Model Diagnostics

Stat 230: Applied Regression Analysis

PDF version of slides

RMarkdown demo

Conditions required for inference

Our model must be valid for inference to be valid

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$
 where $\varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$

Conditions to check:

- Linear relationship is appropriate
- Errors are independent and identically distributed (iid)
- Errors are normally distributed
- Variance of the errors doesn't depend on x

Residuals

Definition:
$$e_i = \widehat{\varepsilon}_i = y_i - \widehat{y}_i$$

Properties:

- sum to zero \implies mean is 0
- **uncorrelated** with x and \hat{y}
- normally distributed

•
$$SD(e_i) = \widehat{\sigma} \sqrt{1 - \frac{1}{n} - \frac{(x_i - \overline{x})^2}{\sum (x_i - \overline{x})^2}}$$

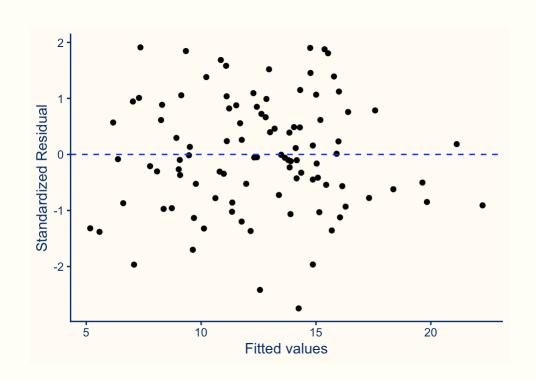
Standardized residuals

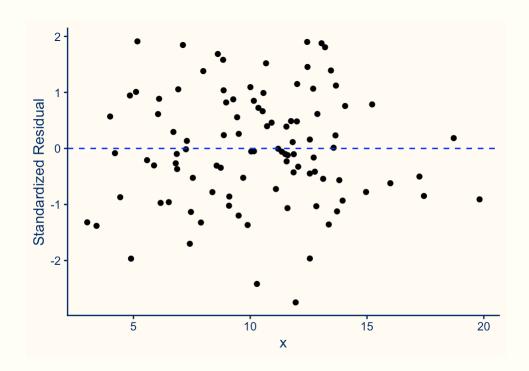
$$r_{i} = \frac{e_{i}}{\widehat{\sigma}\sqrt{1 - \frac{1}{n} - \frac{(x_{i} - \overline{x})^{2}}{\sum(x_{i} - \overline{x})^{2}}}}$$

Properties:

- sum to zero \implies mean is 0
- **uncorrelated** with x and \hat{y}
- normally distributed
- $SD(r_i) = 1$

A "good" residual plot





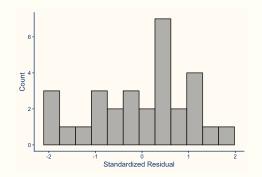
Your turn

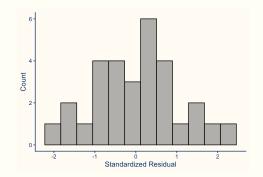
- Work in groups
- On the whiteboards, sketch a plot of *y* vs. *x* and a corresponding residual plot that would indicate a violation of the
 - 1. linearity condition
 - 2. constant variance condition

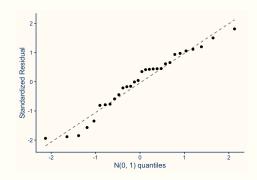
Assessing normality

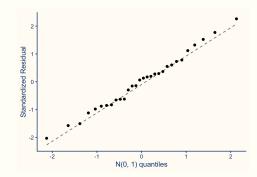
- histogram of residuals
- normal Q-Q plot of residuals

Examples of "good" plots:





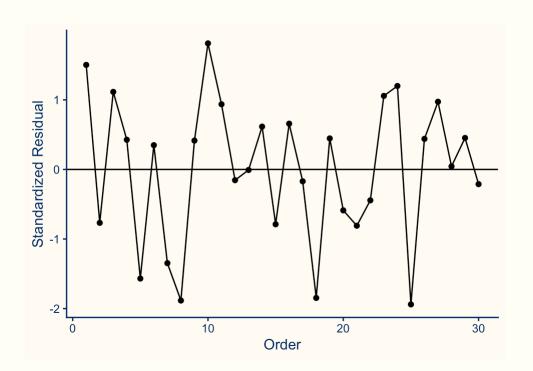


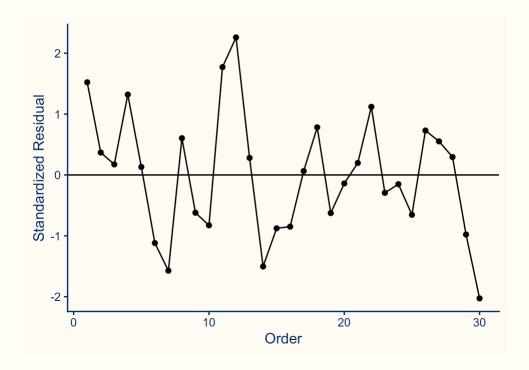


Assessing independence

• plot residuals vs. variable inducing dependence (e.g. time, location, subject ID)

Examples of "good" plots:





Your turn

Work through Example 1 on the worksheet

What happens the conditions aren't valid?

- Linearity: if nonlinear, everything breaks!
- **Independence:** estimates are still unbiased (i.e. we fit the right line) but measures of the accuracy of those estimates (the SEs) are typically too small
- **Normality:** estimates are still unbiased (i.e. we fit the right line), SEs are correct BUT confidence/prediction intervals are wrong (we can't use t-distribution)
- Constant error variance: estimates are still unbiased but standard errors are wrong (and we don't know how wrong)

What do we do if our assumptions are violated?

- 1. Change our assumptions (hard, need more stats)
- 2. Transform *y*, *x*, or both
- 3. Change the type of inference (remember the bootstrap?)

Transforming variables can

- Address non-linear patterns (i.e., linear on transformed scale)
- Stabilize variance
- Correct skew
- Minimize the effects of outliers

Applying transformations

To apply a transformation, we calculate a new variable and use it in place of the original variable in our model

Examples

$$\log(y) = \beta_0 + \beta_1 x + \varepsilon$$
$$y = \beta_0 + \beta_1 \sqrt{x + \varepsilon}$$
$$\log(y) = \beta_0 + \beta_1 \sqrt{x + \varepsilon}$$

Your turn

Work through Example 2 on the worksheet

Review of logarithms

The logarithm $\log_b(x)$ is a function that is the exponent (power) that the base, b, must be raised to produce the value x:

- $\log_{10}(100) = 2$ since $10^2 = 100$
- $\log_{10}(10) = 1$ since $10^1 = 10$
- $\log_2(1) = 0$ since $2^0 = 1$
- $\log_2(0.5) = -1 \text{ since } 2^{-1} = \frac{1}{2}$

Review of logarithms

- Takes in only positive numbers, i.e. x > 0
- The log of products is the sum of the logs

$$\log_b(mx) = \log_b(m) + \log_b(x)$$

The log of quotients is the difference of the logs

$$\log_b\left(\frac{m}{x}\right) = \log_b(m) - \log_b(x)$$

The log of powers is the exponent times the log

$$\log_b(x^p) = p \log_b(x)$$