

# Out-Of-Bag Discriminative Graph Mining

Firstname Lastname  
First Affiliation  
Street Address  
12345 City, Country  
author1@unknown.net

Firstname Lastname  
Second Affiliation  
Street Address  
12345 City, Country  
author2@unknown.net

Firstname Lastname  
Third Affiliation  
Street Address  
12345 City, Country  
author3@unknown.net

## ABSTRACT

In class-labeled graph databases, each graph is associated with one from a finite set of classes, which induces associations between classes and subgraphs occurring in the database graphs. The subgraphs with strong class associations are called discriminative subgraphs. In this work, discriminative subgraphs are repeatedly mined on bootstrap samples of a graph database in order to improve on estimation of subgraph associations. The number of times a subgraph occurs in a graph associated with each class (support values) is recorded over the out-of-bag instances of the bootstrap process. We investigate sample mean and maximum likelihood estimation for the approximation of the true underlying support from these empirical values. It is shown that both significantly improve on the process, compared to single runs of discriminative graph mining, by applying the methods to publicly available toxicological databases, and validating support values, class bias, and class significance. In toxicology, the detection of subgraphs (fragments of chemical structure) that induce toxicity is a major goal. Apart from the subgraph associations being statistically validated, the number of subgraphs created by the proposed methods are much lower than for ordinary discriminative graph mining, which is often a bottleneck in the application of computational models to such databases, and hinders interpretation of the results.

## Categories and Subject Descriptors

H.2.8 [Database Applications]: Data Mining, Statistical Databases

## Keywords

Bootstrap Sampling, Graph Mining, Out-of-bag Estimation

## 1. INTRODUCTION

Given a class-labeled graph database, discriminative graph mining is a supervised learning task with the goal to extract

graph fragments (subgraphs) with strong associations to the classes, according to statistical constraints set by the user. For example, the subgraphs may be molecular fragments that induce toxicity. Indeed, finding such fragments is a major goal in toxicology [6]. However, discriminative graph mining yields large result sets, even for very high thresholds on subgraph class associations [5]. Moreover, it is an unstable process, i.e., slight changes in the sample may lead to substantially different subgraph sets.

In statistical learning, a bagged predictor consists of several aggregate predictors, each trained on a dedicated bootstrap sample of the training instances. As bootstrapping draws instances with replacement,  $1/e$  (roughly 37 %) of instances are not drawn in each sample, on average. These instances (referred to as out-of-bag instances) may be used to obtain estimates of the bagged predictor, by letting each predictor in the bag vote on “his” corresponding out-of-bag instances, which are unknown test examples to the predictor. For example, with random forests, out-of-bag estimation of node errors in decision trees could improve on estimates obtained from the training data as a whole [1].

This work extends discriminative graph mining by out-of-bag estimation of subgraph occurrence frequencies in the graphs associated with each class (support values), and by aggregating subgraphs in the output that pass the statistical constraints. Infrequently occurring subgraphs are filtered out, which removes the instability to a large extent. The result is a compact set of statistically validated discriminative subgraphs, which may be useful for diagnostic or predictive modelling, or for expert inspection.

The remainder of this work is structured as follows: We present related work with a focus on discriminative graph mining (section 2), our proposed methods for out-of-bag estimation of support values, as well as algorithmic implementation (section 3). Experiments include validation of subgraph support values,  $p$ -values, and class bias on publicly available databases of various sizes and numbers of classes (section 4). We draw conclusions in section 5. The contributions are summarized as follows:

- Repeated discriminative graph mining on bootstrap samples of a class-labeled graph database is proposed as a means to stabilize estimates for subgraph properties, such as support values per class, compared to single runs of discriminative graph mining. The estimation is performed using the out-of-bag instances of the individual bootstrap samples.
- Two methods for estimation of support values are described, both handling multiple class values. Signifi-

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cance tests are applied to these estimated values, and insignificant subgraphs are removed from the results.

- The estimated support values, class significance values, and class biases are empirically validated on molecular databases of various sizes. The results indicate significant improvements over ordinary discriminative graph mining, where the effect is generally larger the smaller the databases are. We conclude that out-of-bag estimation has potential to generate subgraph sets with more accurate statistical properties, which may be useful for statistical models, or expert inspection.

