

5.7: Variance of a Continuous Random Variable

The variance of a continuous random variable X is defined by

$$\begin{aligned}\text{Var}(X) &= E[(X - \mu)^2] \\ &= E[X^2] - (E[X])^2\end{aligned}$$

where $\mu = E[X]$

Comments

- This is the same definition that we saw for discrete random variables.

- Recall

$$E[X] = \int_{-\infty}^{+\infty} x f(x) dx \quad \text{where } f \text{ is the prob. density function for } X$$

- To compute $\text{Var}(X)$ we'll need to be able to compute the expected value of a function $g(X) = (X - \mu)^2$ (or $g(X) = X^2$)

- we can do this by finding the prob. density function for $Z = (X - \mu)^2$ or use the following result...

Proposition 2.1 (p. 181, Ross)

If X is a continuous random variable with probability density function $f(x)$, then for any real valued function g ,

$$E[g(X)] = \int_{-\infty}^{+\infty} g(x) f(x) dx$$

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A proof of Proposition 2.1 is given in the textbook (p.181).

EXAMPLE (see previous example p. (154) - (155) notes)

• Let X have prob. density function

$$f_X(x) = \begin{cases} 0 & x < 0 \\ e^{-x} & x \geq 0 \end{cases}$$

(exponential with $\lambda=1$)

• Let $Y = g(X) = X^2$

• Recall that we found $f_Y(y) = \begin{cases} 0 & y < 0 \\ \frac{e^{-\sqrt{y}}}{2\sqrt{y}} & y > 0 \end{cases}$

Find the expected value of Y $E[Y]$ (i.e. $E[Y^2]$).

Method 1: use pdf f_Y

$$\begin{aligned} E[Y] &= \int_{-\infty}^{\infty} y f_Y(y) dy = \int_0^{\infty} y \frac{e^{-\sqrt{y}}}{2\sqrt{y}} dy \\ &= \frac{1}{2} \int_0^{\infty} \sqrt{y} e^{-\sqrt{y}} dy \end{aligned}$$

$$\begin{cases} \text{Let } s = \sqrt{y} \\ ds = +\frac{1}{2} y^{-1/2} dy = +\frac{1}{2\sqrt{y}} dy & dy = 2\sqrt{y} ds = 2s ds \end{cases}$$

$$= \frac{1}{2} \int_0^{\infty} \underbrace{s^2}_{u} \underbrace{e^{-s}}_{dv} ds$$

↓ Integrate by parts

$$= -s^2 e^{-s} \Big|_0^{\infty} + \int_0^{\infty} 2s e^{-s} ds$$

$$= 0 + 2 \int_0^{\infty} s e^{-s} ds$$

$$= 2 \left[-s e^{-s} \Big|_0^{\infty} + \int_0^{\infty} e^{-s} ds \right] = 2 \left[0 - e^{-s} \Big|_0^{\infty} \right] = \boxed{2} = E[Y]$$

$$u = s^2 \quad du = 2s ds$$

$$dv = e^{-s} ds \quad v = -e^{-s}$$

$$u = s \quad du = ds$$

$$dv = e^{-s} ds \quad v = -e^{-s}$$

Method 2 use pdf f_X and Proposition 2.1

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_0^{\infty} x^2 e^{-x} dx$$

we just did this integral!

$$\text{So } \boxed{E[X^2] = 2} = \boxed{E[I]} \quad \boxed{= 2}$$

So for this random variable X with $f_X(x) = \begin{cases} 0 & x < 0 \\ e^{-x} & x \geq 0 \end{cases}$

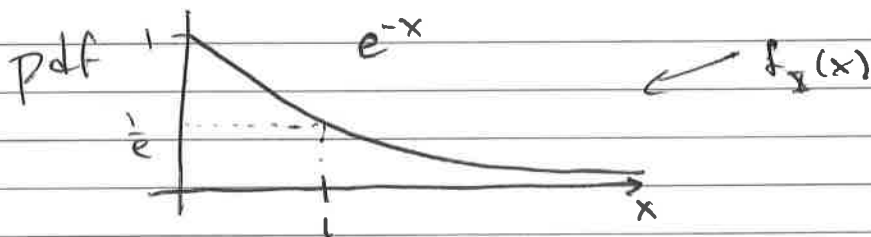
$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^{\infty} x e^{-x} dx$$

$$= x(-e^{-x}) \Big|_0^{\infty} - \int_0^{\infty} -e^{-x} dx$$

$$= 0 + -e^{-x} \Big|_0^{\infty} = 1$$

$$\text{So } E[X] = 1$$

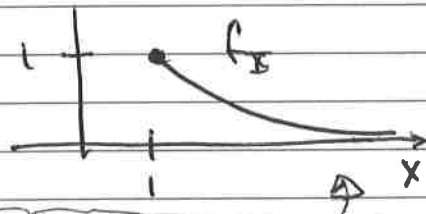
$$\text{Var}(X) = E[X^2] - (E[X])^2 = 2 - 1 = \boxed{1 = \text{Var}(X)}$$



