Fast Bayesian A/B and multivariate testing

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Introduction (1/2)

- Frequentist inference (classical inference / classical point estimation)
 - ▶ Base on asymptotic performance ⇒ Central Limit Theorem (CLT).
 - P-value: the probability of seeing a result at least as extreme as a real result after a A/A test of the same size [Stu15]. It is defined as

$$\text{p-value} = P[t \geq t_e | H_0], \quad H_0: B \equiv A$$

- ▶ Hypothesis testing based on rejecting H_0 . p-value $\neq P[B > A]!$
- Confidence interval: if we repeated the same experiment used to construct an interval for an unobserved value n times $(n \to \infty)$, (100δ) % of the intervals would contain the true value. This is not a credible interval!
- Set and fix stopping rule and sample size (power test calculation).
- Bayesian inference
 - Bayes Theorem: given a prior distribution $p(\theta)$, update belief based on sample x

$$p(\theta|x) = \frac{f(x|\theta)p(\theta)}{f(x)}$$

- ▶ Choose of prior parameters ⇒ update ⇒ posterior parameters.
- Calculation of posterior distribution.
- Calculation of predictive posterior distribution.

Introduction (2 / 2)

- ► Bayesian inference: advantages
 - Ease interpretability.
 - ▶ Sample size is not fixed in advance ⇒ repeated/streaming testing.
 - ▶ Account for uncertainty; points estimates ⇒ random variables.
 - ► Immune to data peeking
- Bayesian inference: disadvantages
 - Analytical tractability
 - Computational cost

Bayesian testing metrics: probability to beat

▶ **A/B testing**: the error probability or probability of $X_B > X_A$

$$E(B) = P[X_B > X_A]$$

```
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>> (xb > xa).mean()
```

▶ Multivariate testing: the probability to beat all

$$E(X_i) = P\left[X_i > \max_{j \neq i} X_j\right]$$

```
>>> import numpy as np
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>> xc = stats.beta(5, 60).rvs(size=int(1e6))
>>> xd = stats.beta(7, 90).rvs(size=int(1e6))
>>> maxall = np.maximum.reduce([xa, xc, xd])
>>> (xb > maxall).mean()
```

Bayesian testing metrics: expected loss

► A/B testing: the expected value of the loss function

$$EL(B) = \mathrm{E}[\max(X_A - X_B, 0)]$$

```
>>> import numpy as np
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>> np.maximum(xa - xb, 0).mean()
```

▶ Multivariate testing: the expected loss function vs all

$$EL(X_i) = \mathbb{E}[\max(\max_{j \neq i} X_j - X_i, 0)]$$

```
>>> import numpy as np
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>> xc = stats.beta(5, 60).rvs(size=int(1e6))
>>> xd = stats.beta(7, 90).rvs(size=int(1e6))
>>> maxall = np.maximum.reduce([xa, xc, xd])
>>> np.maximum(maxall - xb, 0).mean()
```

Bayesian testing metrics: expected relative loss

► A/B testing: the expected value of the relative loss function

$$ERL(B) = E[(X_A - X_B)/X_B]$$

```
>>> import numpy as np
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>>((xa - xb) / xb).mean()
```

► Multivariate testing: the expected relative loss function vs all

$$ERL(X_i) = E[(\max_{j \neq i} X_j - X_i)/X_i]$$

```
>>> import numpy as np
>>> from scipy import stats
>>> xa = stats.beta(2, 10).rvs(size=int(1e6))
>>> xb = stats.beta(3, 12).rvs(size=int(1e6))
>>> xc = stats.beta(5, 60).rvs(size=int(1e6))
>>> xd = stats.beta(7, 90).rvs(size=int(1e6))
>>> maxall = np.maximum.reduce([xa, xc, xd])
>>> ((maxall - xb) / xb).mean()
```

Computation of credible intervals

Definition: A credible interval is a region with a particular probability to contain an unobserved value. Bayesian equivalent of the confidence interval. Given a significance level δ :

▶ Equally-tailed Interval (ETI): Credible interval using the quantile method, with quantile function $Q = F^{-1}$, solving $F(z) = \delta/2$ and $F(z) = 1 - \delta/2$, satisfying

$$P(Q(\delta/2) < Z < Q(1-\delta/2)) = 1-\delta$$

- Assumption: distribution is symmetric.
- Highest Density Interval (HDI): Solving

$$P(I < Z < u) = 1 - \delta$$

for I and u, being the lower and upper bound of the interval.

