

Ch. 7 Properties of Expectations

Recall the definitions of expected value of a random variable X ...

- X is discrete:

$$E[X] = \sum_x x p(x)$$

where $p(x)$ is the probability mass function of X

also

$$E[g(X)] = \sum_x g(x) p(x)$$

expected value of a function of X .

(see proposition 4.1, p. 122 textbook)

- X is continuous:

$$E[X] = \int_{-\infty}^{+\infty} x f(x) dx$$

where $f(x)$ is the probability density function of X

also

$$E[g(X)] = \int_{-\infty}^{+\infty} g(x) f(x) dx$$

expected value of a function of X

(see Proposition 2.1, p. 181 textbook)

What about jointly distributed random variables X, Y ?

7.2 Expectation of Sums of Random Variables

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Here we have the following results generalizing the previous ones ~~the~~ just listed.

Proposition 2.1

discrete case

$$\left\{ \begin{array}{l} \text{If } X \text{ and } Y \text{ have a joint pmf } p(x, y) \text{ then} \\ E[g(X, Y)] = \sum_y \sum_x g(x, y) p(x, y) \end{array} \right.$$

continuous case

$$\left\{ \begin{array}{l} \text{If } X \text{ and } Y \text{ have a joint pdf } f(x, y) \text{ then} \\ E[g(X, Y)] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} g(x, y) f(x, y) dx dy \end{array} \right.$$

(Proof, see textbook, p. 281)

EXAMPLE (important special case)

Suppose X and Y are jointly continuous with ^{joint} pdf $f(x, y)$.

What is $E[X + Y]$? (i.e. $g(X, Y) = X + Y$)

$$\begin{aligned} E[X + Y] &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x + y) f(x, y) dx dy \\ &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} x f(x, y) dx dy + \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} y f(x, y) dx dy \\ &= \int_{-\infty}^{+\infty} x \left[\int_{-\infty}^{+\infty} f(x, y) dy \right] dx + \int_{-\infty}^{+\infty} y \left[\int_{-\infty}^{+\infty} f(x, y) dx \right] dy \\ &\quad = f_X(x) \text{ marginal pdf of } X \quad = f_Y(y) \text{ marginal pdf of } Y \end{aligned}$$

So

$$E[X+Y] = \int_{-\infty}^{+\infty} x f_X(x) dx + \int_{-\infty}^{+\infty} y f_Y(y) dy$$

$$= E[X] + E[Y]$$

(assuming $E[X]$ and $E[Y]$ are finite)

(see p. 282 textbook)

see also
notes
p. 145
Corollary
9.2

More generally... for random variables X_1, X_2, \dots, X_n

$$E[X_1 + X_2 + \dots + X_n] = E[X_1] + E[X_2] + \dots + E[X_n]$$

extend
previous
argument
using
induction

This result (*) holds for ~~discrete~~ jointly continuous X_1, \dots, X_n as well as for jointly discrete X_1, \dots, X_n .

EXAMPLE (2c, p. 283 textbook)

Suppose that X_1, X_2, \dots, X_n are independent and identically distributed random variables each having ~~an~~ cumulative distribution function, F , and expected value $E[X_i] = \mu$ for $i=1, 2, \dots, n$.

Define the sample mean as

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

(i.e. \bar{X} is ~~the~~ a function of random variables X_1, X_2, \dots, X_n)

Compute $E[\bar{X}]$.

$$E[\bar{X}] = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right)$$

$$= \frac{1}{n} E\left(\sum_{i=1}^n X_i\right)$$

by our result on previous page

$$= \frac{1}{n} \sum_{i=1}^n E[X_i]$$

$$= \frac{1}{n} \sum_{i=1}^n \mu$$

$$= \frac{1}{n} (n \cdot \mu) = \mu$$

The expected value of the sample mean is the mean of the distribution.

EXAMPLE (Expectation of Binomial Random Variable)

Let X be a binomial R.V. with parameters n and p

recall, this means

$$p(i) = \binom{n}{i} p^i (1-p)^{n-i} \quad \text{for } i=0,1,\dots,n$$

$$= P\{X=i\}$$

i = # of successes
 p = probability of a success

n = total # of trials

Let

$$X = X_1 + X_2 + \dots + X_n$$

where

$$X_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ trial is a success} \\ 0 & \text{if } i^{\text{th}} \text{ trial is a failure} \end{cases}$$

think of this as a clever way to count up # of successes.

Bernoulli R.V.

