

Bayesian regression

Stat 340: Bayesian Statistics

Example

- Partial census data for the Dobe area !Kung San, a foraging population
- Compiled from Nancy Howell's interviews

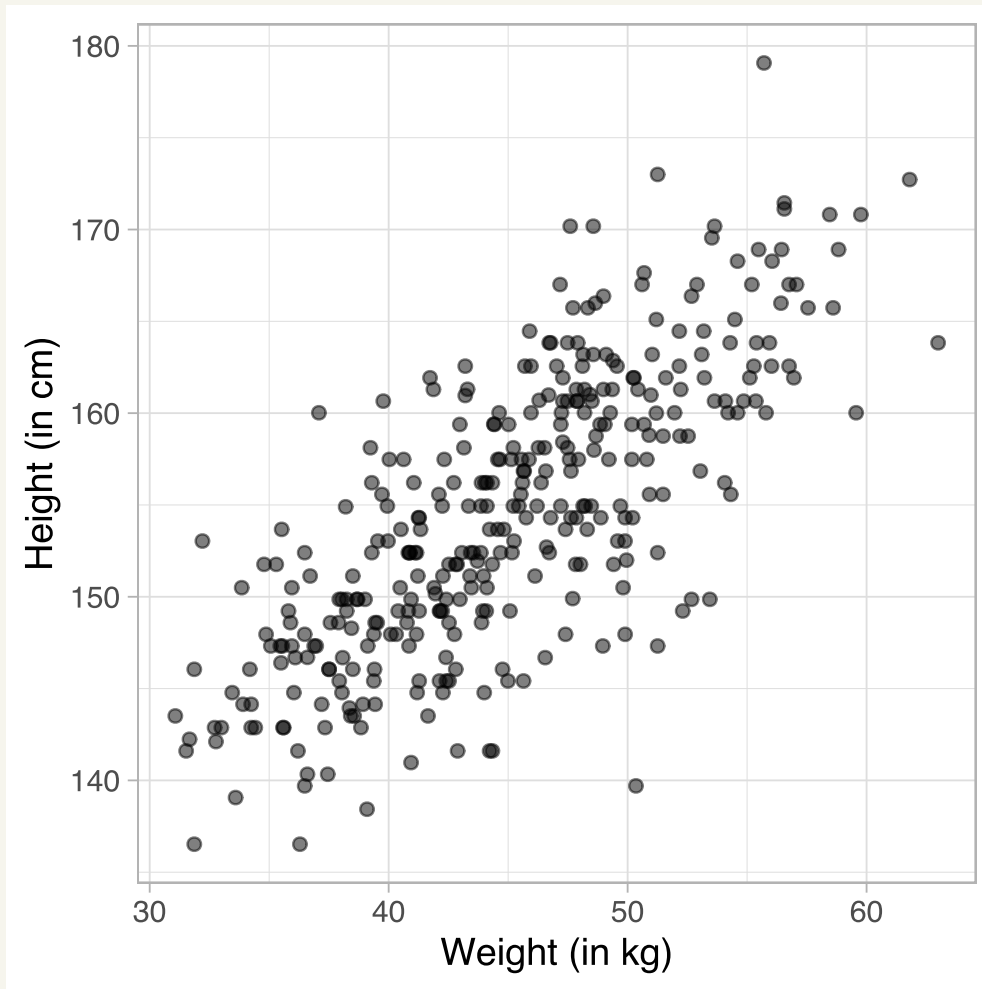
| ## | | mean | sd | 5.5% | 94.5% | histogra |
|----|--------|-----------|------------|----------|-----------|----------|
| ## | height | 154.59709 | 7.7423321 | 142.8750 | 167.00500 | |
| ## | weight | 44.99049 | 6.4567081 | 35.1375 | 55.76588 | |
| ## | age | 41.13849 | 15.9678551 | 20.0000 | 70.00000 | |
| ## | male | 0.46875 | 0.4997328 | 0.0000 | 1.00000 | |



Life Histories of the **DOBE !KUNG**

FOOD, FATNESS, AND WELL-BEING OVER THE LIFE-SPAN

NANCY HOWELL



How can we write a **Bayesian** model to relate weight and height of the Kalahari foragers?

