

Factor Zoo (.zip)

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KEY FINDINGS

- An iterative factor selection approach reveals that 15 factors are sufficient to explain the available factor zoo alpha.
- This evidence suggests that many factors are redundant, but also that merely using a handful of factors, as in common asset pricing models, is insufficient.
- The selected style factor clusters are stable, but new factor definitions often supersede older ones within clusters, emphasizing the relevance of continuous factor innovation.

ABSTRACT

The number of factors allegedly driving the cross-section of stock returns has grown steadily over time. The authors explore how much this “factor zoo” can be compressed, focusing on explaining the available alpha rather than the covariance matrix of factor returns. Our findings indicate that about 15 factors are enough to span the entire factor zoo. This evidence suggests that many factors are redundant, but also that merely using a handful of factors, as in common asset pricing models, is insufficient. While the selected factor styles remain persistent, the specific style representatives vary over time, underscoring the importance of continuous factor innovation.

To explain the cross-section of stock returns, the asset pricing literature advocates factor models that comprise the factors deemed most representative and relevant. The capital asset pricing model (CAPM) developed by Sharpe (1964), Lintner (1965), Mossin (1966), and Treynor (1961), is one of the earliest factor models, a one-factor model built around the equity market factor. Despite its theoretical appeal, the CAPM does not perform well in explaining cross-sectional differences in average stock returns. Most prominently, it fails to account for the proper pricing of size (Banz 1981) and value effects (Basu 1977 or Rosenberg, Reid, and Lanstein 1985), leading Fama and French (1993) to propose a three-factor model consisting of market, size, and value factors. For many years, this model was the industry standard, sometimes augmented with the momentum factor of Jegadeesh and Titman (1993), as in Carhart (1997).

However, over the last 25 years, hundreds of factors have emerged in the literature, all of which allegedly offer a unique new source of return. Cochrane (2011) aptly characterizes the state of play as a “zoo of factors” that needs to be tamed and structured. The existing factor models of Fama and French (1993) and Carhart (1997) cannot explain many of the new factors, so Fama and French (2015, 2018) extend these models to five- and six-factor models by adding investment

(Cooper, Gulen, and Schill 2008) and profitability (Novy-Marx 2013) factors. These factor models compete with alternative four-factor models such as the Hou, Xue, and Zhang (2015) *q*-factor model and the Stambaugh and Yuan (2017) mispricing model or the revised six-factor model of Barillas et al. (2020). Although these models use different factors, there seems to be a consensus among leading academics that most of the factor zoo can be explained by parsimonious models consisting of just four to six factors, (Bartram et al. 2021).

Our study of the factor zoo investigates the question of how many factors it takes to compress the factor zoo, that is, substantially reducing the number of factors without losing (much) information about the tangency portfolio of the entire zoo. To this end, we iteratively identify factors that capture most of the available alpha in the factor zoo. Specifically, the first iteration of our identification strategy augments the CAPM with that factor for which the resulting two-factor model reduces the remaining candidate factor alphas most. The subsequent iteration augments this model further to a three-factor model that captures most of the remaining factor alphas. Sequentially adding factors, we ultimately arrive at a factor model that eliminates all remaining factor alphas. Note that our procedure echoes the approach followed in the early literature, where the CAPM was initially extended by size, value, and momentum factors and later with investment and profitability factors. In contrast, we systematically consider all available candidate factors documented to date until the factor zoo is sufficiently compressed, leading to alternative paths and insights.

A challenge in analyzing and structuring the factor zoo is to completely reconstruct the existing factors in the literature. To this end, Chen and Zimmermann (2022) as well as Jensen, Kelly, and Pedersen (2023) replicate the vast majority of existing factors and publish open-source databases to facilitate further research. Both studies document that many (if not most) of the proposed factors with high statistical relevance can indeed be replicated, challenging the often-claimed replication crisis in modern finance (Hou, Xue, and Zhang 2020). Given that Jensen et al. (2023) also provide international factors, our study sources factor data from their database.

Our main findings can be summarized as follows. First, using a comprehensive set of 153 US equity factors, we find that a factor model consisting of 15 factors spans the entire factor zoo. The selected 15 factors originate from 8 out of the 13 factor style clusters, speaking to the heterogeneity of the factor set. Second, iterative factor models also beat common academic models when they contain the same number of factors by selecting alternative value, profitability, investment, or momentum factors or including alternative factor style clusters, such as seasonality or short-term reversal. When comparing the existing academic models, we find that the Barillas et al. (2020) revised six-factor model explains most of the available alpha in the factor zoo. Third, when repeating the factor selection to factors as they become available over an expanding window, we likewise recover a diverse set of selected factors. Specifically, newly published factors sometimes supersede older factor definitions, emphasizing the relevance of continuous factor innovation based on new insights or newly available data. Fourth, using equal-weighted factors as opposed to capped value-weighted factors requires more than 30 factors to span the factor zoo, indicating that equal-weighted factors exhibit stronger and more diverse alphas. Finally, applying our factor selection strategy to a set of global factors results in a similar set of selected factors. Although the factor models selected based on global data shrink the alpha for US and World ex US subuniverses, they perform better for the US market, implying that international factors exhibit larger and more diverse alpha.

We contribute to the literature in several ways. First, we propose a simple yet effective method to identify the important alpha contributors in the factor zoo. The resulting factor sets are relevant from a practitioner's perspective, as they represent the available factor zoo alpha with the minimum number of factors. Our approach

differs from previous work on variation in factor returns, for example Bessembinder, Burt, and Hrdlicka (2021) and Kozak, Nagel, and Santosh (2018), who mainly investigate the covariance structure in factor returns. These statistical factor studies based on principal component analysis (PCA) methods typically identify latent factors that describe the covariance structure rather than information about the means (that is, the factors' return level). For instance, consider a hypothetical factor that generates a 1% return every month at zero variance. While this factor would not be considered relevant from a PCA perspective, it is genuinely relevant from a factor premium perspective. In that vein, Lettau and Pelger (2020) develop an alternative risk premium PCA (RP-PCA) approach that incorporates information in the first and second moments of data. Yet, rather than identifying new latent factors in the factor zoo, we are eager to learn about the most relevant factors from an alpha perspective.

Second, we contribute to the debate about the ideal factor model size by consistently identifying 10 to 20 factors over time, depending on the selected statistical significance level. This contrasts with leading academic factor models, which typically only comprise between three and six factors (Barillas et al. 2020; Fama and French, 1993, 2015, 2018; Hou, Xue, and Zhang 2015; Stambaugh and Yuan 2017). Interestingly, our results are more in line with the results of studies that apply cross-sectional regressions. For instance, Green, Hand, and Zhang (2017) find that 12 out of 94 characteristics are reliably independent determinants of return among non-microcap stocks, and that 11 of the 12 independent characteristics lie outside prominent benchmark models. Similarly, Jacobs and Müller (2018) find a high degree of dimensionality in international stock returns. Also, the recent evidence from machine-learning models indicates that many characteristics matter for predicting individual stock returns (see Gu, Kelly, and Xiu 2020; Hanauer and Kalsbach 2023; Tobek and Hronec 2021).

Third, we contribute to the literature on global versus local pricing. Griffin (2002), Fama and French (2012), and Hanauer and Linhart (2015) document that local factors dominate global factors in explaining local return patterns. In contrast, Tobek and Hronec (2021) and Hanauer and Kalsbach (2023) find that the out-of-sample performance of machine-learning models for non-US markets is better for global models than for local models. We emphasize the regional impact on factor selection and model construction. While it takes about 6 to 15 factors to span the US factor zoo regardless of the significance level, the global factor zoo is characterized by a similarly sized set of highly significant factors, but cannot be compressed to less than 25 to 30 factors at a lower significance level. Lastly, we document a set of global factors that spans the US factors while it needs more factors to span World ex US factors.

METHODOLOGY

Identifying Factors That Compress the Factor Zoo

Our goal is to determine the minimum number of factors to explain all factor alphas. From the perspective of a systematic investor, it is worthwhile to identify a factor model that captures as much alpha as possible because this factor model could guide portfolio allocation for harvesting the underlying factor premiums.

Given the large number of factors put forward in the literature, evaluating competing models is challenging. Valuing the contribution of individual factors vis-à-vis existing alternative factors and quantifying the incremental value added of (potentially) non-nested (i.e., not all of the factors in one model are contained in the other model), competing factor models is still an open challenge. Previous work typically differentiates between left-hand-side (LHS) and right-hand-side (RHS) approaches. The former evaluates models by their intercepts (alphas) in time-series regressions of LHS test

portfolios' excess returns. Prominent examples for test assets are two-way 5×5 sorts of stocks on size and either book-to-market sorts, momentum, or mispricing (see, e.g., Fama and French 2015, 2016; Stambaugh and Yuan 2017) or decile portfolios using various characteristics (see, e.g., Hou, Xue, and Zhang 2020). However, one limitation of this approach is that the inferences are dependent on the LHS test portfolios and might vary across different test sets (see Barillas and Shanken 2017).

Conversely, Barillas and Shanken (2017) demonstrate that the key in comparing models is how well models price the factors not included in the model and that, surprisingly, the choice of test assets is irrelevant. For nested models, the RHS approach is based on spanning regressions. Specifically, new candidate factors are regressed against the existing model factors to test whether they increase the opportunity set. If the corresponding intercept is non-zero, the tested factor contains unexplained information and therefore extends the efficient portfolio frontier. An early proof of the RHS approach is shown, for instance, in Fama (1998), and the approach is applied for model comparisons in Barillas et al. (2020) and Hanauer (2020).

