

Econometric Theory I

Computer Lab Class II

Juergen Amann juergen.amann@nottingham.ac.uk Wednesday 12:00 - 13:00, C42 SCGB

Last week

 Using earnings.xls we analysed the effect of education and work experience on earnings.

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- We estimated

$$earnings_i = \beta_1 + \beta_2 educ_i + \beta_3 workexp_i + u_i$$

where

- earnings; earnings of individual i in USD per hour
- educ_i : education of individual *i* as in years
- workexp;: work experience of individual i in years
- u_i : idiosyncratic error of i.

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where

- earnings; earnings of individual i in USD per hour
- educ_i : education of individual i as in years
- workexp; : work experience of individual *i* in years
- u_i : idiosyncratic error of i.
- We found educ and workexp to be highly significant and positive.

Two exercises:

• Exercise 1: Examine the evidence of a gender pay gap.

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 - Dummy variables.

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 - Joint parameter hypothesis testing of coefficients; e.g. $H_0: \beta_2 = \beta_3$.
 - Restricted Least Squares.

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$$female_i = \begin{cases} 1 & \text{if individual } i \text{ is female} \\ 0 & \text{otherwise} \end{cases}$$

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$$female_i = \begin{cases} 1 & \text{if individual } i \text{ is female} \\ 0 & \text{otherwise} \end{cases}$$
• $male_i = \begin{cases} 1 & \text{if individual } i \text{ is male} \\ 0 & \text{otherwise} \end{cases}$

- We examine evidence of a gender pay gap.
- Our data set earnings2.xls contains the same variables as earnings.xls plus two dummy variables:

 We want to know if there is a a significant difference in earnings between men and women controlling for education and work experience.

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Type: regress EARNINGS EDUC WORKEXP FEMALE

Stata Code

. regress EARNINGS EDUC WORKEXP

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EARNINGS	Coef.	Std. Err.	t	P> t	[95% Conf	. Interval]
EDUC	.5624327	.2336497	11.46	0.000	2.219146	3.137105
WORKEXP		.1285136	4.38	0.000	.3099816	.8148837
_cons		4.27251	-6.20	0.000	-34.87789	-18.09213

Stata Code

. regress EARNINGS EDUC WORKEXP

EARNINGS		Std. Err.	t			. Interval]
EDUC WORKEXP	2.678125	.2336497 .1285136 4.27251	11.46 4.38 -6.20	0.000 0.000 0.000	2.219146 .3099816 -34.87789	3.137105 .8148837 -18.09213

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. regress EARNINGS EDUC WORKEXP FEMALE

EARNINGS	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
EDUC	2.591137	.2285497	11.34	0.000	2.142174	3.0401
WORKEXP	.4056773	.1288199	3.15	0.002	.1526236	.658731
FEMALE	-5.90905	1.113972	-5.30	0.000	-8.097337	-3.720764
_cons	-19.69195	4.36076	-4.52	0.000	-28.25822	-11.12567

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 - More specifically:

```
\widehat{earnings}_i = \begin{cases} (-19.69 - 5.91) + 2.59 * \text{educ} + 0.41 * \text{workexp} & \text{if female}_i = 1 \\ -19.69 & + 2.59 * \text{educ} + 0.41 * \text{workexp} & \text{if female}_i = 0 \end{cases}
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• Structural stability in cross-section setting:

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- Structural stability in cross-section setting:
 - A different relationship exists for different cross-sectional groups, i.e. there is a significant difference in earnings between men and women.

How important are controls in this example?
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EARNINGS	I	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
EDUC WORKEXP FEMALE _cons		2.591137 .4056773 -5.90905 -19.69195	.2285497 .1288199 1.113972 4.36076	11.34 3.15 -5.30 -4.52	0.000 0.002 0.000 0.000	2.142174 .1526236 -8.097337 -28.25822	3.0401 .658731 -3.720764 -11.12567

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EARNINGS					[95% Conf.	Interval
EDUC WORKEXP FEMALE	2.591137 .4056773 -5.90905 -19.69195	.2285497 .1288199 1.113972	11.34 3.15 -5.30	0.000 0.002 0.000	2.142174 .1526236 -8.097337 -28.25822	.65873 -3.72076
EARNINGS		Std. Err.			[95% Conf.	Interval
	-6.956519	1.205095	-5.77	0.000	-9.323786 21.44057	

gress EARN	INGS EDUC WO	RKEXP FEMALE				
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regress EARN	INGS FEMALE					
EARNINGS					[95% Conf.	
+					[95% Conf.	
FEMALE	-6.956519	1.205095	-5.77	0.000		-4.589251

• The gender pay gap is bigger if we do **not** control for educ and workexp.

How important are controls in this example?

