# SidhuAmandeep\_HW4

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## Question 1:

Build a complete pipeline with a data set of your choice and a tree-based model of your choice in R (using tidymodels) or Python (using scikit-learn). For each step, include a paragraph explaining why you did that step the way you did (what components were included and, possibly, what you decided not to do).

```
library(tidymodels)
library(BART)
```

## Step 1: a brief description of where the data came from

```
#read in the data set
df_wdbc <- read.csv("data.csv")</pre>
```

The data set being used is called the 'Breast Cancer Wisconsin (Diagnostic) Data Set' and was obtained at the following link: https://www.kaggle.com/datasets/uciml/breast-cancer-wisconsin-data?resource=download. The important features involved is the diagnosis attribute (which says whether a patient has a malignant or benign cancer is present). The remaining attributes are simply measurements about the imaging that was recorded by the medical professionals. The measurements were computed from an imagine of a fine needle aspirate (FNA) of a breast mass sampled. The class distribution overall is 357 benign and 212 malignant cancer. This data set was first used by professionals in San Jose to linearly program breast cancer diagnosis and still to this day is used for machine learning algorithm testing.

# Step 2: some initial investigation of the data [which textual or graphical summaries did you investigate? Did you find anything unusual?]

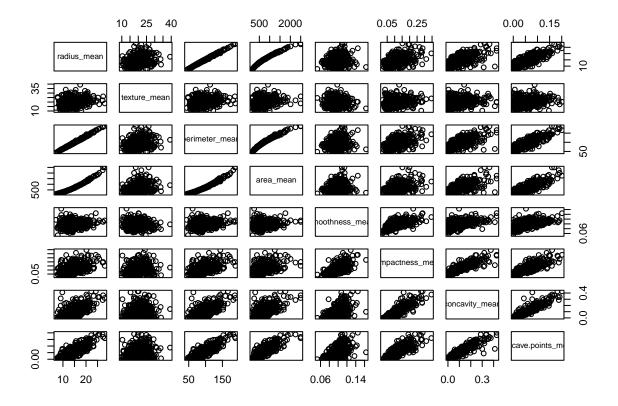
Let us run some preliminary investigation on the data.

```
#for sake of space I will only choose the first few values to display summary(df_wdbc[,c(3:10)])
```

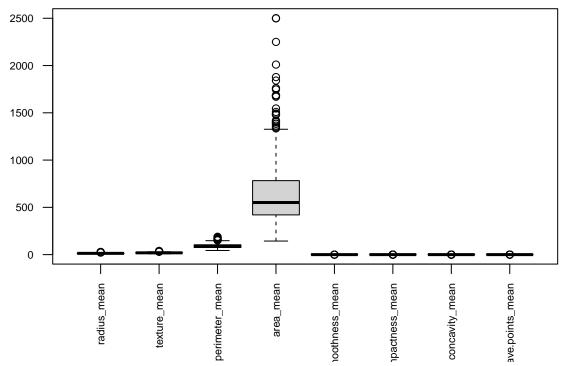
```
##
    radius_mean
                      texture_mean
                                     perimeter_mean
                                                         area_mean
##
   Min.
           : 6.981
                            : 9.71
                                     Min.
                                           : 43.79
                                                       Min.
                                                             : 143.5
                     Min.
                                     1st Qu.: 75.17
                                                       1st Qu.: 420.3
   1st Qu.:11.700
                     1st Qu.:16.17
   Median :13.370
                     Median :18.84
                                     Median: 86.24
                                                       Median : 551.1
##
   Mean
           :14.127
                     Mean
                            :19.29
                                     Mean
                                             : 91.97
                                                       Mean
                                                              : 654.9
##
   3rd Qu.:15.780
                                     3rd Qu.:104.10
                     3rd Qu.:21.80
                                                       3rd Qu.: 782.7
           :28.110
                            :39.28
                                     Max.
                                             :188.50
                                                              :2501.0
   smoothness_mean
                      compactness_mean concavity_mean
                                                           concave.points_mean
```

```
## Min.
          :0.05263
                     Min. :0.01938
                                      Min. :0.00000
                                                        Min.
                                                               :0.00000
   1st Qu.:0.08637
                     1st Qu.:0.06492
                                      1st Qu.:0.02956
                                                        1st Qu.:0.02031
##
  Median :0.09587
                     Median :0.09263
                                      Median :0.06154
                                                        Median :0.03350
  Mean
          :0.09636
                     Mean
                           :0.10434
                                      Mean
                                             :0.08880
                                                        Mean
                                                               :0.04892
##
                     3rd Qu.:0.13040
                                      3rd Qu.:0.13070
                                                        3rd Qu.:0.07400
##
   3rd Qu.:0.10530
##
   Max.
          :0.16340
                     Max.
                           :0.34540
                                      Max.
                                             :0.42680
                                                        Max.
                                                               :0.20120
```