## 2. RELATED WORK

Out-of-bag methods have been used to robustly estimate node probabilities and node error rates in decision trees [1] as well as the generalization error and accuracy of bagged predictors. In the work by Bylander [3], generalization error was well modeled by out-of-bag estimation, and its already small bias could be further reduced by a correction method, where, in order to correct the prediction of a given instance, similar out-of-bag instances with the same class label were employed. However, the method of out-of-bag estimation is not confined to these examples, and may be used to estimate other statistical properties in supervised learning.

Discriminative graph mining is often employed as a pre-processing step to statistical learning because discriminative subgraphs may be useful as descriptors [2]. At present, there is a variety of well-known statistical learning algorithms available “off the shelf”, which makes a workflow attractive where subgraphs are extracted from the data, represented in a unified format, and fed into a machine learning algorithm [7]. Usually, discriminative subgraphs are mined from a graph database using a predefined threshold with regard to target class associations. A subgraph qualifies for the result set if it passes a corresponding significance test. However, such approaches tend to produce huge sets of very similar subgraphs, even with very high thresholds on discriminative potential, which would prevent machine learning methods from building models in acceptable time, and post-processing would be required to lower redundancy and eliminate the vast majority of subgraphs [5]. Moreover, the result is not stable with regard to perturbations of the database.

Subgraph boosting [9] is an integrated approach to build models on small sets of subgraphs, alternating between graph mining and model building. The method presented here is clearly different from boosting because it calculates subgraphs independently from a specific model. It is similar in that it calculates a small collection of discriminative subgraphs, but it is also stable against perturbations of the database, which is generally not the case for boosting.

## 3. METHODS

### 3.1 Basic Graph Theory

A graph database is a tuple  $(G, \Sigma, a)$ , where  $G$  is a set of graphs,  $\Sigma \neq \emptyset$  is a totally ordered set of labels and  $a : G \rightarrow I$ ,  $I \subset \mathbb{N}$ , is a surjective function that assigns one from a finite set of class values to every graph in the database. The set  $I$  consists of at least two values, and for all  $i \in I$ , the set that contains all  $g \in G$  with  $a(g) = i$  is called  $G^i$ . We consider labeled, undirected graphs, i.e., tuples  $g =$

	$g$	$all$
class 1	$k_1$	$ G^1 $
class 2	$k_2$	$ G^2 $
...	...	...
class $ I $	$k_{ I }$	$ G^{ I } $
$\Sigma$	$k$	$ G $

Table 1: Contingency table for subgraph  $g$ .

$(V, E, \Sigma, \lambda)$ , where  $V \neq \emptyset$  is a finite set of nodes and  $E \subseteq V = \{\{v_1, v_2\} \in \{V \times V\}, v_1 \neq v_2\}$  is a set of edges and  $\lambda : V \cup E \rightarrow \Sigma$  is a label function. An ordered set of nodes  $\{v_1, \dots, v_m\}$  is a *path* between  $v_1$  and  $v_m$ , if  $\{v_i, v_{i+1}\} \in E, i \in \{1, \dots, m-1\}$  and  $v_i \neq v_j$  for all  $i, j \in \{1, \dots, m\}$ . We only consider connected graphs here, i.e., there is a path between each two nodes in the graph.

A graph  $g' = (V', E', \Sigma', \lambda')$  is subgraph-isomorphic to  $g$  if  $V' \subseteq V$  and  $E' \subseteq E$  with  $V' \neq \emptyset$  and  $E' \neq \emptyset$ ,  $\lambda'(v_1) = \lambda(v_2)$  whenever  $v_1 = v_2$ , and  $\lambda'(e_1) = \lambda(e_2)$  whenever  $e_1 = e_2$ , for all nodes and edges in  $g'$ . In this case, graph  $g'$  is also referred to as a subgraph of  $g$ . The subset of the database instances  $G$  that  $g'$  is a subgraph of is called the occurrences of  $g'$ , and its size as (total) support of  $g'$ , denoted by  $support(g')$ . As a special case, the size of the subset of occurrences with  $a(g) = i$ , for any  $g$  in the occurrences, is referred to as the support of  $g'$  for class  $i$ . Thus, any subgraph has associated support values per class, ranging each between 0 and the  $support(g')$ , and summing up to  $support(g')$ .

Here, the subgraphs mined from the graph databases are free subtrees. We define a tree as a graph with exactly  $n - 1$  edges that connect its  $n$  nodes (thus it cannot contain cycles). A free tree is a tree without a designated root node. For an introduction to tree mining and the terminology used there, see the overview by Chi *et al.* [4].

