Analytic Estimation

Modern Theory of Detection and Estimation. Block-1: Estimation

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- Introduction to estimation Problems
 - Definitions and notation
 - Different estimation scenarios
- Elements of an analytical estimation problem
 - Statistical relationship between observations and estimated variables
 - Cost Functions
 - Expected Cost
- Bayesian estimation
 - Examples of Bayesian Estimators
 - Minimum absolute deviation estimation
 - Maximum a posteriori estimation
- 4 Estimation of Maximum Likelihood

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Estimation problem definition

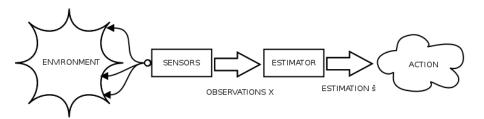
Goal in a estimation problem

You receive a scalar observation (x) or a vector observation (x), and your task is to find out (estimate) the value of a real number s. The variable you want to find out can be either deterministic or random.

Examples:

- Estimate the number of stocks of a certain instrument that will be traded tomorrow.
- Estimate the daily energy consumption in a building as a function of the weather, season of year, day of week.
- Estimate the price of a house given the surface, number of rooms, size of the yard, distance to the centre of the town, etc.
- Guess the foot length of people based on their height.

Estimation problem setup



- Sensors sample information about the environment and construct observations \mathbf{x}^k , k = 1, ..., l.
- Estimator constructs estimated variable \hat{s} as a function of the observations. There **must be** a statistical relationship between the observations and the variable we want to estimate, s.
- Then we use the estimation \hat{s} instead of s to take actions. Remember we **never** have access to s (that's why we need to estimate it!!)

Notation

- X: observations. Random in the analytical case setup. Could be scalars or vectors.
- S: desired output. This is the variable we want to estimate. It can be random (S) or deterministic (s). It can be a vector or an scalar, most of the time this course it will be an scalar.
- \hat{S} : **estimator**. $\hat{S} = f(\mathbf{X})$ is our approximation to the value of S given the observation. In the analytic estimation case \hat{S} is random since it is a deterministic function (f()) of a random variable (\mathbf{X}) , therefore its pdf can be obtained applying random variable transformation to the pdf of the observations.
- $p_{X,S}(\mathbf{x}, s)$: statistical relationship between observations and estimated variable. We can estimate S using \mathbf{X} as input because they are related. We express this relationship as a joint pdf.

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Random variable vs. Deterministic parameter

The variable we want to estimate can be either random S, or deterministic s. Examples:

- Deterministic: I've thought the mean of a Gaussian distribution m. I draw K samples from the distribution and ask you to estimate m. Note that m is unknown but does not follow any prior distribution p(m); m is a fixed [unknown] constant.
- Random: I pick at random one student from the class, I tell you his/her foot length (X = x) and you have to guess his/her height (S = s). S is random because I can pick any student in the class.

Analytic estimation vs. machine estimation

• Machine estimation or regression: the problem is defined in terms of a set of examples, $\{\mathbf{x}^k, s^k\}_{k=1}^l$. We use the examples to fit an estimation model, for example a linear regressor:

$$\hat{s} = f(\mathbf{x}) = \mathbf{w}^T \mathbf{x} + w_0$$

we learn **w** and w_0 with the training set $\{\mathbf{x}^k, s^k\}_{k=1}^l$. Notice in the machine estimation setting the training observations $\{\mathbf{x}^k, s^k\}_{k=1}^l$ are not random (we do know them). The test observations can be considered random.

- Analytic estimation: the problem is defined in terms of a complete statistical characterisation: there exists a joint pdf $p_{\mathbf{X},S}(\mathbf{x},s)$ and we know how to calculate it for any pair (\mathbf{x}, s) . In this setting **X** is random.
- Semi-analytic estimation: The problem is defined in terms of a data set, as in the regression case. We do not use this data to fit a regression model but to learn the statistics, i.e., estimate $p_{\mathbf{X},S}(\mathbf{x},s)$. Once we have an estimation of the joint pdf (or any other useful pdf) we can apply analytic estimation.

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Likelihood

We rely on different pdfs that capture the relationship between \mathbf{x} and s.

Likelihood

$$p_{\mathbf{X}|S}(\mathbf{x}|s)$$

It's the key pdf. You will have a likelihood in both random and deterministic variable estimation cases.