In order to identify a factor model that spans the whole factor zoo from an alpha perspective, we follow a very intuitive and effective nested model approach: We iteratively add new factors to an extending factor model until all remaining alphas in the cross-section of equity factors are rendered insignificant. Our starting point is the CAPM, and we add that factor for which the resulting two-factor model reduces the remaining factor alphas most, measured by the lowest GRS (Gibbons, Ross, and Shanken 1989) statistic. Please note that this selection criterion is equivalent to selecting the factor with the largest alpha t -statistic for the existing model. Once identified, the factor is permanently added to the factor model, and we repeat the procedure based on the resulting augmented factor models until there are no significant contributors left. Formally, the selection strategy can be stated as shown in the following subsection.

Factor Selection Steps

Step 1. Set $l := 0$ and start spanning the factor zoo using the CAPM

$$f_i = \alpha_i + \beta_m r_m + \varepsilon_i \quad i = 1, \dots, N \quad (1)$$

where r_m is the excess market return and N the size of the factor zoo beyond the market.

Step 2. Test $N - l$ different augmented factor models that each add one of the remaining factors, labeled f^{test} , to the model from the previous iteration:

$$f_i = \alpha_i + \beta_m r_m + \sum_{k=1}^l \beta_k f_k + \beta^{\text{test}} f^{\text{test}} + \varepsilon_i \quad i = 1, \dots, N - l \quad (2)$$

Step 3. Sort the tested factor models based on their explanatory power (as quantified by their GRS statistic, see next section) and select the strongest model.

Step 4. Set $l := l + 1$ and calculate the number of remaining factor alphas $n(\alpha)_{t>x}$ based on the augmented factor model as

$$n(\alpha)_{t>x} = |\{\alpha_i | t(\alpha_i) > x\}| \quad i = 1, \dots, N - l \quad (3)$$

where x is the selected significance threshold.

Step 5. Stop if $n(\alpha)_{t>x} = 0$, that is, if the remaining factors are statistically indif-ferent from zero. Continue with Step 2 otherwise.

A few things need to be considered when following this iterative nested approach. First, how does one measure the value-added of a tested factor and compare different nested models. Given the linear nature of factor models, it is intuitive to follow a regression-based approach to classify the individual factors' strengths. In the next section, we discuss different metrics used in the literature and rationalize our choice. But note that the approach discussed here also allows for alternative methods to evaluate nested factor models.

Second, a stopping criterion needs to be chosen to effectively pinpoint the number of factors needed to explain all alphas in the factor zoo. We use a straightforward criterion that requires the total number of remaining significant factor alphas to be zero. That is, once a new factor model is identified, we test all remaining factors against this model and determine the alphas for the remaining candidate factors. If the newly added factors are of significance, the number of remaining significant factor alphas should decrease during the process. Alternative criteria could be the significance level of the newly added factor based on the statistical test to identify that factor. That is, if the new factor does not pass a significance threshold, it should not be considered a strong factor and therefore should not be added to the model. One caveat of this approach is the large number of regressions needed to identify a factor model that spans the factor zoo. Addressing such data-mining concerns and accounting for potential misspecifications, we resort to higher statistical thresholds. Harvey, Liu, and Zhu (2016) deem a t -statistic of 3.00 appropriate to account for resulting biases and data-mining concerns. Therefore, we run our analysis using the standard thresholds of $t > 1.96$ as well as a more conservative one where $t > 3.00$.

Evaluating Factor Models

When looking to span the whole factor zoo, one is dealing with nested models. A common metric in this field is the GRS statistic of Gibbons, Ross, and Shanken (1989), which produces a test of whether candidate factors help to improve a given model's explanation of expected returns. Specifically, the GRS test investigates whether the alphas of the test assets are jointly different from zero. The GRS test is widely used in empirical finance and has become a standard tool for evaluating the performance of asset pricing models as in, for instance, Fama and French (1996, 2015) and Stambaugh and Yuan (2017). Formally, the empirical GRS statistic is given as follows: Consider an asset pricing model (2) with K factors, N test assets, and τ return observations for each time series. We follow Fama and French (2018) and define the maximum squared Sharpe ratio for the intercepts as

$$Sh^2(\alpha) = \alpha^\top \Sigma^{-1} \alpha \quad (4)$$

where $\Sigma = e^\top e / (\tau - K - 1)$ is the covariance matrix of the regression residuals e . The maximum squared Sharpe ratio for the factors of the given model is defined as

$$Sh^2(f) = \bar{f}^\top \Omega^{-1} \bar{f} \quad (5)$$

where \bar{f} is the model's average factor returns and $\Omega = (f - \bar{f})^\top (f - \bar{f}) / (\tau - 1)$ is the covariance matrix of the model's factors. The GRS test statistic is calculated as

$$F_{GRS} = \frac{\tau(\tau - N - K)}{N(\tau - K - 1)} \frac{Sh^2(\alpha)}{(1 + Sh^2(f))} \quad (6)$$

with $F_{GRS} \sim F(N, \tau - N - K)$. The null hypothesis of the GRS test is that all test assets' alphas are strictly equal to zero. If the GRS test statistic is greater than the critical value of the F -distribution at a given significance level, then the null hypothesis is

rejected, indicating that the factor model does not adequately explain the variation in test asset returns.

Note that the GRS statistic is crucially determined by the ratio of $Sh^2(\alpha)$ and $Sh^2(f)$. When evaluating factor models, the goal is to identify a model that observes the smallest maximum squared Sharpe ratios for the alphas and thus captures most of the return variation through its systematic components. Given the relation in Equation (6), Barillas and Shanken (2017) propose to use the factors' maximum squared Sharpe ratio to evaluate the power of a set of candidate models. Fama and French (2018) further analyze the resulting implications and conclude that the model reducing $Sh^2(\alpha)$ the most is also the model with the highest $Sh^2(f)$, consistent with Barillas and Shanken (2017). Thus, our empirical analysis will not only report GRS statistics and their associated p -values but also $Sh^2(f)$ s and average absolute alphas, labeled $Avg|\alpha|$.

COMPRESSING THE FACTOR ZOO

Our empirical study is based on the global factor data of Jensen, Kelly, and Pedersen (2023, hereafter JKP), covering 153 factors using data from 93 countries.¹ The set of factors extends the set of factors in Hou, Xue, and Zhang (2020), is similar to that of Chen and Zimmermann (2022), and is thus a meaningful representation of the factor zoo.² The JKP database provides one-month holding period factor returns based on the most recent accounting data at a given point in time.

To enable covering all 153 factors in our study, we start our investigation of US capped value-weighted factors in November 1971, and our sample period ends in December 2021. In capped value-weighted factors, stocks are sorted into characteristic terciles each month and the capped value-weighted tercile returns are calculated as the market equity-weighted portfolio returns capped at the NYSE 80th percentile. The factor return is then defined as the high-tercile minus low-tercile return (see JKP). This factor construction is designed to create tradable yet balanced portfolios that are dominated by neither mega nor tiny caps. Note that we check for robustness of our main analysis with respect to alternative factor weighting schemes in a later section.

Exhibit 1 provides an overview of all factors' annualized alphas based on monthly CAPM regressions. All factors are clustered into 13 categories as identified by JKP based on hierarchical agglomerative clustering (Murtagh and Legendre 2014), and cluster names are driven by the most representative characteristics. We observe mostly positive annualized alphas for all clusters but the Low Leverage cluster. The average alpha across all factors is 3.51% a year, and alphas are fairly evenly distributed across and within clusters.

