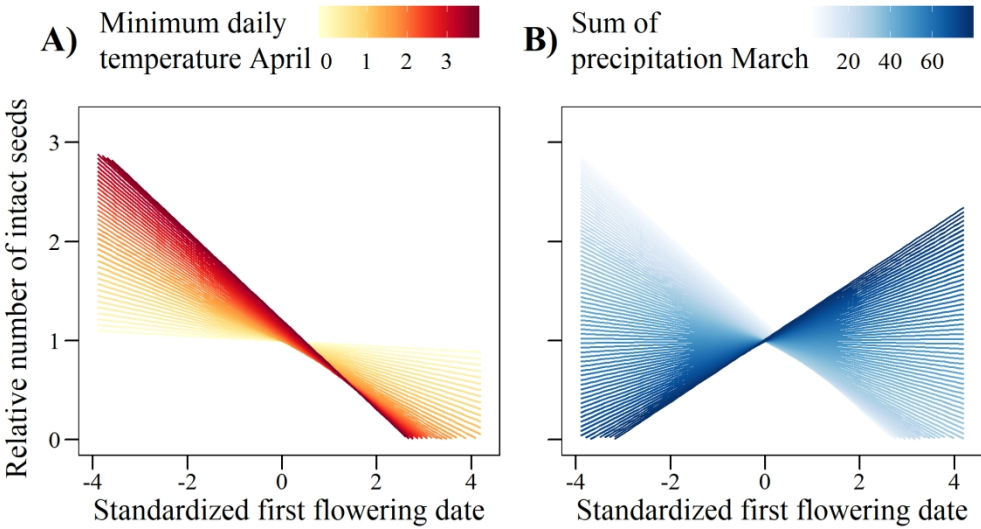
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