pairs(df\_wdbc[,c(3:10)])



 $boxplot(df_wdbc[,c(3:10)], cex.axis = 0.7, las = 2)$ 

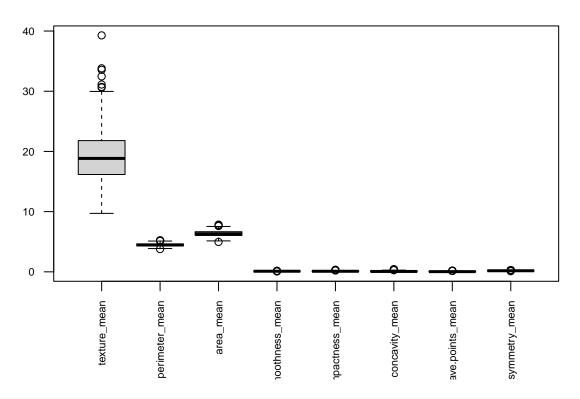


Here, all I am doing is using some summaries to observe the numerical value as well as the trend in feature variables. For the sake of space, I did not display all the features but it is fairly easy to change the parameters of each plot to observe the trends in the respective variables. According the the pairsplot, there appears to be random scatter among the variables present, but there seems to be a strong linear relationship between the radius\_mean and perimeter\_mean variables. Moreover, if we observe the boxplot, there is a value that has much larger values than the others (area\_mean). If we observe the summary() function output, this agrees with this observation. As a result, we may need to either scale or drop this variable entirely as it may skew the data all together. If we do scale the data, it will simply just be taking the log of the entire column.

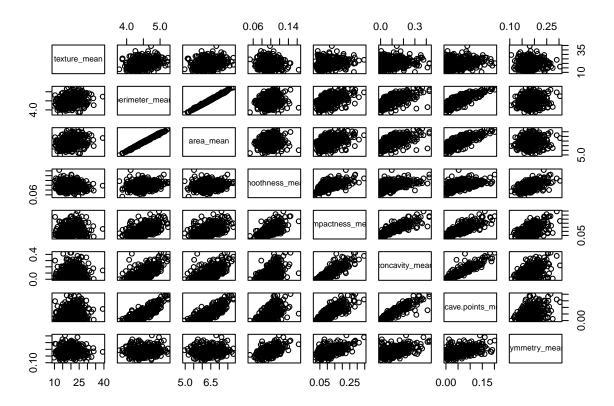
Step 3: preprocessing step(s) [scaling, feature engineering/variable selection {based on predictors only}, lumping or dropping categories from predictors, one-hot encoding, etc.]

```
df_new <- df_new %>%
    select(2:32)

boxplot(df_new[,c(3:10)], cex.axis = 0.7 , las = 2)
```



pairs(df\_new[,c(3:10)])



Next, I will create the recipe to be used in the model.

```
#convert to factor
df_new[, 'diagnosis'] <- as.factor(df_new[, 'diagnosis'])
#First split the model using dplyr for pipelines

set.seed(400076920)
split <-
    df_new %>%
    na.omit %>%
    initial_split()
train_data <- training(split)
test_data <- testing(split)

#recipe cooking
df_rec <- recipe(diagnosis ~., data = train_data)</pre>
```

Later on when we tune the parameters, we will see we need the diagnosis to be a factor. All that was done in the above step was scale certain features/variables by taking the log to get values closer to the other variables or else the magnitude would have been much higher in those parameters. Notice now how much tighter the variables are for the boxplot (the same variables were displayed as in step 2). The pairsplot seems to have gotten better linear relationship with most of the variables displayed in step 2 but there appears to be somewhat of a curve relationship for the first 3 variables. Overall, the scaling implemented here seems to have a positive effect on the data. Thus, we will keep this change. I have also encoded the malignant cancer

types as 1 and the benign as 0 as this will help the classification problem later (i.e. here we are dealing with the binary response variable). We also dropped the patient ID column as this has no effect on the model for classification purposes.