Observation-specific mean

We can adapt our normal model for the mean to use an observation-specific mean, μ_i :

Sampling model: $Y_i | \mu_i, \sigma \stackrel{\text{ind}}{\sim} \mathcal{N}(\mu_i, \sigma)$

Now we need to link μ_i and x_i

A weakly informative prior

We may have limited prior information about the regression coefficients, β_0 and β_1 , and/or the standard deviation σ

Assume independence: $\pi(\beta_0, \beta_1, \sigma) = \pi(\beta_0, \beta_1)\pi(\sigma)$

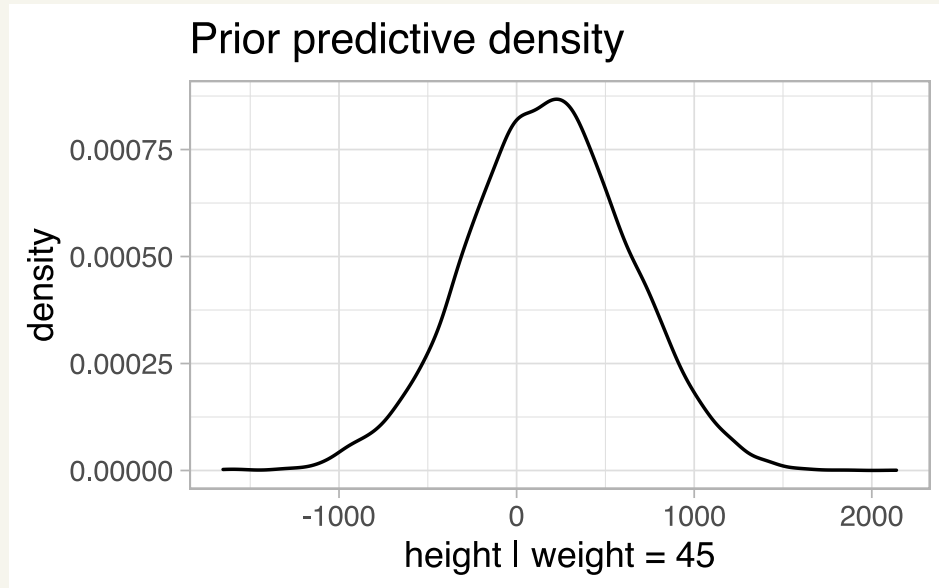
1. Assume β_0 and β_1 are independent

- $\pi(\beta_0, \beta_1) = \pi(\beta_0)\pi(\beta_1)$
- Assign a weakly informative prior each coefficient: $\beta_i \sim \mathcal{N}(m_i, s_i)$
- example: $\mathcal{N}(0, 100)$

2. Assign a weakly informative prior to σ

- example: $1\sigma^2 \sim \text{Gamma}(1, 1)$

Prior predictive as "sanity check"



Simulate the prior predictive:

1. Draw parameters from their prior distributions
2. Draw data from the sampling model plugging in these simulated parameters

Sampling from the prior predictive distribution

```
nsim <- 1e4                                     # no. of simulations

prior.sims <- data_frame(                       # simulate from ind. priors
  beta0 = rnorm(nsim, 178, 100),
  beta1 = rnorm(nsim, 0, 10),
  sigma = runif(nsim, 0, 50)
)
```

```
weight.value <- 45                             # condition on value of x
```

```
prior.pred <- prior.sims %>%
  mutate(
    mu = beta0 + beta1 * weight.value,          # calculate  $E(y|x)$ 
    height = rnorm(n(), mean = mu, sd = sigma) # draw sample from normal
  )
```

Deriving the posterior

Sampling model: $Y_i|x_i, \beta_0, \beta_1, \sigma \sim \mathcal{N}(\beta_0 + \beta_1 x_i, \sigma)$

Joint likelihood:

$$\begin{aligned} L(\beta_0, \beta_1, \sigma) &= \prod_{i=1}^n \left[\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{ -\frac{1}{2\sigma^2} (y_i - \beta_0 - \beta_1 x_i)^2 \right\} \right] \\ &\propto \left(\frac{1}{\sigma^2} \right)^{n/2} \exp\left\{ -\frac{1/\sigma^2}{2} \sum_{i=1}^n (y_i - \beta_0 - \beta_1 x_i)^2 \right\} \end{aligned}$$

