## Mean Estimation from One-bit Measurements

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Summary

Point estimation under communication constraints:

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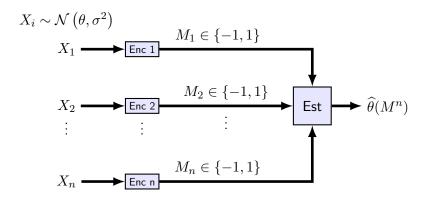
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- (ii) limited bits

#### Relevant scenarios:

- big data
- low-power sensors
- distributed computing / optimization

### This talk:

Estimating the mean  $\theta$  of a normal distribution  $\mathcal{N}(\theta, \sigma^2)$  from one-bit per sample  $(\sigma \text{ is known})$ 



▶ Distributed:  $M_i = f_i(X_i)$ 

$$X_{1} \sim \mathcal{N}\left(\theta,\sigma^{2}\right)$$

$$X_{1} \longrightarrow \underbrace{\text{Enc 1}}_{M_{1} \in \{-1,1\}}$$

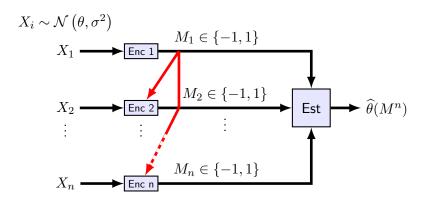
$$X_{2} \longrightarrow \underbrace{\text{Enc 2}}_{\vdots}$$

$$\vdots$$

$$M_{n} \in \{-1,1\}$$

$$X_{n} \longrightarrow \underbrace{\text{Enc n}}_{M_{n} \in \{-1,1\}}$$

- ▶ Distributed:  $M_i = f_i(X_i)$
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- Distributed hypothesis testing under quantization [Tsitsiklis '88] (distributed)
- Remote multiterminal source coding (CEO) [Berger, Zhang, Wiswanathan '96], [Oohama '97] (distributed)

# Consistency

Q: in what setting consistent estimation is possible?

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A: all!

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Q: in what setting consistent estimation is possible?

A: all ! Proof:

$$M_i = \mathbf{1}(X_i > 0), \quad i = 1, \dots, n$$

(as in the distributed setting)

$$\frac{1}{n}\sum_{i=1}^{n} M_i \to \mathbb{P}(X>0) = \Phi(\theta/\sigma)$$

## Efficiency

**Definition:** asymptotic relative efficiency (ARE) of an estimator:

$$ARE(\widehat{\theta}) \triangleq \lim_{n \to \infty} \mathbb{E}\left[\left(\widehat{\theta} - \theta\right)^2\right] / \left(\sigma^2/n\right)$$

(relative to minimax risk without bit constraint)

## Proposition

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- ▶ Encoder is required to describe  $\bar{\theta}$  using n bits
  - divide parameter space  $\Theta$  into  $2^n$  regions of equal size
  - send region index where  $\bar{\theta}$  falls

## Proposition

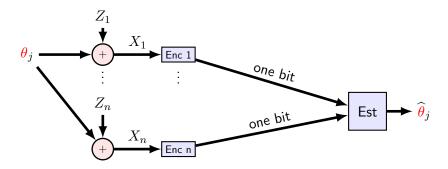
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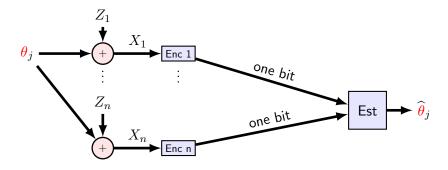
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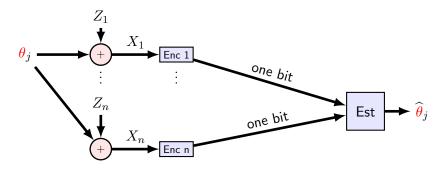
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Note: mean of  $\bar{\theta}$  is unknown – hard to derive a globally optimal strategy

