

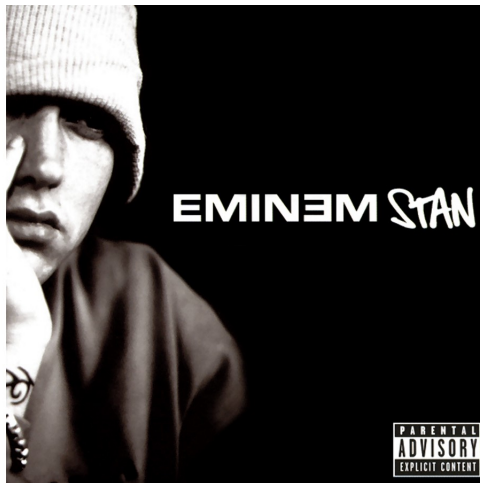
Hierarchical Bayesian analysis using Stan - From a binary logit to advanced models of bounded rationality

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Stan



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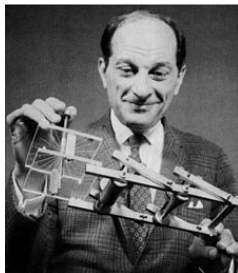


Probabilistic programming
language

Stan



Probabilistic programming
language



Stanislaw Ulam
(1909–1984)
Inventor of the Monte Carlo
method

Stan language

Stan program (with R, Python, Matlab, Julia, Stata, CmdStan interface)

- ▶ declares data and (constrained) parameter variables
- ▶ defines log posterior (or penalized likelihood)

Stan inference

- ▶ MCMC for full Bayes
- ▶ VB for approximate Bayes
- ▶ Optimization for (penalized) MLE

Bayesian workflow

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- Exploratory data analysis

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- ▶ *Prior* predictive checking

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- ▶ Exploratory data analysis
- ▶ *Prior* predictive checking
- ▶ Model fitting and algorithm diagnostics
- ▶ Posterior predictive checking
- ▶ Model comparison (e.g., via cross-validation)

Hierarchical binary logit example: The data

A choice model of buying decisions given price and promotions:
500 consumers with 10 purchase occasions each

	User_ID	Observation	Y	Intercept	Price	Promotion
1	1	1	1	1	0.4242656054	0.429378508
2	1	2	0	1	0.8425126909	0.274047020
3	1	3	1	1	0.3764421751	0.512474872
4	1	4	1	1	0.9115300649	0.355274114
5	1	5	1	1	0.2585383623	0.127846915
6	1	6	1	1	0.2032627692	0.294366778
7	1	7	1	1	0.2246137869	0.161435480
8	1	8	1	1	0.9146362843	0.506660538
9	1	9	1	1	0.8902309975	0.377715006
10	1	10	1	1	0.2795407828	0.185361522
11	2	1	0	1	0.5262798222	0.636969226
12	2	2	0	1	0.9985261350	0.641892214
13	2	3	0	1	0.9664521548	0.662417121
14	2	4	0	1	0.4240157611	0.409965415
15	2	5	0	1	0.2575837581	0.383950177
16	2	6	1	1	0.9943112358	0.431254078
17	2	7	0	1	0.1853664671	0.063719698
18	2	8	0	1	0.7580353320	0.305855863
19	2	9	0	1	0.0594366395	0.969301577

Hierarchical binary logit example: The model

A choice model of buying decisions given price and promotions:

