

Predicting cancer mortality in the United States

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Abstract

Background Cancer is a leading cause of death in the United States [Wang]. Several factors have been linked to cancer risk on individual and ecological scales, including family history, lifestyle habits, median regional income, and others [NIH]. We seek to understand what population attributes can predict cancer mortality per capita (100,000 persons).

Data Data for this study is aggregated from the US Census Bureau, ClinicalTrials.gov, and National Cancer Institute databases. The final dataset, which consists of information for 3047 US counties, includes demographic, cancer incidence, and cancer mortality data.

Methods We evaluate educational, economic, and other population characteristics (e.g. age) and their association with aggregate cancer outcomes on a state level and fit a linear model to fit and predict cancer mortality.

Results Our final model predicts cancer mortality rate based on cancer incidence, percent of the population in poverty, median age, percent of the population that is married, percent of the population with a college degree, percent of the population that is non-white, percent unemployment, and the percent of the population with health insurance.

Discussion Our final model is not the one with the lowest MSE, but it accounts for over 50% of the variability in the outcome and uses what we believe is the fewest number of significant and/or practically meaningful covariates. Additionally, limitations may exist in limiting prediction to linear models. Further exploration on this topic using alternative and non-linear methods may produce higher accuracy in prediction.

Introduction

Cancer is a leading cause of death in the United States [Wang]. Large amounts of resources are poured into cancer research and treatment in an effort to better understand the biological pathways, risk factors, and improve outcomes [NIH]. Several factors have been linked to cancer risk on individual and ecological scales, including family history, lifestyle habits, median regional income, and others [NIH]. In the last twenty years, cancer incidence has increased. However, it's unclear the extent to which this rise in incidence ought to be attributed to improved detection, poorer lifestyles compared to earlier populations, excess exposure to risk factors, etc. [NIH]. And for many malignancies, prognosis has not much improved despite increased knowledge of biological mechanisms and technological/medical developments. For example, glioblastoma is one of the most aggressive cancers, and despite the large amount that is known about the biology, there are few effective therapies [2]. Here, we seek to understand what factors can predict cancer mortality per capita (100,000 persons) using linear regression methods.

Methods

Data was aggregated across multiple sources - including the American Community Survey, US Census, ClinicalTrials.gov, and the National Institutes of Health (NIH).

Several groups of covariates were observed to be associated with each other or had skewed distributions. The estimates for the employment rate, percent of people with private insurance alone, and the percent of people between ages and 18-24 who attended some college had much missing data, and so we excluded them from consideration. This was inconsequential since they showed an association with other variables that did not have any missing data.

Among correlated covariates and to reduce risk of multicollinearity, we chose the single possible predictor that is most associated with the outcome. Namely, we recoded counties as their home state; we looked across age groups at the percent of people with college degrees; we transformed the number of new cases of cancer to a logarithmic scale; and we looked at the percent of the population that was non-white, as opposed to looking at the makeup of specific races (Table 1). Ultimately, we examined the significance of the cancer incidence rate, percentage of the population in poverty, per capita number of cancer-related clinical trials, median age, average number of people in a household, percent of the population that is married, education level, and health care coverage in predicting cancer mortality rate across states.

We fit a linear model using a combination of stepwise and Lasso methods. We use the Akaike Information

Criterion (AIC) in stepwise model selection. In the Lasso model, a variety of λ values ranging from 10^{-2} to 10^5 , were tested to determine the ideal weight for the Lasso model. Once a final model was determined, 10-fold cross validation was used to assess predictive ability based on the root mean squared error (RMSE). Figures and tables, as well as additional detail on data cleaning, model building, and cross validation can be found in the Appendix.

Results

Our final model uses cancer incidence, percent of the population in poverty, median age, percent of the population that is married, percent of the population with a college degree, percent of the population that is non-white, percent unemployment, and the percent of the population with health insurance to predict the average cancer death rate per 100,000 across states. State-level distributions of cancer mortality rates per county are given in Figure 1; chosen predictors and their effects are given in Table 2.

Expected cancer mortality rate varies by state, as expected. The adjusted R-squared of the model is 0.55, which is acceptable. There are positive associations with death rate in incidence, age, poverty, unemployment rate, and proportion of minorities. Negative associations were seen with marriage rate, proportion of college graduates, and health insurance status. Marriage rate, percent non-white, and health insurance coverage are not suggested to be significant predictors of cancer mortality rate ($p > 0.05$).

