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Cognitive behavior therapy for externalizing disorders in children and adolescents in routine clinical care: A systematic review and meta-analysis

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A R T I C L E I N F O

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A B S T R A C T

Various Cognitive Behavioral Therapy (CBT) programs for externalizing disorders in children and adolescents are supported by a substantial body of empirical evidence. Most of the research evidence comes from efficacy studies conducted in university settings, but there is less knowledge about the effect of these treatments in routine clinical care. The purpose of this meta-analysis was to investigate the effectiveness of CBT in non-university settings for Attention Deficit Hyperactivity Disorder (ADHD), Conduct Disorder (CD) and Oppositional Defiant Disorder (ODD). Embase OVID, Ovid MEDLINE and PsycINFO were systematically searched for eligible studies published up to May 2020. In total, 51 treatment effectiveness studies involving 5295 patients were included.

The average within-group effect size at post-treatment was significant (*g* = 0.91), and there were large effect sizes for both ADHD (*g* = 0.80) and CD/ODD (*g* = 0.98). At post treatment, remission rates were 38% for ADHD and

48% for CD/ODD, and the overall attrition rate was 14%. Benchmarking against efficacy studies showed that CBT in routine clinical care yields remission rates, within-group effect sizes and attrition rates that are very similar to those found in university settings. The findings support the transportability of CBT for externalizing disorders from university settings to routine clinical care.

PROSPERO registration: CRD42020147524.

Externalizing behavior disorders are among the most common rea- sons for referral to mental health services for children and adolescents and include the diagnoses of Attention Deficit Hyperactivity Disorder (ADHD), Conduct Disorder (CD) and Oppositional Defiant Disorder (ODD). ADHD has an estimated prevalence of 7.2% ([Thomas, Sanders,](#_bookmark81) [Doust, Beller, & Glasziou, 2015](#_bookmark81)) whereas CD and ODD have estimated prevalence rates of 3.5% and 2.8%, respectively ([O’Connell, Boat, &](#_bookmark65) [Warner, 2009](#_bookmark65)). Although ADHD is classified as a neurodevelopmental disorder and CD and ODD as disruptive behavior disorders they are all characterized by problems with aggression, impulse-control and rule- breaking behavior. Children with externalizing disorders have poorer academic and social functioning than other children ([Clark, Prior, &](#_bookmark27) [Kinsella, 2002](#_bookmark27); [DuPaul, Morgan, Farkas, Hillemeier, & Maczuga, 2016](#_bookmark33)) and have a higher risk of negative outcomes in adulthood, including an increased risk of unemployment, criminality, financial problems, and

increased mortality rates ([Franke et al., 2018](#_bookmark46); [Scott et al., 2017](#_bookmark71)). They are also at a higher risk of developing emotional problems, as exter- nalizing disorders are highly comorbid with both anxiety and depres- sion, which appears to lead to more severe impairment than when children are diagnosed with either disorder alone ([Jarrett & Ollendick,](#_bookmark50) [2008](#_bookmark50); [Schatz & Rostain, 2006](#_bookmark69); [Wolff & Ollendick, 2006](#_bookmark87)). Furthermore, this vulnerability for the development of emotional disorders appears to be present in adulthood as well, as childhood ODD has been found to predict depression in young adulthood ([Copeland, Shanahan, Costello,](#_bookmark28) [& Angold, 2009](#_bookmark28)). The relatively high prevalence of externalizing dis- orders and the severe consequences for the individual and the society highlights the necessity of providing evidence-based interventions.

Various treatment programs for externalizing disorders have been supported by findings from numerous randomized controlled trials (RCTs). There is evidence for the efficacy of a wide range of CBT-

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oriented treatments. Programs with well-established efficacy for exter- nalizing disorders include behavior therapy programs for parents, chil- dren and/or teachers and cognitive interventions aimed at increasing children’s self-control and problem solving skills as well as programs addressing the organizational problems of children with ADHD ([Evans,](#_bookmark38) [Owens, & Bunford, 2014](#_bookmark38); [Kaminski & Claussen, 2017](#_bookmark52); [McCart &](#_bookmark61) [Sheidow, 2016](#_bookmark61)). Although it is encouraging that a growing body of methodologically sound trials are carried out, it has been argued that more attention should be directed at investigations of clinical effec- tiveness in routine clinical care ([Tolin, McKay, Forman, Klonsky, &](#_bookmark83) [Thombs, 2015](#_bookmark83)). Efficacy trials from university settings are often carried out with a methodologically stringent procedure which produces results with high internal validity. However, the circumstances in such trials may differ from “real world” clinical settings. For instance, participants are usually actively recruited through advertisements and willing to risk being randomized to a control condition. Also, it is common for efficacy studies to apply more exclusion criteria in order to obtain a more ho- mogenous sample of participants. ADHD is, for example, highly co- morbid with other disorders (e.g. [Owens & Hinshaw, 2013](#_bookmark67); [Steinhausen](#_bookmark79) [et al., 2006](#_bookmark79)) and a comorbid diagnosis of ADHD and ODD or CD has been found to increase the severity of ADHD symptoms ([Efron, Bryson, Lycett,](#_bookmark35) [& Sciberras, 2016](#_bookmark35); [Jerrell, McIntyre, & Park, 2015](#_bookmark51); [Takeda, Ambrosini,](#_bookmark80) [deBerardinis, & Elia, 2012](#_bookmark80)). Similarly, children and adolescents with CD are frequently diagnosed with comorbid depression and typically show greater impairment than children diagnosed with either disorder alone ([Wolff & Ollendick, 2006](#_bookmark87)). This could result in patients in efficacy trials having less severe symptoms and that the sample may not be repre- sentative of patients in routine clinical settings. Furthermore, the ther- apists involved in efficacy trials are often experts in the intervention under investigation, and are dedicated part- or full time to the research project as opposed to clinicians in routine care who carry large caseloads with a variety of disorders. The differences between university settings and routine clinical care suggest that findings from methodologically stringent efficacy trials may not be transferable to clinical practice.

Previous reviews comparing studies from routine clinical care to

efficacy studies from university settings have shown different results for different disorders. A meta-analysis of effectiveness studies for children with autism spectrum disorder found smaller effect-sizes for community- based studies than for university-based clinical trials, indicating a gap between research settings and routine clinical care ([Nahmias, Pellec-](#_bookmark64) [chia, Stahmer, & Mandell, 2019](#_bookmark64)). Conversely, in a recent meta-analysis of CBT for internalizing disorders in children and adolescents in routine

clinical care, [Wergeland, Riise and O](#_bookmark86)¨ [st (2021)](#_bookmark86) reported treatment out-

comes in routine clinical care comparable to those in university settings. To better understand how CBT-oriented treatments for externalizing disorders work in clinical practice, it is important to evaluate how evidence-based treatment programs perform in real world settings. During the last 10 years a substantial number of effectiveness studies in non-university settings have been published, and findings from such studies provide important information to clinicians and decision makers as to whether empirically supported treatments can be expected to yield the intended results when implemented in routine clinical care. Despite the importance of effectiveness studies for dissemination of empirically validated treatments, systematic reviews of treatments for externalizing disorders in routine clinical care are scarce. Several systematic reviews of psychological treatment have been performed in the past decade for ADHD and CD/ODD respectively, but have not focused on effectiveness studies specifically ([Bakker, Greven, Buitelaar, & Glennon, 2017](#_bookmark19); [Burkey](#_bookmark23) [et al., 2018](#_bookmark23); [Fabiano, Schatz, Aloe, Chacko, & Chronis-Tuscano, 2015](#_bookmark42); [Kaminski & Claussen, 2017](#_bookmark52); [Schatz et al., 2020](#_bookmark70); [Sibley, Kuriyan, Evans,](#_bookmark76) [Waxmonsky, & Smith, 2014](#_bookmark76); [Sonuga-Barke et al., 2013](#_bookmark77)). [Lee, Horvath,](#_bookmark57) [and Hunsley (2013)](#_bookmark57) carried out a review of effectiveness studies of child therapy and reported an increased number of effectiveness studies from 2007, high completion rates, and maintenance of treatment gains in follow-up assessments. They also found encouraging evidence for the effectiveness of parent training for disruptive behavior problems in

routine clinical practice. Although a variability in treatment outcome data was found, the authors concluded that positive results could be obtained in various treatment settings with children presenting with multiple problems ([Lee et al., 2013](#_bookmark57)). This review was, however, limited to 13 studies of parent training which did not report inclusion criteria, making it difficult to ascertain whether the results apply to specific disorders such as ADHD, CD or ODD. Therefore, further assessment of effectiveness studies seems called for and a meta-analysis of the effec- tiveness of CBT in the treatment of externalizing disorders in routine clinical care is warranted.

The present meta-analysis is the first to review effectiveness studies of CBT and behavior therapy (BT) for ADHD, CD and ODD in children and adolescents. The included studies investigated the effect of empir- ically supported treatment programs delivered by practising clinicians in non-university settings to patients referred for treatment through usual clinical routes (i.e. not actively recruited to a research project).

The aims of the present meta-analysis are:

1. To evaluate the effectiveness of CBT for children and adolescents with ADHD, CD and ODD in routine clinical care.
2. To investigate moderators of treatment outcome in effectiveness studies.
3. To investigate whether there are differences between effectiveness and efficacy studies in terms of background data or treatment data.
4. To evaluate whether there are differences in the effect of CBT for externalizing disorders between effectiveness and efficacy studies.

# Method

The protocol for this meta-analysis was pre-registered at PROSPERO with ID CRD42020147524. The meta-analysis was conducted according to the PRISMA guidelines ([Liberati et al., 2009](#_bookmark58)), and reported according to AMSTAR 2 ([Shea et al., 2017](#_bookmark74)), see online Supplement S7 and S8. Two independent raters were involved during each step of the study, except for the screening of title and abstract conducted by one rater only.

The meta-analysis was designed according to the PICOS acronym in the following way:

* Population: children and adolescents with ADHD, ODD and/or CD
* Intervention: CBT or BT delivered in routine clinical care
* Comparison: within-group change. i.e. pre vs. post-data
* Outcome: primary continuous measure and remission
* Study design: RCTs and open trials
  1. *Literature search*

Studies were identified by a systematic and comprehensive literature search of electronic databases and scanning reference lists of articles. The search was applied to Ovid MEDLINE, Embase OVID, and PsycINFO from the start of the data bases to June 11th 2019. An updated search was done May 5th 2020. The list of search terms utilized to identify potential studies were generated by all four authors in collaboration with a university librarian, who conducted the database searches. We used the following search terms to search the databases: (Cognitive therapy OR behav\* therapy OR cognitive behav\* therapy OR anger management therapy OR exposure therapy) AND (attention deficit dis- order OR attention deficit hyperactivity disorder (including the different subtypes) OR conduct disorder OR oppositional defiant disorder OR Hyperkinetic disorder OR Externali\* disorder) AND (open study OR clinical study OR community trial OR intervention study OR Pre post study OR randomized controlled trial) AND (youth OR child\* OR adolescent OR pediatric). For full search strategy for Ovid MEDLINE, Embase OVID and PsychINFO, see the online Supplement, S1.

