# **KU LEUVEN**

# Fitting mixed random regret minimization models using mixrandregret.

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- Ziyue Zhu (C, in)'
- Álvaro A. Gutiérrez-Vargas (, in)'
- Martina Vandebroek
- Research Centre for Operations Research and Statistics (ORSTAT)

- Random Regret Minimization Models
- 2 Differences between RUM and RRM models.
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- Random Regret Minimization Models Random Utility vs Random Regret Classical Regret Function
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# 1 What is Regret and how to use it for Choice Modeling?

- From Utility to Regret.
- Regret: Situation where a non-chosen alternative ends up being more attractive than the chosen one
- Individuals are assumed to minimize regret.

Table: Hypothetical Choice Situation

Attribute \ Route	1	2	3	
Travel Time	23 min.	27 min.	35 min.	
Travel Cost	6 euros	4 euros	3 euros	

If we chose alternative 2:

- ► Alternative 1 is faster...
- Alternative 3 is cheaper...
- ⇒ RRM models will (formalize and) minimize this notion of regret!

# 1 Formalization of the later example.

Attribute $\setminus$ Route	1	2	3	
Travel Time	23 min.	27 min.	35 min.	
Travel Cost	6 euros	4 euros	3 euros	

Table: Hypothetical Choice Situation

### We will denote in the following:

- lacktriangle Decision-makers (referred to n) answer s choice situations
- ▶ They decide among J alternatives (referred to i or j indistinctly).
- ightharpoonup Each alternative is described by M attributes (referred to m).
- $ightharpoonup x_{ins,m}$  The value of attribute m of alternative i for individual n in choice situation s is denoted by .
- $y_{ins}$  is the response variable that takes the value of 1 when alternative i is chosen by individual n in choice situation s and 0 otherwise.

### 1 RUM vs RRM

► Random Utility Maximization (RUM)
Systematic Utility

► Random Regret Minimization (RRM)

Systematic Regret

$$\begin{array}{c} \text{Regret} \leftarrow RR_{ins} = R_{ins} + \varepsilon_{ins} \\ \\ = \sum_{j \neq i}^{J} R_{i \leftrightarrow jns, T} + \sum_{j \neq i}^{J} R_{i \leftrightarrow jns, T} + \varepsilon_{ins} \end{array}$$

- The notion of regret is characterize by the systematic regret  $R_{ins}$ .
- $R_{in}$  is described in terms of **attribute level regret**  $(R_{i\leftrightarrow jns,m})$ .

# 1 The Attribute level regret $R_{i \leftrightarrow jns,m}$

Attribute \ Route	1	2	3	
Travel Time	23 min.	27 min.	35 min.	
Travel Cost	6 euros	4 euros	3 euros	

 $ightharpoonup R_{i\leftrightarrow j,mn}$  describes the pairwise combinations of regret derived from alternatives.

$(x_{jns,m} - x_{\mathbf{i}ns,m})$	$Attribute \setminus Route$	j = 1	j=2	j = 3
$(x_{jns,m} - x_{1ns,T})$	Travel Time	0	4	12
$(x_{jns,m} - x_{1ns,C})$	Travel Cost	0	-2	-3
$(x_{jns,m} - x_{2ns,T})$	Travel Time	-4	0	8
$(x_{jns,m} - x_{2ns,C})$	Travel Cost	2	0	-1
$(x_{jns,m} - x_{3ns,T})$	Travel Time	-12	-8	0
$(x_{jns,m} - x_{3ns,C})$	Travel Cost	3	1	0

**Takeaway**: We will define  $R_{i \leftrightarrow jns,m}$  in terms of the attribute differences.

# 1 Classical RRM (Chorus, 2010)

► (Chorus, 2010) proposed the following attribute level regret:

$$R_{i \leftrightarrow jns,m} = \ln\left[1 + \exp\left\{\beta_{n,m} \cdot (x_{jns,m} - x_{ins,m})\right\}\right]$$

- $ightharpoonup R_{i \leftrightarrow jns,m}$  compares alternative i with alternative j in attribute m.
- $ightharpoonup \sum_{j \neq i} R_{i \leftrightarrow jns,m}$  is the equivalent to  $x_{ins,m} \times \beta_{n,m}$  in an utilitarian model.
- $ightharpoonup eta_{n,m}$  is the taste parameter of attribute m of individual n.