Joint posterior: $\pi(\beta_0, \beta_1, \sigma|\text{data}) \propto \pi(\beta_0, \beta_1, \sigma)L(\beta_0, \beta_1, \sigma)$

JAGS for Bayesian SLR

Write down the model string

```
slr_model <- "model {  
  ## sampling model  
  for (i in 1:N){  
    y[i] ~ dnorm(beta0 + beta1 * x[i], invsigma2)  
  }  
  
  ## priors  
  beta0 ~ dnorm(mu0, g0)  
  beta1 ~ dnorm(mu1, g1)  
  invsigma2 ~ dgamma(a, b)  
  sigma <- sqrt(pow(invsigma2, -1))  
}"
```

JAGS for Bayesian SLR

Define the data and set prior parameters

```
the_data <- list(  
  y = adults$height,      # response variable  
  x = adults$weight,      # explanatory variable  
  N = nrow(adults),       # sample size  
  mu0 = 0,                # prior mean for beta0  
  g0 = 0.0001,            # prior precision for beta0  
  mu1 = 0,                # prior mean for beta1  
  g1 = 0.0001,            # prior precision for beta1  
  a = 1,                  # prior shape for 1/sigma2  
  b = 1,                  # prior scale for 1/sigma2  
)
```

JAGS for Bayesian SLR

Generate samples from the posterior

```
posterior <- run.jags(  
  slr_model,  
  data = the_data,  
  n.chains = 1,  
  monitor = c("beta0", "beta1", "sigma"),  
  adapt = 1000,  
  burnin = 5000,  
  sample = 5000,  
  silent.jags = TRUE  
)
```

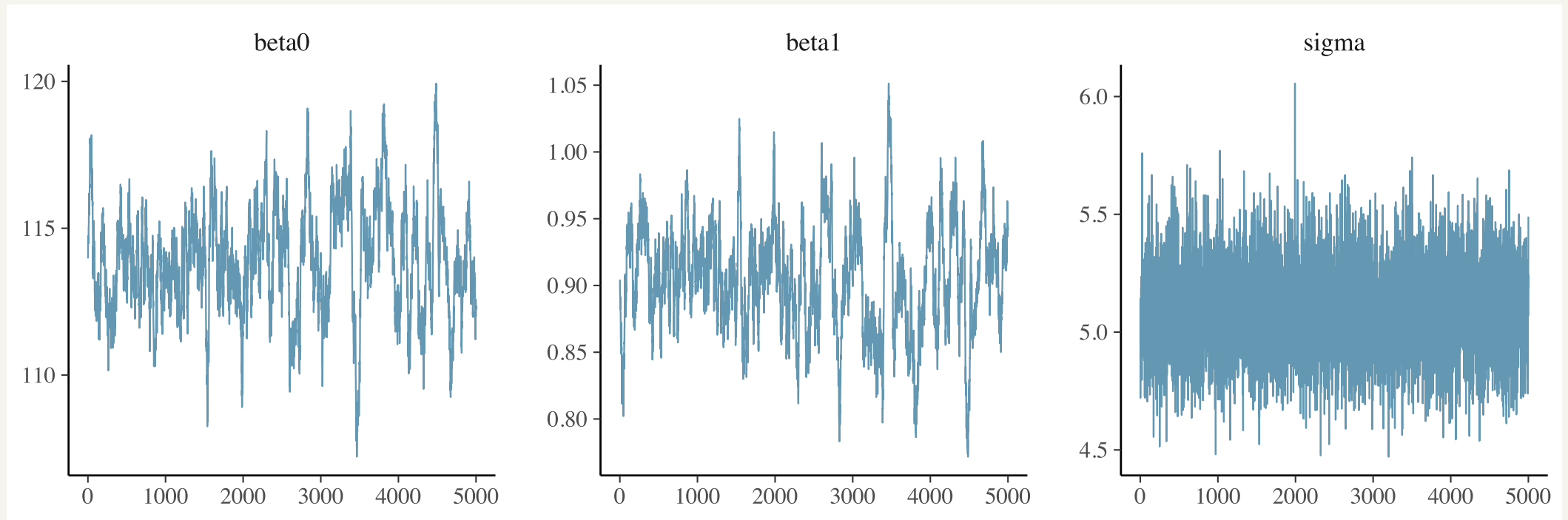
Summary of the fitted model

```
summary(posterior)
```