The first author read the titles and abstract of all the papers from this

initial search to decide whether a study warranted a more detailed reading. When there was an indication of a group of patients receiving the particular cognitive-behavioral treatment in a non-university setting the full-text was retrieved. The reference lists in the retrieved articles were then checked against the database search and any other articles that might fulfill the inclusion criteria were retrieved. Although research articles were the target of the search, review articles were also examined for additional references. Key authors were searched in databases to identify additional publications but were not contacted in personal communication. In total, 730 full-text articles were considered for in- clusion. The final decision for article inclusion was made using a stricter set of inclusion and exclusion criteria detailed below. The full text ar- ticles were read by different pairs of the authors and any disagreements were resolved by consensus discussion. It was determined that 51 arti- cles could be included in the present meta-analysis.

* + 1. *Inclusion criteria*

In order to be included in the review and meta-analysis a study had

to:

1. Be published, or in press, in an English language journal.
2. Have participants diagnosed with ADHD, CD, or ODD according to DSM or ICD, or fulfilling a cut-off score on a validated parent or teacher rating scale.
3. Be testing a form of CBT, cognitive therapy (CT) or behavior therapy (BT).
4. Have participants referred for treatment through usual clinical routes.
5. Be an effectiveness study, i.e. carried out in a non-university setting such as clinical routine care or school health care.
6. Have therapists who are practicing clinicians for whom provision of service is a substantial part of the job ([Shadish, Matt, Navarro, &](#_bookmark73) [Phillips, 2000](#_bookmark73)).
7. Have a treated sample consisting of at least 10 participants.
8. Have a maximum participant age of 18.
9. Provide a measure of the primary disorder treated.
   * 1. *Exclusion criteria*
10. The study is a secondary analysis of a previously published study.
11. The study is an evaluation of a service where the results for indi- vidual disorders cannot be extracted.
12. The study is not testing a form of CBT, CT, or BT.
13. The study is testing a combination of CBT and pharmacological treatment.
    1. *Cut-off scores for applied parent or teacher rating scales*

All ADHD- and 18 of 28 CD/ODD-studies used diagnosis as inclusion criterion, whereas 10 CD/ODD-studies used cut-off scores on the following parent or teacher rating scales: Child Behavior Checklist (CBCL; [Achenbach, 1991](#_bookmark18)), Eyberg Child Behavior Inventory (ECBI;

**Identification**

**Eligibility**

**Screening**

Records after duplicates removed (n = 3471)

Additional records identiﬁed through other sources

(n = 254)

Records identiﬁed through database searching

(n = 4070)

Irrelevant records excluded (n = 2741)

Records screened (n = 3471)

Full-text articles assessed for eligibility

(n = 730)

Studies included in quantitative synthesis (meta-analysis)

(n = 51)

Studies included in qualitative synthesis (n = 51)

Full-text articles excluded (n = 679)

Participants not clinically referred Participants not diagnosed Treatment at research clinics Secondary analyses

Not testing evidence-based CBT No data on outcome measure Combination of CBT and drug tx Not targeted diagnosis

(n=144) (n=136) (n=75)

(n=75)

(n=74)

(n=43)

(n=38)

(n=30)

Dissertations, conference proceedings (n=29)

<10 participants in treatment condition (n=15) Not in english lanuage journal (n=13)

Age > 18 years (n=7)

**Fig. 1.** Flowchart of the inclusion of studies.

**Included**

[Eyberg & Pincus, 1999](#_bookmark40)), and Aversive Behavior Rate (ABR; [Patterson,](#_bookmark68) [Chamberlain, & Reid, 1982](#_bookmark68)). We transformed the various scores to percentiles and the five studies using CBCL had cut-off scores between 84 and 95, the four studies using ECBI had scores between 90 and 95, and the single study applying the behavior observation instrument ABR used a score of 90.

[Fig. 1](#_bookmark7) shows a flowchart of the inclusion of studies in the present meta-analysis. For references to included studies, see online Supplement S2, and for references to studies excluded in the meta-analyses, see online Supplement S3.

* 1. *Potential categorical moderators*

In order to include any potential categorical or continuous moder- ator in the analysis we required that at least 70% of the studies provided information on that variable. With lower proportions it is questionable if the information extracted is representative of the entire body of studies.

* + 1. *Type of study and statistical analysis*

Type of study was either RCT (when a CBT-condition was compared with some kind of control condition) or open trial (when only a CBT- condition was used in the study). Statistical analysis was categorized as intent-to-treat (ITT) if all randomized participants were included in the statistical analysis or completers if dropouts were deleted.

* + 1. *Type of treatment and conditions*

It was not possible to classify the many treatment methods used based on their names since some methods might have different names but only differ slightly from each other. Instead we categorized the target of treatment as: child, parent, or child and parent, and treatment format as: group, individual, or group and individual.

* + 1. *Parental and teacher involvement*

Degree of parental involvement was classified as *low* if parents were not present during sessions but informed about the progress of therapy, *moderate* if parents were present during some therapy sessions full-time or only part-time of all sessions, and *high* if parents were present full- time during all therapy sessions. Teacher involvement was classified as yes or no.

* + 1. *Therapist profession*

The profession that the majority of the therapists within a study belonged to was classified as clinical psychologist, social worker, or various professions.

* + 1. *Continent*

The country in which the study was carried out was categorized as North America, South America, Europe, Asia, Australia, or Africa.

A coding scheme and manual including the variables of interest were developed. The data extraction and categorizations were done inde- pendently by the first and the third author and any disagreements were solved after consensus discussion.

* 1. *Potential continuous moderators*

The following continuous measures on which at least 70% of the studies provided information were used as potential moderators: num- ber of participants in the study, percent declining participation in the study, percent attrition in the study, percent on drug treatment for the principal disorder, percent boys, mean age, pre-treatment severity (calculated as percentage of the maximum score of the rating scale applied), methodology score (see below), risk-of-bias score (see below), number of therapy sessions, and treatment intensity (hours/week). In addition, we also extracted information on a number of other variables, but these variables did not reach the 70% criterion.

* 1. *Methodological quality*
     1. *The Psychotherapy Outcome Study Methodology Rating Scale (POMRS)*

The scale consists of 22 items covering various important aspects of the methodology in psychotherapy outcome research ([O](#_bookmark66)¨ [st, 2008](#_bookmark66)). Each item is rated as 0 = poor, 1 = fair, and 2 = good, and each step has a

verbal description of one or more sentences. The total score can vary from 0 to 44 points. Since all items were not applicable to all studies the total score was recalculated as a percentage of the maximum score possible for the individual study. The internal consistency of the scale was good with a McDonald’s ω of 0.80. The inter-rater reliability of the scale (between the second and the fourth author), based on 20%

randomly selected and blindly rated studies was ICC(3,1) = 0.94 (95%

CI 0.75–0.99), which according to [Cicchetti (1994)](#_bookmark26) is excellent.

* 1. *Risk-of-bias*

The Cochrane Collaboration tool for assessing risk-of-bias ([Higgins,](#_bookmark48) [Altman, & Sterne, 2011](#_bookmark48)) was used, and the following domains were rated: random sequence generation, allocation concealment, blinding of outcome assessment, incomplete outcome data, and selective reporting. Blinding of patients and therapists cannot be used in psychotherapy studies. A high risk-of-bias in a domain was given 1 point, an unclear risk 0.5, and a low risk 0 point. Summarizing over the five domains the total score could vary between 0 and 5, with higher scores indicating higher risk-of-bias. Inter-rater reliability was assessed between the second and

the fourth author based on 20% randomly selected and blindly rated studies. This yielded an intra-class correlation, ICC(3, 1) = 0.95 (95% CI 0.76–0.99), which is also excellent.

* 1. *Effect size measures*

We extracted data on both primary and secondary measures in the studies. This was done independently by the third and the fourth au- thors, and any discrepancies were discussed to reach consensus. Since some studies used proportion of remitted participants as their primary outcome measure, whereas other studies used a continuous rating scale we decided to include both in this meta-analysis.

* + 1. *Remission*

Below follows a description of the criteria of remission used in the different studies.

*ADHD*: 11 out of 23 studies (48%) provided data and in 9 it was that the child’s post- or follow-up score was within the normal range on the primary outcome measure, whereas in 2 it was loss of the principal diagnosis.

*CD/ODD*: 12 out of 28 studies (43%) provided data and in 11 it was that the child’s post- or follow-up score was within the normal range on the primary outcome measure (for 2 of which Reliable Change Index must also be fulfilled), and in 1 it was loss of the principal diagnosis.

* + 1. *Continuous rating scales*

When a study named its primary outcome measure among rating scales we used that. If none was pinpointed we selected measures ac- cording to the following hierarchy: independent assessor or observer rating, teacher report scale, parent report scale, and child self-report scale. All but one of the ADHD-studies and all CD/ODD studies pro- vided data on a continuous rating scale. The various rating scales used for the respective studies are described in the online Supplement S4.

* + 1. *Secondary outcome measures*

Besides symptom measures we also aimed to extract data on func- tioning, broadly defined. However, only 7 out of 23 (30%) ADHD- studies and 10 out of 28 (36%) CD/ODD-studies had used such mea- sures. With these low proportions of studies providing the relevant data

it is questionable if the outcome of a meta-analysis would be represen- tative so we decided not to carry out one.

* 1. *Meta-analysis*

In order to obtain as large as possible a body of effectiveness studies we included both RCTs and open trials in the meta-analysis since within- group ES can be calculated from both types of studies. Within-group ES was calculated as (Mpre – Mpost)/SDpre according to recommendation by [Lakens (2013)](#_bookmark55), since there is good reason to assume that the in- terventions influence not only the means but also the standard de- viations. The mean ES was computed by weighting each ES by the inverse of its variance. Rate of remission, with event rate as the effect measure, was analyzed using mixed effect analysis in the subgroup analysis. In this analysis a random effects model is used to combine studies within each subgroup and a fixed effects model is used to compare subgroups and yield the overall effect. When a study presented intent-to-treat data (49%) these were used, if not completer data (51%) were used.

Before pooling the effect sizes we screened for statistical outliers, defined as being outside M ± 2SD. At the post-treatment assessment four (2.2%) of the ESs were outliers, and at follow-up assessment there was

one (2.3%). For these ESs *winsorising* ([Lipsey & Wilson, 2001](#_bookmark59)) was used by reducing outliers to the exact value of M + 2SD. The software *Comprehensive Meta-Analysis v.3* (CMA; [Borenstein, Hedges, Higgins, &](#_bookmark22)

[Rothstein, 2013](#_bookmark22)) was used for all analyses and to correct for small sample sizes Hedges’s *g* was calculated. A random effects model was used since it cannot be assumed that the ESs come from the same population.