# Step 4: model choice [What model classes did you pick? Why?]

The model choice I have selected is BART (Bayesian Additive Regression Trees). Recall from lecture that this method is a sum-of-trees model to try and approximate an unknown function. We have selected this option because it is a totally acceptable algorithm for classification and regression problems (in this case we are using it for a classification problem as we are trying to classify patients into malignant and benign cancer types). Moreover, it is deemed often as more robust to outliers and missing values in the data set. This is very useful in our scenario just to provide a safety net to missing potential outliers in the data set even though we have scaled variables. The reason I have selected this over CART (classification and regression trees) as it can provide a more flexible approach to the problem (just in case the model is not linear). The computational cost is slightly more in BART than CART but that is the price of being more flexible.

# Step 5: model tuning [What hyperparameters did you tune? How? What loss function did you use and why? What was the range of achieved/minimized loss functions?]

The hyperparameters I will be tuning in this BART model include: number of trees, prior\_terminal\_node\_coef (this sets the precision of this normal distribution, which controls the amount of shrinkage applied to the terminal node coefficients) and prior\_terminal\_node\_exp. These were obtained by reading the R documentation found on https://www.rdocumentation.org/packages/parsnip/versions/1.0.2. I will do this by using the built in tune() function and using the 'roc\_auc' metric as the rsq and rmse are only reserved for regression problems in the tune\_grid() library to determine performance. The value that is deemed the best will be the values of number of trees and tree depth that I will implement. Next, the loss function that I will use is the logistic distribution as the response variable is binary by nature (i.e. only takes 0 and 1 values). As outline here: https://peterroelants.github.io/posts/cross-entropy-logistic/, logistic distribution is a powerful tool for binary classification.

In the next step, I will tune the parameters as discussed above and find the best model (i.e. will state the optimal tree values and the range of the loss function - the probabilities range).

Step 6: determining and fitting the best model

```
#first let us tune with grid search
#implement the bart model here that is to be tuned
df temp <-
  parsnip::bart(
   trees = tune(),
   prior_terminal_node_coef = tune(),
   prior_terminal_node_expo = tune(),
  ) %>%
  set_engine("dbarts") %>%
  set_mode("classification") #as our problem is a classification problem,
#use regression if you want regression
#Next, let us obtain the parameter objects
params <-
  workflow() %>%
  add model(df temp) %>%
  add_recipe(df_rec) %>%
  extract_parameter_set_dials() %>%
  finalize(train_data)
#Here is where we will be actually tuning the parameters based on a random grid
```

```
set.seed(400076920) #initialize
num_of_folds <- vfold_cv(train_data)</pre>
#tune the actual model on the grid, create a workflow and then essentially
#use the folds created above to score and the one with the best score is deemed
#the best
reg_tune_grid <-
 workflow() %>%
  add recipe(df rec) %>%
  add_model(df_temp) %>%
  tune_grid(
   num_of_folds,
   grid = 5,
   param_info = params,
   metrics = metric_set(roc_auc, sens, spec) #if req, use rsq or rmse
  )
#Now, select the bext hyperparameters after all that stuff of tuning
optimal_params <-
  #I used roc_auc to save some time, could use both prob would take
  #double the time to run
  select_best(reg_tune_grid, metric = 'roc_auc') %>%
  select(-.config)
#the optimal params were ntrees = 306, prior_terminal_node_coef = 0.826,
#prior terminal node expo = 2.99
#estimate and get the score after using the final split!
#this may take some time to compute
final flow <-
 workflow() %>%
 add_model(df_temp) %>%
  add_recipe(df_rec) %>%
  finalize_workflow(optimal_params)
set.seed(400076920)#initialize
fit <- final_flow %>%
 last_fit(split)
#metrics- remember we only care about the classification ones and not the
#rsq and rmse reserved for regression in the tuning documentation
collect_metrics(fit)
## # A tibble: 2 x 4
##
     .metric .estimator .estimate .config
             <chr>
                             <dbl> <chr>
     <chr>
## 1 accuracy binary
                             0.972 Preprocessor1 Model1
## 2 roc_auc binary
                            0.998 Preprocessor1_Model1
```

As we can see, the 'roc auc' is 0.9984486 and the accuracy is 0.9720280, which is great performance.