| ## | | Lower95 | Median | Upper95 | Mean | SD | Mode |
|----|-------|-------------|-------------|-------------|--------------|------------|-------------|
| ## | beta0 | 110.185429 | 113.8831812 | 117.5070605 | 113.8926504 | 1.85060652 | 113.6446508 |
| ## | beta1 | 0.823308 | 0.9048906 | 0.9842071 | 0.9046513 | 0.04065837 | 0.9064154 |
| ## | sigma | 4.692724 | 5.0789234 | 5.4558836 | 5.0853508 | 0.19425986 | 5.0721822 |
| ## | | MCerr | MC%ofSD | SSeff | AC.10 | psrf | |
| ## | beta0 | 0.246236363 | 13.3 | 56 | 0.791740214 | NA | |
| ## | beta1 | 0.005487650 | 13.5 | 55 | 0.793173622 | NA | |
| ## | sigma | 0.002747249 | 1.4 | 5000 | -0.004470933 | NA | |

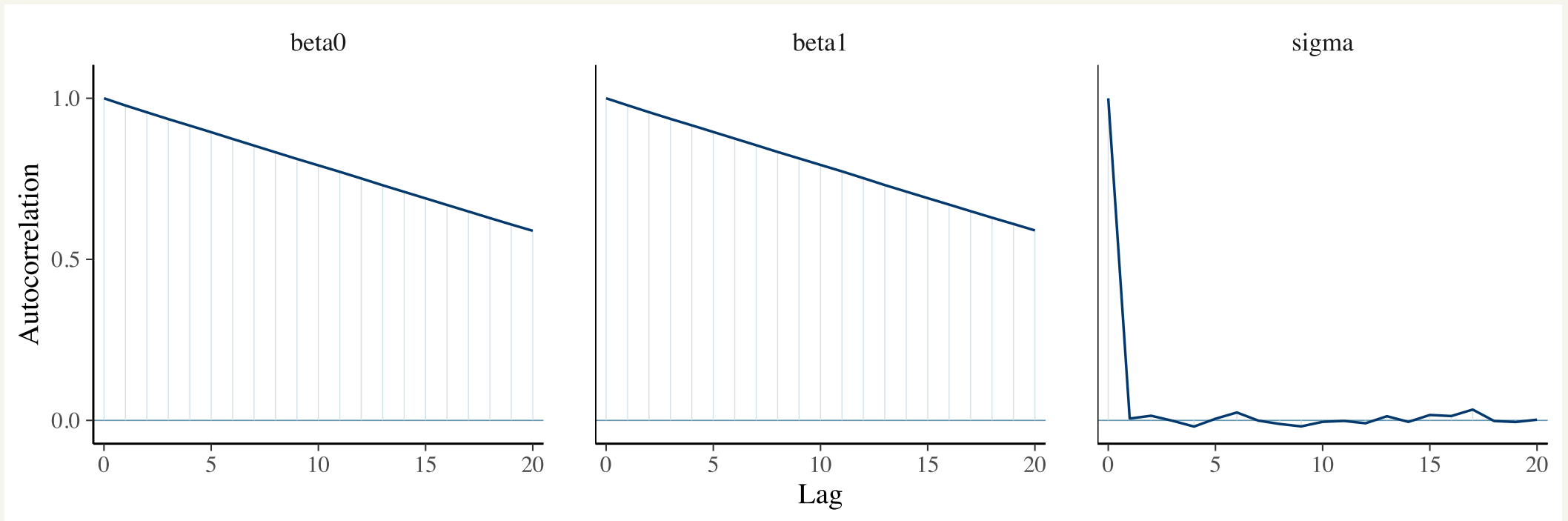
MCMC diagnostics

```
mcmc_trace(posterior$mcmc)
```