We saw positive associations with death rate in incidence, age, poverty, unemployment and proportion of minorities; holding all else constant, each one unit increase of cancer incidence per capita per year, mean cancer mortality is expected to increase by 0.20; with each one percent increase in the percent of population in poverty, mean cancer mortality per capita is expected to increase by 0.68; with each one year increase in median age among men, mean cancer mortalities per capita is expected to increase by 0.23; with each one percent increase in unemployment rate, mean cancer mortality per capita is expected to increase by 0.69; with each one percent increase in the minority population, mean cancer mortality per capita is expected to increase by 0.03.

Negative associations were seen with marriage, proportion of college graduates, and health insurance status. Adjusting for all other covariates in model, with each one percent increase in the percentage of married people, mean cancer mortality per capita is expected to decrease by 0.15; with one percent increase in the percent of people who earned bachelor's degree, mean cancer mortality per capita is expected to decrease by 0.75; and with one percent increase health insurance coverage, mean cancer mortality per capita is expected to decrease by 0.06.

In checking model assumptions, we found that the distribution has long tails, suggesting that there are a substantial number of counties with extremely large or extremely small cancer mortality rates (Figure 2). Only a few of these outliers in Virginia and Nevada showed evidence of influencing the model, and upon inspection proved not to significantly alter the results or improve the diagnostics.

Cross validation of our model showed modest ability to predict cancer mortality, with an RMSE of ~ 18.7 .

Discussion

The associations seen in our model results are not all too surprising considering the variables with the positive association are commonly deemed as disadvantages in society (e.g. unemployment) while those with the negative associations can be deemed as measures of success (e.g. having a college degree). It was interesting to note, however, that some known predictors of cancer outcomes were not statistically significant at the 5% significance level. For example, race (the proportion of the population that is non-white), is not significant here despite known associations with cancer outcomes [Katz]; the same can be said for health insurance coverage [Niu] and marital status [Aizer].

Our model does a modest job of predicting cancer mortality rates. Skewness in the data is a hurdle that we cannot seem to overcome using linear methods; non-linear alternatives or machine learning processes may better handle this issue and predict the outcome. That being said, our final model is not the one with the lowest MSE, but it accounts for over 50% of the variability in the outcome and uses what we believe is the fewest number of significant and/or practically meaningful covariates.

References

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Appendix

Figures and Tables

Measure	Min	1st Q	Mean	Median	3rd Q	Max	Std Dev
Avg household size	0.022	2.37	2.48	2.5	2.63	3.97	0.43
Cancer incidence	201.3	420.3	448.27	453.55	480.85	1206.9	54.56
Percent change in number of new cases	1.79	4.33	5.32	5.14	6.25	10.55	1.43
Median age among men	22.4	36.35	39.57	39.6	42.5	64.7	5.23
Pop. percentage with college degree	2.7	13.2	19.44	17.7	23.75	86.3	8.88
Pop. percentage with insurance	65.4	96.25	100.61	101.3	105.8	131.7	7.39
Pop. percentage non-white	0	4.55	16.35	9.94	22.7	89.8	16.38
Unemployment rate	0.4	5.5	7.85	7.6	9.7	29.4	3.45
Pop. percentage married	23.1	47.75	51.77	52.4	56.4	72.5	6.9
Pop. percentage in poverty	3.2	12.15	16.88	15.9	20.4	47.4	6.41
Avg number of clinical studies	0	0	155.4	0	83.65	9762.31	529.63
Cancer mortality rate (per 100,000)	59.7	161.2	178.66	178.1	195.2	362.8	27.75

Table 2: Predictors of cancer mortality (number of cancer deaths per 100,000 persons) across states

Variable	Estimate (SE)	p-value
(Intercept)	94.71 (11.317)	0.000
Cancer incidence	0.195 (0.007)	0.000
Pop. percentage in poverty	0.675 (0.109)	0.000
Median age among men	0.234 (0.095)	0.014
Pop. percentage married	-0.151 (0.096)	0.113
Pop. percentage with college degree	-0.75 (0.052)	0.000
Unemployment rate	0.685 (0.164)	0.000
Pop. percentage non-white	0.027 (0.038)	0.469
Pop. percentage with insurance	-0.056 (0.085)	0.506

