

# ANCOVA

- ANCOVA = ANalysis of COVariance: regression problems where some predictors are quantitative (i.e., numerical) and some are qualitative (i.e., categorical).
- For simplicity, focus on examples where we have just two predictors:  $X$  (numerical) and  $D$  (categorical).

## A Two-Level Example

- Model the response  $Y$  by two predictors  $X$  and  $D$ , where  $X$  is a numerical variable and  $D$  is categorical with two-levels (such as male or female).
- Code  $D$  as 0 or 1, e.g., 1 for male and 0 for female.

Note: you can code the two levels using any two different values, which will not change  $\hat{y}$ , but the interpretation of the estimated coefficients.

- In general, a factor with  $k$  levels corresponds to  $k - 1$  variables, when there is an additional intercept.

Recall the cats data, where we want to build a model to predict Hwt based on Bwt. For simplicity, assume  $n = 4$  and first two are female.

What are the possible regression models?

1. **Coincident regression line** (the simplest model): the same regression line for both groups, i.e., the categorical variable  $D$  has no effect on  $Y$ .

$$y = \beta_0 + \beta_1 x + e,$$

- 1' **Two-mean model** (another simplest model): the numerical variable  $X$  has no effect on  $Y$ .

$$y = \beta_0 + \beta_2 d + e = \begin{cases} \beta_0 + e, & d = 0 \\ (\beta_0 + \beta_2) + e, & d = 1 \end{cases}$$

2. **Parallel regression lines**: the categorical variable  $D$  **only** changes the intercept, i.e., it produces only an additive effect.

$$y = \beta_0 + \beta_2 d + \beta_1 x + e = \begin{cases} \beta_0 + \beta_1 x + e, & d = 0 \\ (\beta_0 + \beta_2) + \beta_1 x + e, & d = 1 \end{cases}$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} = \begin{pmatrix} 1 & 0 & x_1 \\ 1 & 0 & x_2 \\ 1 & 1 & x_3 \\ 1 & 1 & x_4 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_2 \\ \beta_1 \end{pmatrix} + \mathbf{e}$$

$\beta_2$ : measures the **change** of the additive effect (i.e., difference of the intercept).

Alternative choices for the design matrix (they should give us the same  $\hat{y}$ )

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} = \begin{pmatrix} 1 & 0 & x_1 \\ 1 & 0 & x_2 \\ 0 & 1 & x_3 \\ 0 & 1 & x_4 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_2 \\ \beta_1 \end{pmatrix} + \mathbf{e}$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} = \begin{pmatrix} 1 & 1 & x_1 \\ 1 & 1 & x_2 \\ 1 & 2 & x_3 \\ 1 & 2 & x_4 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_2 \\ \beta_1 \end{pmatrix} + \mathbf{e}$$

3. Regression lines with equal intercepts but different slopes: the categorical variable  $D$  **only** changes the effect of  $X$  on  $Y$ .

$$y = \beta_0 + \beta_1 x + \beta_3(x \cdot d) + e = \begin{cases} \beta_0 + \beta_1 x + e, & d = 0 \\ \beta_0 + (\beta_1 + \beta_3)x + e, & d = 1 \end{cases}$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} = \begin{pmatrix} 1 & x_1 & 0 \\ 1 & x_2 & 0 \\ 1 & x_3 & x_3 \\ 1 & x_4 & x_4 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \\ \beta_3 \end{pmatrix} + \mathbf{e}$$

$\beta_3$ : measures the **change** of the slope.

4. **Unrelated regression lines** (the most general model): the categorical variable  $D$  produces an additive change in  $Y$  and also changes the effect of  $X$  on  $Y$ . **Then should we just divide the data into two sets and run “lm” separately on them?**

$$y = \beta_0 + \beta_1 x + \beta_2 d + \beta_3 (x \cdot d) + e = \begin{cases} \beta_0 + \beta_1 x + e, \\ (\beta_0 + \beta_2) + (\beta_1 + \beta_3)x + e, \end{cases}$$

$$\begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{pmatrix} = \begin{pmatrix} 1 & 0 & x_1 & 0 \\ 1 & 0 & x_2 & 0 \\ 1 & 1 & x_3 & x_3 \\ 1 & 1 & x_4 & x_4 \end{pmatrix} \begin{pmatrix} \beta_0 \\ \beta_2 \\ \beta_1 \\ \beta_3 \end{pmatrix} + \mathbf{e}$$

How to interpret the LS coefficients from model 4?

- The usual “ $\beta_1$  measures the effect of  $X_1$  on  $Y$  when other predictors are held unchanged” does not make much sense for models with interactions. We cannot change  $x$  while holding  $d$  and  $(x \cdot d)$  unchanged.
- Let’s look at the Cathedral Example.



# Which Model to Pick?

You can use  $F$ -test to select the appropriate model.