```
. regress EARNINGS EDUC WORKEXP FEMALE
            Coef. Std. Err. t P>|t| [95% Conf. Interval]
   EARNINGS |
      EDUC | 2.591137 .2285497 11.34 0.000 2.142174
                                                         3.0401
   WORKEXP |
           .4056773 .1288199
                               3.15 0.002
                                            .1526236
                                                        .658731
    FEMALE I
           -5.90905 1.113972
                                -5.30 0.000 -8.097337 -3.720764
                                            -28.25822 -11.12567
     _cons | -19.69195  4.36076  -4.52  0.000
. regress EARNINGS FEMALE
            Coef. Std. Err. t P>|t| [95% Conf. Interval]
   EARNINGS |
           -6.956519 1.205095 -5.77 0.000 -9.323786
    FEMALE I
                                                        -4 589251
           23.11448 .8521306 27.13 0.000 21.44057
     cons
                                                         24.78839
```

- The gender pay gap is bigger if we do **not** control for educ and workexp.
- This is what is meant when you hear 'The gender pay gap shrinks when we take into account ...' in our case education and work experience.
- Also, there's more to the question than it seems! Let's take a look!

- Let's examine the determinants of educational attainment.
- Download education.xls and import it into Stata.

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- Download education.xls and import it into Stata.
- We estimate

$$\texttt{educ_i} = \beta_1 + \beta_2 \texttt{aptitude_i} + \beta_3 \texttt{mothereduc_i} + \beta_4 \texttt{fathereduc_i} + u_i$$

where

- educ_i: education of individual i in years
- aptitude; : test score of individual i attained on aptitude test
- mothereduc_i: years i's mother spent in full-time education
- $fathereduc_i$: years i's father spent in full-time education
- u_i : idiosyncratic error of i.

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      EDUC |
                Coef.
                       Std. Err.
                                   t P>|t|
                                                 [95% Conf. Interval]
   APTITUDE |
              .1257087
                       .0098533 12.76 0.000
                                                 .1063528
                                                           .1450646
 MOTHEREDUC |
              .0492425
                       .0390901 1.26 0.208 -.027546
                                                           .1260309
                       .0309522 3.48 0.001
                                                 .04688
 FATHEREDUC | .1076825
                                                           .1684851
     _cons | 5.370631
                       .4882155 11.00 0.000
                                                  4.41158
                                                           6.329681
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MOTHEREDUC FATHEREDUC		.0492425 .1076825	.0390901 .0309522	1.26 3.48	0.001	027546 .04688	.1684851		
_cons	 	5.370631 	.4882155	11.00	0.000	4.41158	6.329681		

- All coefficients are positive.
- Coefficient for mothereduc is insignificant.

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$$\bullet \ \ F_{1,T-k-1} = \left(\frac{(\hat{\beta}_f - \hat{\beta}_m) - (\beta_f - \beta_m)}{\sqrt{se_{\beta_f}^2 + se_{\beta_m}^2 + 2 \times se_{\beta_f} se_{b_m}}}\right)^2, H_0 : \left(\frac{(\hat{\beta}_f - \hat{\beta}_m)}{\sqrt{se_{\beta_f}^2 + se_{\beta_m}^2 + 2 \times se_{\beta_f} se_{\beta_m}}}\right)^2$$

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- $se^2 = \hat{\sigma}^2$, the sample variance.
- $\hat{\beta}_i$, the estimated coefficient.
- As before: Observed Value Value Predicted under H₀ (here equal to 0) divided by the estimated standard error of the estimator.

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( 1) MOTHEREDUC - FATHEREDUC = 0

F( 1, 536) = 0.90
Prob > F = 0.3440
```

- We **fail** to reject H_0 as F-statistic is smaller than the critical value.
- Important: This is **not** a test of joint significance of both coefficients!
- You can also get the above F-statistic 'by hand'.