# 1 Classical RRM (Chorus, 2010)

• (Chorus, 2010) proposed the following systematic regret:

Attribute level regret. 
$$R_{ins} = \sum_{j \neq i}^{J} \sum_{m=1}^{M} R_{i \leftrightarrow jns,m}^{I} = \sum_{j \neq i}^{J} \sum_{m=1}^{M} \ln\left[1 + \exp\left\{\beta_{n,m} \cdot (x_{jns,m} - x_{ins,m})\right\}\right]$$
 Linear sum of all attribute level regret. (1)

- From our example:  $\mathcal{M} = \{T, C\}$ , J = 3.
- Regret of alternative 1  $(R_{1ns})$  will be described by:

$$\begin{split} R_{1ns} &= \sum_{j \neq i}^{3} \sum_{m \in \mathcal{M}} \ln \left[ 1 + \exp \left\{ \beta_{n,m} (x_{jns,m} - x_{ins,m}) \right\} \right] \\ &= \ln \left[ 1 + \exp \left\{ \beta_{n,T} \left( x_{2ns,T} - x_{1ns,T} \right) \right\} \right] + \ln \left[ 1 + \exp \left\{ \beta_{n,c} \left( x_{2ns,C} - x_{1ns,C} \right) \right\} \right] \\ &+ \ln \left[ 1 + \exp \left\{ \beta_{n,T} \left( x_{3ns,T} - x_{1ns,T} \right) \right\} \right] + \ln \left[ 1 + \exp \left\{ \beta_{n,C} \left( x_{3ns,C} - x_{1ns,C} \right) \right\} \right] \end{split}$$

# 1 Classical RRM (Chorus, 2010): Towards the log-likelihood.

- 1 Defining  $RR_{ins}=R_{ins}+\varepsilon_{ins}$ , where  $\varepsilon_{ins}$  is a type I Extreme Value i.i.d. error.
- 2 Acknowledging that the minimization of the random regret is mathematically equivalent to maximizing the negative of the regret.
- 3 Hence, the probabilities may be derived using the Multinomial Logit:

$$P_{ins} = \frac{\exp(-R_{ins})}{\sum_{j=1}^{J} \exp(-R_{jns})}$$
 for  $i = 1, ..., J$  (2)

4 Consequently, the log-likelihood will be described by:

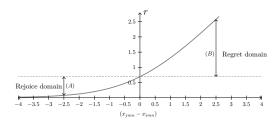
$$\ln L = \sum_{n=1}^{N} \sum_{s=1}^{S} \sum_{i=1}^{J} y_{in} \times \ln (P_{ins})$$

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# 2 Taste Parameter Interpretation in RRM models

- ▶ RUM: parameters are interpreted as the change in utility caused by an increase of a particular attribute level.
- RRM: parameters represent the potential change in regret associated with comparing a considered alternative with another alternative in terms of the attribute, caused by one unit change in a particular attribute level.
  - For instance:  $\widehat{\beta}_m > 0$  suggests that regret increases as the level of that attribute increases in a non-chosen alternative, in comparison to the level of the same attribute in the chosen alternative (e.g. Comfortable level).
  - For instance:  $\widehat{\beta}_m < 0$  suggests that regret decreases as the level of that attribute increases in a non-chosen alternative, in comparison to the level of the same attribute in the chosen alternative (e.g. Total Time).
- ▶ All in all. the parameters in RUM and RRM, are expected to have the same sign, even though their interpretation is dramatically different.