Heterogeneity among ES’s was assessed with the *Q*- and the *I*-square statistic. The possibility of publication bias was analyzed with the trim- and-fill method of [Duval and Tweedie (2000)](#_bookmark34) as well as Egger’s regression intercept ([Egger, Smith, Schneider, & Minder, 1997](#_bookmark36)). Moderator analyses of continuous variables were carried out with meta- regression and for categorical variables with subgroup analysis using the mixed effect model.

* 1. *Efficacy studies for comparison*

In order to obtain the efficacy studies to be used in comparison of the effect of CBT in effectiveness studies we consulted the most recent evi- dence base update reviews of psychosocial treatments published in the *Journal of Clinical Child and Adolescent Psychology* for the respective disorders included in the present meta-analysis. For ADHD it was [Evans](#_bookmark38) [et al. (2014)](#_bookmark38); [Evans, Owens, Wymbs, and Ray (2018)](#_bookmark39), and for CD/ODD [McCart and Sheidow (2016)](#_bookmark61) and [Kaminski and Claussen (2017)](#_bookmark52). In some of these reviews the authors referred to earlier reviews, which we checked in order to get as comprehensive as possible a list of efficacy RCTs. From each of these reviews we listed the RCTs of some kind of cognitive behavioral treatment evaluated as well-established or prob- ably efficacious according to the criteria adopted by the Society of Clinical Child and Adolescent Psychology ([Southam-Gerow & Prinstein,](#_bookmark78) [2014](#_bookmark78)). Then we deleted those RCTs we had already included in the body of effectiveness studies. This resulted in the following number of efficacy RCTs for our comparison: ADHD 29 and CD/ODD 33, for a total of 62 trials. These references are listed in the online Supplement S5.

As for the effectiveness studies we extracted data for the primary

continuous outcome measure and remission rate, separately at post- treatment and follow-up assessment. In order to compare the two cate- gories of studies on background variables we also extracted data on mean age, proportion of boys, percent with comorbid disorders, pro- portion on medication for the principal disorders, pre-treatment severity (calculated as percent of maximum score on the continuous measure), treatment time (60 min. hours), and attrition. Other variables were not reported systematically, or not at all in a large enough proportion of studies, which precluded inclusion as a background variable.

* + 1. *Power analysis*

In the overall comparison of effectiveness and efficacy studies we have the following number of studies and treatment conditions, which is the unit of analysis: effectiveness studies ADHD 23/28, CD/ODD 28/39 for a total of 51/67; efficacy studies ADHD 29/43, CD/ODD 33/62 for a total of 62/105. The total number is then 113 studies with 172 condi- tions with an average of 46 participants per condition. According to the formulas for power analysis in meta-analyses by [Valentine, Pigott, and](#_bookmark84) [Rothstein (2010)](#_bookmark84) we would have 100% power to detect an effect size of

0.20 and 88% for an ES of 0.10, when assuming that the heterogeneity of effect sizes will be high.

# Results

* 1. *Description of the studies*
     1. *Study characteristics*

Background data for the included studies are presented in [Table 1](#_bookmark8). The majority of the 51 studies were done in North America (*n* = 25) and Europe (*n* = 23), whereas only 3 came from Australia, and none from

Africa, Asia, or South America. The total number of participants in these studies was 5295 with 3798 in CBT-conditions, 298 in other treatments,

e.g. pharmacological, and 1199 in control conditions. For ADHD it was 2223 in 23 studies (1675 in CBT-conditions and 548 in control condi- tions) and for CD/ODD 2774 in 28 studies (2123 in CBT-conditions and 651 in control conditions). The proportion of eligible participants that declined the offer of treatment was on average 12.7% (ADHD 10% and CD/ODD 15%).

There was an overall majority of boys (77.7%), both in ADHD (78.0%) and CD/ODD (77.3%). The mean age across all studies was 8.2 years (ADHD 8.7 and CD/ODD 7.9). Proportion of participants with comorbidity was reported by only 24 studies (47%), and in an unsys- tematic fashion. With that in mind, 52.9% of the participants had at least one comorbid disorder (46.7% in ADHD and 63.2% in CD/ODD). Only 33 studies (66.7%) reported what proportion of the participants was on psychotropic medication for their principal disorder at the inclusion to the respective study. The overall mean was 32.5% with a significant

difference (*t*(32) = 2.38, *p* = 0.024) between ADHD (41.6%) and CD/

ODD (19.5%), which would be expected since central stimulants are evidence based for ADHD.

* + 1. *Treatment data*

Treatment data for the included studies are presented in [Table 2](#_bookmark9). The target of treatment was the parent in 25 (49.0%), parent and child in 21 (41.2%), and child only in 5 studies (9.8%). The format of treatment was group in 25 (49.0%), individual in 22 (43.1%), and the combination of group and individual treatment in 4 studies (7.8%). The number of therapists per study was on average 11.1 (range 1–56), which indicates the number of participating therapists working at the routine clinical sites where the studies were done. Treatments were carried out over

15.0 weeks on average (12.0 for ADHD and 17.5 for CD/ODD) and the mean number of sessions was 17.3 (ADHD 17.5, CD/ODD 17.2). The intensity (hours/week) was on average 2.8 (ADHD 3.9, CD/ODD 1.8).

* 1. *Methodological data*
     1. *Methodology ratings*

The research methodology score (% of maximum possible score for the individual study) had an overall mean of 47.4 (*SD* 12.3), which corresponds to a raw score of 20.9 points. The mean for ADHD-studies

was 46.9 (*SD* 14.3) and for CD/ODD studies 47.7 (*SD* 10.8), a non- significant difference (*t*(65) = 0.26, *p* = 0.80). As could be expected the RCTs (*M* 50.8, *SD* 11.2) across the two disorders had a significantly higher methodology score (*t*(65) = 4.44, *p* < 0.001) than did the open trials (*M* 37.3, *SD* 9.5).

**Table 1**

Background data of the included studies.

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Study | Continent | Type of study | Analysis | Methodology | Severity | % Comorb. | % Decliners | N | % Males | Age range | Mean age | % on meds |
| *ADHD* |  |  |  |  |  |  |  |  |  |  |  |  |
| Abikoff, 2015 | North America | RCT | ITT | 0.682 | 0.752 | 41.5 | 0 | 164 | 73.8 | 3–4 | 3.6 | 0 |
| Boyer, 2015 | Europe | RCT | ITT | 0.682 | 0.474 | 31.4 | 0.6 | 159 | 73.5 | 12–17 | 14.4 | 78.0 |
| Breider, 2019 | Europe | RCT | ITT | 0.455 | 0.323 | 76.2 | 29.0 | 21 | 71.4 | 4–12 | 7.8 |  |
| Canu, 2011 | North America | Open trial | ITT | 0.325 | 0.719 | 100 | 0 | 16 | 88 | 4–12 | 8.0 |  |
| Curtis, 2010; 2013 | North America | Open trial | ITT | 0.250 | 0.744 |  | 22.0 | 21 | 71 | 7–10 | 7.8 | 64.0 |
| Daley, 2013 | Europe | RCT | ITT | 0.618 | 0.783 |  | 7.1 | 43 | 81.4 | 4–11 | 7.3 | 0 |
| Do¨pfner, 2004 | Europe | RCT | Compl. | 0.455 | 0.600 | 61.3 |  | 75 | 93.3 | 6–10 | 8.3 | 0 |
| Elkins, 2019 | North America | Open trial | Compl. | 0.350 | 0.770 | 52.0 |  | 67 | 76.0 | 7–10 | 8.0 | 72.0 |
| Fehlings, 1991 | North America | RCT | ITT | 0.523 | 0.563 | 0 |  | 25 | 100 | 7–13 | 9.5 | 0 |
| Hannesdottir, 2017 | Europe | Open trial | ITT | 0.341 | 0.684 |  |  | 41 | 70.7 | 8–10 | 9.2 | 78.0 |
| Heath, 2015 | North America | Open trial | Compl. | 0.375 | 0.740 |  |  | 47 | 79.1 | 7–12 | 8.0 | 74.4 |
| Jerrott, 2010 | North America | Open trial | Compl. | 0.310 | 0.693 | 72.5 |  | 57 | 84.2 | 5–13 | 10.4 | 78.0 |
| Loren. 2015 | North America | Open trial | Compl. | 0.425 |  | 27.6 |  | 241 | 75.6 | 6–12 | 8.6 |  |
| MTA-Group. 1999 | North America | RCT | Compl. | 0.738 | 0.467 | 39.9 | 4.4 | 579 | 80.3 | 7–9.9 | 8.5 | 30.7 |
| Ostberg. 2012 | Europe | RCT | Compl. | 0.310 | 0.200 | 4.3 | 9.8 | 92 | 83.6 | 7–10 | 10.9 | 80.3 |
| Pfiffner. 2013 | North America | Open trial | Compl. | 0.475 | 0.600 |  | 3.2 | 57 | 70.2 | 7–11 | 8.1 | 7.0 |
| Pfiffner. 2016; 2018 | North America | RCT | ITT | 0.477 | 0.626 | 50.4 | 10.0 | 135 | 71 | 7–11 | 8.4 | 8.9 |
| Power. 2012 | North America | RCT | ITT | 0.500 | 0.467 | 27.0 | 23.0 | 199 | 68.3 | 7–12 | 9.5 | 43.0 |
| Power. 2014 | North America | Open trial | Compl. | 0.500 | 0.667 | 36.0 | 28.0 | 72 | 69.4 | 5–10 | 7.0 | 34.7 |
| Sibley. 2018 | North America | RCT | Compl. | 0.571 | 0.467 |  | 21.8 | 218 | 74.2 | 12–15 | 13.5 | 45.9 |
| Tynan. 1999 | North America | Open trial | Compl. | 0.325 | 0.607 |  | 9.7 | 65 | 76.4 | 5–11 | 7.5 | 44.0 |
| Tynan 2004 | North America | Open trial | ITT | 0.250 | 0.641 |  | 9.5 | 67 | 82.4 | 5–11 | 7.9 | 43.1 |
| van den Hoofdakker. 2007 | Europe | RCT | Compl. | 0.477 | 0.933 | 80.9 | 2.0 | 96 | 80.9 | 4–12 | 7.4 | 50.0 |
| *CD/ODD* |  |  |  |  |  |  |  |  |  |  |  |  |
| Axberg. 2012 | Europe | RCT | ITT | 0.524 | 0.635 |  | 6.1 | 62 | 83.9 | 4–8 | 6.0 |  |
| Axelrad. 2009 | North America | Open trial | Compl. | 0.425 | 0.670 |  | 0 | 64 | 84.0 | 2–6 | 4.7 | 1.6 |
| Bjorseth. 2016 | Europe | RCT | ITT | 0.591 | 0.587 | 74.6 | 11.0 | 81 | 70.4 | 2–7 | 5.8 | 12.3 |
| Burke. 2015 | North America | RCT | ITT | 0.429 | 0.753 |  | 22.0 | 252 | 100 |  | 8.9 | 7.5 |
| Chacko. 2015 | North America | RCT | ITT | 0.409 | 0.620 |  |  | 320 | 67.8 | 7–11 | 8.9 | 33.4 |
| Connolly. 2001 | Europe | RCT | Compl. | 0.273 | 0.642 |  |  | 129 | 86.8 | 2–9 | 6.2 | 2.3 |
| Costin. 2007 | Australia | Open trial | Compl. | 0.400 | 0.662 | 76.4 |  | 112 | 80.9 | 5–13 | 9.1 |  |
| David. 2014 | Europe | RCT | Compl. | 0.333 | 0.601 |  | 11.0 | 130 | 48.5 | 4–12 | 6.2 | 0 |
| Gardner. 2006 | Europe | RCT | ITT | 0.575 | 0.606 |  | 28.1 | 76 | 73.7 | 2–9 | 5.9 |  |
| Gavita. 2012 | Europe | RCT | ITT | 0.425 | 0.766 |  | 14.2 | 97 |  | 5–18 | 9.5 |  |
| Goertz-Dorten. 2019 | Europe | RCT | ITT | 0.545 | 0.567 |  | 36.4 | 101 | 100 | 6–12 | 8.7 | 14.3 |
| Greene. 2004 | North America | RCT | Compl. | 0.477 | 0.576 | 66.0 | 36.4 | 50 | 68.1 | 4–12 | 7.2 | 59.6 |
| Hansson. 2012 | Europe | RCT | Compl. | 0.357 | 0.626 |  | 0 | 46 | 60.9 | 2–17 | 15.0 |  |
| Harrington. 2000 | Europe | RCT | ITT | 0.568 | 0.698 |  | 25.0 | 141 | 79.0 | 3–10 | 6.9 |  |
| Helander. 2018 | Europe | RCT | ITT | 0.568 | 0.705 | 67.0 | 11.7 | 120 | 73.0 | 8–12 | 9.3 |  |
| Hutchings. 2002 | Europe | RCT | ITT | 0.429 | 0.696 |  | 10.6 | 41 | 85.4 | 2–10 | 6.0 |  |
| Kolko. 2009 | North America | RCT | ITT | 0.705 | 0.773 | 76.3 | 6.8 | 139 | 84.9 | 6–11 | 8.8 | 51.1 |
| Larsson. 2009 | Europe | RCT | ITT | 0.568 | 0.623 | 35.4 | 8.0 | 136 | 79.5 | 4–8 | 6.6 |  |
| Lees. 2019 | Australia | RCT | Compl. | 0.476 | 0.539 |  | 10.6 | 126 | 68.3 | 3–7 | 5.4 |  |
| McGilloway. 2012 | Europe | RCT | Compl. | 0.525 | 0.621 |  | 4.0 | 149 | 62.8 | 2–7 | 4.8 |  |
| McNeil. 1991 | North America | Open trial | ITT | 0.625 | 0.717 | 90.0 |  | 12 | 100 | 2–7 | 4.5 | 0 |
| Nixon. 2003 | Australia | RCT | ITT | 0.523 | 0.661 |  |  | 63 | 70.4 | 3–5 | 3.9 | 0 |
| Patterson. 1982 | North America | RCT | Compl. | 0.405 |  |  | 34.5 | 19 | 68.4 | 3–11 | 6.8 |  |
| van de Wiel. 2007 | Europe | RCT | Compl. | 0.455 | 0.746 | 62.5 |  | 77 | 88.3 | 8–13 | 10.0 | 40.3 |
| van Manen. 2004 | Europe | RCT | ITT | 0.523 | 0.668 |  |  | 97 | 100 | 9–13 | 11.0 |  |
| Westermark. 2011 | Europe | RCT | ITT | 0.405 | 0.696 |  | 7.9 | 35 | 51.4 | 2–18 | 15.0 |  |
| Woltering. 2011 | North America | Open trial | ITT | 0.342 | 0.325 | 21.0 |  | 71 | 72.0 | 8–12 | 9.5 | 25.4 |
| Woltering. 2015 | North America | Open trial | Compl. | 0.289 | 0.721 |  |  | 39 | 79.5 | 8–12 | 9.5 | 25.6 |

