
Please prepare your solutions using L^AT_EX or another word processing software.

1. (40 points) Consider the binomial sampling distribution $p(x) = \binom{N}{x} p^x (1-p)^{N-x}$. In classical inference, once you observe data (x) , you might write down the 95% confidence interval for p as

$$\hat{p} \pm 1.96 \times \sqrt{\frac{\hat{p}(1-\hat{p})}{N}}, \quad (1)$$

where $\hat{p} = \frac{x}{N}$. This follows from the asymptotic normality assumption of the estimator \hat{p} .

Generate a uniform spacing of p 's from .01 to .99, with a step size of .01. Under each value of p , generate a Binomial random number with $N=30$. Now construct the confidence interval specified above. Since you know the actual value of p , you can confirm if this interval does in fact cover the *true* parameter.

Repeat this 10,000 times for each value of p , and keep track of the frequency of times each of the 10,000 repetitions cover the known parameter. Show a plot of the behavior of the “true” coverage probabilities for the full range of p 's (The y-axis will have frequency on it and the x-axis will be over p).

Repeat the exercise with $N = \{50, 100\}$. Conclude with your thoughts on the experiment. Are you surprised?

2. Consider $X \sim \text{Binomial}(N, \theta)$. Consider the uniform prior $p(\theta) = 1$, where $0 \leq \theta \leq 1$. The posterior sampling distribution should be clear. The prior that you have specified places an equal amount of weight on every possible value of p .

However, some people like to work with the log-odds, which we write as $\Lambda = \log\left(\frac{\theta}{1-\theta}\right)$.

- (a) (5 points) Find $p(\Lambda)$. That is, find the pdf for Λ (this is just a simple transformation problem).
- (b) (5 points) Now, do the problem the other way around. Consider placing a uniform prior on Λ . That is let $p(\Lambda) \propto 1$. What is the implied prior distribution on θ ?
- (c) (10 points) Reflect on the results in part a and part b.

3. (Hoff Exercise 3.3.) Tumor counts: A cancer laboratory is estimating the rate of tumorigenesis in two strains of mice, A and B. They have tumor count data for 10 mice in strain A and 13 mice in strain B. Type A mice have been well studied, and information from other labs suggests that type A mice have tumor counts that are approximately Poisson-distributed with a mean of 12. Tumor counts for type B mice are unknown, but type B mice are related to type A mice. The observed tumor counts for the two populations are:

$$\mathbf{y}_A = (12, 9, 12, 14, 13, 13, 15, 8, 15, 6)$$

$$\mathbf{y}_B = (11, 11, 10, 9, 9, 8, 7, 10, 6, 8, 8, 9, 7)$$

- (a) (15 points) Find the posterior distributions, means, variances, and 95% credible intervals for θ_A and θ_B , assuming a Poisson sampling distribution for each group and the following prior distribution:

$$\theta_A \sim \text{gamma}(120, 10), \theta_B \sim \text{gamma}(12, 1)$$

- (b) (15 points) Compute and plot the posterior expectation for θ_B under the prior distribution $\theta_B \sim \text{gamma}(12 \times n_0, n_0)$ for each value of $n_0 \in \{1, 2, \dots, 50\}$. Describe what sort of prior beliefs about θ_B would be necessary in order for the posterior expectation of θ_B to be close to that of θ_A .

4. (Hoff Exercise 3.12) Jeffreys' prior: Jeffreys suggested a default rule for generating a prior distribution of a parameter θ in a sampling model $p(y|\theta)$. Jeffreys' prior is given by $p_j(\theta) \propto \sqrt{I(\theta)}$, where $I(\theta) = -E[\delta^2 \log p(Y|\theta) / \delta \theta^2 | \theta]$ is the *Fisher information*.

- (a) (10 points) Let $Y \sim \text{binomial}(n, \theta)$. Obtain Jeffreys' prior distribution $p_j(\theta)$ for this model.