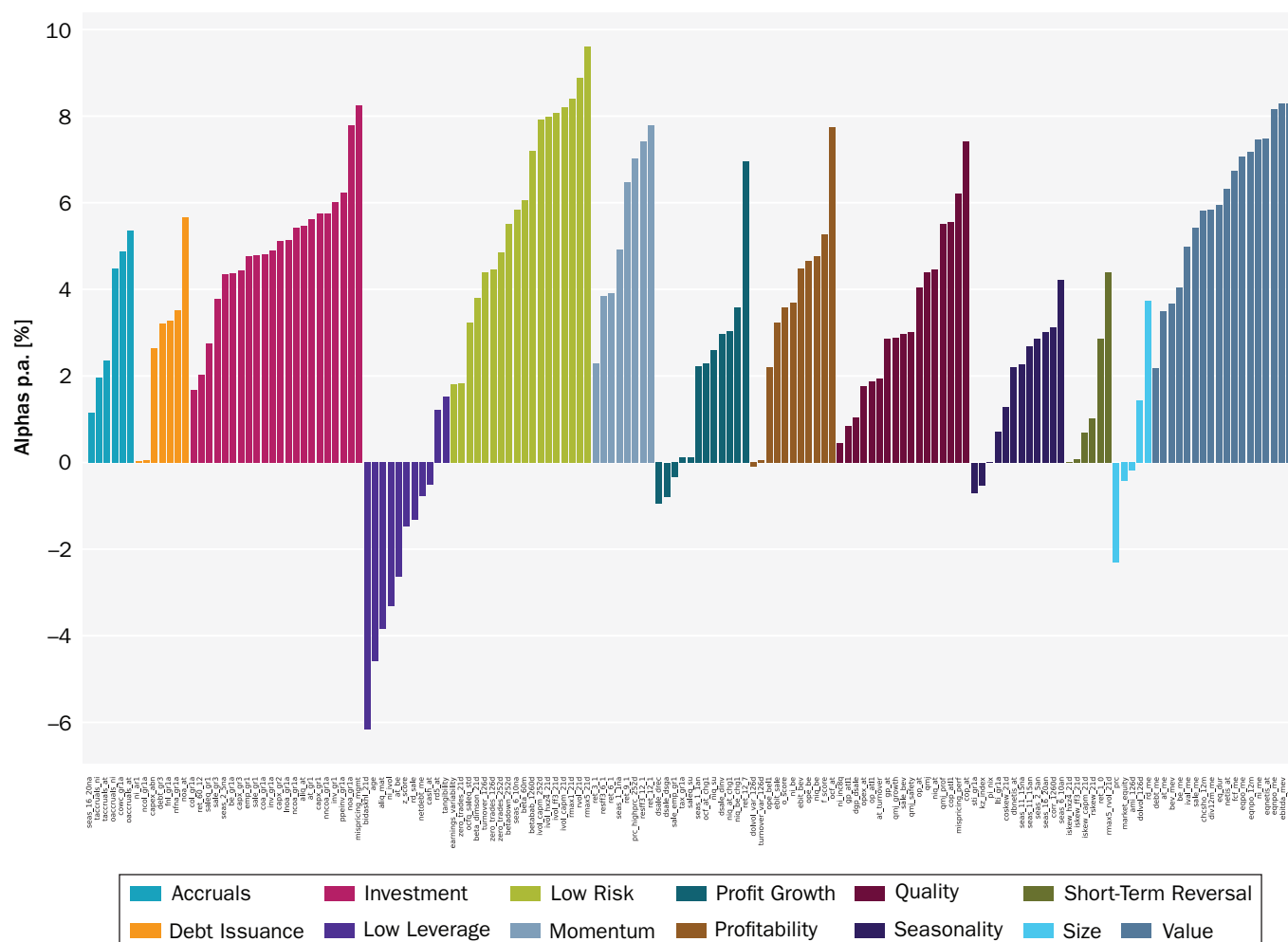
It is not surprising that most of the factors exhibit a significant alpha premium—that is, indicating incremental power beyond the market return—and thus are deemed to be relevant factors to begin with. Blitz (2023) further analyzes this set of factors regarding their market risk, performance cyclicity, and inherent seasonal and momentum effects in the cross-section of factor returns. Yet, the question of the incremental alpha contribution of individual factors within the factor zoo is an open question.

¹The factor data are publicly available at <https://jkpfactors.com/>. An extensive overview of all included factors, their descriptive statistics, as well as the detailed code used to compute them, can be found in JKP.

²There is a variety of choices that a factor researcher must make in empirical research. For instance, Bessembinder, Burt, and Hrdlicka (2022) highlight the impact of weighting methods and of the number of quantile portfolios underlying the factor portfolios. While they report some differences in the resulting factor portfolio returns across the two mentioned databases, they confirm the statistically significant out-of-sample power of both databases' factors to forecast.

EXHIBIT 1

Factor Alphas



NOTES: This exhibit depicts annualized CAPM alphas for the US factor zoo. The underlying excess market return is sourced from https://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html. The factor style clusters are based on Jensen, Kelly, and Pedersen (2023). The sample period is November 1971 to December 2021.

Main Results

Exhibit 2 depicts our main results following the iterative factor selection process described earlier. We report the selected factors, their associated factor style cluster, the GRS statistics and corresponding p -values, $p(\text{GRS})$, the average absolute intercept $\text{Avg}|\alpha|$, the maximum squared Sharpe ratio for the model's factors $\text{Sh}^2(f)$, as well as their Sharpe ratios (SR). Columns 10–11 refer to the number of significant factor alphas after controlling for the specified factor model and thus indicate the incremental explanatory power of the selected factor. The column labels $t > 2$ and $t > 3$ refer to our iterative factor selection with a significance alpha threshold of $t(\alpha) > 1.96$ or $t(\alpha) > 3.00$, respectively. Note that we also report the number of significant factors under common factor models in the subsequent exhibit.

The starting point for our iterative factor selection is the CAPM. Based on this one-factor model, we clearly reject the null of the GRS test of all factors' alphas being statistically indistinguishable from zero (GRS statistic of 4.36, p -value 0.00). The CAPM leaves plenty of significant factor alphas, regardless of the selected threshold

EXHIBIT 2

Iterative Factor Selection

No.	Factor	Description	Cluster	GRS	$p(\text{GRS})$	$\text{Avg} \alpha $	$\text{Sh}^2(f)$	SR	$t > 2$	$t > 3$
	RMRF	Excess market return	Market	4.36	0.00	3.91	0.02	0.14	105	86
1	<i>cop_at</i>	Cash-based operating profits-to-book assets	Quality	3.54	0.00	3.94	0.15	0.39	101	78
2	<i>noa_gr1a</i>	Change in net operating assets	Investment	2.98	0.00	2.15	0.27	0.51	65	34
3	<i>saleq_gr1</i>	Sales growth (1 quarter)	Investment	2.69	0.00	1.51	0.33	0.58	42	10
4	<i>ival_me</i>	Intrinsic value-to-market	Value	2.49	0.00	1.51	0.39	0.62	39	14
5	<i>resff3_12_1</i>	Residual momentum $t-12$ to $t-1$	Momentum	2.31	0.00	1.43	0.44	0.66	35	15
6	<i>seas_6_10an</i>	Years 6–10 lagged returns, annual	Seasonality	2.11	0.00	1.24	0.50	0.71	27	9
7	<i>debt_me</i>	Debt-to-market ratio	Value	1.98	0.00	1.47	0.54	0.74	37	7
8	<i>seas_6_10na</i>	Years 6–10 lagged returns, nonannual	Low Risk	1.87	0.00	1.30	0.58	0.76	25	3
9	<i>zero_trades_252d</i>	Number of zero trades (12 months)	Low Risk	1.78	0.00	0.77	0.61	0.78	13	1
10	<i>cowc_gr1a</i>	Change in current operating working capital	Accruals	1.68	0.00	0.88	0.65	0.81	14	3
11	<i>nncoa_gr1a</i>	Change in net noncurrent operating assets	Investment	1.55	0.00	0.70	0.70	0.84	7	1
12	<i>ocf_me</i>	Operating cash flow-to-market ratio	Value	1.48	0.00	0.62	0.73	0.85	5	1
13	<i>zero_trades_21d</i>	Number of zero trades (1 month)	Low Risk	1.40	0.01	0.80	0.76	0.87	11	1
14	<i>turnover_126d</i>	Share turnover	Low Risk	1.28	0.03	0.77	0.82	0.90	9	2
15	<i>rmax5_vol_21d</i>	Highest 5 days of return scaled by volatility	Short-Term Rev.	1.19	0.09	0.63	0.85	0.92	3	0
16	<i>seas_11_15na</i>	Years 11–15 lagged returns, nonannual	Seasonality	1.16	0.14	0.60	0.87	0.93	2	0
17	<i>o_score</i>	Ohlson O-score	Profitability	1.13	0.18	0.67	0.89	0.94	4	0
18	<i>niq_at</i>	Quarterly return on assets	Quality	1.09	0.26	0.59	0.91	0.95	0	0
19	<i>seas_16_20an</i>	Years 16–20 lagged returns, annual	Seasonality	1.07	0.31	0.56	0.92	0.96	0	0
20	<i>ni_ar1</i>	Earnings persistence	Debt Issuance	1.05	0.36	0.56	0.93	0.97	1	0
21	<i>ivol_ff3_21d</i>	Idiosyncratic volatility FF 3-factor model	Low Risk	1.03	0.42	0.48	0.95	0.97	2	0
22	<i>ni_me</i>	Earnings-to-price ratio	Value	0.99	0.52	0.45	0.97	0.98	0	0
23	<i>dsale_dinv</i>	Change sales minus change inventory	Profit Growth	0.97	0.57	0.42	0.98	0.99	0	0
24	<i>ni_be</i>	Return on equity	Profitability	0.96	0.62	0.46	0.99	0.99	1	0
25	<i>noa_at</i>	Net operating assets	Debt Issuance	0.93	0.69	0.46	1.01	1.00	0	0
26	<i>age</i>	Firm age	Low Leverage	0.91	0.73	0.44	1.01	1.01	0	0
27	<i>ret_12_1</i>	Price momentum $t-12$ to $t-1$	Momentum	0.90	0.76	0.41	1.02	1.01	0	0
28	<i>aliq_mat</i>	Liquidity of market assets	Low Leverage	0.89	0.78	0.39	1.03	1.02	0	0
29	<i>nfna_gr1a</i>	Change in net financial assets	Debt Issuance	0.88	0.80	0.39	1.04	1.02	0	0
30	<i>at_me</i>	Assets-to-market ratio	Value	0.87	0.83	0.39	1.05	1.02	0	0

NOTES: This exhibit reports the results for an iterative factor model construction where the k -th iteration augments the model by the factor in row k . It shows the GRS statistic of Gibbons, Ross, and Shanken (1989) and its p -value, $p(\text{GRS})$; the annualized average absolute intercept $\text{Avg}|\alpha|$ in percentage, the maximum squared Sharpe ratio for the model's factors $\text{Sh}^2(f)$, as well as its Sharpe ratio, SR. Columns 10–11 refer to the number of remaining significant factor alphas after controlling for the specified factor model; $t > 2$ and $t > 3$ control the factor zoo based on the iterative model using a significance alpha threshold of $t(\alpha) > 1.96$ and $t(\alpha) > 3.00$, respectively. The sample period is November 1971 to December 2021.

