



University of  
**Nottingham**  
UK | CHINA | MALAYSIA

Autumn 2018

# Econometric Theory I

## Computer Lab Class III

Juergen Amann

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Wednesday 12:00 - 13:00, C42 SCGB

## Last week

Two exercises:

- **Exercise 1:** Examine the evidence of a gender pay gap.
  - Dummy variables.
  - What does it mean when we '*control for (a) variable(s)*'?
- **Exercise 2:** Examine the determinants of educational attainment.
  - Joint parameter hypothesis testing of coefficients; e.g.  
 $H_0 : \beta_2 = \beta_3$ .
  - Restricted Least Squares.

This week

## **Modelling the birth weights of infants**

- Identify factors that may determine a child's birth weight.

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- Homoskedasticity vs. heteroskedasticity.

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- Our data set `bwght.xls` contains  $N = 1,388$  observations on the following variables:
  - $\text{bwght}_i$ : birth weight in ounces,
  - $\text{faminc}_i$ : family income in USD 1,000,
  - $\text{order}_i$ : birth order of child,
  - $\text{cigs}_i$ : cigarettes smoked per day by mother during pregnancy,
  - $\text{mothereduc}_i$ : number of years mother spent in full-time education.

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  - `orderi`: birth order of child,
  - `cigsi`: cigarettes smoked per day by mother during pregnancy,
  - `motheduci`: number of years mother spent in full-time education.
- What is the expected sign for the variables? Which variables do you expect to be significant?

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Model	9103.2286	2	4551.6143	F(2, 1385)	=	11.15
Residual	565508.491	1,385	408.30938	Prob > F	=	0.0000
				R-squared	=	0.0158
				Adj R-squared	=	0.0144
Total	574611.72	1,387	414.283864	Root MSE	=	20.207

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.1235904	.0290392	4.26	0.000	.0666249	.1805559
order	1.439648	.6086762	2.37	0.018	.2456214	2.633675
_cons	112.7618	1.45622	77.43	0.000	109.9052	115.6185

- Test if variables are significant, interpret signs and magnitudes.

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- All coefficients are found to be highly significant.

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- This test statistic follows an  $F_{(q, n-k)}$  distribution under  $H_0$  where  $q = 2$  (no. of tested parameters) and  $(n - k) = 1,385$ .
- Compare test statistic against critical value (and/or use p-value): We reject  $H_0$  here; variables are 'jointly significant'. ▶ Let's do the same 'by hand'!

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- Only about 1.58% of the variation of the dependent variable about its mean is explained by **regression 1**.
- As before, we can calculate the  $R^2$  using other parts of the regression output. ▶ See how it's done!

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Model	19996.5211	3	6665.50703	F(3, 1384)	=	16.63
Residual	554615.199	1,384	400.733525	Prob > F	=	0.0000
Total	574611.72	1,387	414.283864	R-squared	=	0.0348
				Adj R-squared	=	0.0327
				Root MSE	=	20.018

  

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.0979201	.0291868	3.35	0.001	.040665	.1551752
order	1.616372	.603955	2.68	0.008	.4316058	2.801138
cigs	-.4771537	.091518	-5.21	0.000	-.6566827	-.2976247
_cons	114.2143	1.4693	77.73	0.000	111.3321	117.0966

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- Observe how the regression output has changed!
- Conduct a 1-sided test on the cigs coefficient using statistical tables!



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  - This means `cigs` is considered significantly different from 0 if the test statistic is in the top 2.5% or bottom 2.5% of its probability distribution, resulting in a p-value of less than 0.05.

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  - This means that, depending on the chosen tail,  $cigs$  is considered significantly lower (higher) than 0 if the test statistic is in the bottom (top) 5% of its probability distribution, resulting in a p-value of less than 0.05.



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  - A one-sided test allots **all** of our  $\alpha$  to testing the statistical significance in **one** direction only.
  - This means that, depending on the chosen tail,  $cigs$  is considered significantly lower (higher) than 0 if the test statistic is in the bottom (top) 5% of its probability distribution, resulting in a p-value of less than 0.05.
  - The one-tailed test provides more power to detect an effect in one direction by not testing the effect in the other direction.

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- Above  $H_0 : \beta_{cigs} = 0$ ,  $H_a : \beta_{cigs} \neq 0$  (two-sided).
- We want to test for the one-sided alternative that  $\beta_{cigs} < 0$ . Note that the effect has the correct 'predicted sign'.
- To get the p-value, simply divide  $p - Val_{cigs}/2$  because the effect is going in the 'predicted' direction. This is  $P(|-5.21|)$ .
- We reject  $H_0$  in favour of the one-sided alternative indicating that smoking has a significantly negative impact on birth weight.