### 3.2 Significance Test

For a given subgraph  $g$ , we seek a  $|I| \times 2$  contingency table that lists the support values per class in the first column and the overall distribution of target classes in the second column, as in Table 1. These data serve to check whether  $g$ 's support values differ significantly from the overall class distribution. The  $\chi_d^2$  function for distribution testing, defined as

$$\chi_d^2(x, y) = \sum_{i=1}^{|I|} \frac{(k_i - E(k_i))^2}{E(k_i)}, \quad (1)$$

where  $E(k_i) = \frac{|G^i|k}{|G|}$  is the expected value of  $k_i$ , calculates the sum of squares of deviations from the expected support for all target classes. The function value is then compared against the  $\chi^2$  distribution function to conduct a significance test with  $|I| - 1$  degrees of freedom and obtain a  $p$ -value  $p(g)$ . Additionally, the bias of  $g$  is determined as  $\arg \max_i (\frac{k_i}{k} / \frac{G^i}{G})$ , to designate the dominant class for  $g$ .

### 3.3 Out-Of-Bag Discriminative Graph Mining

In ordinary discriminative graph mining, the task is to identify all subgraphs  $g$  with  $support(g) > minsup$  and  $p(g) < \alpha$ , where  $minsup$  and  $\alpha$  are set by the user. Out-of-bag discriminative graph mining is different in that bootstrapping is repeatedly performed on  $G$ . In each bootstrap sample  $G'$  it is ensured that  $|G^i|$  graphs are associated with

class  $i$ , by drawing  $|G^i|$  times with replacement and uniform probability inside class  $i$ , for all  $i$  in  $I$  (stratification). Ordinary discriminative graph mining is applied to the database induced by  $G'$ , with *minsup* and  $\alpha$  thresholds. Finally, subgraph support values are assessed by performing isomorphism tests on the out-of-bag instances ("matching").

More specifically, bootstrapping and graph mining are repeated  $N$  times, where in each iteration tuples of subgraphs and out-of-bag support values per class  $(g, k_1, \dots, k_{|I|})$  are produced, meaning that subgraph  $g$  occurs in  $k_i$  out-of-bag graphs associated with class  $i$ . Due to the instability of discriminative graph mining, perturbations to the database yield almost always a different, but overlapping, selection of subgraphs. Therefore, if a subgraph is not produced in a specific bootstrap sample, the corresponding  $(k_1, \dots, k_{|I|})$  are encoded as missing values through a special symbol. The results are recorded over the  $N$  bootstrap samples such that for each  $g$ , the list of support values is a tuple  $(\mathbf{k}_1 \dots \mathbf{k}_{|I|})$ , where  $\mathbf{k}_i$  is a vector  $(k_i^1 \dots k_i^N)^T$ , containing all the support values.

To cope with the variety of rare subgraphs, after the end of bootstrapping a fixed threshold removes all subgraphs with less than  $[0.3 * N]$  entries in the  $\mathbf{k}_i$ . For the remaining ones, the total support is determined from the class specific support values by summing up vectors  $\mathbf{k}_i$  across classes:  $\mathbf{k} = \sum_{i=1}^{|I|} \mathbf{k}_i$ . Here, for indices with missing values, the result of the addition is also encoded as missing value.

From the recorded results, support values are estimated. Two methods are described in the next sections, based either on the sample mean support per class, and using data of the current subgraph only, or on a maximum likelihood estimate, involving some of the other subgraphs. Then the significance test from section 3.2 is run on the estimated support values, yielding estimated class significance values ( $p$ -values), and biases.

### 3.3.1 Sample Mean Method

We set the value of  $k_i$  in Table 1 to  $\bar{\mathbf{k}}_i$ , i.e., the sample mean across the entries of vector  $\mathbf{k}_i$  (ignoring missing values), for all  $i \in \{1, \dots, |I|\}$ , with  $k$  being the sum of the  $k_i$ .

