

Bayesian estimation for collisional thermometry

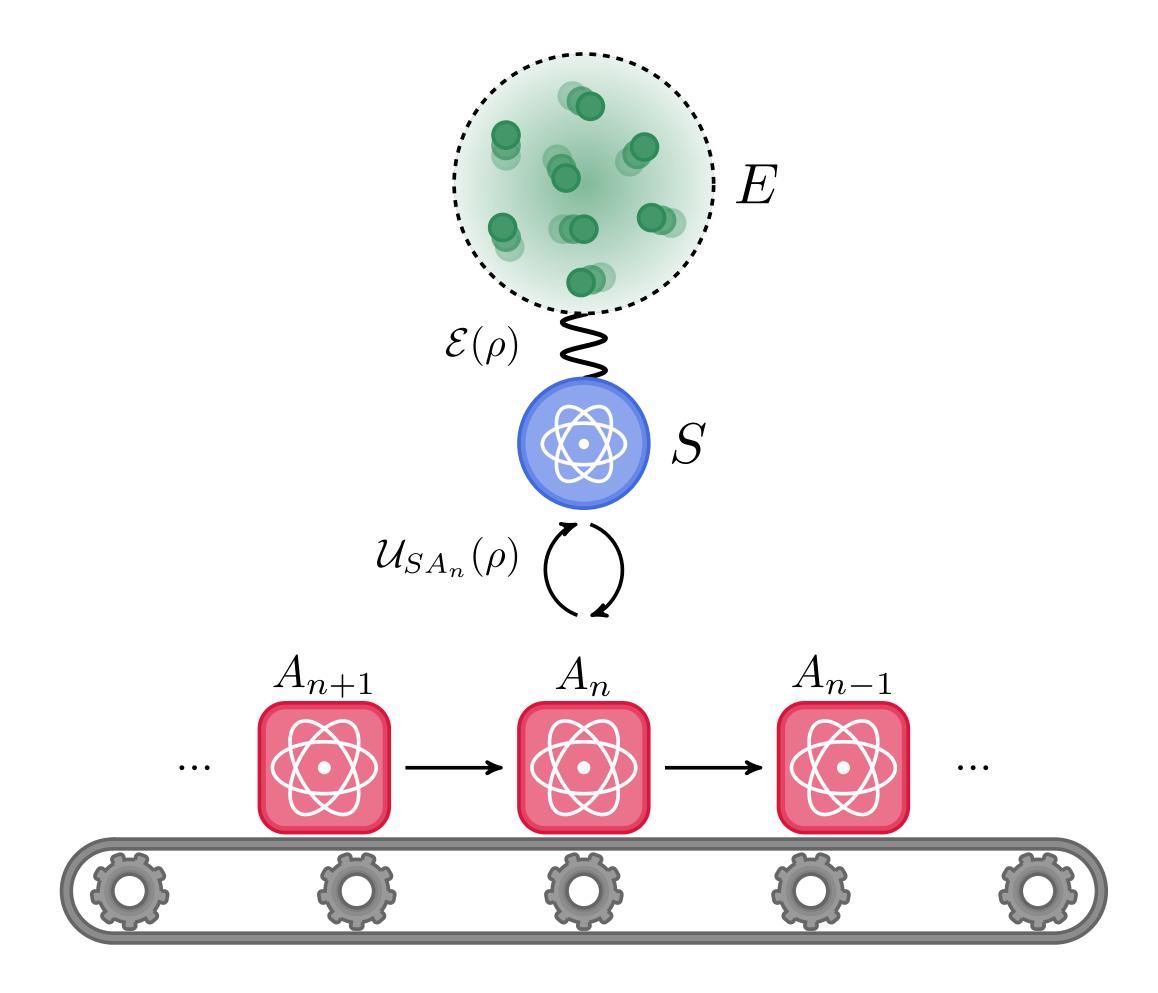
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Introduction

We consider a qubit collisional model as a thermometric platform. It's been shown that the out-of-equilibrium steady state dynamics of the collisional model can be used, for instance, to enhance precision, surpassing the thermal Fisher information [1]. Here we construct a framework for analyzing collisional thermometry using Bayesian inference. In particular, we explicitly plot an estimator and compare the results with the Cramér-Rao and the Van Trees-Schützenberger bounds.



