



University of
Nottingham
UK | CHINA | MALAYSIA

Autumn 2018

Econometric Theory I

Computer Lab Class III

Juergen Amann

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Wednesday 12:00 - 13:00, C42 SCGB

Last week

Two exercises:

- **Exercise 1:** Examine the evidence of a gender pay gap.
 - Dummy variables.
 - What does it mean when we '*control for (a) variable(s)*'?
- **Exercise 2:** Examine the determinants of educational attainment.
 - Joint parameter hypothesis testing of coefficients; e.g.
 $H_0 : \beta_2 = \beta_3$.
 - Restricted Least Squares.

This week

Modelling the birth weights of infants

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- Our data set `bwght.xls` contains $N = 1,388$ observations on the following variables:
 - `bwghti`: birth weight in ounces,
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 - `orderi`: birth order of child,
 - `cigsi`: cigarettes smoked per day by mother during pregnancy,
 - `motheduci`: number of years mother spent in full-time education.

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- What is the expected sign for the variables? Which variables do you expect to be significant?

Regression 1

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Source	SS	df	MS	Number of obs	=	1,388
Model	9103.2286	2	4551.6143	F(2, 1385)	=	11.15
Residual	565508.491	1,385	408.30938	Prob > F	=	0.0000
Total	574611.72	1,387	414.283864	R-squared	=	0.0158
				Adj R-squared	=	0.0144
				Root MSE	=	20.207

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.1235904	.0290392	4.26	0.000	.0666249	.1805559
order	1.439648	.6086762	2.37	0.018	.2456214	2.633675
_cons	112.7618	1.45622	77.43	0.000	109.9052	115.6185

- Test if variables are significant, interpret signs and magnitudes.

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Results

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- All coefficients are found to be highly significant.

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this means $H_0 : \beta_2 = \beta_3 = 0$.

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- This test statistic follows an $F_{(q, n-k)}$ distribution under H_0 where $q = 2$ and $(n - k) = 1,385$.
- Compare test statistic against critical value (and/or use p-value): We reject H_0 here; variables are 'jointly significant'. ▶ Let's do the same 'by hand'!

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- As before, we can calculate the R^2 using other parts of the regression output. [▶ See how it's done!](#)

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Source	SS	df	MS	Number of obs	=	1,388
Model	19996.5211	3	6665.50703	F(3, 1384)	=	16.63
Residual	554615.199	1,384	400.733525	Prob > F	=	0.0000
Total	574611.72	1,387	414.283864	R-squared	=	0.0348
				Adj R-squared	=	0.0327
				Root MSE	=	20.018

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.0979201	.0291868	3.35	0.001	.040665	.1551752
order	1.616372	.603955	2.68	0.008	.4316058	2.801138
cigs	-.4771537	.091518	-5.21	0.000	-.6566827	-.2976247
_cons	114.2143	1.4693	77.73	0.000	111.3321	117.0966

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- Observe how the regression output has changed!
- Conduct a 1-sided test on the cigs coefficient using statistical tables!

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One- vs two-sided tests?

- Two-sided:
 - Testing for the possibility of the relationship in both directions.

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 - Given a fixed significance level of, say, $\alpha = 0.05$, a two-sided test allots half of our α to testing the statistical significance in **one** direction and the other half of α to test in the **other** direction.
 - This means `cigs` is considered significantly different from 0 if the test statistic is in the top 2.5% or bottom 2.5% of its probability distribution, resulting in a p-value of less than 0.05.

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 - A one-sided test allots **all** of our α to testing the statistical significance in **one** direction only.
 - This means that, depending on the chosen tail, $cigs$ is considered significantly lower (higher) than 0 if the test statistic is in the bottom (top) 5% of its probability distribution, resulting in a p-value of less than 0.05.

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 - This means that, depending on the chosen tail, $cigs$ is considered significantly lower (higher) than 0 if the test statistic is in the bottom (top) 5% of its probability distribution, resulting in a p-value of less than 0.05.
 - The one-tailed test provides more power to detect an effect in one direction by not testing the effect in the other direction.

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- Above $H_0 : \beta_{cigs} = 0$, $H_a : \beta_{cigs} \neq 0$ (two-sided).
- We want to test for the one-sided alternative that $\beta_{cigs} < 0$. Note that the effect has the correct 'predicted sign'.
- To get the p-value, simply divide $p - Val_{cigs}/2$ because the effect is going in the 'predicted' direction. This is $P(|-5.21|)$.
- We reject H_0 in favour of the one-sided alternative indicating that smoking has a significantly negative impact on birth weight.

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Residual	554580.264	1,383	400.998022	Prob > F	=	0.0000
				R-squared	=	0.0349
				Adj R-squared	=	0.0321
Total	574611.72	1,387	414.283864	Root MSE	=	20.025

bwght	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
faminc	.0937846	.0323843	2.90	0.004	.0302569	.1573122
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cigs	-.4730827	.0925814	-5.11	0.000	-.6546978	-.2914676
mothereduc	.076161	.2580337	0.30	0.768	-.4300187	.5823408
_cons	113.317	3.376943	33.56	0.000	106.6925	119.9415

- Interpret results!

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From the lecture [Slides 4, pp. 14]:

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- **Regression 3**: *mothereduc* probably irrelevant; **probably over-specified**.
- **Regression 2**: note that $adj. R^2_{reg.3} < adj. R^2_{reg.2}$; **probably best fit**.

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⇒ **Hypothesis testing is invalid!**

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. estat imtest, white
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White's test for H_0 : homoskedasticity
against H_a : unrestricted heteroskedasticity

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chi2(9)      =      0.63  
Prob > chi2  =      0.9999
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- We fail to reject H_0 of homoskedasticity.

Homo- vs. heteroskedasticity

- Re-run [regression 2](#) and check for heteroskedasticity.

```
. regress bwght faminc order cigs
```

- Visual analysis first; e.g. for faminc

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. rvppplot faminc
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- We fail to reject H_0 of homoskedasticity.
- No statistical evidence of heteroskedasticity.
- This is in line with our visual analysis.

Thank you and see you
next time!

Juergen Amann

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Wednesday 12:00 - 13:00, C42 SCGB

Overall relation 'by hand' [◀ Go back](#)

There are two ways how we can recover the previous statistic ourselves:

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( 2) faminc = 0

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$$F^* = \frac{R^2/(q)}{(1 - R^2)/(n - k)}.$$

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- 2 Use one of the results from lecture [Slides 4, p.11], e.g.

$$F^* = \frac{R^2/(q)}{(1 - R^2)/(n - k)}.$$

```
. display 0.0158/(1 - 0.0158) * 1385/2  
11.117151
```

R^2 'by hand' [◀ Go back](#)

From the lecture slides [Slides 3, pp.20]:

- The R^2 measure of goodness-of-fit is defined by $R^2 = \frac{ESS}{TSS} = 1 - \frac{RSS}{TSS}$.

R^2 'by hand' [◀ Go back](#)

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Source	SS	df	MS	Number of obs	=	1,388
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Model	9103.2286	2	4551.6143	Prob > F	=	0.0000
Residual	565508.491	1,385	408.30938	R-squared	=	0.0158
-----+-----				Adj R-squared	=	0.0144
Total	574611.72	1,387	414.283864	Root MSE	=	20.207

```
. display 9103.2286 / 574611.72  
0.0158
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$$adj.R^2 = 1 - \left[\left(\frac{T-1}{T-k} \right) (1 - R^2) \right]$$

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```
. display 1 - ((1388-1)/(1383)*(1-20031.4556/574611.72))  
0.0144
```