*Note:* RCT = randomized controlled trial, ITT = intent-to-treat, Compl. = completers, Blank fields = data not provided, Severity = percentage of the maximum score on the primary outcome measure, % Comorb. = proportion having any psychiatric comorbid disorder, % Decliners = proportion of participants who were offered treatment in the study but declined (irrespective of reason), % on meds. = proportion on any psychotropic medication.

* + 1. *Risk of bias*

The risk of bias categorization is presented in the online Supplement, [Table 6](#_bookmark13). If we delete the 16 open trials from the evaluation of the first two domains we find the following proportions of a low risk-of-bias: Random sequence 83%, allocation concealment 46%, blinding of as- sessors 29%, incomplete data 59%, and selective reporting 84%. A high risk-of-bias was found in: random sequence 6%, allocation concealment 9%, blinding of assessors 51%, incomplete data 35%, and selective reporting 0%. Thus, it was much more common that these studies had a low than a high risk-of-bias regarding the evaluated domains.

In order to score the risk-of-bias a low risk was given 0, an unclear risk 0.5, and a high risk 1 point, which means that the total score could vary from 0 to 5 points. The total mean score was 1.87 (*SD* 1.52) and the studies within each disorder had the following mean (*SD*): ADHD 2.11

(1.67) and CD/ODD 1.69 (1.41), a non-significant difference (*t*(65) = 1.10). As was found for methodological quality the RCTs (*M* 1.23, *SD* 1.15) had a significantly lower risk-of-bias (*t*(65) = 8.38, *p* < 0.001) than the open trials (*M* 3.74, *SD* 0.75).

* 1. *Meta-analysis*
     1. *Attrition*

Forty-seven of the studies (92.2%) provided information on the number of participants who dropped out of the treatments. Using treatment condition (k = 63) as the unit of analysis the overall attrition

rate was 14.4% (95% CI 11.5–18.0, *z* = 13.36, *p* < 0.0001). ADHD-

studies (k = 27) had a dropout rate of 12.0% and CD/ODD-studies (k

= 36) 16.4%, which was not a significant difference (Qbetween (1 df) =

1.72, *p* = 0.19).

* + 1. *Primary continuous measure*

[Table 3](#_bookmark10) displays the mean effect sizes of the primary continuous measure for all studies at post-treatment and follow-up assessment, which was done on average 10.8 months after the end of therapy. At post-treatment the average ES was large (*g* = 0.91) and significantly different from zero. Heterogeneity was significant and large as indicated by the Q- and *I2*-values. CD/ODD-studies (*g* = 0.98) had a somewhat

higher ES than ADHD-studies (*g* = 0.80) but the difference was not

significant.

At follow-up the mean ES (*g* = 1.01) was insignificantly higher than at post-treatment (Qbetween (1 df) = 0.75, *p* = 0.39), but also significantly

heterogeneous. The comparison between disorders was still not signifi- cant. Thus, the treatment effects seen at post-assessment were main- tained at follow-up.

*Publication Bias.* The possibility of publication bias was investigated, using Duval and Tweedie’s trim-and-fill method and Eggers regression intercept. Regarding the post-treatment data the trim-and-fill method suggested trimming 17 conditions to the left of the mean which would lower the ES from 0.91 to 0.72 (95% CI 0.66–0.75). The regression

intercept had a significant *t*(60) = 4.66, *p* < 0.01, indicating that pub-

lication bias probably is an issue for this body of studies.

* + 1. *Remission*

The remission rates at post-treatment and follow-up are presented in [Table 4](#_bookmark11). At post-treatment 43.8% of the participants had remitted, which was significantly heterogeneous. A subgroup analysis indicated that CD/ODD-studies had about 10 percentage points higher remission rate than ADHD-studies but the difference was not significant.

At follow-up assessment the overall remission rate had increased to 51.4% but the difference to the post-assessment rate was not significant (Qbetween (1 df) = 1.38, *p* = 0.24). The subgroup analysis also showed

that the disorders did not differ significantly from each other. Thus, as

was the case for continuous measures the effect sizes seen at post- assessment were maintained at follow-up.

Publication Bias. Eggers regression intercept did not yield a signifi- cant *t*-value (1.15, *p* = 0.26). The trim-and-fill method suggested trim- ming 7 studies, which would have reduced the remission rate from

43.8% to 37.4% (95% CI 30.6–44.8). Thus, publication bias does not seem to be an important problem regarding the remission rate.

* + 1. *Moderator analyses*

As the mean ES was significantly heterogeneous we followed up with moderator analyses. [Table 5](#_bookmark12) shows the results for categorical variables using subgroup analysis with the results for effect size in the left column and remission rate in the right. There was no significant difference be- tween RCTs and open trials. Regarding statistical analysis studies with intent-to-treat analysis yielded non-significantly lower ES than studies using completer analysis. There was no significant differences depend- ing on target or format of therapy, degree of parental involvement, teacher involvement or therapist profession. The continent at which the study was carried out was associated with a significant difference; studies from Australia had the highest ES. However, since the ES for Australia is made up of only 5 conditions this difference should be interpreted with caution.

Continuous variables on which at least 70% of the studies provided information were analyzed with the meta-regression module in the CMA program using the fixed effects analysis (see [Table 6](#_bookmark13)). Since 11 variables were included we used the Holm-Bonferroni correction (see [Jaccard &](#_bookmark49) [Guilamo-Ramos, 2002](#_bookmark49)). For effect size there were four negative mod- erators: mean age of the participants in the condition, proportion of participants on drug treatment for the principal disorder, percent attri- tion, and risk-of-bias, i.e. the higher the values on these moderators the

lower the ES. There was one positive moderator, pre-treatment severity,

i.e. higher severity was associated with higher ES. For remission rate there were three negative moderators: mean age, percent on drug treatment, and attrition rate. In addition, there was one positive moderator: number of sessions, i.e. higher number of sessions was associated with higher remission rate.

* 1. *Efficacy-effectiveness comparison*

In the following section ([Tables 7-10](#_bookmark14)) we compare data for the effectiveness studies reviewed so far with data for the efficacy studies obtained from the evidence base update reviews on ADHD and CD/ODD published in the *Journal of Clinical Child and Adolescent Psychology*.

* + 1. *Background and treatment variables*

[Table 7](#_bookmark14) presents a comparison between efficacy and effectiveness studies on some background variables and one treatment variable. Since there are 7 variables and 14 *t*-tests the Holm-Bonferroni correction was used. The only significant difference emerged on percent comorbidity for the CD/ODD studies where effectiveness studies had a higher pro- portion than efficacy studies. However, this result must be treated with caution since only 43.8% of the effectiveness and 46.7% of the efficacy studies provided data on this variable. There were no significant dif- ference between the two types of studies regarding mean age, proportion of participants on medication for the principal disorder, and percent attrition. If we use a *p*-value of 0.05 effectiveness studies had higher proportion of boys in both disorders. a higher pre-treatment severity in CD/ODD, and a lower treatment time in ADHD-studies, but these dif- ferences can only be regarded as trends since they don’t fulfill the Holm- Bonferroni correction criterion. Thus, judging from the background treatment variables which could be extracted the effectiveness studies do not consist of participants who are easier to treat than do the efficacy studies.