- First test whether the interaction term is significant.

$$H_0 : \text{model 2} \quad H_a : \text{model 4.}$$

If reject the null, stop and take model 4.

Otherwise, decide whether you can further reduce model 2 to model 1 or model 1'.

- What if  $\beta_3$  (the interaction) is significant, but,  $\beta_1$  or  $\beta_2$ , is not significant? What about model 3?

The **Hierarchical Rule** for interactions: an interaction term will be included in a model only if all its main effects have been included. Due to this rule, we would include both  $\beta_1$  and  $\beta_2$ , once  $\beta_3$  is significant.

In practice we could test  $\beta_1 = 0$  or  $\beta_2 = 0$ . We just need to understand what the model looks like when  $\beta_1$  or  $\beta_2$  equals zero.

- when  $\beta_1 = 0$  (doesn't mean  $X$  is not significant)

$$y = \begin{cases} \beta_0 + e, & d = 0 \\ (\beta_0 + \beta_2) + \beta_3 x + e, & d = 1 \end{cases}$$

- when  $\beta_2 = 0$  (gives us model 3; doesn't mean  $D$  is not significant)

$$y = \begin{cases} \beta_0 + \beta_1 x, & d = 0 \\ \beta_0 + (\beta_1 + \beta_3)x, & d = 1 \end{cases}$$

## A Multi-Level Example

- Model the response  $Y$  by two predictors  $X$  and  $D$ , where  $X$  is a numerical variable and  $D$  is categorical with  $k$  levels .
- We need to generate  $k - 1$  dummy variables,  $D_2, \dots, D_k$  where

$$D_i = \begin{cases} 0, & \text{if not level } i \\ 1, & \text{if level } i. \end{cases}$$

Level 1 is the reference level.

The main purpose of the analysis is to decide which of the following models fits the data.

- Model 0:  $Y \sim 1$
- Model 1:  $Y \sim X$
- Model 1':  $Y \sim D$
- Model 2:  $Y \sim D + X$
- Model 4:  $Y \sim D + X + D : X$

The major tool is  $F$ -test. Note that when  $D$  has more than two levels, the difference, in terms of number of parameters, between models may not be one, so  $t$ -test is no longer appropriate.

1) If the interaction  $D : X$  is significant, stop.

$$H_0 : Y \sim D + X, \quad H_a : Y \sim D + X + D : X$$

2) If  $X$  is significant, keep  $X$ .

2') If  $D$  is significant, keep  $D$ .

3) If neither  $X$  nor  $D$  is significant, report the intercept model  $Y \sim 1$ .

2) and 2') are a little tricky.

2) Is  $X$  is significant?

Test the marginal contribution of  $X$

$$H_0 : Y \sim 1, \quad H_a : Y \sim X$$

Test the contribution of  $X$  in addition to  $D$

$$H_0 : Y \sim D, \quad H_a : Y \sim X + D$$

2') Is  $D$  is significant?

$$H_0 : Y \sim 1, \quad H_a : Y \sim D$$

$$H_0 : Y \sim X, \quad H_a : Y \sim X + D$$

# The Sequential ANOVA

The sequence of  $F$ -tests given by `anova(lm(Y ~ X + D + X:D))`

$H_0$	$H_a$
$Y \sim 1$	$Y \sim X$
$Y \sim X$	$Y \sim X + D$
$Y \sim X + D$	$Y \sim X + D + X : D$



The sequence of  $F$ -tests given by `anova(lm(Y ~ D + X + X:D))`

$H_0$	$H_a$
$Y \sim 1$	$Y \sim D$
$Y \sim D$	$Y \sim X + D$
$Y \sim X + D$	$Y \sim X + D + X : D$

Here is the catch: Some of the  $F$ -stats and  $p$ -values from the sequential ANOVA table are different from the ones we calculated based on usual  $F$ -test (we learned) for comparing two nested models.

Suppose we want to compare

$$H_0 : Y \sim X, \quad H_a : Y \sim X + D$$

- The usual  $F$ -stat

$$\frac{(\text{RSS}_0 - \text{RSS}_a)/(k-1)}{\text{RSS}_a/(n-p_a)} = \frac{(\text{RSS}_0 - \text{RSS}_a)/(k-1)}{\hat{\sigma}_a^2}$$

which follows  $F_{k-1, n-1-p}$  under the null.

- The  $F$ -stat from the sequential ANOVA table

$$\frac{(\text{RSS}_0 - \text{RSS}_a)/(k-1)}{\text{RSS}_A/(n-p_A)} = \frac{(\text{RSS}_0 - \text{RSS}_a)/(k-1)}{\hat{\sigma}_A^2}$$

which follows  $F_{k-1, n-p_A}$  under the null, where  $\text{RSS}_A$  denotes the RSS from the biggest model  $Y \sim X + D + X : D$  and  $p_A = 2k$ .