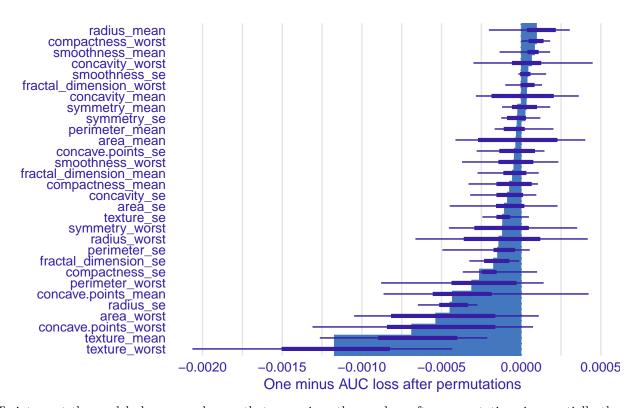
# Step 7: evaluate and explain the results of the model [partial dependence plots, variable importance, etc.]

As previously mentioned, the roc\_auc and accuracy scores are very close to 1 which indicates our performance is fairly high (ideal is 1 for both metrics!). These metrics were used as only regression

```
#to the TA grading, if you try to run this section it may take some time
#as it needs to permuate through all those options
#First, let us define a workflow to fit a bart model given our optimal parameters
library(DALEXtra)
#this will be useful for the training of the data set and determining importance
temp imp data <-
  df_rec %>%
 prep() %>%
 bake(new_data = NULL)
#next, we are going to use the optimal params and use it to create a model as
#follows
optimal_specs<-
 parsnip::bart(
   trees = 306,
   prior_terminal_node_coef = 0.826,
   prior_terminal_node_expo = 2.99,
  ) %>%
  set_engine("dbarts") %>%
  set_mode("classification")
#now, as discussed in lecture we are going to use the dalex package to create
#a explainer object that essentially explains the importance of the vars
explainer <-
  explain_tidymodels(
   optimal_specs %>%
      fit(diagnosis~., data = train_data),
   data = train_data %>% select(-diagnosis),
   y = as.numeric(train_data$diagnosis),
    #need to run it as numeric here or wont work
    verbose = FALSE
  )
#the final step is calculating the variable-importance measure and using this to
#plot the graph and discuss the differences
set.seed(400076920)
imp_plot_df <-</pre>
 model_parts(
   explainer = explainer,
   B = 10, #can change this to whatever number of permutations you want
   type = "difference",
   label = ""
  )
```

```
plot(imp_plot_df) +
  ggtitle("Average Var-Imp over 10 Permutations", "")
```

# Average Var-Imp over 10 Permutations



To interpret the model above, we observe that one minus the auc loss after permutations is essentially the deviation away from the ideal model. It is important to keep in mind that while there are some negative values, look at the magnitude. They are relatively very small so even removing the variables might have little to no effect on them. As a result, we can declare that almost all (an argument can be made to remove some of the variables at the bottom to improve performance) are relatively important. However, no one variable stands alone in providing a positive importance to the model. Thus, we keep all.

# Question 2:

(a) Derive the optimal weight values for both the MSE (L2 norm) and binomial deviance loss functions.

# Derive weights for the MSE

First, let us start with the given information from ESL II, In other words, recall that to find the optimal constants  $\gamma_{jm}$  in a specific region can be obtained using the following:

$$\hat{\gamma_{jm}} = argmin_{\gamma_{jm}} \sum_{x_i \in R_{jm}} L(y_i, f_{m-1}(x_i) + \gamma_{jm})$$

The above equation can be found on page 357 denoted by (10.30). For squared error loss, (i.e. the mean square error in this first part of the question), the regression tree is selected as the best one that predicts the

current residuals  $y_i - f_{m-1}(x_i)$ , whereby the  $\gamma_{jm}$  essentially acts as the mean of those calculated residuals in each region (or leaf in our case).