▶ No assumptions, appropriate for symmetric and skewed distributions.

Computation of credible intervals: HDI – Monte Carlo sampling

The HDI computes the narrowest of the infinite intervals satisfying $P(I < Z < u) = 1 - \delta$. R code in [Kru15]. NumPy implementation to compute HDI given MC samples

```
>>> import numpy as np
>>> n = len(x)
>>> xsorted = np.sort(x)
>>> n_included = int(np.ceil(interval_length * n))
>>> n_ci = n - n_included
>>> ci = xsorted[n_included:] - xsorted[:n_ci]
>>> j = np.argmin(ci)
>>> hdi_min = xsorted[j]
>>> hdi_max = xsorted[j + n_included]
```

Computation of credible intervals: HDI – mathematical optimization

The HDI computes the narrowest interval by solving the minimization problem [CS99],

$$\min_{l < u} (|f(u) - f(l)| + |F(u) - F(l) - (1 - \delta)|).$$

Reformulation: remove absolute values and add term u - I

$$\begin{aligned} & \min_{u,l,t,w} & t + w + u - l \\ & \text{s.t.} & -t + f(u) - f(l) \ge 0 \\ & & t + f(u) - f(l) \ge 0 \\ & & -w + F(u) - F(l) - (1 - \delta) \ge 0 \\ & & w + F(u) - F(l) - (1 - \delta) \ge 0 \\ & & u - l - \epsilon \ge 0 \\ & & l \in [l_{\min}, l_{\max}]] \\ & u \in [u_{\min}, u_{\max}] \end{aligned}$$

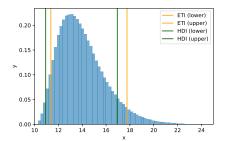
where $\epsilon > 0$. Parameters I_{\min} , I_{\max} , u_{\min} and u_{\max} denote the bounds for the interval limits I and u.

Computation of credible intervals: scipy.optimize

```
def func(x):
   return x[3] + x[2] + x[1] - x[0]
def obj_f(x):
   return f.pdf(x[1]) - f.pdf(x[0])
def obi F(x):
   return f.cdf(x[1]) - f.cdf(x[0])
epsilon = 1e-6
cons = (
    {'type': 'ineq', 'fun': lambda x: x[1] - x[0] - epsilon},
    {'type': 'ineq', 'fun': lambda x: -x[2] + obj_f(x)},
    {'type': 'ineq', 'fun': lambda x: x[2] + obj_f(x)},
   {'type': 'ineq', 'fun': lambda x: -x[3] + obj F(x) - interval length},
   {'type': 'ineq', 'fun': lambda x: x[3] + obj F(x) - interval length}
res = optimize.minimize(func, (*x0, 0, 0), method="SLSQP",
                        constraints=cons, bounds=[*bounds, (0, 1). (0. 1)])
```

Computation of credible intervals: example 1

```
>>> from scipy import stats
>>> from cprior.cdist import ci_interval
>>> x = stats.gamma(4, 10).rvs(size=int(1e6), random_state=42)
>>> ci_interval(x=x, interval_length=0.9, method="ETI")
array([11.36321512, 17.75748775])
>>> ci_interval(x=x, interval_length=0.9, method="HDI")
array([10.92933934, 16.94237247])
```

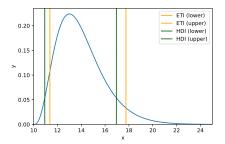


Timings (%timeit)1: ETI: 18 ms, HDI: 107 ms

¹Intel(R) Core(TM) i5-3317 CPU at 1.70GHz.