The likelihood models the generation of observations for every possible value of the target variable S. If S changes, we get different statistics in the observations. That is why we can guess the value of S.

If s is deterministic there is no point in conditioning in the value of s since it is always the same. Anyway we use $p_{\mathbf{X}|s}(\mathbf{x}|s)$ to unify notations.

S is random

If S is a random variable we have three other pdfs:

- **Prior** or marginal of $S: p_S(s)$ gives information about how are the values of S distributed without access to the observations.
- Joint distribution of S and X: $p_{X,S}(\mathbf{x},s)$
- **Posterior** of S given the observation $\mathbf{X} = \mathbf{x}$: $p_{S|\mathbf{X}}(s|\mathbf{x})$ gives the values of S that concentrate a higher probability density for each particular observed value of \mathbf{X} . It is a key pdf in the design of estimators since to know $\mathbf{X} = \mathbf{x}$ narrows the uncertainty about S expressed in the prior (I mean $p_{S|\mathbf{X}}(s|\mathbf{x})$ as a function of s is "narrower" than $P_S(s)$).

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Cost Functions

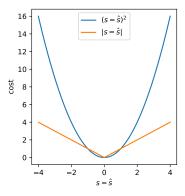
- An estimator is a function of the observations: $\hat{s} = f(\mathbf{x})$.
- The of design an estimator involves an optimisation to pick the best estimation function $f^*()$ from within a family or set of candidates.
- We need a criterion to select this best estimation function from the set of potential candidates. This criterion is the **cost function**.

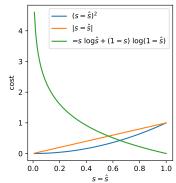
Cost function

Function of the estimator and of the estimated variable: $C(S, \hat{S})$. It compares both quantities and determines the penalty in which we incur if we approximate S=s with $\hat{S}=\hat{s}$. In most cases exact estimations $(\hat{s}=s)$ yield zero cost: $C(s, \hat{s}=s)=0$. The cost is always positive $C(S, \hat{S})\geq 0$.

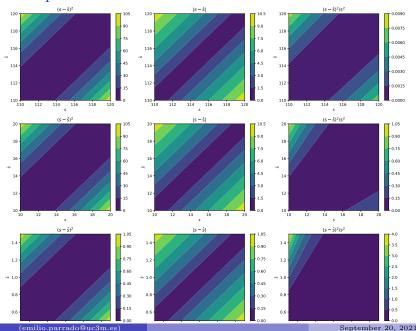
Examples of cost function

- Quadratic Cost: $c(e) = e^2$.
- Absolute Value of the error: c(e) = |e|.
- Relative Square Error: $c(s,\hat{s}) = \frac{(s-\hat{s})^2}{s^2}$
- Log loss: $c(s, \hat{s}) = -s \ln \hat{s} (1 s) \ln (1 \hat{s})$, for $s \in \{0, 1\}$, and $\hat{s} \in [0, 1]$





Examples of cost function



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Expected Cost

Since we plan to use the estimator a large number of times, we are more interested in the expected cost, that takes into account all the possible values of S and \hat{S} .

Expected cost

$$\mathbb{E}\{c(S, \hat{S})\} = \int_{\mathbf{X}} \int_{s} c(s, \hat{s}(\mathbf{x})) \mathbf{p}_{\mathbf{S}, \mathbf{X}}(\mathbf{s}, \mathbf{x}) \mathbf{ds} \mathbf{dx}$$

Notice that $\hat{S} = \hat{S}(\mathbf{X})$ is a deterministic function of \mathbf{X} , therefore the statistics of \hat{S} can be univocally defined in terms of the statistics of \mathbf{X} (transformation of random variable).