Qubit Collisional Model

The system undergoes alternating and piecewise interactions. First through a system-environment interaction for time τ_{SE} :

$$\frac{d\rho_S}{dt} = \mathcal{L}(\rho_S) = \gamma(\bar{n}+1)\mathcal{D}[\sigma_-^S] + \gamma\bar{n}\mathcal{D}[\sigma_+^S],$$

which implies $\mathcal{E}(\rho_S) = e^{\tau_{SE}\mathcal{L}}(\rho_S)$, and then through partial-swap interactions with the ancillas:

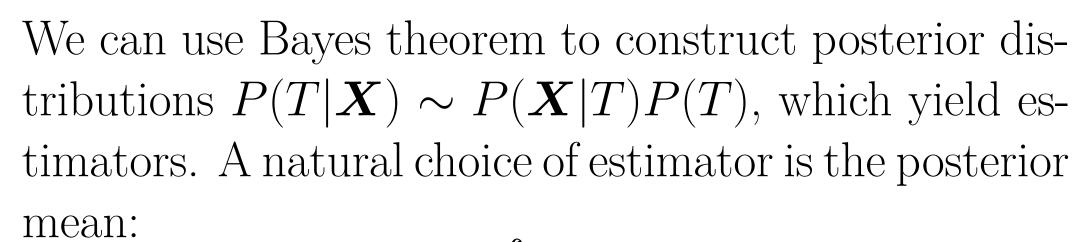
$$U_{SA_n} = \exp\left\{-i\tau_{SA}g(\sigma_+^S\sigma_-^{A_n} + \sigma_-^S\sigma_+^{A_n})\right\}$$

This results in a stroboscopic map:

$$\rho_S^n = \operatorname{tr}_{A_n} \{ \mathcal{U}_{SA_n} \circ \mathcal{E}(\rho_S^{n-1} \otimes \rho_A^0) \}$$

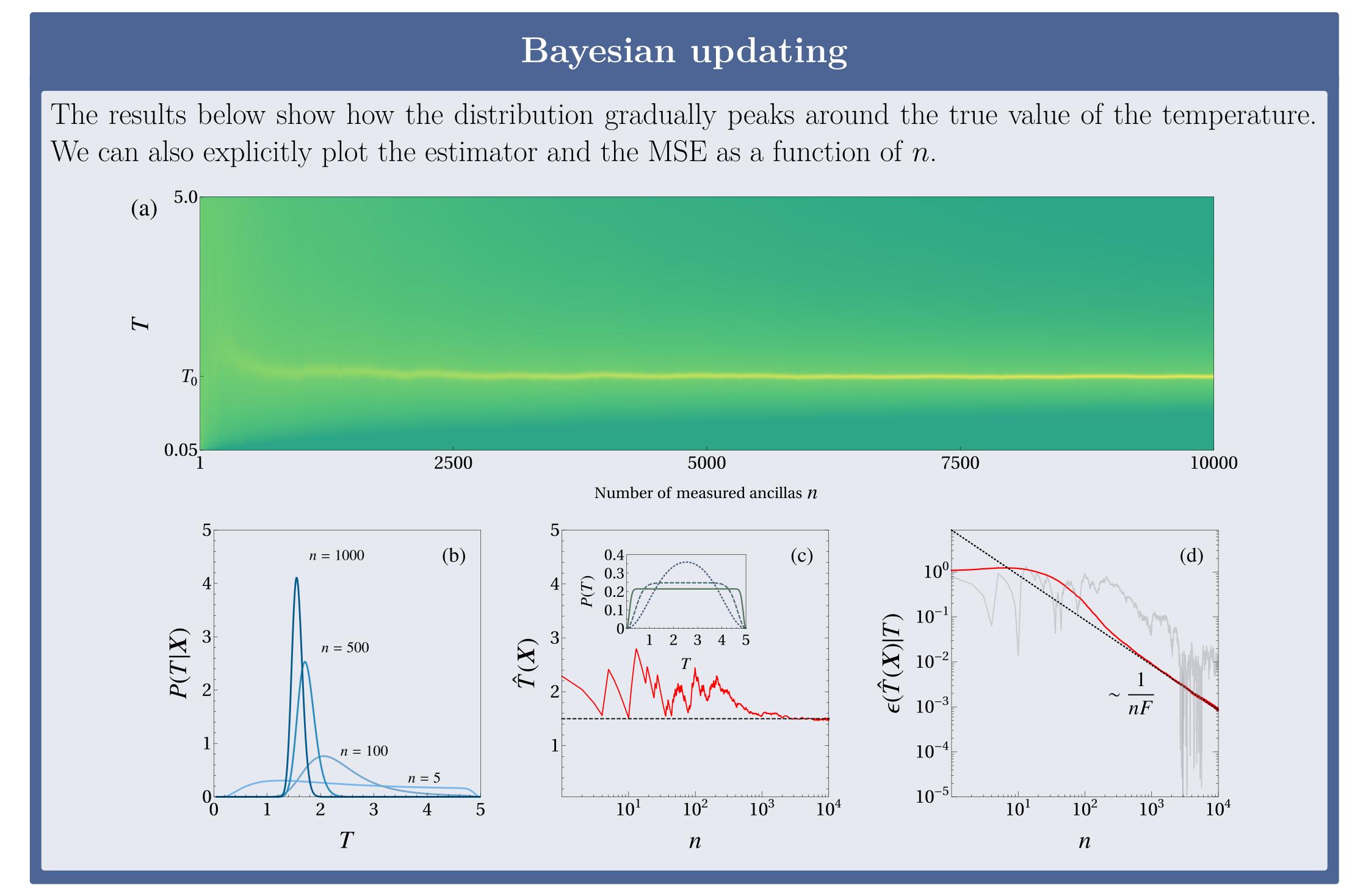
We consider local measurements on the ancillas. The measurements are performed in the computational basis and at the steady state, which is calculated from the map above.