Computation of credible intervals: example 2

```
>>> import numpy as np
>>> from scipy import stats
>>> from cprior.cdist import ci_interval
>>> dist = stats.gamma(4, 10)
>>> ci_interval_exact(dist=dist, interval_length=0.9, method="ETI")
array([11.3663184 , 17.75365653])
>>> bounds = [(0, np.inf), (0, np.inf)]
>>> ci_interval_exact(dist=dist, interval_length=0.9, method="HDI", bounds=bounds)
array([10.93729501, 16.94611345])
```



Timings (%timeit): ETI: 0.2 ms, HDI: 45 ms

The CPrior library

- ▶ Python/C++ library, open source (LGPL-3.0)
- ► Github: https://github.com/guillermo-navas-palencia/cprior
- Documentation: http://gnpalencia.org/cprior/
- ► Technical notes [NP19]
- Support several conjugate prior distributions
 - Beta distribution
 - Gamma distribution
 - ▶ Pareto distribution ✓
 - Normal-inverse-gamma distribution
 - ▶ Others: beta-binomial, inverse gamma, multivariate distributions...
- Fast and accurate results:
 - Development of closed-forms in terms of special functions
 - Fast Monte Carlo methods
 - Median Latin Hypercube Sampling
 - Parallel crude Monte Carlo
 - Numerical integration
 - Streaming Bayesian testing
- ightharpoonup ~15000 lines of code

CPrior testing metrics: probability to beat

Let us consider probability distributions X_i with support \mathbb{R} .

▶ A/B testing: the error probability or probability of $X_B > X_A$

$$P[X_B > X_A] = \int_{-\infty}^{\infty} \int_{x_A}^{\infty} f(x_A, x_B) dx_B dx_A,$$

where $f(x_A, x_B)$ is the joint probability distribution, under the assumption of independence, i.e. $f(x_A, x_B) = f(x_A)f(x_B)$.

Multivariate testing: the probability to beat all

$$P\left[X_i > \max_{j \neq i} X_j\right] = \int_{-\infty}^{\infty} f(x_i) \prod_{j \neq i} F_{X_j}(x_i) dx_i.$$

Given $X_{max} = \max\{X_1, \dots, X_n\}$. The cumulative distribution function is

$$F_{X_{max}}(z) = P\left[\max_{i=1,\ldots,n} X_i \leq z\right] = \prod_{i=1}^n P[X_i \leq z] = \prod_{i=1}^n F_{X_i}(z),$$

where $F_{X_i}(z)$ is the cdf of each random variable X_i .

CPrior testing metrics: expected loss

Let us consider probability distributions X_i with support \mathbb{R} .

▶ A/B testing: the expected loss function if variant X_B is chosen

$$EL(X_B) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \max(x_A - x_B, 0) f(x_A, x_B) dx_B dx_A.$$

▶ Multivariate testing: the expected loss function vs all, taking $Y = \max_{j \neq i} X_j$

$$EL(X_i) = \int_{-\infty}^{\infty} yf(y)F_{X_i}(y) dy - \int_{-\infty}^{\infty} f(y)F_{X_i}^*(y) dy,$$

where $F_{X_i}^*(y) = \int_{-\infty}^y x_i f(x_i) dx_i$. The probability density function is obtain after derivation of $F_{X_{max}}(z)$

$$f_{X_{max}}(z) = \sum_{i=1}^{n} f_{X_i}(z) \prod_{j \neq i} F_{X_j}(z),$$

where $f_{X_i}(z)$ is the pdf of each random variable X_i .