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Residual	554580.264	1,383	400.998022	Prob > F	=	0.0000
Total	574611.72	1,387	414.283864	R-squared	=	0.0349
				Adj R-squared	=	0.0321
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bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.0937846	.0323843	2.90	0.004	.0302569	.1573122
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cigs	-.4730827	.0925814	-5.11	0.000	-.6546978	-.2914676
mothereduc	.076161	.2580337	0.30	0.768	-.4300187	.5823408
_cons	113.317	3.376943	33.56	0.000	106.6925	119.9415

- Interpret results!

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From the lecture [Slides 4, pp. 14]:

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- **Regression 2**: note that  $adj. R^2_{reg.3} < adj. R^2_{reg.2}$ ; **probably best fit**.

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⇒ **Hypothesis testing is invalid!**

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- Visual analysis first; e.g. for faminc

## Homo- vs. heteroskedasticity

- Re-run [regression 2](#) and check for heteroskedasticity.

```
. regress bwght faminc order cigs
```

- Visual analysis first; e.g. for faminc

```
. rvppplot faminc
```

## Homo- vs. heteroskedasticity

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- White test for heteroskedasticity:



## Homo- vs. heteroskedasticity

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- White test for heteroskedasticity:

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```

## Homo- vs. heteroskedasticity

- Re-run [regression 2](#) and check for heteroskedasticity.

```
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```

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```
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- White test for heteroskedasticity:

```
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```

White's test for  $H_0$ : homoskedasticity  
against  $H_a$ : unrestricted heteroskedasticity

```
chi2(9)      =      0.63  
Prob > chi2  =      0.9999
```

- We fail to reject  $H_0$  of homoskedasticity.

## Homo- vs. heteroskedasticity

- Re-run [regression 2](#) and check for heteroskedasticity.

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- We fail to reject  $H_0$  of homoskedasticity.
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## Homo- vs. heteroskedasticity

- Re-run [regression 2](#) and check for heteroskedasticity.

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chi2(9)      =      0.63  
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```

- We fail to reject  $H_0$  of homoskedasticity.
- No statistical evidence of heteroskedasticity.
- This is in line with our visual analysis.

Thank you and see you  
next time!

Juergen Amann

[juergen.amann@nottingham.ac.uk](mailto:juergen.amann@nottingham.ac.uk)

Wednesday 12:00 - 13:00, C42 SCGB

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## Overall relation 'by hand' [◀ Go back](#)

There are two ways how we can recover the previous statistic ourselves:

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. test faminc = order = 0
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. test faminc = order = 0

( 1) faminc - order = 0
( 2) faminc = 0

      F( 2, 1385) =    11.15
      Prob > F =    0.0000
```

## Overall relation 'by hand' [◀ Go back](#)

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$$F^* = \frac{R^2/(q)}{(1 - R^2)/(n - k)}.$$

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F( 2, 1385) = 11.15  
Prob > F = 0.0000
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- 2 Use one of the results from lecture [Slides 4, p.11], e.g.

$$F^* = \frac{R^2/(q)}{(1 - R^2)/(n - k)}.$$

```
. display 0.0158/(1 - 0.0158) * 1385/2  
11.117151
```

## $R^2$ 'by hand' [◀ Go back](#)

From the lecture slides [Slides 3, pp.20]:

- The  $R^2$  measure of goodness-of-fit is defined by  $R^2 = \frac{ESS}{TSS} = 1 - \frac{RSS}{TSS}$ .

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Source	SS	df	MS	Number of obs	=	1,388
				F(4, 1383)	=	12.49
Model	20031.4556	4	5007.8639	Prob > F	=	0.0000
Residual	554580.264	1,383	400.998022	R-squared	=	0.0349
				Adj R-squared	=	0.0321
Total	574611.72	1,387	414.283864	Root MSE	=	20.025

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. display 20031.4556 / 574611.72  
0.03486085
```

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$$adj.R^2 = 1 - \left[ \left( \frac{T-1}{T-k} \right) (1 - R^2) \right]$$



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$$adj.R^2 = 1 - \left[ \left( \frac{T-1}{T-k} \right) (1 - R^2) \right]$$

```
. display 1 - ((1388-1)/(1383)*(1-20031.4556/574611.72))  
0.03206942
```