* + 1. *Effect size on primary outcome measure*

[Table 8](#_bookmark15) presents the subgroup analyses comparing the within-group effect size (Hedges’ *g*) for effectiveness and efficacy studies within each disorder. Neither at post-treatment assessment (upper part) nor at follow-up (lower part) were there any significant differences between the two types of studies. For both types the ESs were significant and the effects were maintained, or somewhat higher, at follow-up, which was done on average 9.9 months after post-assessment for ADHD- and 11.3

months for CD/ODD-studies (*t*(63) = 0.45, *p* = 0.66).

* + 1. *Remission*

[Table 9](#_bookmark16) contains subgroup analyses comparing the remission rates at post-treatment and follow-up assessment. Neither the post-treatment (upper part of [Table 9](#_bookmark16)) nor the follow-up (lower part) remission rates differed between study types for ADHD or CD/ODD. When combining the remission rates for the different disorders the post-treatment means were 43.8% and 46.1% for effectiveness and efficacy studies, respec- tively. The corresponding follow-up means were 51.4% and 53.7%, which means that the increase from post- to follow-up assessment was

7.6 percentage points for both categories of studies.

* + 1. *Comparison of RCTs only*

Since the outcomes presented in [Tables 8 and 9](#_bookmark15) might have been unduly influenced by open trials we repeated the analyses using only RCT effectiveness studies. [Table 10](#_bookmark17) summarizes the results across dis- orders and there were no significant differences between effectiveness and efficacy studies regarding effect size or remission rates at post- treatment or follow-up assessment.

# Discussion

The primary aim of the present meta-analysis was to investigate the

**Table 2**

Treatment data of the included studies.

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Study | Treatment program | CBT method | Profession | Tx target | Tx form. | PMT | Parent involvem. | Teacher involvem. | No. of sessions | Hrs/ week | %  Attrition | F-  Up |
| *ADHD* |  |  |  |  |  |  |  |  |  |  |  |  |
| Abikoff. 2015 | New Forrest Parenting Package | PBT (with | Psychologist | P | I | Y | High | N | 8 | 1 | 9.0 | 6 |
|  | (NFPP) | child) + TI |  |  |  |  |  |  |  |  |  |  |
|  | Helping the non-compliant child | PBT (with | Psychologist | P | I | Y | High | N | 8 | 1 | 6.3 | 6 |
|  | (HNC) | child) |  |  |  |  |  |  |  |  |  |  |
| Boyer, 2015 | Plan my Life (PML) | TI | Psychologist | C | I | N | Moderate | N | 10 | 1.1 | 4.8 | 3 |
|  | Solution focused therapy (SFT) | TI | Psychologist | C | I | N | Low | N | 10 | 1.1 | 5.3 | 3 |
| Breider, 2019 | Blended parent training | PBT | Psychologist | P | G | Y | High | N | 16 | 0.9 | 90.9 | 0 |
|  | Face-to-face parent training | PBT | Psychologist | P | I | Y | High | N | 4 | 1 | 40.0 | 0 |
| Canu, 2011 | Defiant children | PBT | Psychologist | P | G | Y | High | N | 6 | 1.5 | 0 | 0 |
| Curtis, 2010; 2013 | Family Stars | PBT + ChBT | Various | C, P | I | Y | Moderate | N | 10 | 1 | 0 | 0 |
| Daley, 2013 | New Forest Parenting Package | PBT | NA (Self- | P | I | Y | High | N | 1 | 0.3 | 12.5 | 0 |
|  | Self-help |  | help) |  |  |  |  |  |  |  |  |  |
| Do¨pfner, 2004 | Behavior therapy + | PBT + ChBT |  | C, P | I | Y | Moderate | Y | 18 | 3 | 2.9 | 0 |
|  | Psychoeducation | + TT |  |  |  |  |  |  |  |  |  |  |
| Elkins, 2019 | Family Stars | PBT + ChBT |  | C, P | I | Y | Moderate | N | 10 | 1 | 17.9 | 0 |
| Fehlings, 1991 | CBT | ChBT |  | C | I | N | Moderate | N | 20 | 1.8 | 0 | 5 |
| Hannesdottir, 2017 | OutSMART | ChBT + TI | Psychologist | C | G | N | Low | N | 10 | 4 | 6.3 | 3 |
|  | Parent training program | PBT | Psychologist | P | G | Y | High | N | 7 | 2 | 18.2 | 3 |
| Heath, 2015 | Family Stars | PBT + ChBT | Psychologist | C, P | I | Y | Moderate | N | 10 | 1 | 8.5 | 0 |
| Jerrott, 2010 | CBT | PBT + ChBT | Various | C, P | G | Y | Moderate | Y | 32 | 2 |  | 39 |
| Loren, 2015 | Parent group program | PBT | Psychologist | P | G | Y | High | N | 8 | 2 | 31.5 | 0 |
| MTA-Group, 1999 | Behavioral treatment | PBT + ChBT |  | C, P | G, I | Y | Moderate | Y | 91 | 6.9 | 2.1 | 82 |
|  |  | + TT |  |  |  |  |  |  |  |  |  |  |
| Ostberg, 2012 | Strategies in Everyday Life (SEL) | PBT + TT |  | P | G | Y | Moderate | Y | 18 | 2.8 | 16.7 | 3 |
| Pfiffner, 2013 | Collaborative Life Skills Program | PBT + ChBT | Various | C, P | G | Y | Moderate | Y | 22 | 1.5 | 5.0 | 0 |
|  | (CLS) | + TT |  |  |  |  |  |  |  |  |  |  |
| Pfiffner, 2016; | Collaborative Life Skills Program | PBT + ChBT | Various | C, P | G | Y | Moderate | Y | 23 | 1.5 | 0 | 7 |
| 2018 | (CLS) | + TT |  |  |  |  |  |  |  |  |  |  |
| Power, 2012 | Family School Success (FSS) | PBT + ChBT | Psychologist | P | G, I | Y | Moderate | Y | 12 | 1.3 | 4.3 | 3 |
|  |  | + TT |  |  |  |  |  |  |  |  |  |  |
| Power, 2014 | Partnering to Achieve School | PBT + TT | Psychologist | P | I | Y | High | Y | 9 | 0.5 | 24.0 | 0 |
|  | Success (PASS) |  |  |  |  |  |  |  |  |  |  |  |
| Sibley, 2018 | High Intensity skills-based | PBT + ChBT | Various | C, P | G | Y | Moderate | N | 40 | 45 |  | 8 |
|  | summer intervention |  |  |  |  |  |  |  |  |  |  |  |
|  | Low Intensity skills-based | PBT + ChBT | Various | C, P | G | Y | Moderate | N | 8 | 1.5 |  | 8 |
|  | summer intervention |  |  |  |  |  |  |  |  |  |  |  |
| Tynan, 1999 | Parent training and child social- | PBT + ChBT |  | C, P | G | Y | Moderate | N | 8 | 2 | 15.4 | 0 |
|  | skills training |  |  |  |  |  |  |  |  |  |  |  |
| Tynan 2004 | Parent training and child social- | PBT + ChBT |  | C, P | G | Y | Moderate | N | 8 | 2 | 23.9 | 0 |
|  | skills training |  |  |  |  |  |  |  |  |  |  |  |
| van den | Behavioral parent training | PBT | Psychologist | P | G | Y | High | N | 12 | 1.2 | 12.5 | 6 |
| Hoofdakker, 2007 |  |  |  |  |  |  |  |  |  |  |  |  |
| *CD/ODD* |  |  |  |  |  |  |  |  |  |  |  |  |
| Axelrad, 2009 | Brief Behavioral Intervention | PBT |  | P | I | Y | High | N | 5 | 0.83 | 14.1 | 12 |
|  | (BBI) |  |  |  |  |  |  |  |  |  |  |  |
| Bjorseth, 2016 | Parent-Child Interaction Therapy | PBT (with | Psychologist | P | I | Y | High | N | 21.1 | 1 | 12.5 | 12 |
|  | (PCIT) | child) |  |  |  |  |  |  |  |  |  |  |
| Burke, 2015 | Stop Now And Plan (SNAP) | PBT + ChBT |  | C, P | G, I | Y | Moderate | N | 24 | 2 |  | 12 |
| Chacko, 2015 | Multiple Family Group (MFG) | PBT (with | Various | C, P | G | Y | High | N | 16 | 1.8 | 34.2 | 12 |
|  |  | child) + FI |  |  |  |  |  |  |  |  |  |  |
| Connolly, 2001 | Incredible Years (IY) | PBT |  | P | G | Y | High | N | 12 | 1.8 | 50.0 | 6 |
|  | Incredible Years (IY) + Child | PBT |  | P | G | Y | Moderate | N | 12 | 1.8 | 45.1 | 6 |
|  | Intervention |  |  |  |  |  |  |  |  |  |  |  |
| Costin, 2007 | Skilled Parenting Program | PBT |  | P | G | Y | High | N | 8 | 2 | 5.3 | 5 |
| David, 2014 | Rationale Positive Parenting | PBT | Various | P | G | Y | High | N | 10 | 1.5 | 20.0 | 1 |
|  | Program |  |  |  |  |  |  |  |  |  |  |  |
|  | CBT Standard | PBT | Various | P | G | Y | High | N | 10 | 1.5 | 14.9 | 1 |
| Gardner, 2006 | Incredible Years (IY) | PBT | Various | P | G | Y | High | N | 14 | 2 | 11.6 | 12 |
| Gavita, 2012 | Short cognitive-behavioral group | PBT |  | P | G | Y | High | N | 4 | 1.3 | 21.4 | 3 |
|  | parenting program |  |  |  |  |  |  |  |  |  |  |  |
| Goertz-Dorten, | Treatm. Progr. for Children with | ChBT | Psychologist | C | I | N | Low | N | 24 | 0.8 | 0 | 0 |
| 2019 | Aggr. Beh. (THAV) |  |  |  |  |  |  |  |  |  |  |  |
| Greene, 2004 | Collaborative Problem Solving | PBT | Psychologist | P | I | Y | High | N | 11 | 1 | 6.7 | 4 |
|  | (CPS) |  |  |  |  |  |  |  |  |  |  |  |
|  | Parent Training (PT) | PBT | Psychologist | P | I | Y | High | N | 10 | 1 | 5.0 | 4 |
| Hansson, 2012 | Multidimensional Treatment | MCT | Various | C, P | I | Y | Moderate | Y |  |  | 15.8 | 12 |
|  | Foster Care (MTFC) |  |  |  |  |  |  |  |  |  |  |  |
| Harrington, 2000 | Parent training – community | PBT |  | P | G | Y | High | N |  |  | 48.6 | 9 |
|  | Parent training – hospital | PBT |  | P | G | Y | High | N |  |  | 40.6 | 9 |
| Helander, 2018 | KOMET | PBT | Psychologist | P | G | Y | High | N | 11 | 2.5 | 8.0 | 0 |
|  |  | PBT + ChBT | Psychologist | P | G | Y | Moderate | N | 26 | 4.3 | 17.2 | 0 |