Beta distribution: A/B testing

▶ Probability to beat: given two distributions $X_A \sim \mathcal{B}(\alpha_A, \beta_A)$ and $X_B \sim \mathcal{B}(\alpha_B, \beta_B)$, $P[X_B > X_A]$ is given by²

$$P[X_B > X_A] = 1 - \frac{B(\alpha_A + \alpha_B, \beta_A)}{\alpha_B B(\alpha_A, \beta_A) B(\alpha_B, \beta_B)} \, {}_3F_2\left(\begin{matrix} \alpha_B, \alpha_A + \alpha_B, 1 - \beta_B \\ 1 + \alpha_B, \alpha_A + \alpha_B + \beta_A \end{matrix}; 1\right),$$

where B(a, b) is the beta function and ${}_{3}F_{2}(a, b, c; d, e; z)$ is the generalized hypergeometric function.

- ► Implementation hypergeometric series (C++) https://github.com/guillermo-navas-palencia/cprior/blob/ master/cprior/_lib/src/beta.cpp
 - ▶ Special cases in terms of the regularized incomplete beta function $I_x(a, b)$.
- ► Timings

%timeit abtest.probability(variant="B", method="exact") 10.2 μ s \pm 58.2 ns per loop (mean \pm std. dev. of 7 runs, 100000 loops each)

²http://gnpalencia.org/cprior/formulas_conjugate_beta.html

Beta distribution: Multivariate testing - MLHS

Probability to beat:

$$P\left[X_{i} > \max_{j \neq i} X_{j}\right] = \int_{0}^{1} \frac{x^{\alpha_{i}-1}(1-x)^{\beta_{i}-1}}{B(\alpha_{i},\beta_{i})} \prod_{j \neq i} I_{x}(\alpha_{j},\beta_{j}) dx$$
$$= E\left[\prod_{j \neq i} I_{x}(\alpha_{j},\beta_{j})\right], \quad X \sim \mathcal{B}(\alpha_{i},\beta_{i}).$$

Median Latin Hypercube Sampling (MLHS)

```
r = np.arange(1, mlhs_samples + 1)
np.random.shuffle(r)
v = (r - 0.5) / mlhs_samples
x = self.models[variant].ppf(v)
```

1.
$$a \leftarrow 0, b \leftarrow 1$$

- 2. vector of indexes: π_i , $i = 1, \ldots, n$
- 3. random shuffle of π_i

4.
$$v_i = (b-a)\frac{\pi_i - 0.5}{n} + a$$

5.
$$x_i = F^{-1}(v_i)$$

Beta distribution: Multivariate testing - numerical integration

Probability to beat all:

$$P\left[X_i > \max_{j \neq i} X_j\right] = \int_0^1 \frac{x^{\alpha_i - 1} (1 - x)^{\beta_i - 1}}{B(\alpha_i, \beta_i)} \prod_{j \neq i} I_x(\alpha_j, \beta_j) dx.$$

```
def func_mv_prob(x, a, b, variant_params):
   pdf = (a - 1) * np.log(x) + (b - 1) * np.log(1 - x) - special.betaln(a, b)
   g = np.prod([special.betainc(a, b, x) for a, b in variant_params], axis=0)
   return np.exp(pdf) * g
```

integrate.quad(func=func_mv_prob, a=0, b=1, args=(a, b, variant_params))[0]

► Benchmark (5 variants)

Method	Samples	Rel. error	time
Monte Carlo	1e4	8e-2	50 ms
	1e5	1e-2	137 ms
	1e6	2e-3	1160 ms
MLHS	1e2	3e-3	2 ms
	1e3	3e-4	8 ms
	1e4	3e-5	65 ms
Quad	-	-	26 ms

Gamma distribution: A/B testing

▶ Probability to beat: given two distributions $X_A \sim \mathcal{G}(\alpha_A, \beta_A)$ and $X_B \sim \mathcal{G}(\alpha_B, \beta_B)$, $P[X_B > X_A]$ is given by³

$$P[X_B > X_A] = 1 - \frac{\beta_A^{\alpha_A} \beta_B^{\alpha_B}}{(\beta_A + \beta_B)^{\alpha_A + \alpha_B}} \frac{{}_{2}F_{1}\left(1, \alpha_A + \alpha_B; \alpha_B + 1; \frac{\beta_B}{\beta_B + \beta_A}\right)}{\alpha_B B(\alpha_A, \alpha_B)}$$
$$= I_{\frac{\beta_A}{\beta_A + \beta_B}}(\alpha_A, \alpha_B),$$

where ${}_{2}F_{1}(a,b;c;z)$ is the Gauss hypergeometric function and $I_{x}(a,b)$ is the regularized incomplete beta function.