### 3.3.2 Maximum Likelihood Estimation Method

Here, we employ some of the other subgraphs to form estimates for the  $k_i$ . The first step is to extract the subgraphs with the same class bias as  $g$ . Local ties are broken in favor of the dominant global class. In case of a further tie on the global level, one of the globally dominant classes is chosen with uniform probability. In a second step, the subgraphs with the same class bias as  $g$  are used to correct  $g$ 's local frequencies by weighting. This approach has some similarity to the work by Bylander [3], however, his aim is to correct instance predictions, and his correction employs similar out-of-bag instances, whereas our correction is applied across bootstraps, and on the subgraphs (not instances) obtained collectively from all the bootstrap samples. For each class, we model the event that each  $k_i^j \in \mathbf{k}_i$  would occur for each of the subgraphs with the same class bias as  $g$  as a multinomial selection process. More specifically, we determine the class probabilities for each subgraph  $g'$  with the same class bias as  $g$  with a maximum likelihood estimator, defined as

the smoothed vector of relative class specific support values:

$$\alpha_{g'} = \left( \frac{1 + |\mathbf{k}_1|_1}{|I| + |\mathbf{k}|_1}, \dots, \frac{1 + |\mathbf{k}_{|I|}|_1}{|I| + |\mathbf{k}|_1} \right) \quad (2)$$

where the  $\mathbf{k}_i$  and  $\mathbf{k}$  pertain to  $g'$ , and  $|\cdot|_1$  is the one-norm (the sum of the vector elements, ignoring missing values). Subsequently, for each non-missing tuple  $(k_1^j, \dots, k_{|I|}^j)$  pertaining to  $g$ , a probability distribution is determined from this collection of multinomials:

$$p((k_1^j, \dots, k_{|I|}^j)) = \frac{\sum_{g'} p((k_1^j, \dots, k_{|I|}^j); \alpha_{g'})}{\sum_{g'} 1} \quad (3)$$

Finally, the  $k_i^j$  values pertaining to  $g$  are corrected in a weighted average based on this probability distribution:

$$\bar{\mathbf{k}}_i = \frac{\sum_j k_i^j p((k_1^j, \dots, k_{|I|}^j))}{\sum_j p((k_1^j, \dots, k_{|I|}^j))} \quad (4)$$

Again, we set the value of  $k_i$  in Table 1 to  $\bar{\mathbf{k}}_i$ .

## 3.4 Algorithm

Ordinary discriminative graph mining (ODGM) may be implemented using a discriminative graph mining algorithm of choice as "mother algorithm", bounded by minimum total support and significance threshold  $\alpha$ . Here, we employ backbone refinement class mining (BBRC) [8]. To prevent exhausting the memory due to excessive amounts of different subgraphs produced across the sampling process, it cuts down on the number of subgraphs by filtering structurally most similar ones. To this end, it partitions the subgraph search space into non-disjoint sets with the same longest paths (backbones), and selects one representative from each partition, chosen to minimize the  $p$ -value (section 3.2). Apart from this, there is no difference to the notion of ODGM. In the BBRC output, two isomorphic subgraphs are always represented by the same string identifier. The strings are SMARTS, a kind of regular expression to encode molecular fragments as strings<sup>1</sup>. This approach allows to store out-of-bag support values  $\mathbf{k}_i$  in a hash structure, using SMARTS as keys.

Algorithm 1 switches between out-of-bag estimation and ODGM, where the significance parameter  $\alpha$  was kept fixed at 0.05 at all times. We describe out-of-bag estimation first. Line 4 creates a hash table to gather results from mining bootstrap samples (line 7). The resulting subgraphs are matched on the out-of-bag instances (line 8), and the results are stored in the hash. On termination of the loop, each hash entry  $\mathbf{k}_i$  has at most  $N$  non-missing support values. Post-processing the results incurs negligible overhead. It consists of removing subgraphs that (line 11) do not have enough non-missing entries in the  $\mathbf{k}_i$ , specified by *minHashLength*, or which (line 13) do not significantly deviate from the overall distribution of classes, as assessed by the significance test (section 3.2), with contingency table calculated according to mean (section 3.3.1) or maximum likelihood estimation (section 3.3.2) method (line 12).

If *method* = ODGM, the data (support values,  $p$ -values, and biases) is obtained from a single BBRC run on the full training data, without bootstrapping and out-of-bag estimation (line 2).