(*continued on next page*)

**Table 2** (*continued* )

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Study | Treatment program | CBT method | Profession | Tx target | Tx form. | PMT | Parent involvem. | Teacher involvem. | No. of sessions | Hrs/ week | %  Attrition | F-  Up |
|  | KOMET + The Coping Power |  |  |  |  |  |  |  |  |  |  |  |
|  | Program |  |  |  |  |  |  |  |  |  |  |  |
| Hutchings, 2002 | Standard parent training | PBT | Various | P | I | Y | High | N |  |  | 31.6 | 0 |
|  | Intensive parent training | PBT (with | Psychologist | P | I | Y | High | N |  |  | 4.5 | 0 |
|  |  | child) |  |  |  |  |  |  |  |  |  |  |
| Kolko, 2009 | Modular treatment – community | MCT (comm) | Social | P, C | I | Y | Moderate | Y | 21.6 | 1.4 | 2.9 | 36 |
|  |  |  | worker |  |  |  |  |  |  |  |  |  |
|  | Modular treatment – clinic | MCT (clinic) | Social | P, C | I | Y | Moderate | Y | 15.4 | 1.1 | 20.9 | 36 |
|  |  |  | worker |  |  |  |  |  |  |  |  |  |
| Larsson, 2009 | Incredible Years (IY) | PBT | Various | P | G | Y | High | N | 13 | 2 | 4.3 | 12 |
|  | IY + Dinosaur School Program | PBT + ChBT | Various | P, C | G | Y | Moderate | N | 31 | 3.4 | 0 | 12 |
| Lees, 2019 | IY | PBT | Various | P | G | Y | High | N | 13 | 2 | 34.9 | 6 |
|  | IY + Home Parent Support (HPS) | PBT (+ HPS) | Various | P | G, I | Y | High | N | 23 | 2.8 | 17.5 | 6 |
| McGilloway, 2012 | IY | PBT | Various | P | G | Y | High | N | 14 | 2 | 4.0 | 0 |
| McNeil 1991 | Parent-Child Interaction Therapy | PBT (with | Psychologist | P | I | Y | High | N | 14 | 1 | 0 | 18 |
|  | (PCIT) | child) |  |  |  |  |  |  |  |  |  |  |
| Nixon, 2003 | PCIT standard | PBT (with | Psychologist | P | I | Y | High | N | 12 | 1.3 | 22.7 | 6 |
|  |  | child) |  |  |  |  |  |  |  |  |  |  |
|  | PCIT abbreviated | PBT (with | Psychologist | P | I | Y | High | N | 10 | 0.9 | 13.0 | 6 |
|  |  | child) |  |  |  |  |  |  |  |  |  |  |
| Patterson, 1982 | Parent Management Training | PBT (with | Various | P | I | Y | High | N | 17 | 1 | 0 | 0 |
|  | Oregon (PMTO) | child) |  |  |  |  |  |  |  |  |  |  |
| van de Wiel, 2007 | Utrecht Coping Power Program | PBT + ChBT | Psychologist | P, C | G | Y | Moderate | N | 38 | 1 | 10.5 | 0 |
|  | (UCPP) |  |  |  |  |  |  |  |  |  |  |  |
| van Manen, 2004 | Social Cognitive Intervention | ChBT | Psychologist | C | G | N | Low | N | 11 | 1.2 | 0 | 12 |
|  | Program (SCIP) |  |  |  |  |  |  |  |  |  |  |  |
|  | Social Skills Training Program | ChBT | Psychologist | C | G | N | Low | N | 11 | 1.2 | 0 | 12 |
|  | (SST) |  |  |  |  |  |  |  |  |  |  |  |
| Westermark, 2011 | Multidimensional Treatment | MCT | Various | P, C | I | Y | Moderate | Y | 40 | 1 | 10.0 | 12 |
|  | Foster Care (MTFC) |  |  |  |  |  |  |  |  |  |  |  |
| Woltering, 2011 | Stop Now And Plan (SNAP) | PBT + ChBT | Various | C, P | I | Y | Moderate | N | 14 | 3 |  | 0 |
| Woltering, 2015 | Stop Now And Plan (SNAP) | PBT + ChBT | Various | C, P | I | Y | Moderate | N | 14 | 3 |  | 12 |

*Note:* PBT = parent behavior therapy, TI = training intervention, ChBT = child behavior therapy, TT = teacher training, Tx target = target of treatment, C = child, P = parent, Tx form. = treatment format, G = group, I = individual, PMT = parent management training, Y = yes, N = no, Parent involvem. = degree of parental involvement, % Attrition = proportion dropping out of those participating in at least one session, F-Up = follow up in months.

effectiveness of CBT for ADHD, CD and ODD for children and adoles- cents in routine clinical care. The overall within-group effect size across the disorders was large and statistically significant, thereby demon- strating that the treatments are effective in real-world settings. The re- sults showed that 44% of children and adolescents with externalizing disorders achieved remission after treatment in routine clinical care and that more than half of the participants were remitted at follow-up. Furthermore, the results showed that 86% of the participants with externalizing disorders completed CBT. The completion rate is slightly higher than what was reported in a review of treatment effectiveness studies for psychological problems in children and adolescents ([Lee](#_bookmark57) [et al., 2013](#_bookmark57)), which found that 75% of participants completed parent training for disruptive disorders in real world settings. [Lee et al. (2013)](#_bookmark57) included 13 studies and reported significant effect sizes across studies for disruptive disorders, concluding that positive results from parent training can be obtained in a range of treatment settings. Our results extends the findings of [Lee et al. (2013)](#_bookmark57) by including 51 studies and providing meta-analytical evidence of treatment outcome as well as updated remission and attrition rates for CBT for child and adolescent externalizing disorders in routine care. The results are encouraging and suggest that CBT for externalizing disorders are effective in routine clinical care, acceptable to clinically referred participants, and that a majority of the patients achieves remission at follow-up.

Regarding the magnitude of the effect size for ADHD (*g* = 0.80), our

finding is similar to that of [Fabiano et al. (2015)](#_bookmark42) who conducted a sys- tematic review of published meta-analyses on psychosocial treatments for ADHD. Of note, the meta-analyses were largely based on studies from university settings. [Fabiano et al. (2015)](#_bookmark42) concluded that the effect sizes across meta-analyses were generally moderate to large for behavioral outcomes. However, direct comparisons of effect size estimates to spe- cific studies are not applicable since many previous meta-analyses have

investigated subtypes of cognitive behavioral interventions including school-based treatments ([DuPaul & Eckert, 1997](#_bookmark31); [DuPaul, Eckert, &](#_bookmark32) [Vilardo, 2012](#_bookmark32)) and parent training ([Charach et al., 2013](#_bookmark25); [Corcoran &](#_bookmark29) [Dattalo, 2006](#_bookmark29); [Lee, Niew, Yang, Chen, & Lin, 2012](#_bookmark56); [Zwi, Jones, Thor-](#_bookmark88) [gaard, York, & Dennis, 2011](#_bookmark88)). There are meta-analyses that have included a wider range of cognitive behavioral treatments for ADHD, but most of them have excluded open trials and reported between-group ES only ([Klassen, Miller, Raina, Lee, & Olsen, 1999](#_bookmark54); [Sonuga-Barke et al.,](#_bookmark77) [2013](#_bookmark77); [Van der Oord, Prins, Oosterlaan, & Emmelkamp, 2008](#_bookmark85)) making comparisons to the current ES estimates difficult. However, a meta- analysis by [Fabiano et al. (2009)](#_bookmark41) was similar in study inclusion criteria and effect size calculation. That meta-analysis investigated the effect of a wide range of behavioral interventions for ADHD and included both RCTs and open trials. Unlike the present meta-analysis they included studies regardless of their status as efficacy or effective- ness studies. Interestingly, they found a within-group effect size of 0.70, which is only slightly lower than the effect size for ADHD in the present study.

Most meta-analyses on treatment of disruptive behavior disorders have also reported between-group ES only, typically demonstrating small to moderate effect sizes ([Bakker et al., 2017](#_bookmark19); [Erford, Paul, Oncken,](#_bookmark37) [Kress, & Erford, 2014](#_bookmark37); [McCart, Priester, Davies, & Azen, 2006](#_bookmark62)). How- ever, two meta-analyses have investigated the effect of psychosocial treatments for conduct problems and included both RCTs and open trials and reported large within-group effect sizes of 0.95 ([Fossum, Hande-](#_bookmark44) [gård, Martinussen, & Mørch, 2008](#_bookmark44)) and 1.05 ([Fossum, Handegård,](#_bookmark45) [Adolfsen, Vis, & Wynn, 2016](#_bookmark45)). Considering that these two meta-analyses

included a large number of studies from research settings it is encour- aging that the ES for CD/ODD in the present meta-analysis (*g* = 0.98) was similar to those reported by [Fossum et al. (2016, 2008)](#_bookmark45). Thus, the

current results are congruent with previous meta analyses, which have

Within-group effect size (Hedges’ *g*) for all studies (RCTs and open trials) divided by disorder with treatment condition as unit of analysis.

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Disorder | k | *g* | 95% CI | z-value | Q-value | *p*-value *I2* | |
| *Post-treatment*  Both disorders | 65 | 0.91 | 0.81–1.00 | 18.27a | 208.96a |  | 69.4 |
| ADHD | 26 | 0.80 | 0.67–0.93 | 11.76a | 3.50\* | 0.06 |  |
| CD-ODD | 39 | 0.98 | 0.85–1.12 | 14.31a |  |  |  |
| *Follow-up*  Both disorders | 45 | 1.01 | 0.87–1.14 | 15.00a | 188.38a |  | 76.6 |
| ADHD | 14 | 0.88 | 0.70–1.07 | 9.52a | 1.93\* | 0.17 |  |
| CD-ODD | 31 | 1.06 | 1.89–1.24 | 11.87a |  |  |  |

Note: k = number of treatment conditions. a *p* < 0.0001. \* Comparison between the disorders.

**Table 4**

Rates of remission for all studies divided by disorder.

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Disorder | k | % | 95% CI | z-value† | Q-b | *p*-value *I2* | |
| *Post-treatment*  Both disorders | 31 | 43.8 | 36.5–51.4 | 1.61 | 214.34a |  | 86.0 |
| ADHD | 12 | 37.7 | 26.7–50.2 | 1.93 | 1.51\* | 0.22 |  |
| CD-ODD | 19 | 47.6 | 38.0–57.3 | 0.49 |  |  |  |
| *Follow-up*  Both disorders | 17 | 51.4 | 41.3–61.3 | 0.26 | 106.44a |  | 85.0 |
| ADHD | 5 | 38.2 | 21.8–57.9 | 1.18 | 2.68\* | 0.10 |  |
| CD-ODD | 12 | 56.8 | 46.5–66.5 | 1.29 |  |  |  |

Note: k = number of treatment conditions. a *p* < 0.0001. † Test if significantly different from 50%. \* Comparison between the two disorders.

reported within-group ES, but slightly higher than previous meta ana- lyses on disruptive behavior disorders which have only reported between-group ES.