Expected loss:

$$EL(X_B) = \frac{\alpha_A}{\beta_A} I_{\frac{\beta_B}{\beta_A + \beta_A}} (\alpha_B, \alpha_A + 1) - \frac{\alpha_B}{\beta_B} I_{\frac{\beta_B}{\beta_A + \beta_A}} (\alpha_B + 1, \alpha_A).$$

▶ Implementation $I_x(a, b)$: scipy.special.betainc.

³http://gnpalencia.org/cprior/formulas_conjugate_gamma.html

Gamma distribution: Multivariate testing - MLHS

Expected loss vs all:

$$\mathit{EL}(X_i) = \mathrm{E}\left[\mathit{YP}(\alpha_i, \beta_i \mathit{Y}) - rac{lpha_i}{eta_i}\mathit{P}(lpha_i + 1, eta_i \mathit{Y})\right], \quad \mathit{Y} \sim \max_{j \neq i} \mathcal{G}(lpha_j, eta_j),$$

where $P(\alpha, \beta)$ is the regularized lower incomplete gamma function.

Benchmark (5 variants)

Method	Samples	Rel. error	time
	1e4	2e-3	37 ms
MC	1e5	2e-4	100 ms
	1e2	9e-3	31 ms
MLHS	1e3	1e-3	255 ms
Quad	-	-	54 ms

Gamma distribution: Multivariate testing - MLHS

Expected relative loss vs all:

► Benchmark (5 variants)

Method	Samples	Rel. error	time
	1e4	4e-3	35 ms
MC	1e5	5e-4	90 ms
	1e2	6e-3	4 ms
MLHS	1e3	8e-4	16 ms
Quad	-	-	48 ms

Bayesian experiment: Bernoulli distribution (1/5)

A Bayesian multivariate test with control and 3 variants. Data follows a Bernoulli distribution with distinct success probability.

Generate control and variant models and build experiment. Select stopping rule and threshold (epsilon).

Bayesian experiment: Bernoulli distribution (2 / 5)

Check experiment description.

```
>>> experiment.describe()
```

```
Experiment: CTR
  Bayesian model:
                                  bernoulli-beta
  Number of variants:
                                                4
  Options:
    stopping rule
                              probability_vs_all
    epsilon
                                          0.99000
    min_n_samples
                                             1000
    max_n_samples
                                          not set
```

Priors:

beta	alpha	
1	1	Α
1	1	В
1	1	C
1	1	D

Bayesian experiment: Bernoulli distribution (3 / 5)

- Generate or pass new data and update models until a clear winner is found.
- ► The stopping rule will be updated after a new update.

```
with experiment as e:
    while not e.termination:
        data_A = stats.bernoulli(p=0.0223).rvs(size=25)
        data_B = stats.bernoulli(p=0.1128).rvs(size=15)
        data_C = stats.bernoulli(p=0.0751).rvs(size=35)
        data_D = stats.bernoulli(p=0.0280).rvs(size=15)

        e.run_update(**{"A": data_A, "B": data_B, "C": data_C, "D": data_D})

print(e.termination, e.status)
True winner B
```

Bayesian experiment: Bernoulli distribution (4 / 5)

▶ Reporting: experiment summary.

>>> experiment.summary()

	name	probability	expected_loss	improvement	probability_vs_all	expected_loss_vs_all	improvement_vs_all	n_samples
Α	control	-	-	-	0.00%	0.0881716	-572.15%	1675
В	variation	100.00%	1.30573e-27	84.43%	99.94%	1.63007e-06	32.00%	1005
C	variation	100.00%	5.95894e-21	76.97%	0.06%	0.0339692	-49.16%	2345
D	variation	97.89%	4.26579e-05	40.01%	0.00%	0.0764664	-288.51%	1005

▶ Reporting: statistics collected data throughout the experiment.