EXAMPLE

Let X be a random variable with

$$f_X(x) = \begin{cases} 0 & x < 1 \\ \frac{1}{x^2} & x \geq 1 \end{cases}$$



However, Note:

$$E[X] = \int_{-\infty}^{+\infty} x f_X(x) dx$$

$$= \int_1^{\infty} x \frac{1}{x^2} dx = \int_1^{\infty} \frac{1}{x} dx = \ln x \Big|_1^{\infty} \rightarrow +\infty$$

Note: $\int_{-\infty}^{\infty} f(x) dx = \int_1^{\infty} \frac{1}{x^2} dx$
 $= -\frac{1}{x} \Big|_1^{\infty} = 1$

So the Expected Value of X does not exist.

Lemma 2.1 (Ross, p. 181)

For a nonnegative random variable I (range of $I \geq 0$)

$$E[I] = \int_0^{\infty} P\{I > y\} dy$$

the
result

cumulative dist. function
recall $P\{I \leq y\}$
 $= \int_{-\infty}^y f_I(t) dt$

$$= \int_0^{\infty} \left[\int_y^{\infty} f_I(t) dt \right] dy$$

integrate by
parts

$$\begin{cases} dv = dy \\ u = \int_y^{\infty} f_I(t) dt \end{cases} \quad \begin{cases} v = y \\ du = -f_I(y) dy \end{cases}$$

$$= \left[y \int_y^{\infty} f_I(t) dt \right]_{y=0}^{\infty} - \int_0^{\infty} y (-f_I(y)) dy$$

$$= + \int_0^{\infty} y f_I(y) dy = E[I].$$

to
see
...

5.3 The Uniform Random Variable

The continuous random variable X is a uniform random variable on (a, b) if it has probability density function

$$f_X(x) = \begin{cases} \frac{1}{b-a} & a < x < b \\ 0 & \text{otherwise} \end{cases}$$

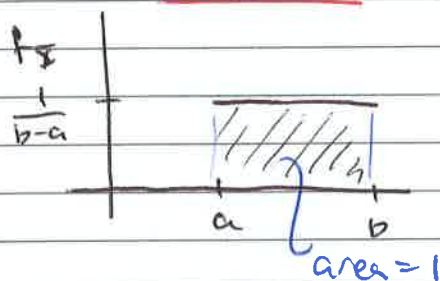
i.e. X is uniformly distributed on (a, b)

(equivalently, we can write $[a, b]$ and $a \leq x \leq b$ since the probability of one-point sets for continuous random variables is zero).

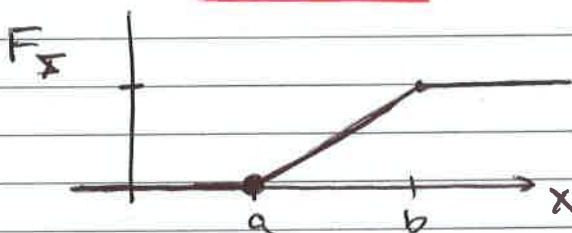
• Cumulative distribution function

$$F_X(x) = P\{X \leq x\} = \int_{-\infty}^x f_X(x) dx = \begin{cases} 0 & x < a \\ \int_a^x \frac{1}{b-a} dx = \frac{x-a}{b-a} & a \leq x \leq b \\ 1 & x > b \end{cases}$$

pdf of X



cdf of X



• Expected Value

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_a^b \frac{x}{b-a} dx = \frac{\frac{1}{2} x^2}{(b-a)} \Big|_a^b = \frac{1}{2} \frac{b^2 - a^2}{b-a} = \frac{1}{2} (b+a)$$

$$E[X^2] = \int_a^b \frac{x^2}{b-a} dx = \frac{1}{3} \frac{b^3 - a^3}{b-a} = \frac{1}{3} \frac{(b-a)(b^2 + ab + a^2)}{(b-a)} = \frac{b^2 + ab + a^2}{3}$$

↑
midpoint
of segment
(a, b)