The fact that the effect size for CD/ODD (*g* = 0.98) was slightly higher than for ADHD (*g* = 0.80) is interesting. It can be argued that this finding is contrary to expectations, considering that the comprehensive

review of meta-analyses for ADHD ([Fabiano et al., 2015](#_bookmark42)) reported moderate to large effect sizes for behavioral outcomes, whereas meta- analyses on disruptive behavior disorders have reported small to mod- erate effect sizes. However, the findings from a meta-analysis of CBT for externalizing disorders ([Battagliese et al., 2015](#_bookmark20)) found a moderate and statistically significant effect size, and in accordance with our findings, they found better outcomes for CD and ODD than for ADHD, with large effect sizes for CD and ODD and moderate effect sizes for ADHD.

Parents received treatment in 90% of the effectiveness studies; either as the only recipient (49%) or together with their child (41%). Parent management training is well-established as a stand-alone treatment for ADHD ([Evans et al., 2018](#_bookmark39)) and childhood disruptive behavior ([Kaminski](#_bookmark52) [& Claussen, 2017](#_bookmark52)) and as part of multicomponent treatments for ado- lescents with disruptive behavior ([McCart & Sheidow, 2016](#_bookmark61)). Further- more, some studies suggest that parenting behavior mediates the effect

**Table 5**

Subgroup analysis of the effect size and remission for all studies at post- treatment.

Effect size Remission

Variable k *g* 95% CI k *%* 95% CI

*Type of study* (Qb = 0.02, *p* = 0.90)\* (Qb = 0.006, *p* = 0.94)\*

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| RCT | 50 | 0.90 | 0.79–1.02 | 25 | 43.9 | 35.2–53.1 |
| Open Trial | 15 | 0.92 | 0.77–1.06 | 6 | 43.5 | 35.7–51.5 |

*Statistical analysis* (Qb = 1.15, *p* = 0.29)\* (Qb = 3.07, *p* = 0.08)\* Intent-to-treat 33 0.86 0.73–0.99 16 38.4 27.6–50.4

Completers 32 0.97 0.82–1.12 15 50.2 44.6–55.8

*Target of treatment* (Qb = 0.46, *p* = 0.80)\* (Qb = 1.51, *p* = 0.47)\* Child 7 0.82 0.56–1.09 5 32.4 11.5–63.9

Parent 37 0.91 0.77–1.06 15 48.9 38.9–59.1

Child and Parent 21 0.93 0.77–1.08 11 42.0 32.0–52.8

*Treatment format* (Qb = 3.70, *p* = 0.16)\* (Qb = 0.80, *p* = 0.37)\* Group 32 0.81 0.68–0.94 19 44.9 36.6–53.6

Individual 29 0.98 0.84–1.11 11 37.5 25.2–51.6

Group + Individual 4 1.11 0.62–1.59 –

*Parental involvement* (Qb = 1.25, *p* = 0.54)\* (Qb = 0.24, *p* = 0.89)\*

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| of parent training ([Forehand, Lafko, Parent, & Burt, 2014](#_bookmark43)). Accordingly, Low | | 5 | 0.84 | 0.46–1.22 | 4 | 42.4 | 17.5–71.9 |
| it might be expected that higher parental involvement would positively Moderate | | 27 | 0.86 | 0.72–1.00 | 12 | 42.1 | 33.7–51.1 |
| affect outcome. However, whether the treatment targeted parents, High | | 33 | 0.97 | 0.82–1.12 | 15 | 45.9 | 34.0–58.2 |
| children or both, or the degree of parental involvement did not moderate *Teacher involvement* (Qb = 0.52, *p* = 0.48)\* (Qb = 0.88, *p* = 0.35)\* | | | | | | | |
| effect size significantly. These results are in congruence with previous Yes | | 12 | 0.84 | 0.62–1.05 | 7 | 48.2 | 42.0–54.5 |
| studies specifically assessing the potential effect of involving parents or No | | 53 | 0.93 | 0.82–1.04 | 24 | 42.7 | 33.6–52.4 |
| children in treatment, which have typically not found significant effects *Therapist profession* (Qb = 2.51, *p* = 0.29)\* (Qb = 1.74, *p* = 0.42)\* | | | | | | | |
| ([Corcoran & Dattalo, 2006](#_bookmark29); [Lundahl, Risser, & Lovejoy, 2006](#_bookmark60); [Thulin,](#_bookmark82) Psychologist | | 25 | 0.90 | 0.74–1.07 | 9 | 37.2 | 21.2–56.5 |
| [Svirsky, Serlachius, Andersson, & O](#_bookmark82)¨ [st, 2014](#_bookmark82)). It might be that the con- Social worker | | 7 | 0.79 | 0.55–1.03 | 6 | 50.6 | 43.9–57.2 |
| Various professions | | 16 | 1.06 | 0.83–1.30 | 7 | 52.4 | 30.0–73.8 |
| tent of the parent intervention is of greater importance than the quan-  tity. For instance, [Kaminski, Valle, Filene, and Boyle (2008)](#_bookmark53) found that *Continent* (Qb = 6.77, *p* = 0.04)\* (Qb = 2.36, *p* = 0.31)\*  Australia 5 1.37 0.99–1.75 4 59.9 37.6–78.8 | | | | | | | |
| treatment programs with components that focused on increasing posi- | North America | 28 | 0.91 | 0.77–1.05 | 11 | 42.7 | 32.0–54.3 |
| tive parent–child interactions and emotional communication skills, | Europe | 32 | 0.83 | 0.69–0.97 | 16 | 40.6 | 30.6–51.4 |

teaching parents to use time out and the importance of parenting con- sistency, and requiring parents to practice new skills with their children during parent training sessions were associated with better outcomes. On the other hand, they found that treatments that focused on teaching parents problem solving, teaching parents to promote children’s

Note: k = number of treatment conditions, Qb = Q between subgroups. \* The statistic in parenthesis tests if the subgroups within the individual category differ significantly from each other.

Meta-regression analysis of the effect size and remission for all studies at post-treatment.

|  |  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | | | Effect Size |  |  |  |  | Remission |  | | | |
|  | Variable |  | k | Point | *z*-value | *p*-value |  | k | Point | *z*-value | *p*-value |  |
|  | Mean age |  | 65 | 0.053 | 5.53 | 0.00001\* |  | 31 | 0.173 | 6.10 | 0.00001\* |  |
|  | % drug treatment |  | 40 | 0.006 | 4.84 | 0.00001\* |  | 19 | 0.018 | 6.64 | 0.00001\* |  |
|  | Attrition |  | 59 | 0.010 | 4.91 | 0.00001\* |  | 30 | 0.027 | 6.72 | 0.00001\* |  |
|  | Pre-treatment severity |  | 64 | 0.643 | 3.05 | 0.0023\* |  | 30 | 0.337 | 0.51 | 0.61 |  |
|  | Risk-of-bias |  | 65 | 0.045 | 2.81 | 0.0047\* |  | 31 | 0.064 | 1.65 | 0.099 |  |
|  | Percent boys |  | 64 | 0.007 | 2.81 | 0.0049 |  | 29 | 0.008 | 1.60 | 0.11 |  |
|  | Intensity of treatment |  | 60 | 0.009 | 2.76 | 0.006 |  | 29 | 0.043 | 2.65 | 0.008 |  |
|  | # of sessions |  | 60 | 0.004 | 2.71 | 0.007 |  | 29 | 0.026 | 3.10 | 0.002\* |  |
|  | 1 outlier deleted |  | 59 | 0.006 | 1.85 | 0.064 |  | – |  |  |  |  |
|  | Percent declining |  | 47 | 0.007 | 2.26 | 0.024 |  | 23 | 0.065 | 0.64 | 0.52 |  |
|  | Methodology score |  | 65 | 0.037 | 1.50 | 0.13 |  | 31 | 0.804 | 1.72 | 0.086 |  |
|  | # of participants |  | 65 | 0.002 | 2.95 | 0.003 |  | 31 | 0.009 | 8.35 | 0.0001 |  |
|  | 1 outlier deleted |  | 64 | 0.000 | 0.11 | 0.91 |  | 30 | 0.014 | 3.34 | 0.008 |  |

Note: k = number of treatment conditions, Point = point estimate. \* Significant using the Holm-Bonferroni correction.

**Table 7**

Some background and treatment data (M and SD) for effectiveness and efficacy studies in the different disorders.

|  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
| Disorder | k | Age (years) | % Boys | % Comorb. | % on meds. | Severity | % Attrition | Tx time |
| *ADHD* |  | *p* = 0.95 | *p* = 0.05 | *p* = 0.26 | *p* = 0.96 | *p* = 0.62 | *p* = 0.28 | *p* = 0.03 |
| Effectiveness | 25 | 8.9 (2.6) | 77.7 (7.7) | 41.1 (25.8) | 42.4 (31.8) | 60.3 (16.2) | 14.3 (18.9) | 18.2 (8.4) |
| Efficacy | 43 | 8.9 (3.3) | 73.0 (10.1) | 49.6 (25.8) | 42.0 (27.6) | 58.5 (14.2) | 10.2 (12.3) | 23.5 (17.4) |
| *CD-ODD* |  | *p* = 0.82 | *p* = 0.03 | *p* = 0.001 | *p* = 0.10 | *p* = 0.04 | *p* = 0.15 | *p* = 0.87 |
| Effectiveness | 39 | 7.6 (2.6) | 76.9 (14.2) | 62.3 (19.9) | 21.1 (22.4) | 65.0 (8.2) | 15.4 (14.5) | 24.4 (13.7) |
| Efficacy | 62 | 7.5 (3.2) | 70.2 (14.6) | 18.0 (18.3) | 11.7 (11.4) | 60.6 (10.6) | 11.4 (10.0) | 23.7 (21.3) |

*Note*: k = number of treatment conditions, % Comorb. = proportion having any psychiatric comorbid disorder, % on meds. = proportion on any psychotropic medication, Severity = percentage of the maximum score on the primary outcome measure. % Attrition = proportion dropping out of those participating in at least one session. Tx time = number of 60 min therapy hours.

**Table 8**

Effect sizes (Hedges’ g) for effectiveness and efficacy studies within the different disorders.

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Disorder | Study type | k | *g* | 95% CI | *z*-value | Qb† | *p*-value |
| ADHD | *Post-treatment*  Effectiveness | 26 | 0.80 | 0.67–0.93 | 11.76a |  |  |
|  | Efficacy | 40 | 0.74 | 0.61–0.87 | 11.05a | 0.40 | 0.53 |
| CD-ODD | Effectiveness | 38 | 0.98 | 0.85–1.12 | 14.10a |  |  |
|  | Efficacy | 58 | 1.07 | 0.92–1.21 | 14.77a | 0.65 | 0.42 |
| ADHD | *Follow-up*  Effectiveness | 14 | 0.88 | 0.70–1.07 | 9.52a |  |  |
|  | Efficacy | 19 | 1.06 | 0.85–1.28 | 9.69a | 1.51 | 0.22 |
| CD-ODD | Effectiveness | 30 | 1.06 | 0.88–1.24 | 11.58a |  |  |
|  | Efficacy | 38 | 1.10 | 0.94–1.26 | 13.79a | 0.09 | 0.77 |

Note: k = number of comparisons. a *p* < 0.0001. Qb = Q between, † Comparison Efficacy vs. Effectiveness within the respective disorders.