- 500 consumers with 10 purchase occasions each

$$y_{ijt} \sim \mathcal{B}(p_{ijt})$$

$$\text{logit}(p_{ijt}) \sim \mathcal{N}(x_{ijt}\beta_{ij})$$

$$\beta_i \sim \text{MultiNormal}(z_i\delta, \Sigma)$$

$$\delta \sim \mathcal{N}(\delta_0, \sigma)$$

Non-conjugate prior

$$\Sigma = \text{diag}(\tau) \Omega \text{diag}(\tau)$$

$$\tau \sim \text{gamma}(a, b)$$

$$\Omega \sim \text{LKJcorr}(\nu)$$

```
data {
  int<lower=1> nvar; // number of parameters in the logit regression
  int<lower=0> N; // number of observations
  int<lower=1> nind; // number of individuals
  int<lower=0,upper=1> y[N];
  int<lower=1,upper=nind> ind[N]; // indicator for individuals
  row_vector[nvar] x[N];
}
```

```
parameters {
  vector[nvar] delta;
  vector<lower=0>[nvar] tau;
  vector[nvar] beta[nind];
  corr_matrix[nvar] Omega; // Vbeta - prior correlation
}
```

```
model {
  to_vector(delta) ~ normal(0, 5);
  to_vector(tau) ~ gamma(2, 0.5);
  Omega ~ lkj_corr(2);
```

```
  for (h in 1:nind)
    beta[h]~multi_normal(delta, quad_form_diag(Omega, tau));
```

```
  for (n in 1:N)
    y[n] ~ bernoulli_logit(x[n] * beta[ind[n]]);
}
```

Noncentered (Re)Parameterization - the "Matt Trick"

Consider a model with a diagonal variance-covariance matrix

- ▶ Assume our intercept model: $\beta_i \sim \mathcal{N}(\delta, \tau)$
- ▶ We can decompose that into: $\mathcal{N}(\delta, \sigma) \stackrel{d}{=} \delta + \tau \mathcal{N}(0, 1)$
- ▶ The trick applies to other distributions in the location-scale family
- ▶ The transformation:
 1. declare α_i in the parameters block and β_i in the transformed parameters block
 2. draw $\alpha_i \sim \mathcal{N}(0, 1)$ & $\tau \sim \text{gamma}(a, b)$
 3. compute $\beta_i = \delta + \tau \alpha_i$

Noncentered (Re)Parameterization - the multivariate case

- ▶ Assume our intercept model: $\beta_{\mathbf{i}} \sim \text{MultiNormal}(\delta, \Sigma)$
 - ▶ If Σ_{kk} is small, then β_{ik} needs to fall into a small range, NUTS needs a small step size
 - ▶ If Σ_{kk} is large, then β_{ik} can fall into a wide range, NUTS needs a large step size/ lots of small steps
- ▶ The transformation:
 1. declare α_i in the parameters block and β_i in the transformed parameters block
 2. draw $\alpha_i \sim \mathcal{N}(0, 1)$ & $\tau \sim \text{gamma}(a, b)$
 3. compute $\beta_{\mathbf{i}} = \delta + \tau \mathbf{L} \alpha_{\mathbf{i}} \sim \mathcal{N}(\delta, \tau^2 \mathbf{L} \mathbf{L}^T)$
 4. where $\tau \mathbf{L}$ is the Cholesky factor of $\Sigma = \tau^2 \mathbf{L} \mathbf{L}^T$, and τ is the standard deviation of the errors.

Noncentered (Re)Parameterization - the implementation

$$y_{ijt} \sim \mathcal{B}(p_{ijt})$$

$$\text{logit}(p_{ijt}) \sim \mathcal{N}(x_{ijt}\beta_{ij})$$

$$\beta_{\mathbf{i}} = \delta + \tau \mathbf{L} \alpha_{\mathbf{i}}$$

$$\alpha_i \sim \mathcal{N}(0, 1)$$

$$\delta \sim \mathcal{N}(\delta_0, \sigma)$$

Non-conjugate prior

$$\tau \sim \text{gamma}(a, b)$$

$$\Omega \sim \text{LKJcorr}(\nu)$$

```
data {
  .....
}

parameters {
  matrix[nvar, nind] alpha; // nvar*H parameter matrix
  row_vector[nvar] delta;
  vector<lower=0>[nvar] tau;
  cholesky_factor_corr[nvar] L_omega;
}

transformed parameters{
  row_vector[nvar] beta[nind];
  matrix[nind,nvar] Vbeta_reparametrized;
  Vbeta_reparametrized = (diag_pre_multiply(tau, L_omega)*alpha)';

  for (h in 1:nind)
    beta[h]=delta+Vbeta_reparametrized[h];
}

model {
  .....
}
```

