



10. 解:

(1)

$$\text{令 } \varsigma^2 = c \sum_{i=1}^{n-1} (X_{i+1} - X_i)^2$$

$$\begin{aligned} E(\varsigma^2) &= cE\left[\sum_{i=1}^{n-1} (X_{i+1} - X_i)^2\right] = c \sum_{i=1}^{n-1} E(X_{i+1}^2 - 2X_{i+1}X_i + X_i^2) \\ &= c \sum_{i=1}^{n-1} (D(X_{i+1}) + [E(X_{i+1})]^2 - 2E(X_{i+1})E(X_i) + D(X_i) + [E(X_{i+1})]^2) \\ &= c(n-1)(\sigma^2 + \mu^2 - 2\mu^2 + \sigma^2 + \mu^2) = \sigma^2 \end{aligned}$$

故有,

$$c = \frac{1}{2(n-1)}$$

(2)

$$E(\overline{X}^2 - cS^2) = E(\overline{X}^2) - cE(S^2) = \left(\frac{\sigma^2}{n} + \mu^2\right) - c\sigma^2 = \mu^2$$

故,

$$c = \frac{1}{n}$$

13. 解:

(1)

由题知, $E(\hat{\theta}) = \theta, D(\hat{\theta}) > 0, \hat{\theta}^2 = (\hat{\theta})^2$

故,

$$E(\hat{\theta}^2) = E((\hat{\theta})^2) = D(\hat{\theta}) + E(\hat{\theta})^2 = D(\hat{\theta}) + \theta^2 > \theta^2$$

(2)

似然函数为,

$$L(\theta) = \begin{cases} \frac{1}{\theta^n} & 0 < x_1, x_2, \dots, x_n \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

故最大似然估计量 $\hat{\theta}$ 为,

$$\hat{\theta} = \max\{x_1, x_2, \dots, x_n\}$$

总体 X 分布函数为,

$$F(X) = \begin{cases} 0 & x < 0 \\ \frac{x}{\theta} & 0 \leq x < \theta \\ 1 & x \geq \theta \end{cases}$$

故,

$$F_{\hat{\theta}}(x) = F_{\max}(x) = (F(x))^n = \begin{cases} 0 & x < 0 \\ (\frac{x}{\theta})^n & 0 \leq x < \theta \\ 1 & x \geq \theta \end{cases}$$

故,

$$f_{\hat{\theta}}(x) = \begin{cases} \frac{n}{\theta} (\frac{x}{\theta})^{n-1} & 0 \leq x \leq \theta \\ 0 & \text{otherwise} \end{cases}$$

故,

$$E(\hat{\theta}) = \int_{-\infty}^{+\infty} x f_{\hat{\theta}}(x) dx = \int_0^{\theta} n (\frac{x}{\theta})^n dx = \frac{n\theta}{n+1} \neq \theta$$

15. 解:

$$E(\hat{\theta}) = E\left(\sum_{i=1}^n a_i X_i\right) = \sum_{i=1}^n a_i E(X_i) = \theta \sum_{i=1}^n a_i$$

故有,

$$\sum_{i=1}^n a_i = 1$$

$$D(\hat{\theta}) = \sum_{i=1}^n a_i^2 D(X_i) = \sum_{i=1}^n a_i^2 \sigma_i^2$$

作函数

$$g(a_1, a_2, \dots, a_n, \lambda) = \sum_{i=1}^n a_i^2 \sigma_i^2 + \lambda \left(\sum_{i=1}^n a_i - 1 \right)$$

令

$$\begin{aligned} \frac{\partial g}{\partial a_i} &= 2\sigma_i^2 a_i + \lambda = 0 \\ \frac{\partial g}{\partial \lambda} &= \sum_{i=1}^n a_i - 1 = 0 \end{aligned}$$

故有 $a_i = -\frac{\lambda}{2\sigma_i^2}$, 代入得,

$$-\frac{\lambda}{2} \sum_{i=1}^n \frac{1}{\sigma_i^2} = 1$$

故,

$$\begin{aligned} \lambda &= -\frac{2}{\sum_{i=1}^n \frac{1}{\sigma_i^2}} \\ a_i &= \frac{1}{\sigma_i^2 \sum_{j=1}^n \frac{1}{\sigma_j^2}} \end{aligned}$$