<sup>1</sup>See <http://www.daylight.com/dayhtml/doc/theory/theory.smarts.html>

**Data:** *dataBase, numBoots, minHashLength, method, minSup*  
**Result:** *[subgraphs, values]* ; *values* include support, *p*-value, and bias

```

1 if method = ODGM then
2   [subgraphs, values] ← BBRC(dataBase, minSup, 0.05);
3 else
4   hashTable ← {};
5   for i := 1 → numBoots (Parallel Processing) do
6     sample, OOB ← drawBsSample(dataBase);
7     [subgraphs, values] ← BBRC(sample, minSup, 0.05);
8     insert(hashTable, match(subgraphs, OOB));
9   [subgraphs, values] ← [];
10  for subgraph ∈ keys(hashTable) do
11    if length(hashTable[subgraph]) ≥ minHashLength then
12      [candidateSupportValues, candidatePValue, candidateBias] ← method(hashTable[subgraph]);
13      if candidatePValue < 0.05 then
14        subgraphs ← subgraphs ∪ subgraph;
15        values ← values ∪ [candidateSupportValues, candidatePValue, candidateBias];

```

**Algorithm 1: Calculation of subgraph significance on out-of-bag instances**

## 4. EXPERIMENTS

### 4.1 Setup

Three methods were compared by their ability to estimate the discriminative potential of subgraphs, by assessing the deviations between the class specific support values, *p*-values, and biases of subgraphs, as a) estimated by the respective method, and b) obtained by matching the subgraphs onto an independent test database. The methods are provided by Algorithm 1:

1. MLE: Out-of-bag estimation with Algorithm 1, according to section 3.3.2, with *numBoots* = 100.
2. MEAN: Out-of-bag estimation with Algorithm 1, according to section 3.3.1, with *numBoots* = 100.
3. ODGM: Ordinary discriminative graph mining.

Setting *numBoots* = 100 means a minimum hash entry length of 30 for each subgraph, which ensures a statistically reliable number of support values. Finally, the whole process was repeated 100 times for the three methods.

The whole procedure is shown in Algorithm 2. Line 1 initializes residual vectors. Inside the main loop, stratified splitting (such that proportions of classes inside each split equal overall proportions) generates equally large training and test databases. The training database is treated by the selected method, which returns a vector of subgraphs and a vector *values*<sup>B</sup> of values, including support values, *p*-values, and bias (line 4). The subgraphs are matched on the test database, yielding analogous values *values*<sup>M</sup> (line 5). Finally, residual vectors *E*<sub>1</sub> – *E*<sub>5</sub> capture the differences between *values*<sup>B</sup> and *values*<sup>M</sup>.

Five different error measures were assessed to compare the methods:

1. *E*<sub>1</sub>, the mean of  $p^B - p^M$  over subgraphs, i.e., the bias of *p*-value errors.
2. *E*<sub>2</sub>, the mean of  $|p^B - p^M|$  over subgraphs, i.e., the absolute *p*-values errors.
3. *E*<sub>3</sub>, the mean of  $\left( \frac{1}{|I|} \sum_{i=1}^{|I|} \left| \frac{k_i^B}{k^B} - \frac{k_i^M}{k^M} \right| \right)$  over subgraphs, i.e., the relative support value errors.

4. *E*<sub>4</sub>, One minus the mean of  $\delta(p^B \leq \alpha, p^M \leq \alpha)$  over subgraphs, i.e., the fraction wrongly recognized as significant.
5. *E*<sub>5</sub>, One minus the mean of  $\delta(bias^B, bias^M)$  over subgraphs, i.e., the fraction with wrongly recognized class bias.

In the above,  $\delta(\cdot, \cdot)$  is the Kronecker function, which returns 1, if the arguments are equal, and 0 otherwise.

Six molecular, class labeled databases were used in the experiments. Four were drawn from the carcinogenic potency database (CPDB)<sup>2</sup>, namely “Combined Carcinogenicity and Mutagenicity” (MUL, 4 classes, 677 compounds), “Mouse Carcinogenicity” (MOU, 2 classes, 914 compounds), “Multi-Cell Call” (MCC, 2 classes, 1050 compounds), and “Rat Carcinogenicity” (RAT, 2 classes, 1128 compounds). MUL’s labels consist of the four cross-combinations of binary carcinogenicity and mutagenicity labels from the CPDB, for all chemicals with both values for both labels present. A rather small database, describing human intestinal absorption (INT, 3 classes, 458 compounds) [10], as well as the rather large Kazius/Bursi mutagenicity database (KAZ, 2 classes, 4069 compounds) [6], were also used. An appropriate value for minimum global support (parameter *minSup* in Algorithm 2) was determined a priori for each database.

### 4.2 Results

Table 2 details the results in the form of mean values and standard deviations for all error measures. Figure 1 shows win/loss statistics. It compares out-of-bag methods MLE and MEAN with ODGM by pooling the former (OOB). This is possible, since they showed uniform win/loss behaviour against ODGM. Numbers indicate the total amount of wins, numbers in brackets the subset of significant wins, as obtained by Wilcoxon signed rank tests ( $n = 100, p = 0.001$ ).