• Variance

$$\begin{aligned} \text{Var}(X) &= E[X^2] - (E[X])^2 = \frac{1}{3} (b^2 + ab + a^2) - \frac{1}{4} (b+a)^2 = \frac{4b^2 + 4ab + 4a^2 - 3(b^2 + 2ab + a^2)}{12} \\ &= \frac{1}{12} (b^2 - 2ab + a^2) = \left(\frac{1}{12} (b-a)^2 \right)^{1/2} = \text{Var}(X) \end{aligned}$$

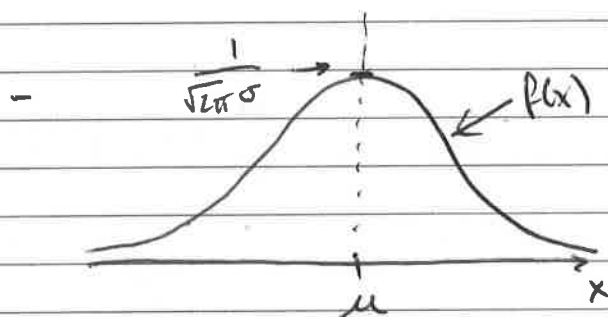
← function
of width.

5.4 Normal Random Variable

A random variable X is a normal random variable (or X is normally-distributed) with parameters μ and σ^2 if its probability ~~mass~~ density function is

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad -\infty < x < \infty$$

- recall, then $P\{a \leq X \leq b\} = \int_a^b f(x) dx$



- "bell-shaped" curve

$E[X] = \mu$
 $Var(X) = \sigma^2$

} — see p. (169) in notes
 we'll check these in the context of standard normal then generalize...

* Let's first check some properties of the standard normal random variable X with

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \quad (\text{i.e. } \sigma=1, \mu=0)$$

$-\infty < x < \infty$

- First, let's confirm that

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

is equal to one (is it?)

$$\text{Let } I = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

← wifty trick

Note:

$$\begin{aligned} I^2 &= \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \right) \left(\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \right) \\ &= \frac{1}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x^2+y^2)/2} dx dy \end{aligned}$$

convert to polar coordinates

$$\begin{aligned} dx dy &\rightarrow r dr d\theta \\ r^2 &= x^2 + y^2 \end{aligned}$$

$$I^2 = \frac{1}{2\pi} \int_0^{2\pi} \left[\int_0^{\infty} e^{-r^2/2} r dr \right] d\theta$$

$$= \frac{1}{2\pi} 2\pi \int_0^{\infty} r e^{-r^2/2} dr$$

$$= -e^{-r^2/2} \Big|_0^{\infty} = 0 - (-1) = 1 \quad \checkmark$$

- Next, let's compute the expected value $E[X]$.

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x e^{-\frac{x^2}{2}} dx$$

$$= \frac{1}{\sqrt{2\pi}} \left[\int_{-\infty}^0 x e^{-\frac{x^2}{2}} dx + \int_0^{\infty} x e^{-\frac{x^2}{2}} dx \right]$$

$$\int_0^{\infty} x e^{-\frac{x^2}{2}} dx = \int_0^{-\infty} -e^u du = -e^u \Big|_0^{-\infty} = (+1)$$

$u = -\frac{x^2}{2}$
 $du = -x dx$

$$\int_{-\infty}^0 x e^{-\frac{x^2}{2}} dx = \int_{-\infty}^0 -e^u du = -e^u \Big|_{-\infty}^0 = (-1)$$

so both of these integrals are finite and it follows

$$\boxed{E[X] = 0}$$

• Next, let's compute the Variance (X = standard normal)

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$E[X^2] = \int_{-\infty}^{\infty} \frac{x^2}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} x^2 e^{-\frac{x^2}{2}} dx$$

$$\begin{aligned} u &= x & du &= dx \\ dv &= x e^{-x^2/2} dx & v &= -e^{-x^2/2} \end{aligned}$$

$$= \frac{1}{\sqrt{2\pi}} \left[-x e^{-x^2/2} \Big|_{-\infty}^{\infty} - \int_{-\infty}^{\infty} -e^{-x^2/2} dx \right]$$

$$= \boxed{\frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-x^2/2} dx} = \int_{-\infty}^{\infty} f(x) dx = 1$$

$= 1$

So $\text{Var}(X) = 1 - 0 = 1$

∴ For standard normal distribution

$$P\{X \in B\} = \int_B f(x) dx$$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$$

$$-\infty < x < \infty$$

$$E[X] = 0$$

$$\text{Var}(X) = 1$$

Cumulative Distribution Function

common tabulated quantity - see P. 190

(167)

$$F(x) = P\{X \leq x\} = \int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2} dt$$

$$\equiv \Phi(x) \text{ often}$$

$\leftarrow X = \text{standard normal distribution}$

Note:

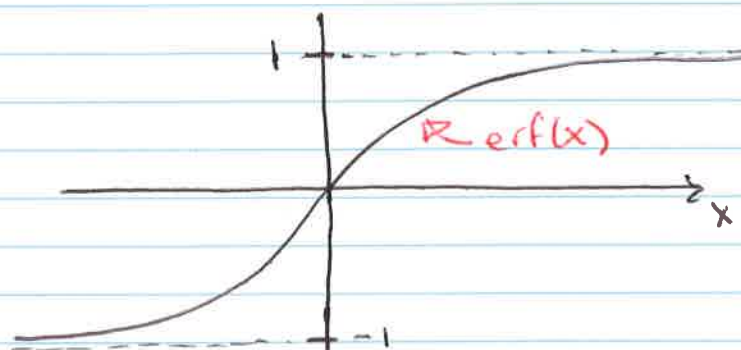
$$\left\{ \begin{aligned} \text{erf}(x) &= \frac{2}{\sqrt{\pi}} \int_0^x e^{-s^2} ds = \text{error function} \end{aligned} \right.$$

$$\left\{ \begin{aligned} \text{erfc}(x) &= \frac{2}{\sqrt{\pi}} \int_x^{\infty} e^{-s^2} ds = \text{complementary error function} \end{aligned} \right.$$

$$\text{erf}(0) = 0$$

$$\text{erf}(x \rightarrow \infty) = 1$$

$$\text{erf}(x) + \text{erfc}(x) = 1$$



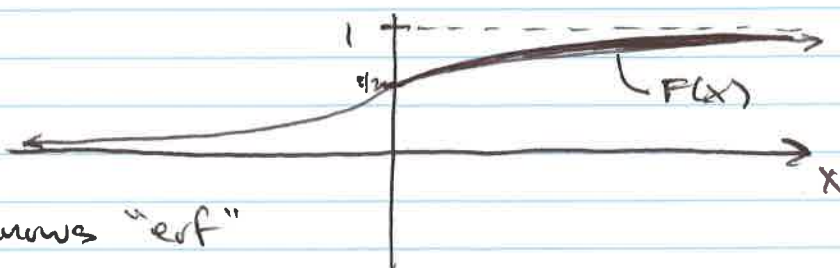
$$\text{let } t = \sqrt{2} s$$

$$dt = \sqrt{2} ds$$

$$\text{Then } F(x) = \int_{-\infty}^{-\frac{x}{\sqrt{2}}} \frac{1}{\sqrt{2\pi}} e^{-s^2} \sqrt{2} ds = \frac{1}{\sqrt{\pi}} \int_{-\infty}^{-\frac{x}{\sqrt{2}}} e^{-s^2} ds = \frac{1}{\sqrt{\pi}} \int_{-\frac{x}{\sqrt{2}}}^{\infty} e^{-s^2} ds$$

$$\text{So } F(x) = \frac{1}{2} \text{erfc}\left(-\frac{x}{\sqrt{2}}\right)$$

$$= \frac{1}{2} \left(1 - \text{erf}\left(-\frac{x}{\sqrt{2}}\right) \right) = \frac{1}{2} \left(1 + \text{erf}\left(\frac{x}{\sqrt{2}}\right) \right)$$



Matlab knows "erf"

The table in book P. 190 shows values for this

We can connect normally distributed X with standard normal random variable Z as follows.

- Let X be ~~random variable~~ a normal random variable with parameters μ and σ^2 . So

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- Let $Z = \frac{X-\mu}{\sigma}$. Then Z is often called $Z = \frac{X-\mu}{\sigma}$.

$$F_Z(y) = P\{Z \leq y\} = P\left\{\frac{X-\mu}{\sigma} \leq y\right\} = P\{X \leq y\sigma + \mu\}$$

$$= \int_{-\infty}^{y\sigma + \mu} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

but then

$$f_Z = \frac{d}{dy} F_Z(y) = \frac{d}{dy} \int_{-\infty}^{y\sigma + \mu} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \bigg|_{x=y\sigma + \mu} \cdot \sigma \cdot \frac{d}{dy}(y\sigma + \mu)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-y^2/2}$$

← standard normal prob. density function.