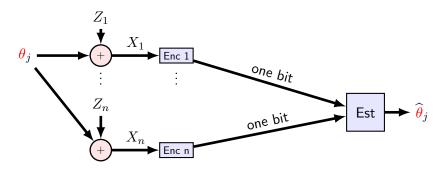




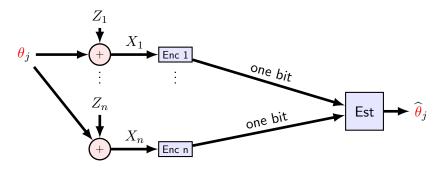
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- ▶ Block encode  $X_{j,1}, \ldots, X_{j,n}$  using k bits,  $j = 1, \ldots, k$
- $D_{CEO} = \inf_{k,\widehat{\theta}} \frac{1}{k} \sum_{j=1}^{k} \mathbb{E} \left( \theta_j \widehat{\theta}_j \right)^2$

Quadratic Gaussian CEO under optimal rate allocation [Chen, Zhang, Berger, Wicker '04] :

$$D_{CEO} \ge \frac{4}{3} \frac{\sigma^2}{n} + o(1/n)$$

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▶ In fact [K., Rini, Goldsmith '17]:

$$D_{CEO} \le \frac{4}{3} \frac{\sigma^2}{n} + \frac{\sigma_0^2}{3n} + o(1/n)$$

(ARE of 4/3 is achievable with coding over blocks)

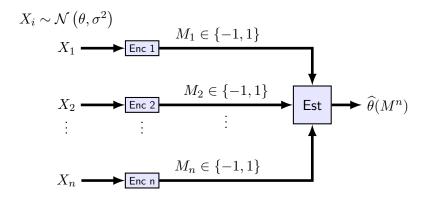
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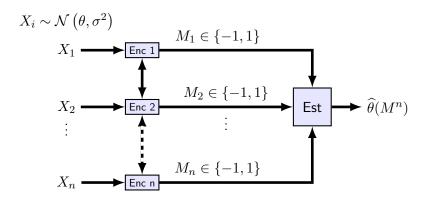
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Summary



▶ Distributed:  $M_i = f_i(X_i)$ 

 $\mathsf{ARE} \geq 4/3$ 

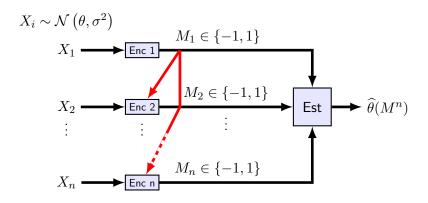


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- ▶ Centralized:  $M^n = f(X_1, ..., X_n)$

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# Three Encoding Scenarios



- ▶ Distributed:  $M_i = f_i(X_i)$
- ▶ Centralized:  $M^n = f(X_1, ..., X_n)$
- ▶ Adaptive / Sequential:  $M_i = f_i(X_i, M^{i-1})$

 $ARE \ge 4/3$ 

ARE = 1

 $\mathsf{ARE} = \pi/2$ 

# Main Results (adaptive encoding)

Theorem (achievability)

There exists an estimator with ARE  $\pi/2$ 

Theorem (converse)

No estimator have ARE lower than  $\pi/2$ 

existence of an estimator with ARE  $=\pi/2$ 

(i) For 
$$X \sim \mathcal{N}(\theta, \sigma^2)$$
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$$\operatorname{med}(X) = \underset{m}{\operatorname{argmin}} \ \mathbb{E} \left| X - m \right|$$

(iii) Stochastic gradient descent on  $\mathbb{E}|X - \theta|$ :

$$\theta_n = \theta_{n-1} + \gamma_n \mathrm{sign}(X_n - \theta_n)$$

$$\widehat{\theta}_n = \frac{1}{n} \sum_{i=1}^n \theta_i$$

existence of an estimator with ARE  $=\pi/2$ 

- (i) For  $X \sim \mathcal{N}(\theta, \sigma^2)$ ,  $\text{med}(X) = \theta$
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From [Polyak & Juditsky '92] (under conditions on  $(\gamma_n)$ ):

$$\sqrt{n}(\theta - \widehat{\theta}_n) \to \mathcal{N}\left(0, \sigma^2 \pi/2\right)$$