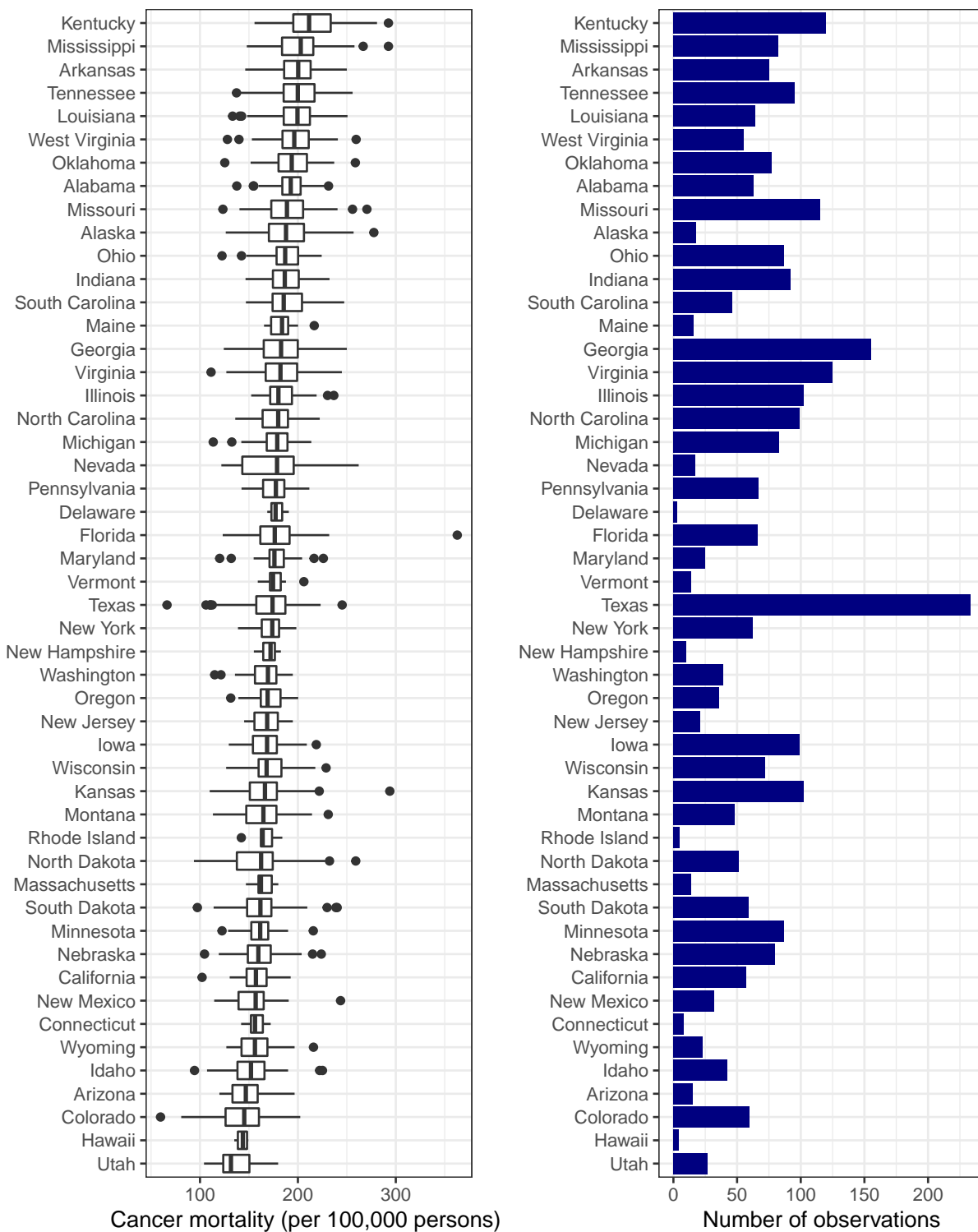


Figure 1: State-level observations

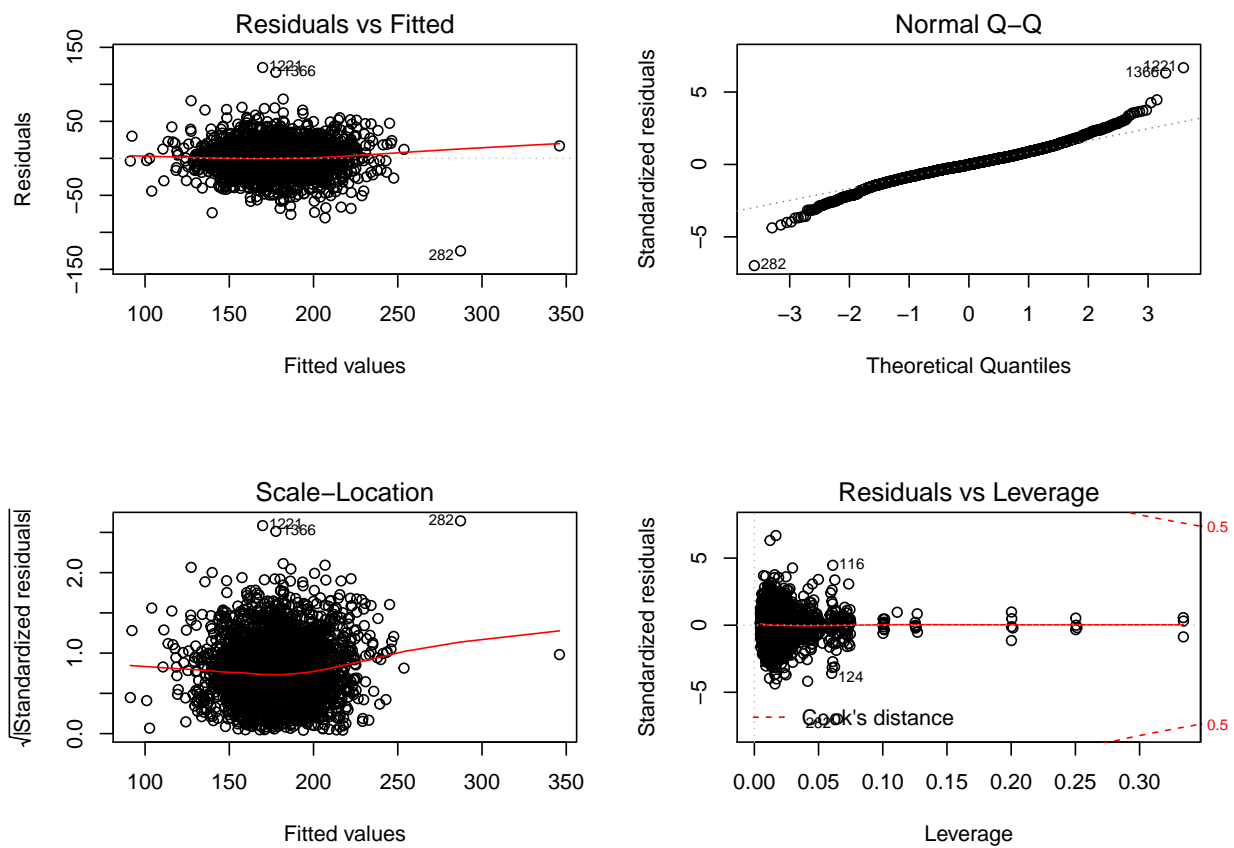


Figure 2: Model diagnostics

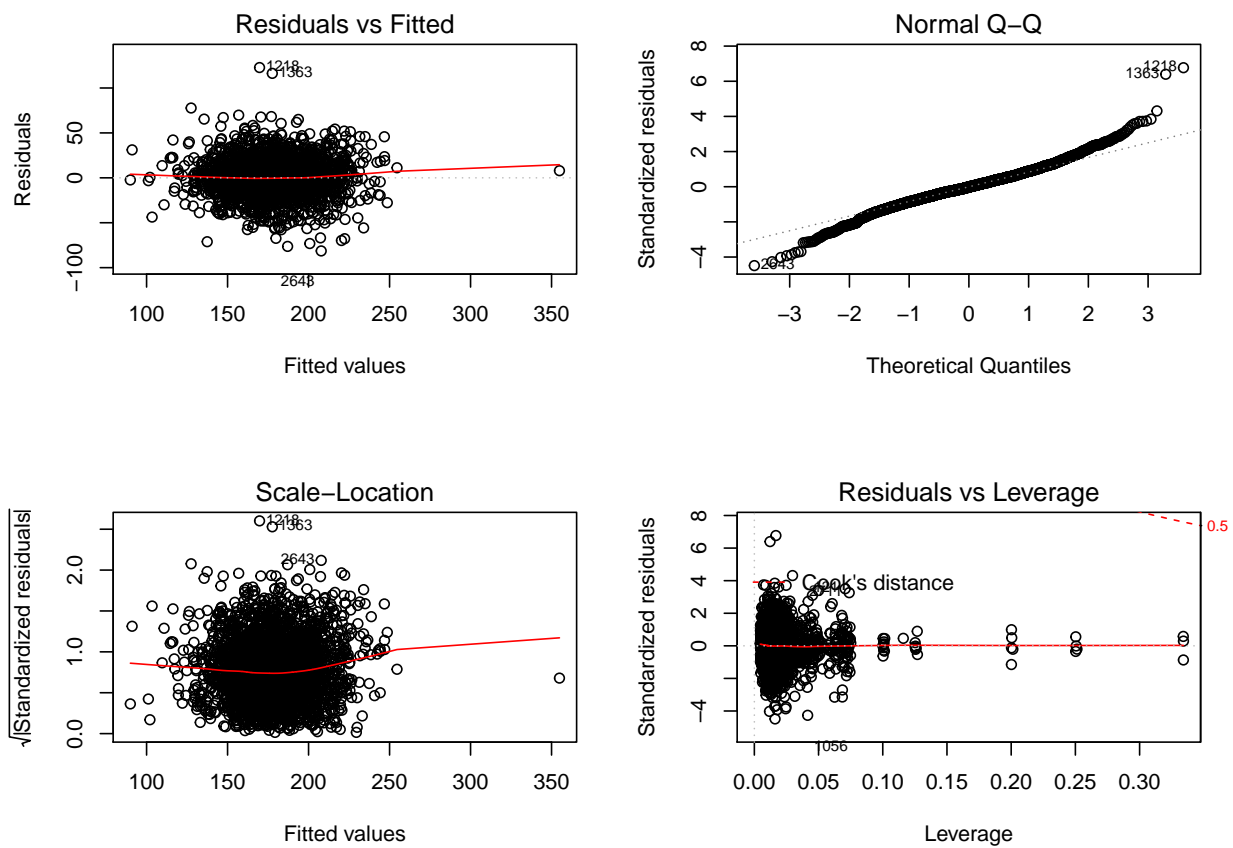


Figure 3: Diagnostics of model excluding potential influential points

Table 3: 10-fold Cross Validation

MSE	RMSE	SE
349.8994	18.70556	0.1255092

Table 4: Cross Validation across metrics

SSE	PRESS	RMSE	Adjusted Rsq
1025830	1068340	18.5257	0.5543684

Detailed Methods

Data exploration and cleaning

We first examined variables for any skewness, and then tested associations between covariates to identify potential sources of collinearity. We found that several sets of variables were correlated. For those groups of correlated covariates, we generally chose to include the single covariate that was most highly associated with the outcome, cancer mortality rate.

We saw that median age was highly skewed, to the point that we think there is some sort of data error. We opt instead to use the median age in males as a proxy, since the distribution (discarding the outliers in median age) is comparable to median age overall and median age in females. It’s also interesting to note that median age between males and females are highly correlated with each other, but not much correlated with target death rate. In fact, median age in males is negatively correlated with death rate, while median age for females is positively correlated.