So if $E[Y] = 0$

$$\text{Var}(Y) = E[Y^2] - (E[Y])^2 = 1 - 0 = 1$$

Then

$$E[X] = E[\mu + \sigma Y]$$

$$= \int_{-\infty}^{+\infty} (\mu + \sigma y) f_Y(y) dy$$

$$= \underbrace{\mu \int_{-\infty}^{+\infty} f_Y(y) dy}_{=1} + \underbrace{\sigma \int_{-\infty}^{+\infty} y f_Y(y) dy}_{=0 = E[Y]} = \mu \quad \checkmark$$

and

$$E[X^2] = E[(\mu + \sigma Y)^2]$$

$$= E[\mu^2 + 2\mu\sigma Y + \sigma^2 Y^2]$$

$$= \int_{-\infty}^{+\infty} (\mu^2 + 2\mu\sigma y + \sigma^2 y^2) f_Y(y) dy$$

$$= \mu^2 \underbrace{\int_{-\infty}^{+\infty} f_Y(y) dy}_{=1} + 2\mu\sigma \underbrace{\int_{-\infty}^{+\infty} y f_Y(y) dy}_{=E[Y]=0} + \sigma^2 \underbrace{\int_{-\infty}^{+\infty} y^2 f_Y(y) dy}_{=E[Y^2]=1}$$

$$= \mu^2 + \sigma^2$$

So

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$= \mu^2 + \sigma^2 - \mu^2$$

$$= \sigma^2 \quad \checkmark$$

EX

Suppose X is a normal random variable
with $\mu = -3$ and $\sigma^2 = 9$

Find $P\{-4 \leq X \leq 0\}$

write $Z = \frac{X - \mu}{\sigma}$

then $X = \mu + \sigma Z$

$$P\{-4 \leq X \leq 0\} = P\{-4 \leq \mu + \sigma Z \leq 0\}$$

$$= P\left\{-\frac{4 - \mu}{\sigma} \leq Z \leq \frac{0 - \mu}{\sigma}\right\}$$

$$= P\left\{-\frac{4 - (-3)}{\sqrt{9}} \leq Z \leq \frac{0 - (-3)}{\sqrt{9}}\right\}$$

$$= P\left\{-\frac{1}{3} \leq Z \leq 1\right\}$$

$$= P\{Z \leq 1\} - P\{Z \leq -\frac{1}{3}\}$$

by symmetry

$$= P\{Z \geq \frac{1}{3}\}$$



$$= P\{Z \leq 1\} - P\{Z \geq \frac{1}{3}\}$$

$$= P\{Z \leq 1\} - (1 - P\{Z < \frac{1}{3}\})$$

$$= \cancel{P\{Z \leq 1\}} - 1 + \cancel{P\{Z < \frac{1}{3}\}}$$

$$= F(1) - 1 + F(\frac{1}{3})$$

$$= \frac{1 + \text{erf}(\frac{1}{\sqrt{2}})}{2} - 1 + \frac{1 + \text{erf}(\frac{1}{3\sqrt{2}})}{2} = 0.8413 - 1 + 0.6306 = 0.4719$$

see p. 167

where $F(x) = \frac{1 + \text{erf}(\frac{x}{\sqrt{2}})}{2}$

Note:

$$P\left\{-\frac{1}{3} \leq Z \leq 1\right\}$$

$$= P\{Z \leq 1\} - P\{Z \leq -\frac{1}{3}\}$$

$$=$$

$$\frac{1}{2} \left(1 + \operatorname{erf}\left(\frac{1}{\sqrt{2}}\right) \right) - \frac{1}{2} \left(1 + \operatorname{erf}\left(-\frac{1/3}{\sqrt{2}}\right) \right)$$

$$= \frac{1}{2} + \frac{1}{2} \operatorname{erf}\left(\frac{1}{\sqrt{2}}\right) - \frac{1}{2} - \frac{1}{2} \operatorname{erf}\left(-\frac{1/3}{\sqrt{2}}\right)$$

$$= \frac{1}{2} \left[\operatorname{erf}\left(\frac{1}{\sqrt{2}}\right) - \operatorname{erf}\left(-\frac{1/3}{\sqrt{2}}\right) \right]$$

Note $\operatorname{erf}\left(-\frac{1/3}{\sqrt{2}}\right) = -\operatorname{erf}\left(\frac{1/3}{\sqrt{2}}\right)$

$$= \frac{1}{2} \left[\operatorname{erf}\left(\frac{1}{\sqrt{2}}\right) + \operatorname{erf}\left(\frac{1/3}{\sqrt{2}}\right) \right]$$

$$=$$

$$= \frac{1}{2} \left[0.6827 + 0.2611 \right] = \underline{\underline{0.4719}}$$