(105 factors for $t > 2$ and 86 factors for $t > 3$). In the next step, our approach identifies *cash-based operating profits-to-book assets* (*cop_at*) as the strongest factor in the factor zoo. Adding this quality factor to the market factor still yields a highly significant GRS statistic of 3.54, but the absolute GRS value is clearly reduced. Yet, there are 101 ($t > 2$) or 78 factors ($t > 3$) with an average absolute alpha of 3.94% a year in this two-factor model.

The second iteration identifies *change in net operating assets* (*noa_gr1a*) as the strongest factor among the remaining factor zoo contenders. The resulting three-factor model leaves 65 ($t > 2$) or 34 ($t > 3$) significant factor alphas. The average absolute alpha drops to 2.15% a year, while the GRS statistic still remains highly significant (2.98 at a p -value of 0.00).

Iterating further, the factor model increases by construction. While Exhibit 2 documents the impact of adding the 30 most relevant factors, it only takes half of that number to span the whole factor zoo. Adding the 15th factor (*highest five days of return*

scaled by volatility, *rmax5_rvol_21d*) the number of remaining significant alphas drops to zero ($t > 3$). Even with the less-strict threshold of $t > 2$, it only takes a total of 18 iterations to render the remaining alphas insignificant. These cut-off numbers are in line with the alternative stopping criterion based on the significance level of the GRS statistic: *highest five days of return scaled by volatility (rmax5_rvol_21d)* is also the first factor exceeding the 5% significance level with a p -value of 0.09. Overall, these results indicate that it only takes 15 to 18 additional factors to span the factor zoo from an alpha perspective, regardless of the stopping criterion.

Against this backdrop, we wonder how our iterative factor selection compares with classic academic factor models. Exhibit 3 reports the number of significant factor alphas under well-known models (measured at a significance threshold of $t > 3$). Here, columns FF5 and FF6 refer to the Fama and French (2015) five-factor model, where FF6 augments the latter by a momentum factor. HXZ, BS, and SY refer to the Hou, Xue, and Zhang (2015) q -factor model, the Barillas et al. (2020) revised six-factor model, and the Stambaugh and Yuan (2017) mispricing model, respectively.³

Relative to the CAPM (86 alphas with $t > 3.00$), the Fama and French five- and six-factor models, the q -factor model, and the mispricing model still leave between 58 and 69 alphas significant. However, the revised six-factor model of Barillas et al. (2020) substantially reduces the number of significant alphas to 33. This superior explanatory power stems from one main difference with the other models, namely, the inclusion of the *cash-based operating profits-to-book assets (cop_at)* factor, which also emerged as a key factor in our iterative factor selection approach. The power of the iterative factor model approach is revealed when we compare the academic factor models with the iterative factor models that contain the same number of factors. While the iterative model with 4 factors merely leaves 10 significant factors, the 4-factor models of Hou, Xue, and Zhang (2015) and Stambaugh and Yuan (2017) still leave 60 and 55 significant alphas, respectively. Similarly, 14 and 15 remain significant for the iterative model with 5 and 6 factors, while for the Fama and French 5- and 6-factor models and the Barillas et al. (2020) revised 6-factor model, 65, 53, and 29 alphas remain significant, respectively. These results reinforce that the selected factors do carry information over and above classic academic factors. More specifically, *cash-based operating profits-to-book assets (cop_at)* is the only factor from the 15 selected factors that is also contained in one of the common academic factor models, namely the Barillas et al. (2020) revised 6-factor model. All the other selected factors represent either alternative value, profitability, investment, or momentum definitions or stem from alternative factor style clusters, such as seasonality or short-term reversal, that offer alpha beyond common factor models (see Blitz et al. 2023).

Note that the 15 selected factors emerge from 8 out of the 13 defined factor style categories and no factor from the remaining five categories is considered; see the highlighted factor bars in Exhibit 4. Moreover, the selected factors are not necessarily those with the highest CAPM alpha in a given factor style cluster; in fact, this only applies to the value, quality, short-term reversal, and seasonality clusters. Notably, while five of the eight represented factor clusters merely feature a single factor, the value, low risk, and investment clusters are represented by three to four factors.

³For consistency, we base these models on those capped value-weighted factors from the JKP database that are most similar to the actual factors used in the original factor model literature. Specifically, the proxies for the SMB, HML, RMW, and CMA factors of Fama and French (2015) are *market equity (market_equity)*, *book-to-market equity (be_me)*, *operating profits-to-book equity (ope_be)*, and *asset growth (at_gr1)* from the JKP database. For the models of Hou, Xue, and Zhang (2015) and Barillas et al. (2020) we use *quarterly return on equity (niq_be)* and *cash-based operating profits-to-book assets (cop_at)* as profitability factors, respectively. The two mispricing factors for the Stambaugh and Yuan (2017) model are *management (mispricing_mgmt)* and *performance (mispricing_perf)*.

EXHIBIT 3

Factor Relevance in Alternative Models

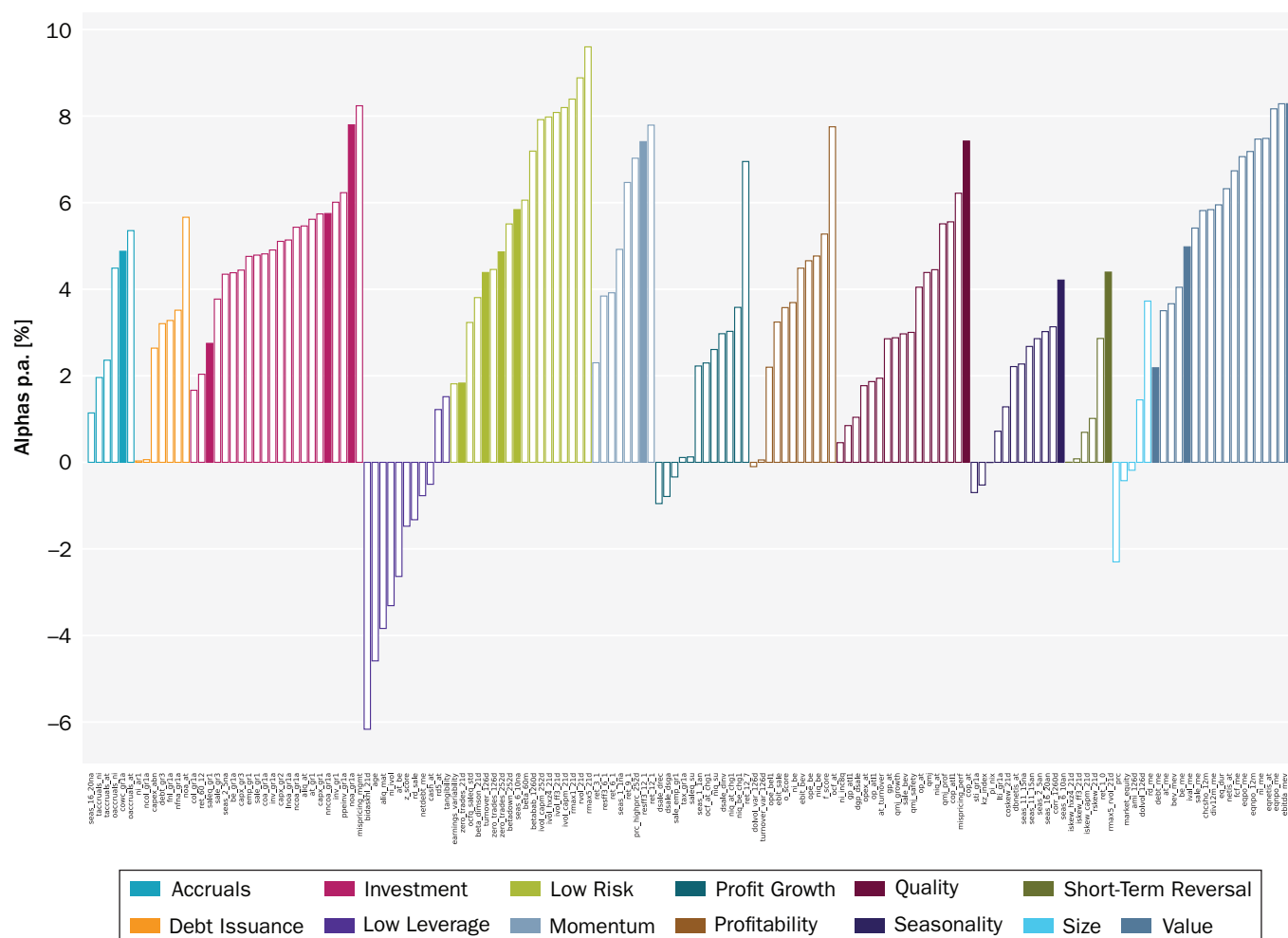
No.	Factor	Cluster	$n(\alpha)$							
			$t > 2$	$t > 3$	FF5	FF6	HXZ	BS	SY	C13
	RMRF	Market	105	86	69	58	63	33	58	6
1	<i>cop_at</i>	Quality	101	78	68	57	62	33	57	6
2	<i>noa_gr1a</i>	Investment	65	34	67	56	61	32	56	6
3	<i>saleq_gr1</i>	Investment	42	10	66	55	60	31	55	5
4	<i>ival_me</i>	Value	39	14	65	54	60	30	55	4
5	<i>resff3_12_1</i>	Momentum	35	15	64	53	59	29	54	4
6	<i>seas_6_10an</i>	Seasonality	27	9	63	52	58	28	53	4
7	<i>debt_me</i>	Value	37	7	62	51	57	27	52	4
8	<i>seas_6_10na</i>	Low Risk	25	3	61	51	56	27	52	4
9	<i>zero_trades_252d</i>	Low Risk	13	1	61	51	56	27	52	4
10	<i>cowc_gr1a</i>	Accruals	14	3	60	50	55	26	51	4
11	<i>nncoa_gr1a</i>	Investment	7	1	59	49	54	25	51	4
12	<i>ocf_me</i>	Value	5	1	59	49	54	25	51	4
13	<i>zero_trades_21d</i>	Low Risk	11	1	59	49	54	25	50	4
14	<i>turnover_126d</i>	Low Risk	9	2	59	49	54	25	50	4
15	<i>rmax5_rvol_21d</i>	Short-Term Rev.	3	0	58	49	53	25	50	4
16	<i>seas_11_15na</i>	Seasonality	2	0	57	48	52	24	49	4
17	<i>o_score</i>	Profitability	4	0	57	48	52	24	48	4
18	<i>niq_at</i>	Quality	0	0	56	47	52	24	48	4
19	<i>seas_16_20an</i>	Seasonality	0	0	55	46	51	23	47	4
20	<i>ni_ar1</i>	Debt Issuance	1	0	55	46	51	23	47	4
21	<i>ivol_ff3_21d</i>	Low Risk	2	0	55	46	51	23	47	4
22	<i>ni_me</i>	Value	0	0	55	46	51	23	47	4
23	<i>dsale_dinv</i>	Profit Growth	0	0	54	45	50	22	46	4
24	<i>ni_be</i>	Profitability	1	0	54	45	49	22	45	4
25	<i>noa_at</i>	Debt Issuance	0	0	53	44	48	21	44	4
26	<i>age</i>	Low Leverage	0	0	53	43	47	21	43	4
27	<i>ret_12_1</i>	Momentum	0	0	52	43	46	21	42	4
28	<i>aliqu_mat</i>	Low Leverage	0	0	52	42	46	21	42	4
29	<i>nfna_gr1a</i>	Debt Issuance	0	0	51	41	45	20	41	4
30	<i>at_me</i>	Value	0	0	50	40	44	20	41	4