• Re-estimate the regression imposing the constraint that

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 - $educ_i = \beta_1 + \beta_2 aptitude_i + \beta_3 (mothereduc_i + fathereduc_i) + u_i$

parentseduc;

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 $parentseduc_i$

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 - If $\beta_3 = \beta_4$ then our model becomes:
 - $\operatorname{educ}_{\mathtt{i}} = \beta_1 + \beta_2 \operatorname{aptitude}_{\mathtt{i}} + \beta_3 \underbrace{\left(\operatorname{mothereduc}_{\mathtt{i}} + \operatorname{fathereduc}_{\mathtt{i}} \right)}_{} + u_i$

 $parentseduc_i$

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 - $\operatorname{educ}_{i} = \beta_{1} + \beta_{2}\operatorname{aptitude}_{i} + \beta_{3}\underbrace{\left(\operatorname{mothereduc}_{i} + \operatorname{fathereduc}_{i}\right)}_{:} + u_{i}$

 $parentseduc_i$

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EDUC		Std. Err.	t	P> t	[95% Conf.	Interval]
APTITUDE PARENTSEDU _cons	.1253106	.0098434 .0164247 .4817972	12.73 5.04 10.99	0.000 0.000 0.000	.1059743 .0505722 4.349731	.1446469 .1151014 6.242608

Thank you and see you next week!

Juergen Amann

juergen.amann@nottingham.ac.uk Wednesday 12:00 - 13:00, C42 SCGB

Exercise 1: Gender differences 1 Go back In our data set men have (on average) higher education and more work experience:

. tabstat EARNINGS EDUC WORKEXP, stat(mean sd) long by(FEMALE)

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In our data set men have (on average) higher education and more work experience:

```
. tabstat EARNINGS EDUC WORKEXP, stat(mean sd) long by(FEMALE)
FEMALE
          stats | EARNINGS
                                EDUC
                                      WORKEXP
          mean | 23.11448 13.72222 17.87201
             sd | 16.05073 2.575381 3.993107
           mean | 16.15796 13.62222 15.9287
             sd | 11.59666 2.297135 4.641399
Total
           mean | 19.63622 13.67222 16.90036
             sd | 14,41566 2,438476 4,433377
```

There is also notably more variation in earnings and work experience for men (check standard deviation).

Exercise 1: Gender differences (Go back)
In our data set men have (on average) higher education and more work experience:

```
. tabstat EARNINGS EDUC WORKEXP, stat(mean sd) long by(FEMALE)
FEMALE stats | EARNINGS
                               EDUC
                                      WORKEXP
         mean | 23.11448 13.72222 17.87201
             sd | 16.05073 2.575381 3.993107
         mean | 16.15796 13.62222 15.9287
            sd | 11.59666 2.297135 4.641399
           mean | 19.63622 13.67222 16.90036
Total
             sd | 14,41566 2,438476 4,433377
```

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Lastly, do the blue numbers look familiar? Compare them with the output when running

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EDUC	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
APTITUDE	.1257087	.0098533	12.76	0.000	.1063528	.1450646
MOTHEREDUC	.0492425	.0390901	1.26	0.208	027546	.1260309
FATHEREDUC	.1076825	.0309522	3.48	0.001	.04688	.1684851
_cons	5.370631	.4882155	11.00	0.000	4.41158	6.329681

. regress EDUC APTITUDE MOTHEREDUC FATHEREDUC

EDUC	Coef.	Std. Err.	t	P> t	L95% Conf.	Interval]
APTITUDE	.1257087	.0098533	12.76	0.000	.1063528	.1450646
MOTHEREDUC	.0492425	.0390901	1.26	0.208	027546	.1260309
FATHEREDUC	.1076825	.0309522	3.48	0.001	.04688	.1684851
_cons	5.370631	.4882155	11.00	0.000	4.41158	6.329681

. vce

. regress EDUC APTITUDE MOTHEREDUC FATHEREDUC

```
EDUC | Coef. Std. Err. t P>|t| [95% Conf. Interval]

APTITUDE | .1257087 .0098533 12.76 0.000 .1063528 .1450646

MOTHEREDUC | .0492425 .0390901 1.26 0.208 -.027546 .1260309

FATHEREDUC | .1076825 .0309522 3.48 0.001 .04688 .1684851

_cons | 5.370631 .4882155 11.00 0.000 4.41158 6.329681
```

. vce

Covariance matrix of coefficients of regress model

FATHEREDUC | -.00006315 -.00066072 .00095804

_cons | -.00320754 -.00529709 -.00044575 .23835441
Remember under
$$H_0: \left[(\hat{\beta}_f - \hat{\beta}_m) \times \left(\sqrt{se_{\hat{\beta}_f}^2 + se_{\hat{\beta}_m}^2 + 2 \times se_{\beta_f} se_{\beta_m}} \right)^{-1} \right]^2$$

```
. display ((.0492425 - .1076825 ) / (sqrt(.00152803 + .00095804 + 2 * .00066072)))^2 .89697298 <- this is the F-statistic F( 1, 536) = 0.90!
```