>>> experiment.stats()

	A	В	C	D
count	1675.000000	1005.000000	2345.000000	1005.000000
mean	0.019104	0.111443	0.073774	0.028856
std	0.136933	0.314836	0.261458	0.167484
min	0.000000	0.000000	0.000000	0.000000
25%	0.000000	0.000000	0.000000	0.000000
50%	0.000000	0.000000	0.000000	0.000000
75%	0.000000	0.000000	0.000000	0.000000
max	1.000000	1.000000	1.000000	1.000000

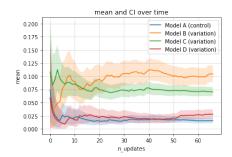
Bayesian experiment: Bernoulli distribution (5 / 5)

- ▶ Reporting: visualize stopping rule metric over time (updates).
- ▶ Reporting: visualize statistics over time (updates).

>>> experiment.plot_metric()

probability vs all 10 0.8 Model A (control) 0.6 Model B (variation) Model C (variation) 0.4 Model D (variation) 0.2 0.0 10 50 60 20 n updates

>>> experiment.plot_stats()



Bayesian experiment: normal distribution (1/4)

A Bayesian multivariate test with control and 3 variants. Data follows a normal distribution with distinct mean and standard deviation.

Generate control and variant models and build experiment. Select stopping rule and threshold (epsilon).

```
from scipy import stats
from cprior.models import NormalModel
from cprior.models import NormalMVTest
from cprior.experiment.base import Experiment
modelA = NormalModel(name="control")
modelB = NormalModel(name="variation")
modelC = NormalModel(name="variation")
modelD = NormalModel(name="variation")
mvtest = NormalMVTest({"A": modelA, "B": modelB, "C": modelC, "D": modelD})
experiment = Experiment(name="GPA", test=mvtest,
                        stopping rule="probability vs all", epsilon=0.99,
                        min n samples=500, max n samples=None,
                        nig metric="mu")
```

Bayesian experiment: normal distribution (2 / 4)

Check experiment description.

```
>>> experiment.describe()
```

Experiment: GPA Bayesian model: normal-normalinversegamma Number of variants: 4 Options: stopping rule probability_vs_all epsilon 0.99000 min_n_samples 500 max_n_samples not set

Priors:

	loc	variance_scale	shape	scale	
Α	0.001	0.001	0.001	0.001	
В	0.001	0.001	0.001	0.001	
C	0.001	0.001	0.001	0.001	
D	0.001	0.001	0.001	0.001	

Bayesian experiment: normal distribution (3 / 4)

- ▶ Generate or pass new data and update models until a clear winner is found.
- ▶ The stopping rule will be updated after a new update.

```
with experiment as e:
    while not e.termination:
        data_A = stats.norm(loc=8, scale=3).rvs(size=10)
        data_B = stats.norm(loc=7, scale=2).rvs(size=25)
        data_C = stats.norm(loc=7.5, scale=4).rvs(size=12)
        data_D = stats.norm(loc=6.75, scale=2).rvs(size=8)

        e.run_update(**{"A": data_A, "B": data_B, "C": data_C, "D": data_D})

print(e.termination, e.status)
```

True winner A

Bayesian experiment: normal distribution (4 / 4)

- ▶ Reporting: visualize stopping rule metric over time (updates).
- ▶ Reporting: visualize statistics over time (updates).

>>> experiment.plot_metric()

probability_vs_all

0.8

0.6

0.4

0.4

0.0

0.10

0.20

30

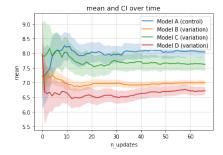
40

50

60

n updates

>>> experiment.plot_stats()



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Thank you!

https://github.com/guillermo-navas-palencia/cprior