Bayesian Inference



$$\hat{T}(\boldsymbol{X}) = \int TP(T|\boldsymbol{X})dT$$

The quantity above minimizes the mean-squared error $\epsilon(\hat{T}(\boldsymbol{X})|T) = \int (T-\hat{T})^2 P(\boldsymbol{X}|T) d\boldsymbol{X}$ and saturates the CRB asymptotically.



Acknowledgements

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References

[1] Stella Seah, Stefan Nimmrichter, Daniel Grimmer, Jader P. Santos,
 Valerio Scarani, and Gabriel T. Landi.
 Collisional Quantum Thermometry.
 Physical Review Letters, 123(18):180602, oct 2019.

[2] Gabriel O. Alves and Gabriel T. Landi.
Bayesian estimation for collisional thermometry, 2021.

Additionally, the posterior distribution converges to a Gaussian, with variance proportional to the Fisher information calculated for the true temperature:

$$P(T|X) \approx \sqrt{\frac{nF(T_0)}{2\pi}} e^{-\frac{nF_0(T-T_0)^2}{2}}, \quad (n \text{ large}).$$

We can also analyze the problem in terms of a figure of merit which is independent of the temperature. We call it the *Bayesian* error:

$$\epsilon_B(\hat{T}(\boldsymbol{X})) = \int P(T)dT \int (T - \hat{T})^2 P(\boldsymbol{X}|T)d\boldsymbol{X}$$