MCMC diagnostics

```
mcmc_acf(posterior$mcmc)
```

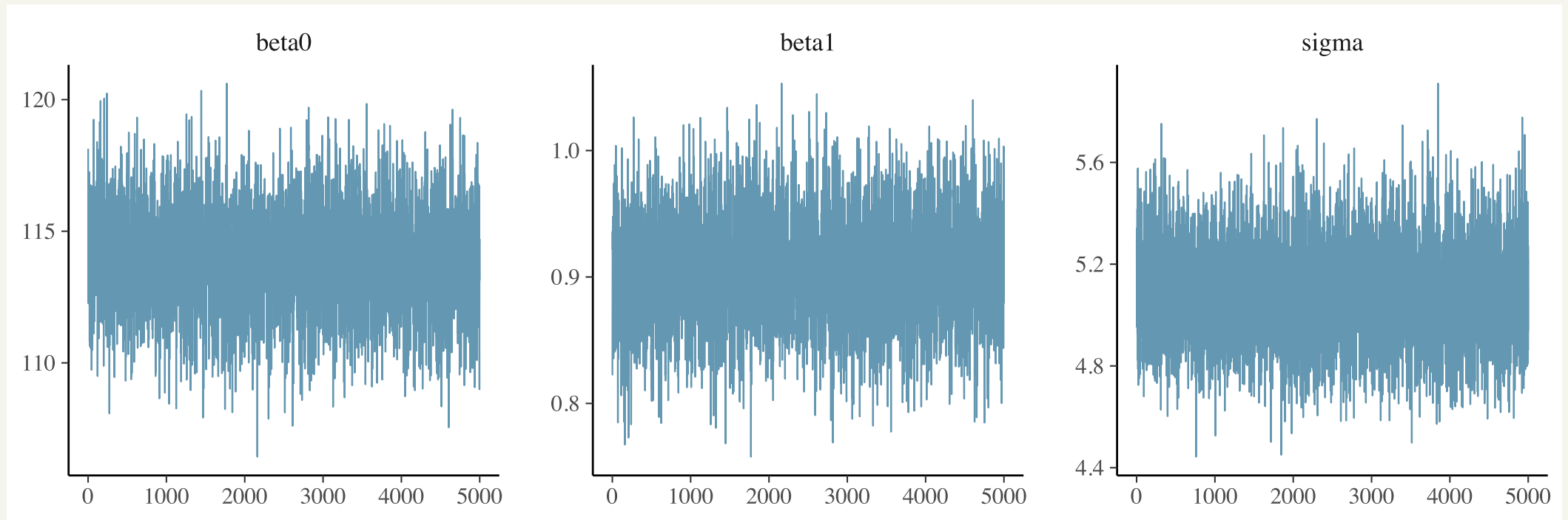


Setting `thin = 50`

```
posterior <- run.jags(  
  slr_model,  
  data = the_data,  
  n.chains = 1,  
  monitor = c("beta0", "beta1", "sigma"),  
  adapt = 1000,  
  burnin = 5000,  
  sample = 5000,  
  thin = 50,  
  silent.jags = TRUE  
)
```

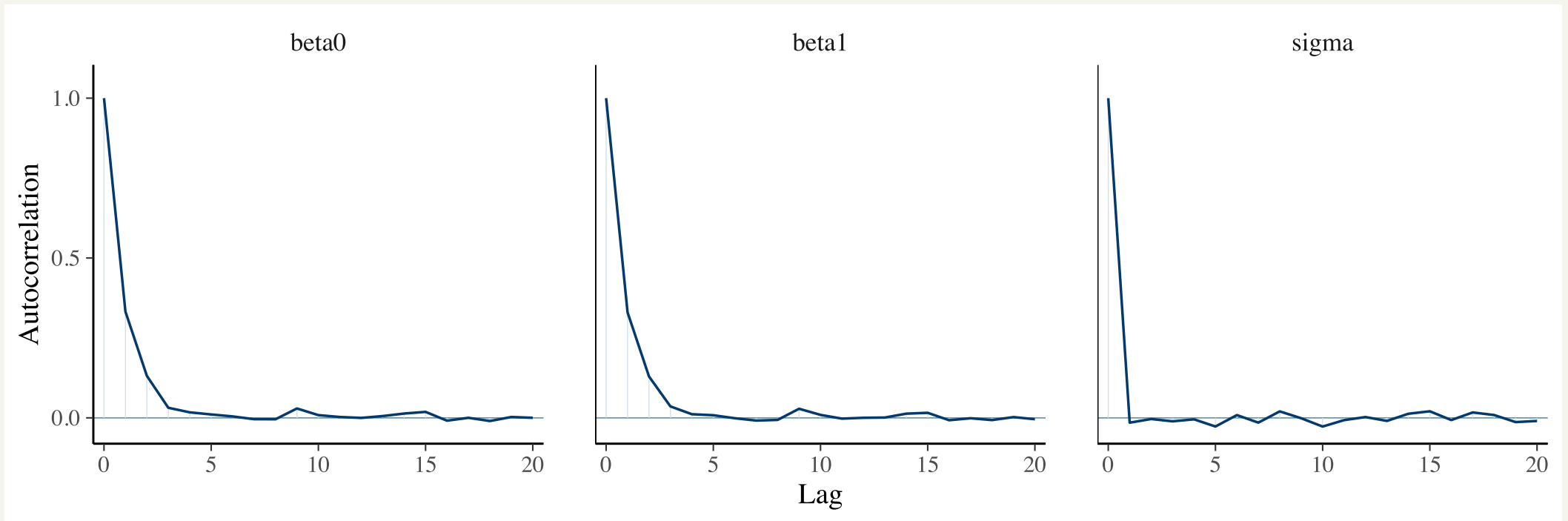
MCMC diagnostics

```
mcmc_trace(posterior$mcmc)
```