NOTES: This exhibit reports the results for an iterative factor model construction where the k-th iteration augments the model by the factor in row k. Columns 4–10 refer to the number of remaining significant factor alphas after controlling for the specified factor model; $t > 2$ and $t > 3$ control the factor zoo based on the iterative model using a significance alpha threshold of $t(\alpha) > 1.96$ and $t(\alpha) > 3.00$, respectively. FF5 refers to the Fama and French (2015) five-factor model while the augmented version include an additional momentum factor (FF6). HXZ refers to the model of Hou, Xue, and Zhang (2015), while BS and SY refer to the models of Barillas et al. (2020) and Stambaugh and Yuan (2017), respectively. C13 is based on the 13 strongest factors in the considered factor style clusters. The red numbers indicate the specifications for which the academic factor models contain the same number of factors as the iterative factor models. The sample period is November 1971 to December 2021.

Against this backdrop, we wonder how important it is to go with these selected factors or whether it is sufficient to determine the strongest factor from each of the 13 categories. The last column in Exhibit 3 reports results for a 13-factor model consisting of the strongest (largest absolute CAPM alpha) factor per cluster. This model virtually spans the whole zoo, leaving just 5 factors unexplained. Out of the first 30 factors, it only fails to explain away the alpha from *sales growth for one quarter* (*saleq_gr1*) and *intrinsic value-to-market* (*ival_me*), highlighting the power of a cluster-spanning model. Nevertheless, the iterative model with 13 additional factors only leaves one alpha significant.

EXHIBIT 4

Selected Alpha Factors



NOTES: This exhibit depicts annualized CAPM-alphas for the US factor zoo. The factors selected by the iterative factor selection process are indicated by full colors; all other factors are whitened. The underlying excess market return is sourced from https://mba.tuck.dartmouth.edu/pages/faculty/ken.french/data_library.html. The factor style clusters are based on Jensen, Kelly, and Pedersen (2023). The sample period is November 1971 to December 2021.

The Relevance of Factors through Time

Having analyzed the full sample evidence, we wonder about the persistence of the individual factors' relevance through time. One caveat when analyzing the relative strengths of individual factors in the factor zoo is that many factors have been published along the way and show a weaker post-publication performance (McLean and Pontiff 2016). To better understand how the set of selected factors evolves through time, we repeat the iterative factor selection but restrict ourselves to the available factors at any point in time. Indeed, many factors that we find to work very well over the whole sample period were not known for many years. For example, the *residual momentum factor* (*resff3_12_1*) of Blitz, Huij, and Martens (2011) was only published in 2011 and would, therefore, not have been viable in the first 25 years of the sample. Specifically, our analysis is based on an expanding window analysis using an initial window of 180 months such that we obtain the first out-of-sample observation in December 1986. Each year, we only consider the already published

factors when annually running the iterative factor selection and collect the information as presented in Exhibit 2. However, note that we stop the iteration at the first occurrence of $n(\alpha)_{t(\alpha)>3} = 0$, that is, when we arrive at the first factor model that renders all remaining factor alphas insignificant at a threshold of $t > 3.00$.

Exhibit 5 highlights the relevant factors through time, colored by their corresponding factor style cluster. That is, whenever a factor is chosen in the corresponding year's factor model, it is highlighted on the timeline. While the vast majority of factors are either never or rarely included, the top factors from the full sample evidence in Exhibit 2 show up prominently, especially over the last 10–15 years.

We observe many factor style clusters to be included in the model for most of the time once a representative factor is published. For instance, the value cluster is present most of the time, but also the momentum cluster is constantly represented since its publication in Jegadeesh and Titman (1993). Other persistent factor categories are accruals, investment, seasonality, and short-term reversal.

However, many factor style clusters see a change in their representative factors. For example, quality factors have been deemed highly relevant throughout the last three decades, but typically only one quality factor was selected at a time. A similar observation applies to the momentum factor cluster, which is represented by four different factors during the sample period. Notably, the introduction of *residual momentum* (*resff3_12_1*) rendered the previously selected classic momentum factor (*ret_12_1*) insignificant (see Blitz, Hanauer, and Vidojevic 2020). Another example is the accruals cluster. Once published, the new factor *change in current operating working capital* (*cowc_gr1a*) replaced the *operating accruals* factor (*oaccruals_at*). These observations emphasize the need for and relevance of continuously adding and innovating factors and their definitions based on new insights or newly available data.

Rolling Window Analysis

Given the long-run relevance of different factor style clusters, we next investigate their relevance at shorter time intervals. Therefore, we run a rolling window analysis based on a window size of 180 months that is updated each year. We follow the same iterative factor selection as before and only report the chosen factors that add maximally to a model in leaving the least amount of alphas unexplained.

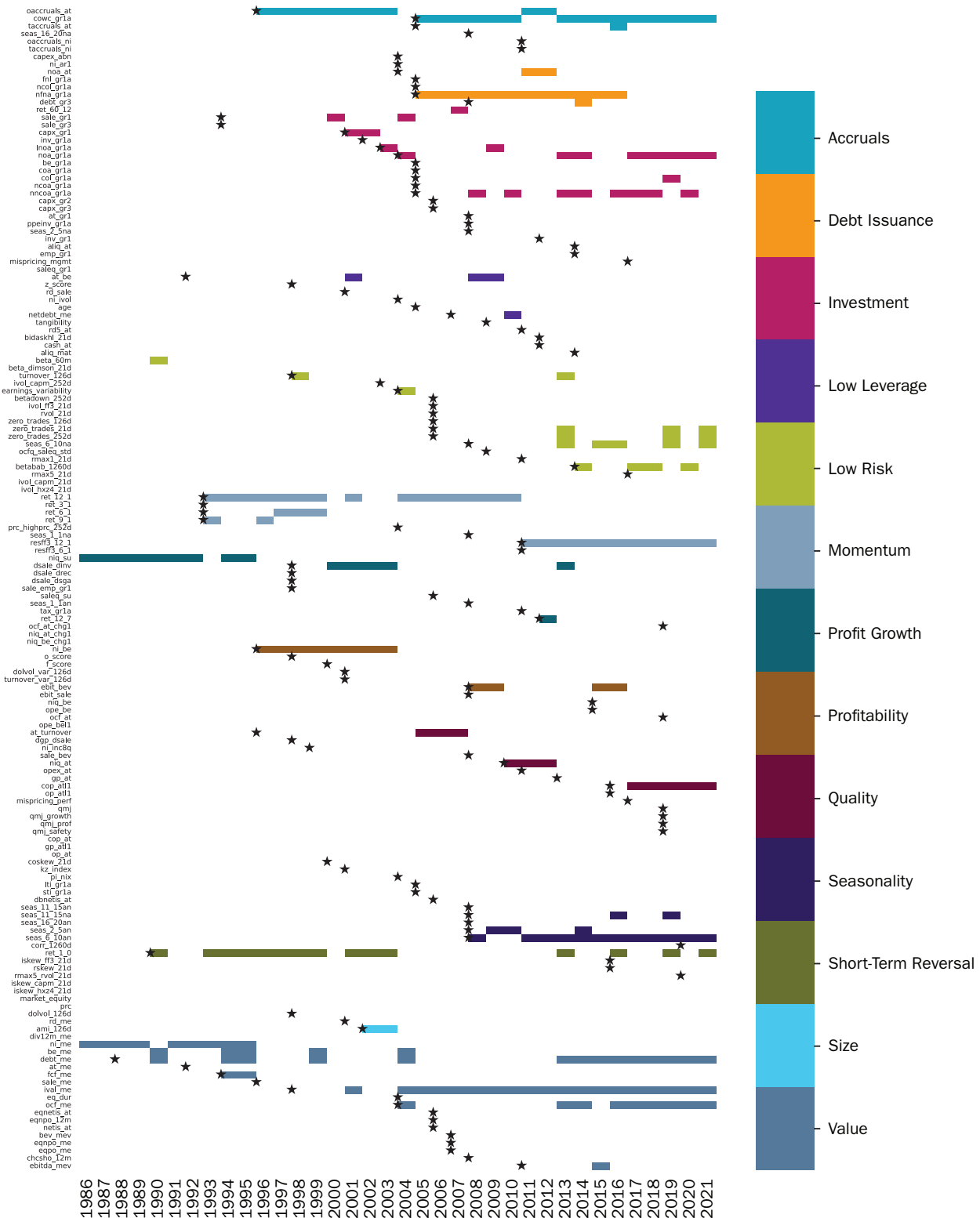
Exhibit 6 depicts the development of the iterative selection process, aggregated at the factor style cluster level. For each year in the rolling window analysis, we collect the selected factors by style factor cluster. The upper panel of Exhibit 6 plots the number of factors included in each year's model using a cutoff of $t > 2$, while the lower panel reports the results for a cutoff of $t > 3$. We report both thresholds to address datamining concerns while simultaneously gauging the relevance of borderline factors. Also, using a window size of only 180 months naturally raises the bar for factors to exceed a given threshold, relative to the full sample evidence.