~~The following is a very messy and incorrect derivation of the expectation of a binomial random variable. It is included here for educational purposes only.~~

Note that ~~each~~ ^{the} X_i 's are independent
 (trial ~~to~~ ^{has} success or failure independent
 of ~~previous~~ other trials outcomes)

Also $E[X_i] = 1 \cdot p + 0 \cdot (1-p) = p$

Therefore,

$$E[X] = E[X_1 + \dots + X_n]$$

$$= E[X_1] + \dots + E[X_n]$$

$$= p + \dots + p$$

$$= np$$

} independence
of X_i

← we had found
this result
previously

(notes p. (120)
(123) 10.18.18)

→ EXAMPLE (Expectation of a Hypergeometric Random Variable)

same
example
in notes
p. 146-147
10.23.18

Let X be a hypergeometric R.V. with parameters N, n, m . Recall

$$P\{X=i\} = p(i) = \frac{\binom{m}{i} \binom{N-m}{n-i}}{\binom{N}{n}} \quad \text{for } i=0,1,\dots,n$$

→ see textbook p.151 (Section 4.8.3)

We found $E[X] = \frac{mn}{N}$

select n balls from an urn with
 m white balls
 $N-m$ black balls
 $X = \#$ of white balls selected
without replacement!

Let $X = X_1 + X_2 + \dots + X_m$

where

$$X_i = \begin{cases} 1 & \text{if } i^{\text{th}} \text{ white ball is selected} \\ 0 & \text{otherwise} \end{cases}$$

Note $P\{X_i=1\} = \frac{n}{N} = p$

n chances out of N for white ball i to be selected = proportion of all balls selected.

then

$$P\{X_i=0\} = 1-p = \frac{\binom{1}{1} \binom{N-1}{n-1}}{\binom{N}{n}} = \frac{n}{N}$$

As in the previous example, X_i are each Bernoulli R.V.

So $E[X_i] = 1 \cdot p + 0 \cdot (1-p) = p$.

again a clever way to add up how many white balls are selected.

Then

$$\begin{aligned}
 E[X] &= E[X_1 + X_2 + \dots + X_m] \\
 &= E[X_1] + E[X_2] + \dots + E[X_m] \\
 &= p + p + \dots + p \\
 &= mp \\
 &= m \left(\frac{n}{N} \right)
 \end{aligned}$$

← See also notes

p. (147)

(140)

(107, 111)

Recall, we had found that the expected value for X was the same for drawing balls with and w/o replacement $\left[\begin{array}{l} m \text{ white balls} \\ N-m \text{ black balls} \end{array} \right]$ select n balls

with replacement

$$P\{X=i\} = \binom{n}{i} p^i (1-p)^{n-i} \quad i=0, \dots, n$$

$$p = \frac{m}{N}$$

$$E[X] = np = \frac{nm}{N}$$

without replacement

$$P\{X=i\} = \frac{\binom{m}{i} \binom{N-m}{n-i}}{\binom{N}{n}} \quad i=0, \dots, n$$

$$E[X] = \frac{nm}{N}$$

same expected value

7.4 Covariance, Variance of Sums, and Correlations

Recall: For a single R.V.

Discrete: $E[X] = \sum_x x p(x)$

$$\text{Var}(X) = E[(X - \mu)^2] = E[X^2] - (E[X])^2$$

$$\mu = E[X]$$

Continuous

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \mu$$

$$\text{Var}(X) = E[(X - \mu)^2] = E[X^2] - (E[X])^2$$

Def: Covariance between X and Y

The covariance between X and Y , denoted by $\text{Cov}(X, Y)$, is

$$\text{Cov}(X, Y) = E[(X - E[X])(Y - E[Y])]$$

Note:
~~Covariance~~

$$\begin{aligned} \text{Cov}(X, Y) &= E[XY - E[X]Y - E[Y]X + E[X]E[Y]] \\ &= E[XY] - E[X]E[Y] - E[Y]E[X] + E[X]E[Y] \end{aligned}$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y]$$

- If X and Y are independent (see for continuous R.V. case)

$$E[XY] = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x,y) dx dy$$

but $f(x,y) = f_X(x) f_Y(y)$ if X and Y are indep.

$$\begin{aligned} \text{so } E[XY] &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_X(x) f_Y(y) dx dy \\ &= \int_{-\infty}^{+\infty} x f_X(x) dx \cdot \int_{-\infty}^{+\infty} y f_Y(y) dy \\ &= E[X] \cdot E[Y]. \end{aligned}$$

~~Therefore~~

- So, if X and Y are independent R.V.

$$\text{Cov} \cancel{\text{Cov}}(X, Y) = E[XY] - E[X] \cdot E[Y]$$

$$= E[X] \cdot E[Y] - E[X] \cdot E[Y]$$

$$\boxed{\cancel{\text{Cov}}(X, Y) = 0} \quad \leftarrow \text{true for continuous } X, Y \text{ and for discrete } X, Y.$$