Choice of prior distribution of the variance components

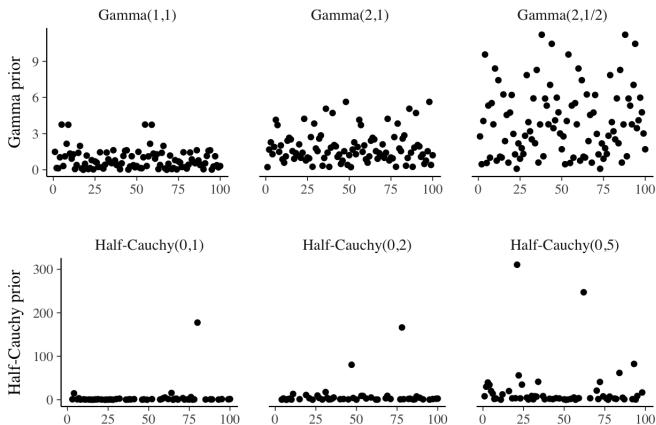


Figure 1: Choice of prior for the variance τ

Choice of the LKJ prior distribution

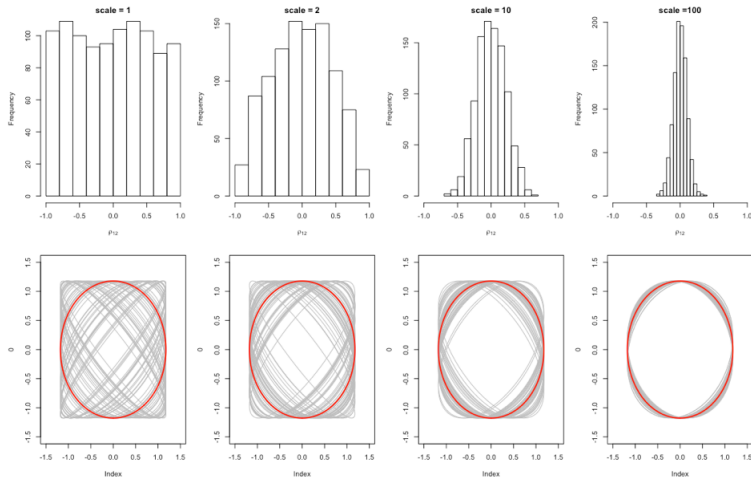


Figure 2: Choice of LKJ prior

Traceplots of model parameters

The noncentered reparametrization helps tremendously

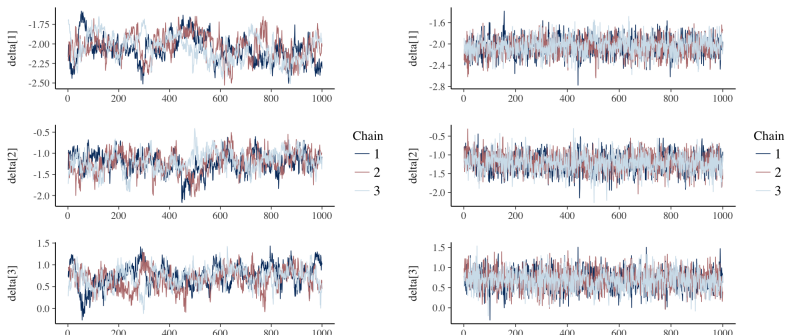


Figure 3: Traceplot of δ parameters (after burnin) under the centered (*left*) vs. the noncentered reparametrization (*right*), using package *bayesplot*

Summary statistics

Effective sample size and convergence properties

```
$summary
```

	mean	se_mean	sd	2.5%	97.5%	n_eff	Rhat
delta[1]	-2.0464282	0.003151008	0.1725878	-2.3917852	-1.7207339	3000	1.001876
delta[2]	-1.2073192	0.004632004	0.2537053	-1.6928554	-0.7322237	3000	1.002893
delta[3]	0.6915687	0.004494100	0.2461520	0.2134696	1.1791871	3000	1.001500

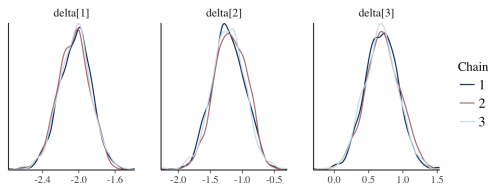


Figure 4: Density plot of δ parameters (after burnin) under the noncentered reparametrization, using package *bayesplot*