Example

Let X be a noisy observation of S, such that

$$X = S + R$$

with S a random variable with mean 0 and variance 1, and R a Gaussian random variable, independent of S, with mean 0 and variance v. Consider the estimator $\hat{S} = X$, the expected quadratic cost is

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The expected absolute cost is

$$\mathbb{E}\{|S - \hat{S}|\} = \mathbb{E}\{|R|\} = \int_{-\infty}^{\infty} |r| \frac{1}{\sqrt{2\pi v}} \exp\left(-\frac{r^2}{2v}\right) dr$$
$$= 2 \int_{0}^{\infty} r \frac{1}{\sqrt{2\pi v}} \exp\left(-\frac{r^2}{2v}\right) dr = \sqrt{\frac{2v}{\pi}}$$

Example: Calculation of Expected Quadratic Cost

Two random variables S and X follow a joint pdf

$$p_{S,X}(s,x) = \begin{cases} \frac{1}{x}, & 0 < s < x, \quad 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Consider two estimators $\hat{S}_1 = \frac{1}{2}X$ and $\hat{S}_2 = X$. Which is the best estimator from the point of view of minimising the quadratic cost?

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Consider two estimators $\hat{S}_1 = \frac{1}{2}X$ and $\hat{S}_2 = X$. Which is the best estimator from the point of view of minimising the quadratic cost? To find it out, we'll compute the mean square error of both estimators. For a general w,

$$\mathbb{E}\{(S - wX)^2\} = \int_0^1 \int_0^x (s - wx)^2 p_{S,X}(s, x) ds dx$$

$$= \int_0^1 \int_0^x (s - wx)^2 \frac{1}{x} ds dx$$

$$= \int_0^1 \left(\frac{1}{3} - w + w^2\right) x^2 dx$$

$$= \frac{1}{3} \left(\frac{1}{3} - w + w^2\right)$$

Example: Calculation of Expected Quadratic Cost (ctd)

If w = 1/2

$$\mathbb{E}\{(S-\hat{S}_1)^2\} = \mathbb{E}\{(S-\frac{1}{2}X)^2\} = \frac{1}{3}\left(\frac{1}{3} - \frac{1}{2} + \frac{1}{4}\right) = \frac{1}{36}$$
 (1)

Alternatively, if w = 1

$$\mathbb{E}\{(S-\hat{S}_2)^2\} = \mathbb{E}\{(S-X)^2\} = \frac{1}{3}\left(\frac{1}{3}-1+1\right) = \frac{1}{9}$$
 (2)

Therefore, from the point of view of the mean square error, \hat{S}_1 is better estimator than \hat{S}_2 .

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Bayesian estimation

Given a cost function and a statistical relationship (joint distribution) between observations and estimated variable, the Bayesian Estimator is the one that minimizes the expected cost.

$$\begin{split} \hat{S}_{\text{Bayes}} &= \arg\min_{\hat{S}} \mathbb{E}\{c(S, \hat{S})\} = \arg\min_{\hat{S}} \int_{\mathbf{x}} \int_{s} c(s, \hat{s}) p_{S, \mathbf{X}}(s, \mathbf{x}) ds d\mathbf{x} \\ &= \arg\min_{\hat{S}} \int_{\mathbf{x}} \int_{s} c(s, \hat{s}) p_{S|\mathbf{X}}(s|\mathbf{x}) ds \; p_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} = \\ &\arg\min_{\hat{S}} \int_{\mathbf{x}} \mathbb{E}\{c(S, \hat{S}) | \mathbf{X}\} p_{\mathbf{X}}(\mathbf{x}) d\mathbf{x} \end{split}$$

Bayesian estimation

If we look carefully into the last integral, we see that $\mathbb{E}\{c(S,\hat{S})|\mathbf{X}\}$ is multiplied by a positive function $(p_{\mathbf{X}}(\mathbf{x}))$. Therefore, to minimize the expected global cost is to minimise the integral and it is achieved minimising $\mathbb{E}\{c(S,\hat{S})|\mathbf{X}\}$. Thus a first result is

$$\hat{S}_{\mbox{\footnotesize Bayes}} = \arg\min_{\hat{S}} \mathbb{E}\{c(S,\hat{S}) | \mathbf{X}\}$$

In summary, to compute a Bayesian Estimator we need to take two steps:

- Select a cost function $c(S, \hat{S})$.