The upper panel of Exhibit 6 documents a decrease in the number of selected factors over time. While it took some 15 factors to span the factor zoo in the early years of the sample period, this number has decreased to about 8 factors in more recent years. Yet, there are specific factor styles that are persistently relevant through time, including low volatility, seasonality, investment, and quality. While momentum, short-term reversal, and value were included almost every year until the early 2010s, their relevance weakens toward the end of the sample.

Conversely, the lower panel with a threshold of $t > 3$ only reports 4 to 6 factors, on average, to span the remaining factor zoo. The most relevant factor styles in recent years are quality, low volatility, as well as seasonality. Generally, we observe a similar trend in the declining size of the factor models, although the starting models are already quite small compared with models based on the threshold of $t > 2$.

EXHIBIT 5

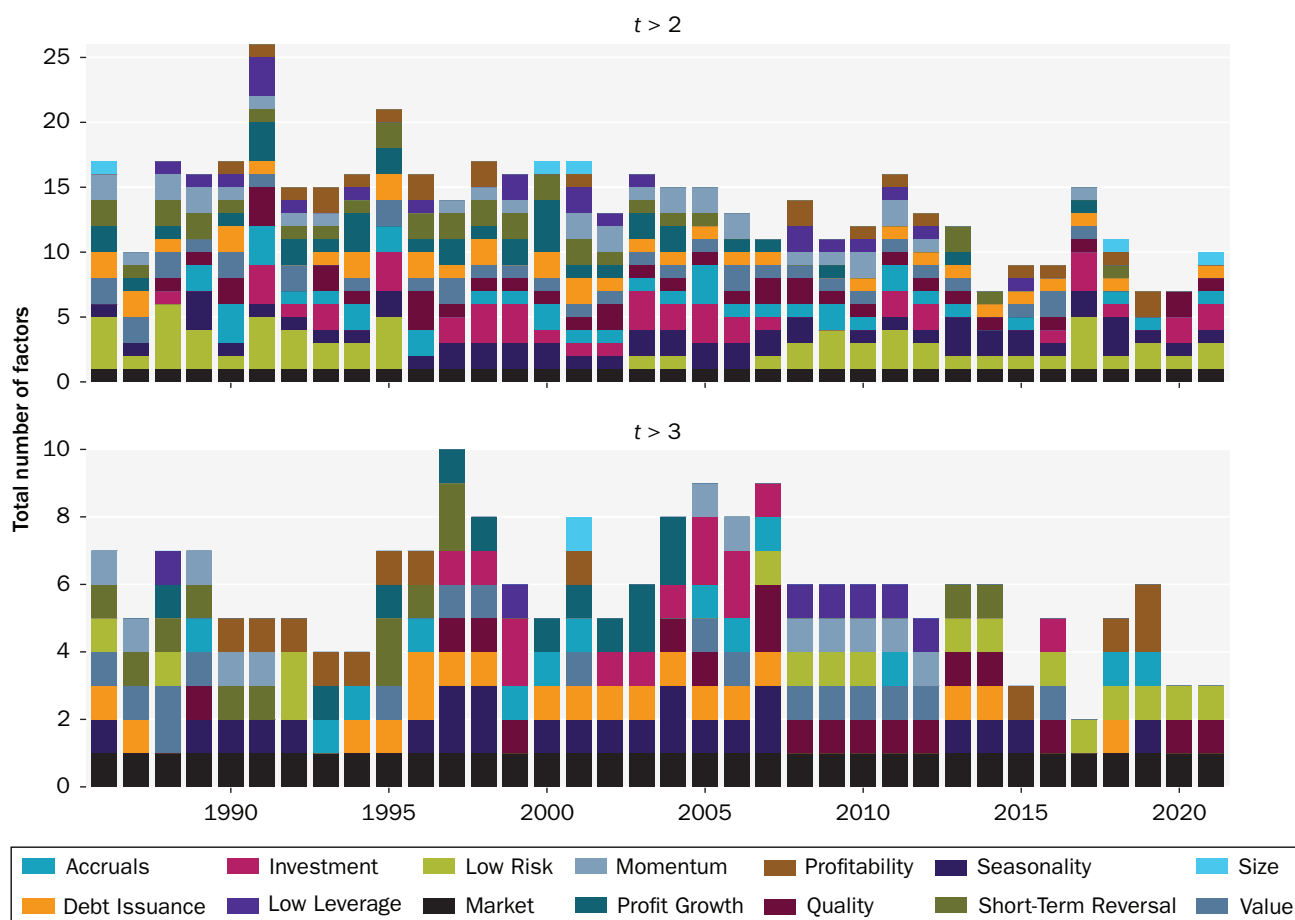
Factor Persistence



NOTES: This exhibit depicts the evolution of the iterative factor selection process over time when using a significance threshold of $t > 3.00$. Factors are added to the factor zoo on an annual basis following their publication. The publication year for each factor is indicated by a star. The initial estimation window size is 180 months and is expanded in subsequent iterations. The sample period is November 1971 to December 2021.

EXHIBIT 6

Rolling Window Factor Selection



NOTES: This exhibit depicts the rolling window outcome of the iterative factor selection using a significance threshold of $t > 1.96$ (upper panel, labelled $t > 2$) and $t > 3.00$ (lower panel, labelled $t > 3$), respectively. The rolling window size is 180 months. The sample period is November 1971 to December 2021.

Overall, the represented factor style clusters are slowly changing over time, and there is typically some factor representative of the low volatility, seasonality, and quality clusters involved. Interestingly, the classic size factor is rarely chosen and does not seem relevant in spanning other factors' alpha.

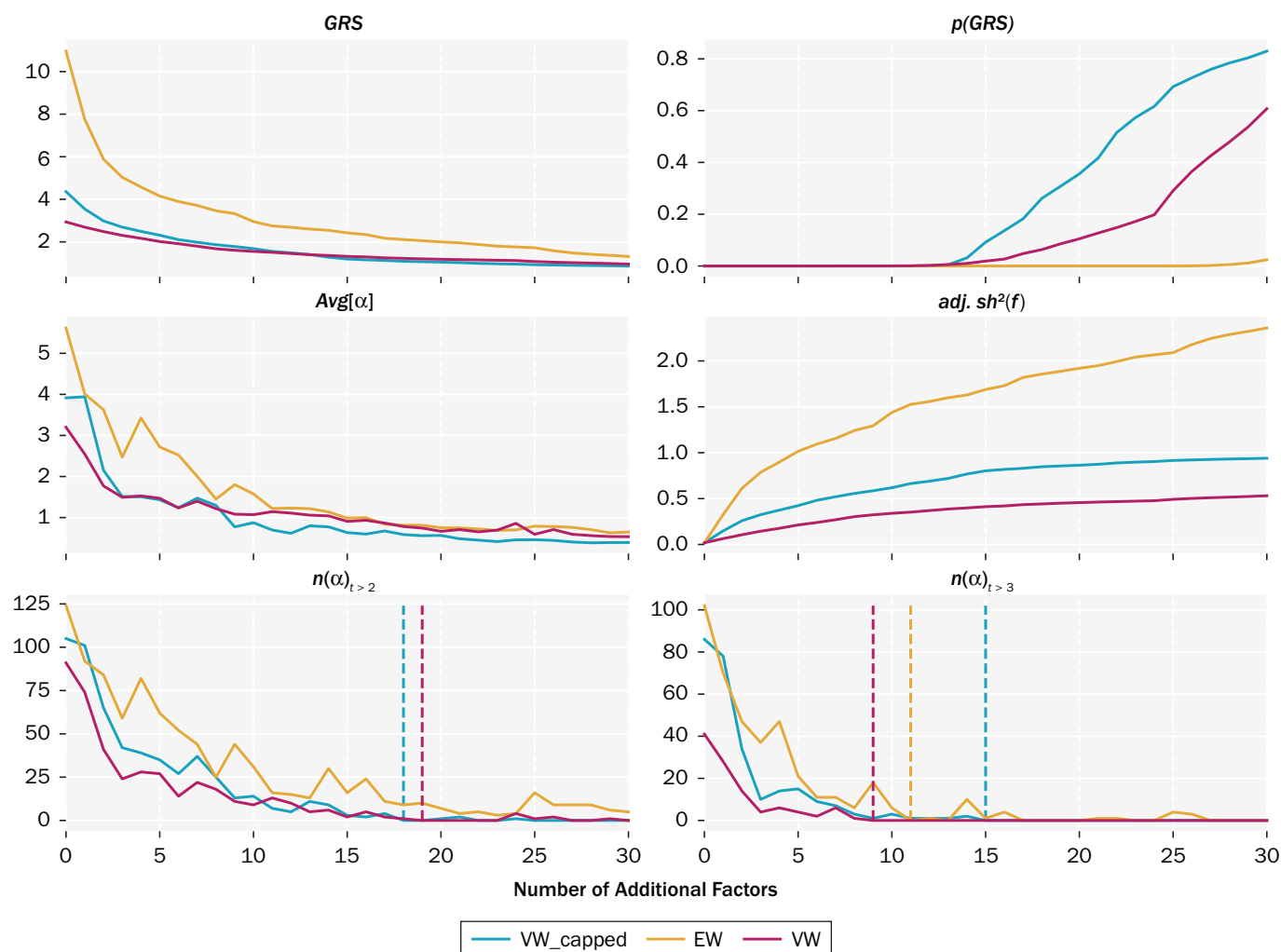
Robustness Regarding Alternative Weighting Schemes