Individual level parameters

Most parameters are within the 95% highest density intervals

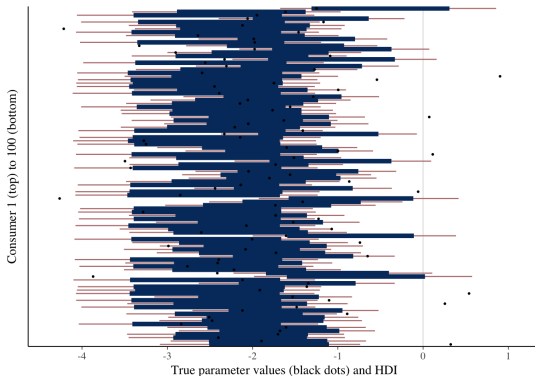


Figure 5: True values (black dots) and the 80% and 95% highest density intervals for the intercept, for the first 100 consumers

Model checking

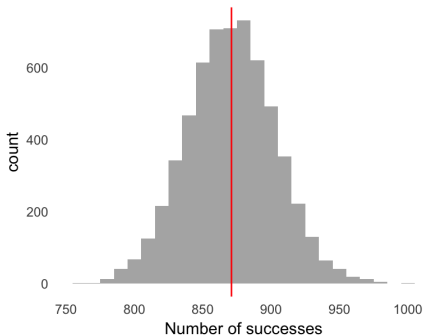


Figure 6: Number of successes: posterior replications vs. true value

Model checking

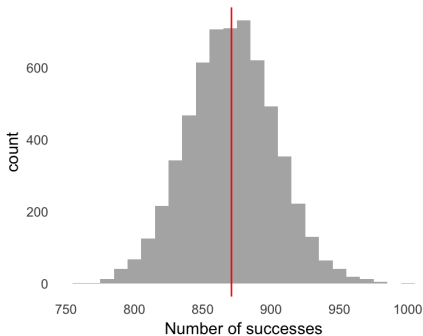


Figure 6: Number of successes: posterior replications vs. true value

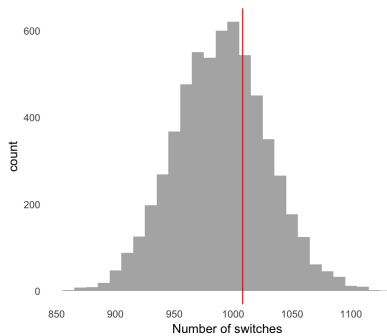


Figure 7: Switches between buying/ not buying: posterior replications vs. true value

Compute hit rates and MSEs based on posterior replications

Model comparison

Likelihood-based measures: Leave-one-out cross-validation

Table 1: Model comparison based on LOO-CV, using package *loo*

	<i>Variance model</i>		<i>Full covariance model</i>		<i>NCP model</i>	
	Estimate	SE	Estimate	SE	Estimate	SE
elpd_loo	-1874.3	43.1	-1871.9	43.2	-1871.2	43
p_loo	398.6	12.5	364.4	11.8	363.7	11.8
looic	3748.6	86.2	3743.7	86.4	3742.4	86.5

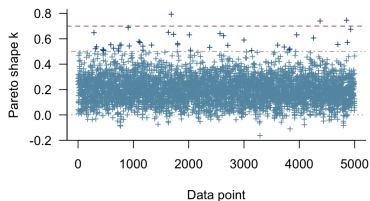


Figure 8: PSIS diagnostic plot, using package *loo*



github.com/alinafere/Dutch-Stan-Meetup

Stan resources

Stan ecosystem

- ▶ lang, math library (C++)
- ▶ interfaces and tools (R, Python, many more)
- ▶ documentation (example model repo, user guide & reference manual, case studies, R package vignettes)
- ▶ online community (Stan Forums on Discourse)

Libraries implementing Stan

- ▶ rstanarm: complex hierarchical models
- ▶ hBayesDM: behavioral decision making models (stan codes on GitHub)
- ▶ bayesplot: data visualization

More Stan resources

StanCon 2018 talks: [▶ Link](#)

Books:

- ▶ Bayesian Data Analysis: aka the Bible:)

[▶ Link](#)

- ▶ Bayesian Statistics using Stan

[▶ Link](#)

- ▶ Statistical Rethinking

[▶ Link](#)