Two random variables S and X follow a joint pdf

$$p_{S,X}(s,x) = \begin{cases} \frac{1}{x}, & 0 < s < x, \quad 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Find the Bayesian estimator that minimizes the quadratic cost

Two random variables S and X follow a joint pdf

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Find the Bayesian estimator that minimizes the quadratic cost

First, find the posterior

$$p_{S|X}(s|x) = \frac{p_{S,X}(s,x)}{p_X(x)}$$

$$p_{S|X}(s|x) = \begin{cases} \frac{1}{x}, & 0 < s < x, \quad 0 < x < 1\\ 0, & \text{otherwise} \end{cases}$$

Where you should apply that

$$p_X(x) = \int_{(s)} p_{S,X}(s,x)ds = \int_0^x \frac{1}{x}ds = 1$$



The expected cost given x:

The expected cost given x:

$$\mathbb{E}\{c(S,\hat{S})|\mathbf{X} = \mathbf{x}\} = \mathbb{E}\{(S-\widehat{s})^2)|\mathbf{X} = \mathbf{x}\}$$

$$= \int_{(s)} (s-\widehat{s})^2 p_{S|X}(s|x) ds$$

$$= \frac{1}{x} \int_0^x (s-\widehat{s})^2 ds$$

$$= \frac{1}{x} \left(\frac{(x-\widehat{s})^3}{3} + \frac{\widehat{s}^3}{3}\right)$$

$$= \frac{1}{3} x^2 - \widehat{s}x + \widehat{s}^2$$

The expected cost given x, as a function of \hat{s} , is a 2nd degree polynomial. We can minimize it by equating its derivative to 0.

The expected cost given x:

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The expected cost given x, as a function of \hat{s} , is a 2nd degree polynomial. We can minimize it by equating its derivative to 0.

$$\frac{d}{d\hat{s}}\mathbb{E}\{c(S,\hat{S})|\mathbf{X}=\mathbf{x}\} = -x + 2\hat{s} = 0$$

And the solution is:

$$\widehat{s}_{\text{Bayes}} = \frac{1}{2}x$$

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Minimum Mean Square Error (MMSE) Estimator

This is the Bayesian Estimator when the cost function is the quadratic error: $c(s,\hat{s}) = e^2 = (s-\hat{s})^2$. Therefore the estimator is the solution to the optimisation problem given by

$$\hat{s}_{\text{MMSE}} = \arg\min \hat{s} \mathbb{E}\{c(S, \hat{s}) | \mathbf{X} = \mathbf{x}\} = \arg\min \hat{s} \int_{s} (s - \hat{s})^{2} p_{S|\mathbf{X}}(s|\mathbf{x}) ds$$

Minimum Mean Square Error (MMSE) Estimator

To minimise, we take the derivative with respect to \hat{s} and make it equal to zero.

$$\frac{\mathrm{d}}{\mathrm{d}\hat{s}} \int_{s} (s-\hat{s})^{2} p_{S|\mathbf{X}}(s|\mathbf{x}) ds = \int_{s} \frac{\partial}{\partial \hat{s}} (s-\hat{s})^{2} p_{S|\mathbf{X}}(s|\mathbf{x}) ds$$
$$= \int_{-\infty}^{\infty} -2(s-\hat{s}) p_{S|\mathbf{X}}(s|\mathbf{x}) ds$$

In the optimum the above integral is equal to zero:

$$\int_{-\infty}^{\infty} -2(s-\hat{s}_{\text{\tiny Bayes}})p_{S|\mathbf{X}}(s|\mathbf{x})ds = 0 \Rightarrow$$

$$\hat{s}_{\text{\tiny Bayes}} = \mathbb{E}\{s|\mathbf{X}\}$$

The intuition behind this result is that the knowledge of \mathbf{X} determines as optimum estimator the one corresponding to choosing the expected value of S for that particular observed value of \mathbf{X} .

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Expected absolute cost

The minimum absolute error estimator corresponds to the cost function $c(e) = |e| = |s - \hat{s}|$. Therefore:

$$\begin{split} \widehat{s}_{\mathrm{MAD}} &= \arg\min_{\widehat{\widehat{s}}} \mathbb{E}\{c(S, \widehat{s}) | \mathbf{X} = \mathbf{x}\} \\ &= \arg\min_{\widehat{\widehat{s}}} \int_{s} |s - \widehat{s}| p_{S|\mathbf{X}}(s|\mathbf{x}) ds \end{split}$$