#### Converse

#### $ARE \geq \pi/2$

The van-Trees inequality (e.g. [Tsybakov '08]) implies

$$\mathbb{E}\left(\theta - \widehat{\theta}\right)^2 \ge \frac{1}{\mathbb{E}I_{\theta}(M^n) + I_{\pi}} \ge \frac{1}{\sum_{i=1}^n I_{\theta}(M_i|M^{i-1}) + I_{\pi}}$$

 $I_{\pi}$  is the location Fisher information with w.r.t. some prior  $\pi(d\theta)$  on  $\Theta$ 

Lemma (K. & Duchi '17)

$$I_{\theta}(M_i|M^{i-1}) \le \frac{2}{\pi\sigma^2}$$

#### Proof:

Stein identity implies that portion maximizing the information is a threshold:  $M_i^{-1}(1)=(\theta,\infty)$ 

The rest follows by induction over number of portions

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## Distributed Encoding

Threshold Detection

We consider only messages of the form

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$$\longleftarrow \qquad \qquad M_i = -1 \qquad \stackrel{t_i}{\longleftarrow} \qquad \qquad M_i = 1 \qquad \longrightarrow$$

Assume:

$$\lambda_n([a,b]) = \frac{1}{n} \mathsf{card}\left([a,b] \cap \{t_i\}\right)$$

converges weakly to a probability distribution  $\boldsymbol{\lambda}$ 

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**Example:**  $t_1, \ldots, t_n$  drawn i.i.d. from a distribution  $\lambda$  on  $\mathbb R$ 

# Main Results (distributed encoding)

#### Theorem

(i) The Maximum likelihood estimator  $\widehat{ heta}_{ML}$  satisfies

$$\sqrt{n}(\theta - \widehat{\theta}_{ML}) \to \mathcal{N}\left(0, \sigma^2/K_{\lambda}(\theta)\right)$$

where:

$$K_{\lambda}(\theta) = \int_{\mathbb{R}} \eta \left( \frac{t - \theta}{\sigma} \right) \lambda(dt)$$
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(ii) For any estimator  $\widehat{\theta}(M_1, \dots, M_n)$ :

$$\liminf_{c \to \infty} \liminf_{n \to \infty} \sup_{\tau : |\tau - \theta| \le \frac{c}{\sqrt{n}}} n \mathbb{E} \left( \widehat{\theta} - \tau \right)^2 \ge \sigma^2 / K_{\lambda}(\theta),$$

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ARE = 
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(equality iff  $\lambda = \delta_{\theta}$ )

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▶ Minimax  $\lambda$  and ARE can be found using a convex program – depends on radius of  $\Theta$ 

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- ARE of MLE characterized by density of threshold values

#### Open question

Is there a distributed encoding scheme with ARE that is both finite and independent of size of parameter space ?

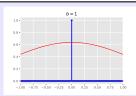
# Minimax threshold density

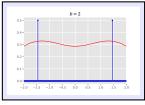
Minimax 
$$\lambda$$
 for  $\theta \in (-b\sigma, b\sigma)$ :

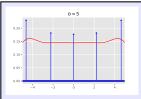
$$\begin{aligned} & \underset{\tau \in (-b,b)}{\inf} \int \eta(t-\tau) \lambda(dt) \\ & \text{subject to} \quad \lambda(dt) \geq 0, \quad \int \lambda(dt) \leq 1. \end{aligned}$$

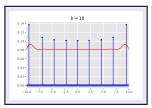
### Minimax $\lambda$

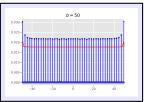
support of optimal threshold density  $\lambda^\star$ 

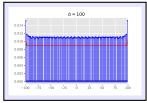








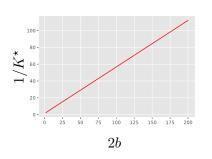


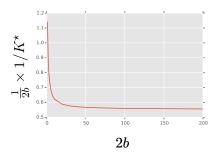


$$K^* = \inf_{\theta} K^*(\theta) = \inf_{\theta} \int \eta(t - \theta) \lambda^*(dt)$$

### Minimax $\lambda$

#### Minimax ARE vs size of parameter space





► ARE increases with size of parameter space