# 2 Semi-compensatory Behavior and the Compromise Effect



- ▶ Attribute level regret  $R_{i\leftrightarrow j,mn}$  with  $\beta_m = 1$ .
- $lackbox{(A)} = rejoice$  and (B) = regret on an equal difference of attribute level.
- ightharpoonup For an equal difference of the attribute levels  $\Rightarrow$  regret >>> rejoice
- ▶ Linear RUM models ⇒ fully-compensatory model.
- ➤ Compromise Effect: Alternatives with "balanced" performance in all attributes are more attractive than alternatives with a severe poor performance in one attribute.

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# 3 Mixed Random Regret Minimization Models

New assumptions of the model.

- $ightharpoonup eta_n$  follows a parametric distribution, denoted by  $f(eta|\varphi)$ .
- $ightharpoonup \varphi$  are the parameters that describe the distribution (e.g., mean and variance of a Normal distribution).
- We define the conditional probability (CP) of the observed sequence of choices of individual n (conditional on knowing  $\beta_n$ ) as:

$$P_n(\beta_n) = \prod_{s=1}^{S} \prod_{i=1}^{J} \{P_{ins}\}^{y_{ins}}$$
 (3)

ightharpoonup The unconditional probability of the observed sequence of choices is the CP integrated over the entire domain of the distribution of eta

$$\ln L(\beta) = \sum_{n=1}^{N} \ln \left[ \int_{\beta} P_n(\beta) f(\beta|\varphi) d\beta \right]$$
 (4)

# 3 Mixed Random Regret Minimization Models

- Given that equation (4) has no closed form we will approximate it using simulations (Train, 2009).
- ► Hence, we will estimate the model using Maximum Simulate Likelihood where we will maximize the following simulated log-likelihood function:

$$SLL(\beta) = \sum_{n=1}^{N} \ln \left\{ \frac{1}{R} \sum_{r=1}^{R} P_n(\boldsymbol{\beta}^r) \right\}$$
 (5)

▶ R is the number of replications and r is the r-th drawn from  $f(\beta|\varphi)$ .

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### 4 Individual Level Parameters

• We can obtain estimates of the individual level parameters  $(\beta_n)$ :

$$\bar{\beta}_n = \frac{\int_{\beta} \beta \times P_n(\beta) f(\beta|\varphi) d\beta}{\int_{\beta} P_n(\beta) f(\beta|\varphi) d\beta}$$
 (6)

We will approximate this individual level parameter by simulation using:

$$\check{\boldsymbol{\beta}}_n = \sum_{r=1}^R \left( \frac{\boldsymbol{\beta}^r \times P_n(\boldsymbol{\beta}^r)}{\sum_{r=1}^R P_n(\boldsymbol{\beta}^r)} \right) \tag{7}$$

► For this estimation we will use the command mixrbeta after estimating the population parameters using mixrandregret.

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# 5 Syntax

mixrandregret is implemented as a Mata-based gf-0 ml evaluator. The command allows the inclusion of normally and log-normally distributed random parameters.

```
mixrandregret depvar [indepvars] [if] [in] group(varname)

alternative(varname) rand(varlist) [, id(varname)

basealternative(string) noconstant ln(string) nrep(string)

burn(string) robust cluster(varname) level(#) maximize_options]
```

The command mixrbeta can be used after mixrandregret to calculate individual-level parameters corresponding to the variables in the specified varname using the method proposed by Train (2009).

```
mixrbeta varlist saving(filename) [, plot nrep(#) burn(#)]
```

### 5 The Data

▶ Data from van Cranenburgh (2018): Stated Choice (SC) experiment.

```
. rename (choice) (choice_w)
. qui reshape long tt tc , i(obs) j(altern)
. generate choice = 0
. replace choice = 1 if choice_w==altern
(1,060 real changes made)
. label define alt_label 1 "First" 2 "Second" 3 "Third"
. label values altern alt_label
. list obs altern choice id tt tc in 1/6, sepby(obs)
```

	obs	altern	choice	id	tt	tc
1.	1	First	0	1	23	6
2.	1	Second	0	1	27	4
3.	1	Third	1	1	35	3
4.	2	First	0	1	27	5
5.	2	Second	1	1	35	4
6.	2	Third	0	1	23	6

- Three unlabeled route alternatives (J = 3).
- Described by Travel Cost (tc) and Travel Time (tt) (M=2).
- Each respondent (id) answered a total of 10 choice situations.