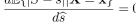
$$\begin{split} \mathbb{E}\{|S-\widehat{s})|\mathbf{X} &= \mathbf{x}\} = \int_{-\infty}^{\widehat{s}} (\widehat{s}-s) p_{S|\mathbf{X}}(s|\mathbf{x}) ds + \int_{\widehat{s}}^{\infty} (s-\widehat{s}) p_{S|\mathbf{X}}(s|\mathbf{x}) ds \\ &= \widehat{s} \Bigg[\int_{-\infty}^{\widehat{s}} p_{S|\mathbf{X}}(s|\mathbf{x}) ds - \int_{\widehat{s}}^{\infty} p_{S|\mathbf{X}}(s|\mathbf{x}) ds \Bigg] + \\ &+ \int_{\widehat{s}}^{\infty} s \, p_{S|\mathbf{X}}(s|\mathbf{x}) ds - \int_{-\infty}^{\widehat{s}} s \, p_{S|\mathbf{X}}(s|\mathbf{x}) ds \end{split}$$

Minimization of the average cost a posteriori

The Fundamental Theorem of Calculus enables to get the derivative of the average cost a posteriori as:

$$\begin{split} \frac{d\mathbb{E}\{|S-\widehat{s}||\mathbf{X}=\mathbf{x}\}}{d\widehat{s}} &= \\ \frac{d}{d\widehat{s}} \left(\widehat{s} \left[\int_{-\infty}^{\widehat{s}} p_{S|\mathbf{X}}(s|\mathbf{x}) ds - \int_{\widehat{s}}^{\infty} p_{S|\mathbf{X}}(s|\mathbf{x}) ds \right] + \\ \int_{\widehat{s}}^{\infty} s \, p_{S|\mathbf{X}}(s|\mathbf{x}) ds - \int_{-\infty}^{\widehat{s}} s \, p_{S|\mathbf{X}}(s|\mathbf{x}) ds \right) \\ \frac{d\mathbb{E}\{|S-\widehat{s}||\mathbf{X}=\mathbf{x}\}}{d\widehat{s}} &= 0 \end{split}$$

and then make



MAD estimator

MAD estimator

$$\hat{s}_{\text{MAD}}(\mathbf{x}) = \text{median}\{S|\mathbf{X} = \mathbf{x}\}\$$

Note: It is usually computed with the following expression:

$$\int_{-\infty}^{\widehat{s}_{\mathrm{MAD}}} p_{S|\mathbf{X}}(s|\mathbf{x}) ds = \int_{\widehat{s}_{\mathrm{MAD}}}^{\infty} p_{S|\mathbf{X}}(s|\mathbf{x}) ds = \frac{1}{2}$$

Example: Find Bayesian estimation with Absolute Error Cost

Two random variables S and X follow a joint pdf

$$p_{S,X}(s,x) = \begin{cases} \frac{1}{x}, & 0 < s < x, \quad 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Find the Bayesian estimator that minimizes the absolute error cost From previous example retrieve the posterior

$$p_{S|X}(s|x) = \begin{cases} \frac{1}{x}, & 0 < s < x, \quad 0 < x < 1 \\ 0, & \text{otherwise} \end{cases}$$

Example: Find Bayesian estimation with Absolute Error Cost

$$\begin{split} \widehat{s}_{\mathrm{MAD}} &\Rightarrow \int_{0}^{\widehat{s}_{\mathrm{MAD}}} p_{S|\mathbf{X}}(s|\mathbf{x}) ds = \frac{1}{2} = \int_{\widehat{s}_{\mathrm{MAD}}}^{x} p_{S|\mathbf{X}}(s|\mathbf{x}) ds \\ &\int_{0}^{\widehat{s}_{\mathrm{MAD}}} \frac{1}{x} ds = \frac{1}{2} \quad \Rightarrow \quad \widehat{s}_{\mathrm{MAD}}(x) = \frac{x}{2} \end{split}$$

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MAP estimator

MAP Estimator

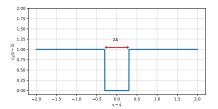
$$\widehat{s}_{\text{MAP}}(\mathbf{x}) = \arg \max_{s} p_{S|\mathbf{X}}(s|\mathbf{x})$$
$$= \arg \max_{s} \ln \left[p_{S|\mathbf{X}}(s|\mathbf{x}) \right]$$

When the distribution a posteiori presents several global maxima, the MAP estimator is not unique!