We looked again at the variables for the percent of the population of a certain race and found that the log transformation was bimodal. It might make sense to make this binary (i.e. >5% of a certain race or not). In the end, we chose to make a single variable measuring the percent of the population that is non-white (considered minority). This variable was skewed, but we did not perform a transformation for the sake of interpretability.

We combined several variables to create aggregate measures. We initially thought to group counties by region (Northeast, South, Midwest, West), but found that state-level groupings provided better fit and predictive

capability. We designated the District of Columbia as being part of Maryland, since it had a single observation. We also created single measures for the percent of the population with a bachelor's degree and percent of the population with health insurance.

Median income and the percent of the population in poverty are correlated with each other as well as the percent of the population with a college degree. Percent in poverty had a slightly higher association with the death rate, and a slightly lower correlation with education (as measured), so we opted to include it over median income.

Model selection

We used a combination of stepwise selection (using AIC) and Lasso to choose our final model. When we tested our original selected variables under both processes, we find the same parameters are downweighted/removed. We also tried fitting Lasso on the full original dataset (with our transformed/grouped variables) and found that many of the covariates appeared meaningful. However, this model seemed to overfit the data, especially since it chose to include some predictors with known correlation (e.g. median income and the percent of the population in poverty), suggesting presence of multicollinearity.

We performed stepwise regression using both AIC and BIC criterion, and found that using BIC resulted in several fewer variables being selected. But due to known associations between the eliminated variables and cancer mortality, we choose to include them in our model. For example, stepwise regression based on both criteria eliminated the percent of the population with health insurance as a predictor of cancer mortality per capita. However, we suspect this may be a practically meaningful covariate, since it has been shown that those with health insurance tend to have improved cancer outcomes over those without health insurance. This was done on an individual level [Ellis] and by relating individual insurance coverage with communities [Niu].

Model diagnostics

We saw that there are long tails on both sides of the distribution (Figure 2). This may be due to outliers at either extreme, so next we examine presence of outliers and potentially influential points.

Skewness still exists without these influential points, but the model's R^2 and Adj R^2 both improve marginally (Figure 3). As a result, we opted to keep these observations in our model.

Model validation

Our 10-fold cross validation shows root MSE of ~ 18.7 (Table 3). Performing leave-one-out (N-fold) cross validation gives a comparable value. We also examine other criteria (SSE, R-squared, and PRESS) (Table 4).