The weighting scheme used to construct factor portfolios can have a big impact, see Bessembinder, Burt, and Hrdlicka (2022) and Soebhag, Van Vliet, and Verwijmeren (2023), among others. We thus check for robustness of our results with respect to the three weighting schemes: capped value weighting (CW), value weighting (VW), and equal weighting (EW). We repeat the analysis of Exhibit 2 for the different weighting schemes and summarize the corresponding outcome in Exhibit 7.

The first row of Exhibit 7 depicts the development of the GRS statistic and its associated p -value when increasing the number of factors. While the starting GRS statistics for capped value-weighted, and value-weighted factor models are low single-digit numbers, the EW factor models come with double-digit numbers.

EXHIBIT 7

Alternative Factor Weighting Schemes



NOTES: This exhibit depicts the key performance statistics for iterative factor models when based on different weighting schemes. We consider equal weighting (EW), value weighting (VW), and capped value weighting (CW). The number of factors refers to the model building process described in the text. Average absolute alphas $\text{Avg}[\alpha]$ are annualized and in percentage. The vertical dashed lines in the last panel mark the minimum amount of factors needed to explain away the available factor zoo alpha in each weighting setting. The sample period is November 1971 to December 2021.

Although the GRS statistic for EW factors quickly declines when increasing the number of factors, p -values suggest significance even at 30 factors, unlike the other two-factor weighting schemes. Comparing the CW and VW factor models, we observe CW to pick up in p -values first and cross the 5% threshold when adding the 15th factor, while VW takes 18 factors.

Whereas the average absolute alphas, $\text{Avg}[\alpha]$, of all three weighting schemes seem to converge toward the same value, the adjusted squared Sharpe ratios, $\text{adj. } Sh^2(f)$, do not. Indeed, CW and VW factor models each converge toward different $\text{adj. } Sh^2(f)$, while the EW factor models do not yet converge when considering 30 factors. These observations, combined with the results of Fama and French (2018), who generalize that the model with the highest $Sh^2(f)$ must be the best model to minimize the remaining alphas, imply that EW factor models are not confined to a small number of factors to span the entire factor zoo. Put differently, EW factors present

a higher and more diverse alpha potential, and thus it takes more factors to span the EW factor zoo.

In the bottom panels of Exhibit 7, we track the number of remaining significant factors for a given factor model. The vertical dashed lines indicate the first occurrence of zero remaining significant factors for a given weighting scheme. While all three weighting schemes reduce the number of remaining significant alphas with increasing size, the EW factor models are sometimes choppy in doing so. That is, increasing the number of factors does not necessarily lead to a decrease in the remaining significant factors in the zoo. Also, it takes 18 and 19 factors for the CW and VW factor models to explain away all factor zoo alphas, while EW factor models take more than 30 factors at a significance threshold of $t > 2$. Notably, this order is almost reversed at a threshold of $t > 3$. Then, VW factor models require 9 additional factors to eliminate significant alphas, followed by 11 and 15 factors for EW and CW, respectively. These results indicate that the selected CW factors are relatively stronger in explaining the CW factor zoo, whereas EW factor models are based on fewer strong factors, and the remainder are more subject to data mining concerns.

INTERNATIONAL EVIDENCE

Global Factor Selection

We next broaden our view and investigate whether the US evidence from Exhibit 2 carries over to global factors, using international data for 93 different countries. Given the limited availability of some stock-specific measures, we shorten our international sample and focus on the period from August 1993 to December 2021. The sample covers 136 common factors for the three regions World, US, and World ex US.⁴

Exhibit 8 documents the iterative factor selection based on global factor data. Notwithstanding the use of global factors and a shorter sample period, we observe a good overlap in the selected factors compared to the US results in Exhibit 2. Out of the 10 selected factors, 3 are identical (*cop_at*, *resff3_12_1*, *cowc_gr1a*), and the 2 selected investment factors are close cousins of the US ones. Also, the selected factor style cluster order is very similar (almost identical) for these top factors.

The selection process reveals 11 global factors to span the global factor zoo when enforcing a threshold of $t > 3$. Even at the lower threshold of $t > 2$, it only takes about two dozen factors to span the factor zoo. From a GRS test perspective, it takes between 1 dozen and 2 dozen factors to reject the null of no significant remaining alphas, and we can clearly observe the monotonic decline in the GRS statistic for the global factors.

However, the GRS statistics for the US and World ex US samples also decrease almost monotonically, indicating that the global factor models also work for these subsamples. While the models in the US sample have generally lower GRS statistics, resulting in a rejection of the null at about six factors, the models for the rest of the world experience higher GRS values. Note that the shorter sample period of this international analysis already induces a reduction in relevant factors by construction; see Exhibit 9 in the next subsection. Although the factor models derived from global data have explanatory power for the World ex US factors, they do a better job on US factors. This observation is likely explained by the much higher alphas that have to be explained as indicated by a GRS statistic of more than 7.03 for international factors compared to a GRS statistic of 2.13 for the US.

⁴ See Jensen, Kelly, and Pedersen (2023) for a detailed overview of the construction of global factors.

EXHIBIT 8

Global Factor Analysis

No.	Factor	Cluster	World				US				World ex US			
			GRS	$p(\text{GRS})$	$t > 2$	$t > 3$	GRS	$p(\text{GRS})$	$t > 2$	$t > 3$	GRS	$p(\text{GRS})$	$t > 2$	$t > 3$
	Market	Market	5.34	0.00	86	50	2.13	0.00	81	44	7.03	0.00	91	56
1	<i>cop_at</i>	Quality	4.11	0.00	57	26	1.77	0.00	71	36	5.43	0.00	73	39
2	<i>ncoa_gr1a</i>	Investment	3.75	0.00	44	14	1.58	0.00	24	7	5.06	0.00	44	15
3	<i>col_gr1a</i>	Investment	3.40	0.00	45	19	1.49	0.00	54	26	4.99	0.00	46	22
4	<i>eq_dur</i>	Value	3.08	0.00	37	16	1.43	0.01	53	19	4.32	0.00	48	17
5	<i>cowc_gr1a</i>	Accruals	2.75	0.00	26	4	1.31	0.04	35	8	4.20	0.00	52	18
6	<i>resff3_12_1</i>	Momentum	2.65	0.00	44	11	1.24	0.08	17	0	3.91	0.00	63	30
7	<i>cash_at</i>	Low Leverage	2.35	0.00	21	3	1.21	0.12	26	8	3.50	0.00	18	5
8	<i>age</i>	Low Leverage	2.19	0.00	18	4	1.21	0.11	27	10	3.46	0.00	32	9
9	<i>dolvol_126d</i>	Size	2.03	0.00	27	4	1.20	0.13	32	12	3.36	0.00	50	19
10	<i>oaccruals_at</i>	Accruals	1.92	0.00	25	4	1.20	0.12	37	19	3.29	0.00	43	18
11	<i>at_be</i>	Low Leverage	1.82	0.00	11	0	1.15	0.18	31	14	3.31	0.00	45	19
12	<i>turnover_var_126d</i>	Profitability	1.75	0.00	8	0	1.16	0.18	32	15	3.05	0.00	42	13
13	<i>nncoa_gr1a</i>	Investment	1.72	0.00	6	0	1.15	0.18	33	14	3.07	0.00	42	13
14	<i>dsale_dinv</i>	Profit Growth	1.67	0.00	6	0	1.16	0.17	34	14	3.09	0.00	43	14
15	<i>iskew_ff3_21d</i>	Short-Term Rev.	1.63	0.00	5	0	1.16	0.17	33	14	3.00	0.00	42	13
16	<i>ret_60_12</i>	Investment	1.59	0.00	3	0	1.13	0.22	33	9	3.02	0.00	42	13
17	<i>rd_sale</i>	Low Leverage	1.54	0.00	3	0	1.13	0.22	36	10	2.61	0.00	20	6
18	<i>mispricing_perf</i>	Quality	1.49	0.01	4	0	1.06	0.35	20	2	2.63	0.00	19	6
19	<i>o_score</i>	Profitability	1.43	0.01	3	0	1.06	0.35	23	2	2.64	0.00	20	5
20	<i>rd5_at</i>	Low Leverage	1.39	0.02	1	0	1.06	0.35	23	2	2.64	0.00	22	5
21	<i>zero_trades_21d</i>	Low Risk	1.34	0.04	9	3	1.06	0.36	17	5	2.57	0.00	19	4
22	<i>zero_trades_126d</i>	Low Risk	1.23	0.10	7	0	0.96	0.58	10	0	2.48	0.00	17	6
23	<i>tangibility</i>	Low Leverage	1.18	0.15	1	0	0.97	0.57	11	0	2.47	0.00	16	6
24	<i>div12m_me</i>	Value	1.13	0.23	1	0	0.96	0.60	5	0	2.44	0.00	17	5
25	<i>be_me</i>	Value	1.07	0.33	1	0	0.96	0.58	6	0	2.22	0.00	8	2
26	<i>prc</i>	Size	1.01	0.47	1	0	0.92	0.68	4	0	2.24	0.00	8	2
27	<i>cop_atl1</i>	Quality	0.98	0.54	0	0	0.91	0.70	6	0	2.21	0.00	6	1
28	<i>coskew_21d</i>	Seasonality	0.96	0.58	0	0	0.90	0.74	6	0	2.23	0.00	6	1
29	<i>qmj_prof</i>	Quality	0.95	0.61	0	0	0.90	0.73	5	0	2.21	0.00	8	1
30	<i>ebit_bev</i>	Profitability	0.93	0.67	0	0	0.88	0.76	4	1	2.23	0.00	8	1