### 5 Fixed Parameter RRM model

- First we estimate a fixed parameters regret model.
- As expected, both parameter estimates are negative.

```
. randregret choice tc tt, gr(obs) alt(altern) rrmfn(classic) ///
> nocons cluster(id) nolog
Fitting Classic RRM Model
RRM: Classic Random Regret Minimization Model
Case ID variable: obs
                                               Number of cases
                                                                         1060
Alternative variable: altern
                                               Number of obs
                                                                         3180
                                               Wald chi2(2)
                                                                       40.41
Log likelihood = -1118.4784
                                               Prob > chi2
                                                                        0.0000
                                 (Std. Err. adjusted for 106 clusters in id)
                             Robust
      choice
                   Coef.
                           Std. Err.
                                               P>|z|
                                                          [95% Conf. Interval]
RRM
                 -.417101
                             .068059
          t.c
                                       -6.13
                                                0.000
                                                         -.5504943
                                                                     -.2837078
          t.t.
                 -.102813
                            .0182526
                                       -5.63
                                                0.000
                                                         -.1385874
                                                                     - .0670386
```

<sup>.</sup> matrix b\_rrm = e(b)

- Mixed model with total time (tt) being Normally distributed.
- ▶ The mean of tt is negative, as expected, however...

```
. mixrandregret choice tc . gr(obs) alt(altern) rand(tt) id(id) ///
> nocons cluster(id) nrep(500) from(init mix rrm) tech(bhhh) nolog
Case ID variable: obs
                                               Number of cases
                                                                           1060
Alternative variable: altern
Random variable(s): tt
                                 (Std. Err. adjusted for 106 clusters in id)
Mixed random regret model
                                                Number of obs
                                                                         3,180
                                                Wald chi2(2)
                                                                        606.11
Log likelihood = -771.05731
                                                Prob > chi2
                                                                         0.0000
                               OPG
                            Std. Err.
                                                P>|z|
                                                          [95% Conf. Interval]
      choice
                    Coef.
Mean
                -1.102133
                            .0449728
                                       -24.51
                                                0.000
                                                         -1.190278
                                                                      -1.013988
          t.c
          t.t.
                - 3579483
                            .0581553
                                        -6.16
                                                0.000
                                                         -.4719307
                                                                      -.243966
SD
          t.t.
                 5068777
                            .0413721
                                     12.25
                                                0.000
                                                           4257899
                                                                       .5879656
```

The sign of the estimated standard deviations is irrelevant: interpret them as being positive

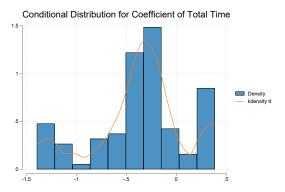
- mixrbeta allows us to compute the individual level parameters of Equation (7).
- ▶ It creates a new data set with one observation per individual (id) and its corresponding parameters estimates.

```
. preserve
. /* Computing Individual Level Parameters */
. qui mixrbeta tt , nrep(500) replace saving("${graphs_route}\mixRRM_normal_idl")
. use "${graphs_route}\mixRRM_normal_idl" , replace
. list id tt in 1/5
```

	id	tt
1. 2. 3. 4.	1 2 3 4 5	.37648493 05517235 .37680846 .38503775 .37615914
3. 4.	3 4	.37680846 .38503775

### 5 Mixed RRM model: Individual Level Parameters

Individual Level Parameters for total time when we assume it as Normally distributed.



- We see some individuals with positive estimates.
- To prevent this from happening we can use a bounded distribution...