MCMC diagnostics

```
mcmc_acf(posterior$mcmc)
```



Summary of the fitted model

```
summary(posterior)
```

```
##              Lower95      Median      Upper95      Mean      SD      Mod
## beta0 110.1454786 113.8402826 117.6245171 113.8453309 1.90365131 113.681470
## beta1  0.8258936  0.9058944  0.9889891  0.9058291 0.04190476  0.908775
## sigma  4.7025891  5.0730595  5.4504958  5.0810728 0.19099318  5.094771
##              MCerr MC%ofSD SSeff      AC.500 psrf
## beta0 0.0381309057      2.0  2492 0.008593566    NA
## beta1 0.0008545736      2.0  2405 0.009331956    NA
## sigma 0.0027010514      1.4  5000 -0.027179755    NA
```

Plotting the fitted model

```
post_means <- apply(  
  posterior$mcmc[[1]], 2, mean  
)
```

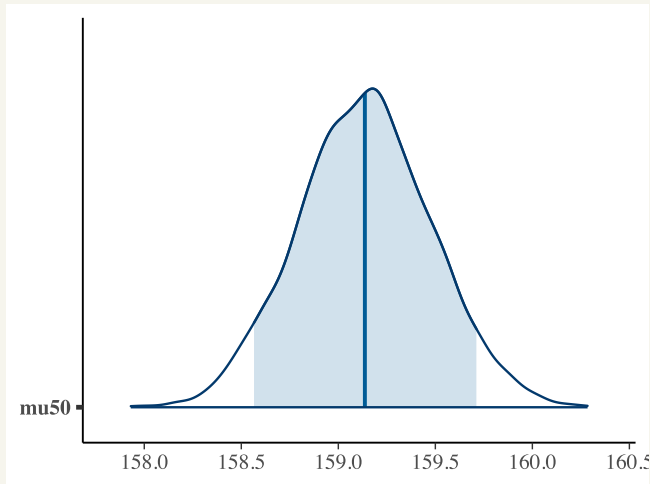
Sampling from the joint posterior

```
post_draws <- as.data.frame(  
  posterior$mcmc[[1]]  
)  
head(post_draws)
```

| ## | | beta0 | beta1 | sigma |
|----|------|----------|-----------|----------|
| ## | 6001 | 118.1336 | 0.8222446 | 5.299407 |
| ## | 6051 | 114.7289 | 0.8955723 | 4.953780 |
| ## | 6101 | 112.5141 | 0.9357981 | 5.153145 |
| ## | 6151 | 113.4607 | 0.9216525 | 5.020973 |
| ## | 6201 | 112.2582 | 0.9389457 | 5.214484 |
| ## | 6251 | 112.8237 | 0.9406860 | 4.993891 |

Generating mean responses

```
mu_at_50 <- with(post_draws,  
                  beta0 + beta1 * 50)
```



```
quantile(mu_at_50, probs = c(0.05, 0.95))  
##           5%           95%  
## 158.5656 159.7114
```