#### Derivation:

Recall that the objective for the MSE is to minimize the sum of squared errors between the predicted and true values. DUring the boosting process, at each step the next step is given by the sum of the prediction from the previous boosting step (i.e.  $f_{m-1}$ ). Thus, let us first start with the loss function of MSE given in ESL of:

$$L = \frac{1}{n} \sum_{i=0}^{n} (y_i - \gamma_i)^2$$

Here,  $y_i$  acts as the observed value and  $\gamma_i$  acts as the predicted value.

Next, as we are required to minimize the sum of squared errors, we must take the derivative with respect to  $\gamma$ .

Basically, we want to fit a tree at each step 'm' of the boosting algorithm. In that senze, we take the  $-\partial L/\partial f(x)$  to obtain our value of  $\gamma_{jm}$  where j is the index of the leaf and m is the m boosting step.

So, let us start with the result of step 2c of the algorithm 10.3.

$$\hat{\gamma_{jm}} = argmin_{\gamma_{jm}} \sum_{x_i \in R_{im}} L(y_i, f_{m-1}(x_i) + \gamma_{jm})$$

Sub in the loss function for MSE.

$$= argmin_{\gamma_{jm}} \sum_{x_i \in R_{im}} (y_i - (f_{m-1}(x_i) + \gamma_{jm}))^2$$

Next, expand the sum of squared.

$$= argmin_{\gamma_{jm}} \sum_{x_i \in R_{im}} (y_i^2 - 2y_i(f_{m-1}(x_i) + \gamma_{jm}) + (f_{m-1}(x_i) + \gamma_{jm})^2)$$

differentiating with respecting  $\gamma$  and simplifying,

$$= \sum_{x_i \in R_{jm}} (y_i - f_{m-1}(x_i) - \gamma_{jm})$$

$$= \sum_{x_i \in R_{jm}} (y_i - f_{m-1}(x_i) - \gamma_{jm}) = 0$$

$$= \sum_{x_i \in R_{jm}} y_i - \sum_{x_i \in R_{jm}} f_{m-1}(x_i) - \sum_{x_i \in R_{jm}} \gamma_{jm} = 0$$

$$\gamma_{jm} \sum_{x_i \in R_{jm}} 1 = \sum_{x_i \in R_{jm}} y_i - \sum_{x_i \in R_{jm}} f_{m-1}(x_i)$$

But note that the term  $\sum_{x_i \in R_{jm}} 1$  is just the magnitude of jm (i.e. |jm|).

Thus, the optimal weights for the MSE  $\gamma_{jm}$  is

$$\gamma_{jm} = \frac{1}{|jm|} \sum_{x_i \in R_{jm}} y_i - f_{m-1}(x_i)$$

#### Derive weights for binomial deviance loss function

Using a similar approach as above, we can derive the weights with a binomial deviance loss function.

Recall the binomial deviance loss function (as a form of the log-loss family) in an equivalent form to ESL pg. 346 as:

$$l(y, f(x)) = ylog(f(x)) + (1 - y)log(1 - f(x))$$

Now, we sub this loss function into the optimal weights equation previously defined.

$$\hat{\gamma_{jm}} = argmin_{\gamma_{jm}} \sum_{x_i \in R_{jm}} L(y_i, f_{m-1}(x_i) + \gamma_{jm})$$

$$= -argmin_{\gamma_{jm}} \sum_{x_i \in R_{jm}} y_i log(f_{m-1}(x_i) + \gamma_{jm}) + (1 - y_i) log(1 - f_{m-1}(x_i) - \gamma_{jm})$$

Take the derivative with respect to  $\gamma_{im}$ :

$$= -\sum_{x_i \in R_{jm}} \frac{y_i}{f_{m-1}(x_i) + \gamma_{jm}} - \frac{1 - y_i}{1 - f_{m-1}(x_i) - \gamma_{jm}}$$