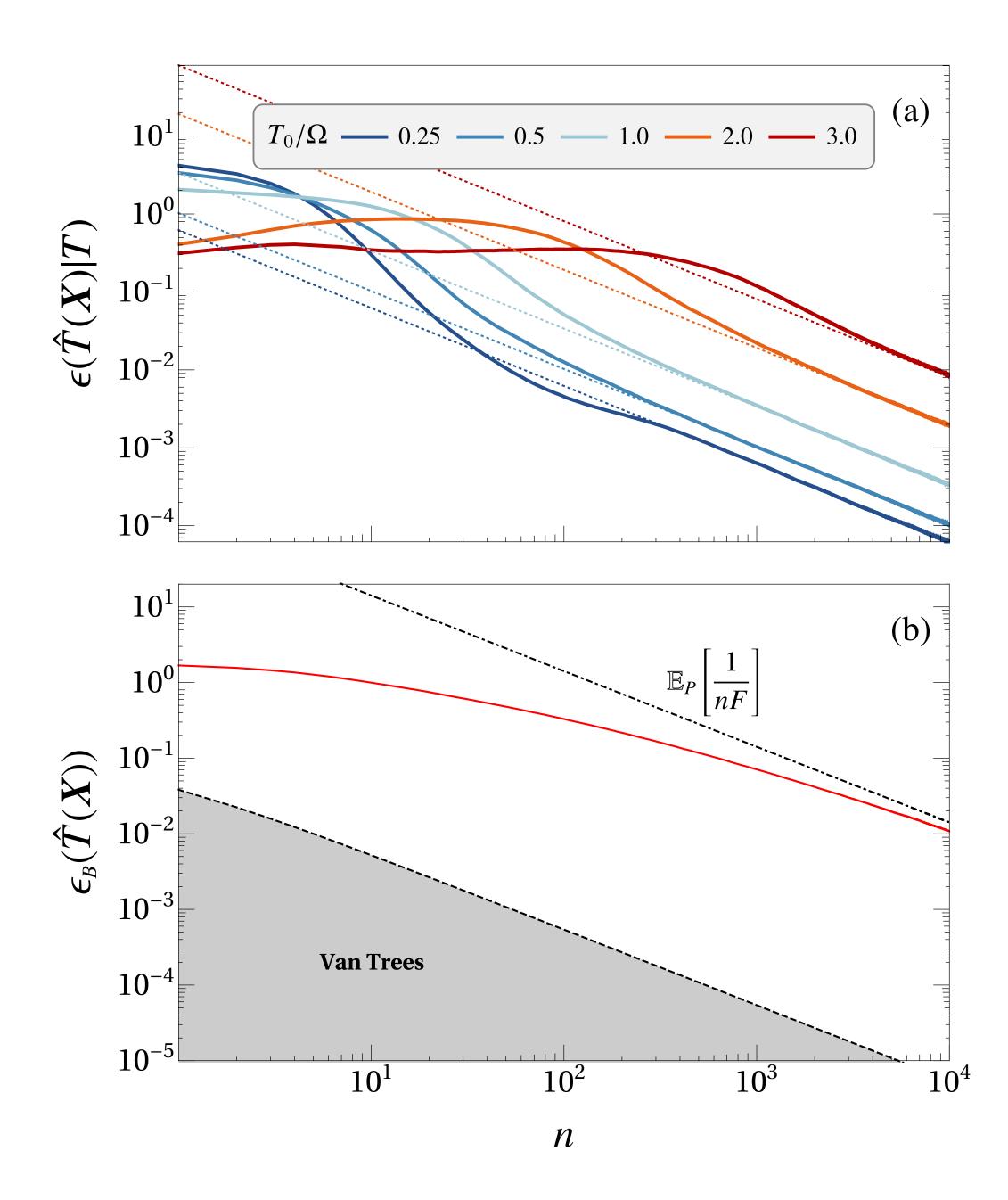
Note that the figure of merit above depends only on the estimator, the prior and the likelihood, and it's not conditioned neither on the outcomes nor on a particular temperature.

Van Trees-Schützenberger Inequality

This inequality establishes a bound for the Bayesian risk defined above:

$$\epsilon_B(\hat{T}(X)) \ge \frac{1}{\mathbb{E}_P[F(T)] + F_P},$$

where $\mathbb{E}_P[F(T)] = \int F(T)P(T)dT$ is the Fisher information averaged over the prior.



The figure above shows (a) the MSE for different temperatures and (b) the Bayesian risk, with the Van Trees-Schützenberger inequality in gray. Notice how the Bayesian risk converges to $\mathbb{E}_P[1/nF(T)]$, the prior-averaged CRB. This provides an asymptotic analysis which does not depend on a particular value of the (unkown) temperature. This fact makes it possible to devise strategies which are, for instance, suited to larger temperature intervals.