MAP estimator from a Bayesian point of view

Strictly speaking, the MAP estimator is not Bayesian because it is not minimizing any expected cost. However, if we consider the following cost function:

$$c_{\Delta}(s-\widehat{s}) = \begin{cases} 1, & \text{if } |s-\widehat{s}| > \Delta \\ 0, & \text{if } |s-\widehat{s}| < \Delta \end{cases}$$



and \hat{s}_{Δ} as the bayesian estimator corresponding to c_{Δ} :

$$\widehat{s}_{\text{MAP}} = \lim_{\Delta \to 0} \widehat{s}_{\Delta}$$

Therefore the MAP estimator can be considered a limit case in a family of bayesian estimators.

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Maximum Likelihood (ML) Estimator

ML estimator

$$\hat{s}_{ML} = \arg\max_{s} p_{\mathbf{x}|s}(\mathbf{x}|s) = \arg\max_{s} \ln(p_{\mathbf{x}|s}(\mathbf{x}|s))$$

- No associated cost function, therefore in general the ML is not a Bayesian estimator. It could happen that the \hat{s}_{ML} coincides with a Bayesian estimator, but by definition the \hat{s}_{ML} does not optimise a cost function
- \hat{s}_{ML} does not take into account the prior distribution $p_S(s)$. Therefore it can be used when s is a **deterministic parameter**.

Relationship between the MAP estimator and the ML estimator

$$\hat{s}_{MAP} = \arg\max_{s} p_{S|\mathbf{X}}(s|\mathbf{x}) = \arg\max_{s} \frac{p_{\mathbf{X}|S}(\mathbf{x}|s)p_{S}(s)}{p_{\mathbf{X}}(\mathbf{x})}$$

If we don't know $p_S(s)$, we can assume it is **uniform** (no preference for any particular value of s). Therefore

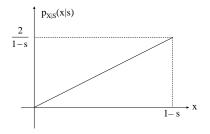
$$\arg\max_{s} \frac{p_{\mathbf{X}|S}(\mathbf{x}|s)p_{S}(s)}{p_{\mathbf{X}}(\mathbf{x})} \sim \arg\max_{s} p_{\mathbf{X}|S}(\mathbf{x}|s) = \hat{s}_{ML}$$

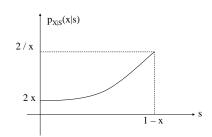
Example: ML estimation of a random variable

One wishes to estimate the value of a random variable S from the observed value of another variable X, statistically related to S through:

$$p_{X|S}(x|s) = \frac{2x}{(1-s)^2}, \ 0 < x < 1-s, \ 0 < s < 1$$

We need to maximise $p_{X|S}(x|s)$ with respect to s, (not to x)





 $p_{X|S}(x|s)$ as a function of x

 $p_{X|S}(x|s)$ as a function of s

Solution: $\hat{s}_{ML} = 1 - X$

Example: Estimation of the parameters of a Gaussian

We have access to a collection of l data samples $\{X^{(k)}\}_{k=1}^l$ drawn **independently** from a Gaussian pdf. Construct estimators for the mean m and the variance v of the Gaussian.

ullet The likelihood of each observation given m and v is

$$p_X(x) = p_{X|m,v}(x|m,v) = \frac{1}{\sqrt{2\pi v}} \exp\left[-\frac{(x-m)^2}{2v}\right]$$

• Since we observe *l* samples, the **joint** pdf will be the product of the individual pdfs.

$$p_{\{X^{(k)}\}|m,v}(\{x^{(k)}\}|m,v) = \prod_{k=1}^{l} p_{X|m,v}(x^{(k)}|m,v)$$
$$= \frac{1}{(2\pi v)^{l/2}} \prod_{k=1}^{l} \exp\left[-\frac{(x^{(k)}-m)^2}{2v}\right]$$

Example: Estimation of the parameters of a Gaussian

• Taking logs and optimising

$$\hat{m}_{\rm ML} = \frac{1}{l} \sum_{k=1}^{l} x^{(k)}$$

• Using $\hat{m}_{\rm ML}$ in the expression for the estimator of the variance

$$\hat{v}_{\mathrm{ML}} = \frac{1}{l} \sum_{k=1}^{l} (x^{(k)} - \hat{m}_{\mathrm{ML}})^2$$