We can use a bounded distribution (negative Log-normal) to force that every individual has a negative parameter for total time.

```
. gen ntt = -1 * tt
. mixrandregret choice tc . gr(obs) alt(altern) rand(ntt ) ln(1) id(id) ///
> nocons cluster(id) nrep(100) tech(bhhh) from(b_mixrrm) nolog
Case ID variable: obs
                                               Number of cases
                                                                           1060
Alternative variable: altern
Random variable(s): ntt
                                 (Std. Err. adjusted for 106 clusters in id)
Mixed random regret model
                                                Number of obs
                                                                          3,180
                                                Wald chi2(2)
                                                                        1224 65
Log likelihood = -785.40416
                                                Prob > chi2
                                                                         0.0000
                               OPG
      choice
                    Coef
                            Std. Err.
                                                P>|z|
                                                           [95% Conf. Interval]
                                           z
Mean
                -1.217628
                            .0443379
                                       -27.46
                                                0.000
                                                          -1.304529
                                                                      -1.130727
          t.c
         ntt
                -1.292678
                            .1549439
                                        -8.34
                                                0.000
                                                          -1.596362
                                                                      -.9889936
SD
                 1.362446
                                        11.55
         ntt
                            .1179392
                                                0.000
                                                           1.13129
                                                                       1.593603
```

The sign of the estimated standard deviations is irrelevant: interpret them as being positive

Similarly, we can compute the individual level parameters for the log-normally distributed variable tt.

```
. /* Computing Individual Level Parameters */
. qui mixrbeta ntt , nrep(100) replace saving("${graphs_route}\mixRRM_ln_idl")
. use "${graphs_route}\mixRRM_ln_idl" , replace
. replace ntt = -1 * ntt /*reverse sign for graph*/
(106 real changes made)
. list id ntt in 1/3

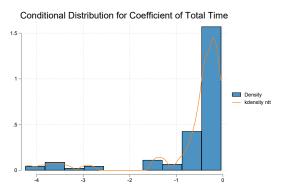
id ntt

1. 1 -.04005948
2. 2 -.08146095
```

3.

-.04124789

Individual Level Parameters when total time is assumed to be Log-normally distributed.



Now we observe that the individual level parameters are all negative.

- The parameters we estimated are the mean  $(\beta_{tt})$  and standard deviation  $(s_{tt})$  of the natural logarithm of the total time coefficient.
- Hence, the mean, median and variance of log-normal distributed parameter are equal to  $\exp(\beta_{tt})$ ,  $\exp(\beta_{tt} + s_{tt}/2)$  and  $\exp(\beta_{tt} + s_{tt}/2) \times \sqrt{\exp(s_{tt}^2) 1}$ , respectively.
- Finally, we can compute them using nlcom.

```
. nlcom (mean_time: -1*exp([Mean]_b[ntt] +0.5*[SD]_b[ntt]^2)) ///
> (med_time: -1*exp([Mean]_b[ntt])) ///
> (sd_time: exp([Mean]_b[ntt] +0.5*[SD]_b[ntt]^2) * sqrt(exp([SD]_b[ntt]^2)-1))
  mean_time: -1*exp([Mean]_b[ntt] +0.5*[SD]_b[ntt]^2)
  med_time: -1*exp([Mean]_b[ntt])
  sd_time: exp([Mean]_b[ntt] +0.5*[SD]_b[ntt]^2) * sqrt(exp([SD]_b[ntt]^2)-1)
```

choice	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
mean_time	6945106	.1603422	-4.33	0.000	-1.008776	3802456
med_time	2745346	.0425375	-6.45	0.000	3579065	1911627
sd_time	1.613861	.6351531	2.54	0.011	.3689834	2.858738

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### 6 Conclusions

- ➤ The mixrandregret command extends its predecessor randregret (Gutiérrez-Vargas et al., 2021) by allowing the inclusion of random coefficients in the regret functions.
- ▶ The parameters are estimated by Simulated Maximum Likelihood.
- The random parameters can follow either a Normal or Log-normal distribution.
- Additionally, we can compute the individual level parameters using the mixrbeta command.
- ► The programs can be downloaded from Ziyue's Github account.
- ▶ The example code used in this presentation is available here.

- Bibliography

# 8 Bibliography

- Chorus, C. G. (2010). A new model of random regret minimization. European Journal of Transport and Infrastructure Research, 10(2):181–196.
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