Simplifying the term and setting it to 0:

$$0 = -\sum_{x_i \in R_{im}} \frac{y_i - f_{m-1}(x_i) - \gamma_{jm}}{(f_{m-1}(x_i) + \gamma_{jm})(1 - f_{m-1}(x_i) - \gamma_{jm})}$$

Solving for  $\gamma_{jm}$ :

$$\gamma_{jm} = \sum_{x_i \in R_{jm}} y_i - (f_{m-1}(x_i))$$

Note, here because we are dealing with the binomial deviance (or equivalently the log-likelihood loss function),  $y_i$  is the observed probability and  $f_{m-1}(x_i)$  is the predicted probability. So,  $\gamma_{jm}$  is simply the difference between the two.

The only sources I used for this problem was ESL II to obtain the loss functions and the papers mentioned by the assignment.

(b) Do the same for Newton boosting (Chen and Guestrin 2016), where we use a second-order rather than a first-order approximation to the loss function.

## Derive the Newton boosting weight assuming a second-order approximation for the MSE Loss

As suggested by : https://www.youtube.com/watch?v=StWY5QWMXCw, we can use a second order taylor approximation to get the newtons boosting estimate.

Basically, we need to use second-order approximations of the two loss functions to get the optimal weights for newton's boosting.

Recall the loss function for the MSE,

$$L = \frac{1}{n} \sum_{i=0}^{n} (y_i - f_{m-1}(x_i))^2$$

A second order approximation would be,

$$L \approx L(y_i, f_{m-1}(x_i)) + L'(y, f_{m-1}(x_i))\gamma + 1/2 * L''(y_i, f_{m-1}(x_i))\gamma^2$$

Taking the derivative with respect to  $\gamma$ 

$$\frac{\partial L}{\partial \gamma} \approx L'(y, f_{m-1}(x_i)) + L''(y_i, f_{m-1}(x_i))\gamma$$

Solving for  $\gamma$ ,

$$\gamma = \frac{L'(y, f_{m-1}(x_i))}{L''(y_i, f_{m-1}(x_i))}$$

Plugging in the appropriate values,

$$\gamma = \frac{-L'(y, f_{m-1}(x_i))}{L''(y_i, f_{m-1}(x_i))}$$

Notice, that the second derivative of the MSE loss function is simply a constant. Thus, the optimal value is

$$\gamma_{jm} = \frac{1}{|jm|} \sum_{i=0}^{m} (y_i - f_{m-1}(x_i))^2$$

.

which is simply the average of residuals at each leaf!

## Derive the Newton Boosting Weight using binomial loss function

So, let us find a second order taylor approximation of the likelihood.

Using the equivalent form of 10.18 on page 346 of ESL II, we observe that the likelihood is:

$$L = y_i log(f_{m-1}(x_i) + \gamma_{im}) + (1 - y_i) log(1 - f_{m-1}(x_i) - \gamma_{im})$$

$$L \approx L(y_i, f_{m-1}(x_i)) + L'(y, f_{m-1}(x_i))\gamma + 1/2 * L''(y_i, f_{m-1}(x_i))\gamma^2$$

Ignoring the first term, we take the derivative with respect to  $\gamma$ .

$$\frac{\partial L}{\partial \gamma} \approx L'(y, f_{m-1}(x_i)) + L''(y_i, f_{m-1}(x_i))\gamma$$

Solving for  $\gamma$ ,

$$\gamma = \frac{L'(y, f_{m-1}(x_i))}{L''(y_i, f_{m-1}(x_i))}$$

$$\gamma = \frac{\frac{y_i - f_{m-1}(x_i)}{(f_{m-1}(x_i))(1 - f_{m-1}(x_i)}}{\frac{f_{m-1}^2(x_i) - 2y_i f_{m-1}(x_i) + y_i}{(f_{m-1}(x_i) - 1)^2 f_{m-1}^2(x_i)}}$$

Simplifying this further, we arrive at the conclusion of:

$$\gamma_{jm} = \frac{y_i - f_{m-1}(x_i)}{f_{m-1}(x_i)(1 - f_{m-1}(x_i))}$$

which is essentially the residual divided by predicted probability multiplied by 1 minus the predicted probability. Recall that the residual is essentially the observed probability minus the predicted probability.