NOTES: This exhibit reports the results for an iterative factor selection where the k -th iteration augments the model by the factor in row k . The factor selection is based on global factors, and the corresponding factor order is then investigated in the two other regions, US and World ex US, using the respective local factors. The exhibit shows the GRS statistic of Gibbons, Ross, and Shanken (1989) and its p -value, $p(\text{GRS})$, as well as the number of remaining significant factor alphas after controlling for the specified factor model. $t > 2$ and $t > 3$ control the factor zoo when based on an iterative model at a significance alpha threshold of $t(\alpha) > 1.96$ and $t(\alpha) > 3$, respectively. The sample period is August 1993 to December 2021 and considers 136 common factors for all three regions.

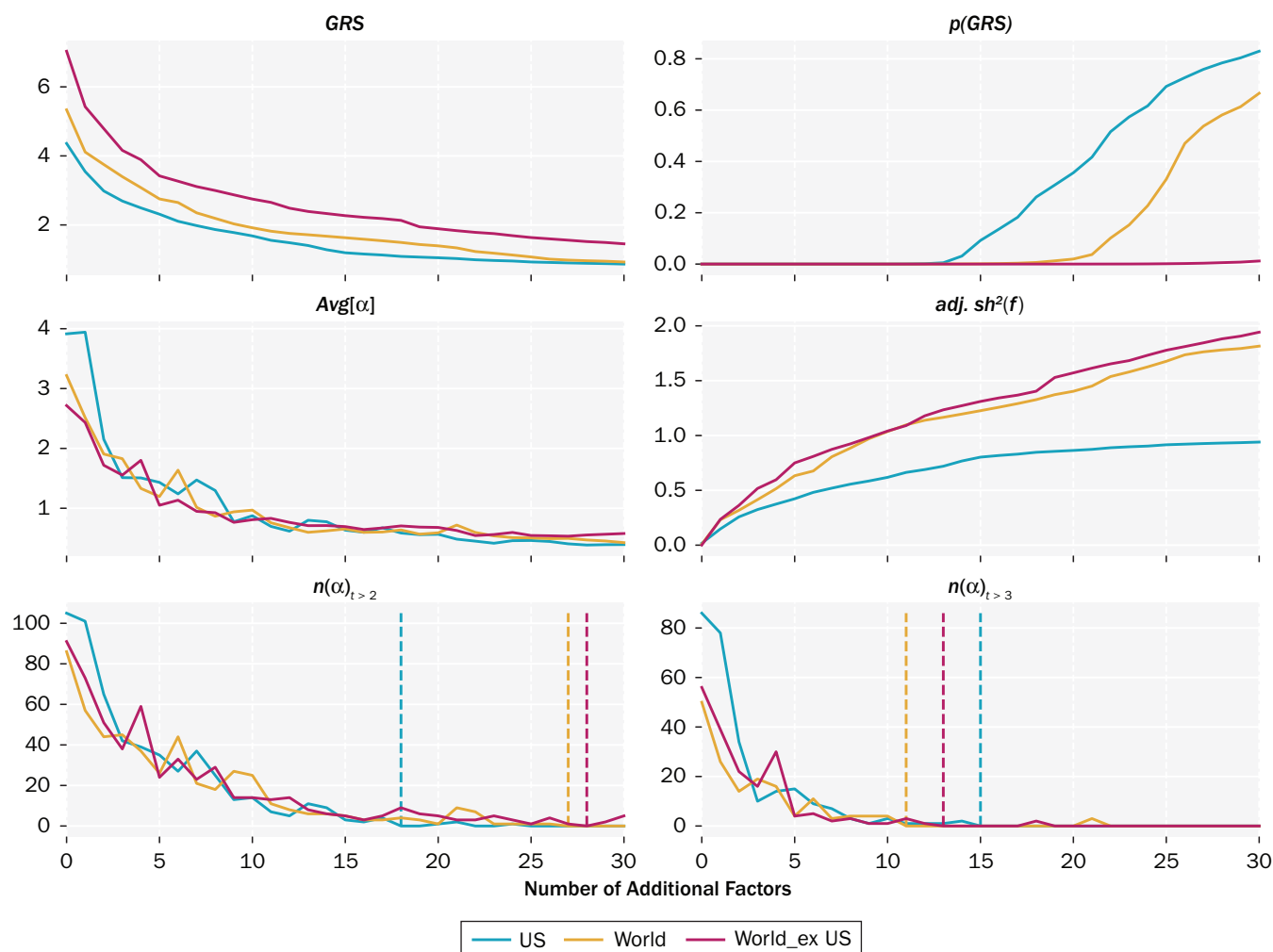
Regional Comparisons

So far, global factors have proven relevant for explaining US factor returns, but they lack explanatory power for the World ex US factors. We wonder whether local factor models are indeed stronger than global ones. We thus determine the iterative factor model within each region separately and juxtapose the relevant statistics in Exhibit 9 by regions.

The GRS statistics are declining for all three regions by design. However, we clearly see that the World ex-US factors' decline in GRS statistic occurs at a higher level than that of the US and the global factors. As a result, it takes more than 30 factors to reject the null at a significance level of 5% for the World ex US sample, while US and global models only take 15 and 22 factors, respectively. The differences are primarily driven by the slower convergence of the adjusted squared Sharpe ratios

EXHIBIT 9

Different Regions



NOTES: This exhibit depicts the key performance statistics for iterative factor models in different regions. The number and choice of factors refers to the model building process described in the text. Average absolute alphas $Avg[\alpha]$ are annualized and in percentage points. The vertical dashed lines in the lower panels mark the minimum amount of factors needed to explain the available factor zoo alpha. The sample period is August 1993 to December 2021 and considers 136 common factors for all three regions.

($adj. Sh^2(f)$) in the different regions as shown in the middle panels of Exhibit 9. While the US factor models seem to converge to some limit in $adj. Sh^2(f)$, the other two models' respective lines still have a positive slope at 30 factors. Conversely, the average absolute alphas seem to converge to a statistically insignificant number for all three regions once 30 factors are considered.

Comparing the required model size for spanning all factor zoo alphas, we note that the US models come with a fairly stable model size in between 6 ($t > 3$) and 12 ($t > 2$) factors, highlighting the genuine relevance of the selected factors. Yet, the other regions' factors vary in relevance. Out of the 27 (28) factors identified in the global (World ex US) zoo at a threshold of $t > 2$, only 11 (13) are deemed relevant for the higher one at $t > 3$. Thus, these locally selected factor models are even slightly stronger than a universal one, which already helps to span the different regions and especially US factors reasonably well.

CONCLUSION

The factor zoo has grown significantly over the last decades as outlined in Cochrane (2011) and Harvey, Liu, and Zhu (2016), highlighting the need to separate sheep from goat factors. To this end, our study investigates the alpha contribution of individual factors in the factor zoo. Specifically, we propose an iterative factor selection strategy to compress the factor zoo, thus substantially reducing the number of factors without losing vital information for the tangency portfolio. The resulting factor sets capture the available alpha in the factor zoo with the minimum number of factors, which is relevant for both practitioners and academic purposes.

Using a comprehensive set of 153 US equity factors, we find that a set of 10 to 20 factors spans the entire factor zoo, depending on the selected statistical significance level. This implies that most candidate factors are redundant but also that academic factor models, which typically contain just three to six factors, are too narrowly defined. When repeating the factor selection to factors as they become available over an expanding window, we find that newly published factors sometimes supersede older factor definitions, emphasizing the relevance of continuous factor innovation based on new insights or newly available data. The identified factor style clusters are quite persistent, however, emphasizing the relevance of diversification across factor styles.

Furthermore, we document that using equal-weighted factors as opposed to capped value-weighted factors requires more than 30 factors to span the factor zoo, indicating that equal-weighted factors exhibit stronger and more diverse alphas. Finally, applying our factor selection strategy to a set of global factors results in a similar set of selected factors. Although the factor models selected based on global data shrink the alpha for US and World ex US sub-universes, they perform better for the US market, implying that international factors exhibit larger and more diverse alpha.

Overall, the proposed method effectively captures the available alpha in the factor zoo across different regions and subperiods, helping investors focus on the most relevant factors and providing academics with inspiration for enhancing asset pricing models.

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