**Table 9**

Remission rates for effectiveness and efficacy studies for the different disorders.

|  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- |
| Disorder | Study type | k | % | 95% CI | *z*-value\* | Qb† | *p*-value |
|  | *Post-treatment* |  |  |  |  |  |  |
| ADHD | Effectiveness | 12 | 37.7 | 26.7–50.2 | 1.93 |  |  |
|  | Efficacy | 17 | 33.4 | 23.2–45.4 | 2.67a | 0.26 | 0.81 |
| CD-ODD | Effectiveness | 19 | 47.6 | 38.0–57.3 | 0.49 |  |  |
|  | Efficacy | 23 | 54.8 | 48.1–61.4 | 1.40 | 1.43 | 0.23 |
|  | *Follow-up* |  |  |  |  |  |  |
| ADHD | Effectiveness | 5 | 38.2 | 21.8–57.9 | 1.18 |  |  |
|  | Efficacy | 2 | 44.3 | 36.5–52.7 | 1.37 | 0.33 | 0.57 |
| CD-ODD | Effectiveness | 12 | 56.8 | 46.5–66.5 | 1.29 |  |  |
|  | Efficacy | 23 | 54.9 | 47.3–62.3 | 1.27 | 0.08 | 0.77 |

Note: k = number of comparisons. a *p* < 0.01. \* Test if significantly different from 50%. Qb = Q between, † Comparison Efficacy vs. Effectiveness within the respective disorders.

**Table 10**

Effect sizes for randomized controlled studies only: ADHD and CD-ODD combined.

|  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- |
| Study type | k | *ES* | 95% CI | *z*-value | Qb† | *p*-value |
| *g-value at post-treatment* | | | | | | |
| Effectiveness | 49 | 0.90 | 0.78–1.02 | 14.86a |  |  |
| Efficacy | 98 | 0.93 | 0.83–1.03 | 18.16a | 0.08 | 0.79 |
| *g-value at follow-up* |  |  |  |  |  |  |
| Effectiveness | 38 | 0.99 | 0.84–1.14 | 13.29a |  |  |
| Efficacy | 57 | 1.09 | 0.96–1.21 | 17.05a | 0.97 | 0.32 |
| *Remission rate at post-treatment* | | | | | | |
| Effectiveness | 25 | 43.9 | 35.2–53.1 | 1.30 |  |  |
| Efficacy | 40 | 46.1 | 39.5–52.9 | 1.12 | 0.14 | 0.71 |
| *Remission rate at follow-up* | | | | | | |
| Effectiveness | 16 | 50.9 | 40.4–61.4 | 0.17 |  |  |
| Efficacy | 25 | 53.7 | 47.0–60.4 | 1.08 | 0.19 | 0.66 |

Note: k = number of comparisons. a *p* < 0.0001. Qb = Q between, † Comparison Effectiveness vs. Efficacy.

cognitive, academic, or social skills, and providing other, additional services were associated with smaller effect sizes.

The majority of the included studies delivered the treatment in groups, but treatment format (i.e. individual, group or combined indi- vidual and group treatment) or involving teachers in treatment did not significantly moderate the effect sizes. Although the involvement of teachers may be beneficial in certain cases, the results suggest that teacher involvement is not a crucial factor for treatment outcome in routine care, and that CBT for externalizing disorders is effective across different formats in routine care. The treatments had a mean number of

17 sessions across a mean of 15 weeks, with a mean intensity of approximately three hours per week. There was no evidence that treatment intensity significantly moderated outcome; however, number of sessions positively affected treatment outcome, i.e. more sessions yielded higher effect sizes. Previous meta-analyses have demonstrated similar findings ([Fossum et al., 2016](#_bookmark45); [Shelleby & Kolko, 2015](#_bookmark75)).

Regarding patient characteristics, lower age was associated with a better outcome of treatment. It has been argued that externalizing problems become increasingly resistant to change with age (e.g. [Ber-](#_bookmark21) [nazzani, Coˆt´e, & Tremblay, 2001](#_bookmark21); [Dekovi´c & Stoltz, 2015](#_bookmark30)), and our re- sults fit well with that assumption. Moreover, most of the studies in this meta-analysis involved parent behavior therapy, which is found to have a stronger effect on the youngest children ([McCart et al., 2006](#_bookmark62)). In line with these findings, a recent meta-analysis on psychosocial treatment for conduct disorder also found larger reductions in conduct problems in younger than in older children ([Fossum et al., 2016](#_bookmark45)). However, the literature is not consistent on this matter; two previous meta-analyses on parent training found no significant association between age and treat- ment outcome ([Cedar & Levant, 1990](#_bookmark24); [Lundahl et al., 2006](#_bookmark60)), whereas another found that older children benefitted more from behavioral parent training ([Serketich & Dumas, 1996](#_bookmark72)). Accordingly, no general conclusions can be drawn regarding the effect of age on treatment outcome. Nevertheless, the findings from the present meta-analysis suggest that, in routine care, younger age is associated with greater symptom reductions and higher remission rates after CBT.

More severe symptoms at pre-treatment were associated with a larger effect size. This seems to be a common finding (e.g. [Lundahl et al.,](#_bookmark60) [2006](#_bookmark60)) and a plausible reason is that it reflects that more severe cases have a larger room for improvement. This explanation is supported by the fact that remission rates were not associated with pre-treatment severity. The proportion of patients on psychotropic drugs at the beginning of treatment was a negative moderator of ES and remission rates. These two findings might seem somewhat contradictory, since it may be assumed that the samples with high pre-treatment severity would also have a high proportion of participants on psychotropic drugs. However, we did not include studies that investigated the effect of

combination treatment. Hence, the participants were, as part of routine care, medicated if indicated and on a stable dose before starting CBT- treatment. Therefore, the participants who were already on psychotro- pic drugs, may in fact have had less severe symptoms at pre-treatment, leaving a smaller room for improvement. It does not, however, explain why a higher proportion of patients on psychotropic drugs was associ- ated with lower remission rates. The finding probably reflects that the ADHD studies had a significantly higher proportion of patients receiving medication while also yielding lower effect size and remission rates than CD/ODD studies (although not significantly lower).

The most important and encouraging finding from the present meta- analysis was that the effectiveness-efficacy comparisons demonstrated no significant differences in ES between effectiveness and efficacy studies. The results are in line with a meta-analysis by [Michelson,](#_bookmark63) [Davenport, Dretzke, Barlow, and Day (2013)](#_bookmark63) investigating the outcome of parent management training (PMT) across different real-world prac- tice contexts. The authors found that whether PMT was delivered to clinically referred or study recruited samples, in service-oriented or research settings, or by non-specialist or specialist therapists, there was no difference in the overall outcome of PMT. The findings from the present meta-analysis lend further support to the transportability of empirically supported CBT-treatments for externalizing disorders from university settings to routine clinical practice. The fact that the effec- tiveness studies in the present meta-analysis seemed to have somewhat more severe CD/ODD samples with higher proportion of comorbid dis- orders provide additional support for the robustness of CBT in routine care.

* 1. *Strengths and limitations*

A methodological strength is that a power analysis indicated that this meta-analysis has a high power to detect a small effect size (based on the number of effect sizes summarized, the mean number of participants in the treatment conditions, and the observed degree of heterogeneity). Statistical power is rarely calculated in meta-analyses, but the impor- tance of power calculations in the interpretation of results has been highlighted ([Hedges & Pigott, 2001, 2004](#_bookmark47)). Another methodological strength is that all extractions of information from the included studies were done in pairs and any disparities were solved in consensus with all authors. A third strength is that the ratings of methodological quality and risk-of-bias was done by one of the authors and independently and blindly by another, yielding excellent inter rater reliability (ICC 0.94; 0.95). Still, the most important strength of the current meta-analysis is that it is the first to investigate the effects of CBT for children and ad- olescents in routine care with a diagnosis of ADHD, CD or ODD. The results therefore give valuable information for clinicians and decision makers. The findings are encouraging and indicate that externalizing disorders in childhood and adolescence respond well to CBT-oriented treatments in routine care.

The present meta-analysis has limitations that should be considered. First, the quality of a meta-analysis is limited by the quality of the original studies and can never be better than the data it summarizes. Hence, an important limitation is that few studies reported measures of functioning or quality of life or comorbidity. Second, the judgment of inclusion was in some cases challenging due to poor descriptions of the treatment setting, the clinicians and/or referral procedures. This might have caused exclusion of some studies that should have been included. Third, titles and/or abstracts were screened by one rater only. The PRISMA guideline recommends the use of two independent raters when screening title and abstract as this procedure may reduce the possibility of rejecting relevant reports. However, by having two independent raters reading the 730 full-text articles and including 7.0% of them in the current meta-analyses the risk of missing out on relevant reports may be low. The inclusion of both RCTs and open trials is a further limitation. However, in the present meta-analysis we wanted to calculate the amount of improvement that can be achieved in routine clinical care and

compare it with efficacy studies. Both RCTs and open trials could contribute with data for calculating within-group effect size and were therefore included. Finally, as only English-language peer-reviewed journal articles were included we may have missed relevant studies not published in peer-reviewed journals or published in languages other than English.

# Conclusion

The findings from the present meta-analysis demonstrate the effec- tiveness of CBT in routine clinical care for children and adolescents with externalizing disorders. The results showed that substantial effects can be obtained in routine care across different formats (i.e. individual or group), with different participants (i.e. parents, child or both), and with varying degrees of parental involvement. Finally, the results suggest that treatment delivered in routine care by practicing clinicians to clinically referred participants is as effective as treatments delivered in university settings.

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# Contributors

ENR elaborated the protocol registration in PROSPERO and wrote the first draft of the introduction and discussion. GJW rated the meth- odology and risk of bias of the included studies. ENR and GJW con- ducted literature searches in collaboration with an academic librarian. UN co-wrote the first draft of introduction and discussion with ENR. L-

GO¨ designed the meta-analysis, wrote the coding scheme, co-rated the

methodology and risk of bias of the studies, meta-analyzed the included studies, and wrote the first draft of methods and results. All authors contributed in the screening process, the extraction of data, and writing of the paper. All authors have approved the final manuscript.

# Declaration of Competing Interest

None of the authors have any conflict of interest to report.

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# Appendix A. Supplementary data

Supplementary data to this article can be found online at [https://doi.](https://doi.org/10.1016/j.cpr.2020.101954) [org/10.1016/j.cpr.2020.101954](https://doi.org/10.1016/j.cpr.2020.101954).

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