*E*<sub>1</sub> and *E*<sub>2</sub> measure the numerical bias and absolute error on the *p*-values. *E*<sub>3</sub>, *E*<sub>4</sub>, and *E*<sub>5</sub> describe errors on relative support values, significance (as a binary attribute), and class bias, respectively. The latter measures are readily interpretable, thus we discuss them separately from *E*<sub>1</sub> and *E*<sub>2</sub>.

<sup>2</sup>See <http://potency.berkeley.edu/cpdb.html>

**Data:** *graphDatabase, method* (MLE, MEAN, or ODGM), *minSup*  
**Result:**  $E_1, E_2, E_3, E_4, E_5$   
1  $E_1 = E_2 = E_3 = E_4 = E_5 = []$ ;  
2 **for**  $i := 1 \rightarrow 100$  **do**  
3    $[trainSet, testSet] \leftarrow splitStratified(graphDatabase, 0.5)$ ;  
4    $[subgraphs, values^B] \leftarrow Alg. 1(trainSet, 100, 30, method, minSup)$ ;  
5    $[subgraphs, values^M] \leftarrow match(subgraphs, testSet)$ ;  
6   **for**  $j := 1 \rightarrow 5$  **do**  
7      $E_j \leftarrow [E_j, errorJ(values^M, values^B)]$ ;

**Algorithm 2:** Calculation of error measures

Method	E1	E2	E3	E4	E5
OOB	6 (6)	6 (5)	5 (4)	6 (5)	6 (5)
ODGM	0 (0)	0 (0)	1 (0)	0 (0)	0 (0)
MLE	2 (1)	2 (0)	6 (4)	5 (3)	4 (0)
MEAN	4 (3)	4 (2)	0 (0)	1 (0)	2 (0)

**Figure 1:** Wins and losses

Assay	Method	E1	E2	E3	E4	E5	Subgraphs
INT	MLE	-0.0107 (0.0610)	0.0414 (0.0487)	0.1086 (0.0365)	0.1626 (0.1948)	0.0029 (0.0137)	15.96
INT	MEAN	-0.0062 (0.0553)	0.0341 (0.0469)	0.1112 (0.0340)	0.1729 (0.1879)	0.0013 (0.0085)	18.36
INT	ODGM	0.0630 (0.0955)	0.0635 (0.0952)	0.1177 (0.0247)	0.4672 (0.1318)	0.1733 (0.1071)	103.53
MUL	MLE	-0.0129 (0.0154)	0.0188 (0.0132)	0.0610 (0.0200)	0.0806 (0.1405)	0.1457 (0.1723)	9.67
MUL	MEAN	-0.0098 (0.0101)	0.0155 (0.0104)	0.0694 (0.0239)	0.1040 (0.1392)	0.1576 (0.1655)	10.93
MUL	ODGM	0.0795 (0.0510)	0.0795 (0.0510)	0.1041 (0.0187)	0.5616 (0.1214)	0.4076 (0.1083)	94.95
MOU	MLE	0.0086 (0.0704)	0.0397 (0.0607)	0.0944 (0.0508)	0.2357 (0.2719)	0.0155 (0.0851)	4.51
MOU	MEAN	0.0161 (0.0593)	0.0353 (0.0526)	0.1050 (0.0504)	0.2976 (0.2761)	0.0132 (0.0716)	5.52
MOU	ODGM	0.1285 (0.1249)	0.1315 (0.1222)	0.1209 (0.0355)	0.5919 (0.1471)	0.1298 (0.1059)	48.17
MCC	MLE	0.0040 (0.0386)	0.0159 (0.0382)	0.0664 (0.0348)	0.1618 (0.1716)	0.0075 (0.0424)	12.71
MCC	MEAN	0.0062 (0.0337)	0.0194 (0.0327)	0.0791 (0.0373)	0.2089 (0.1743)	0.0097 (0.0389)	14.75
MCC	ODGM	0.0636 (0.0511)	0.0663 (0.0485)	0.0971 (0.0225)	0.5327 (0.1137)	0.0583 (0.0524)	79.47
RAT	MLE	-0.0028 (0.0125)	0.0084 (0.0116)	0.0573 (0.0286)	0.0936 (0.1435)	0.0022 (0.0141)	13.00
RAT	MEAN	-0.0017 (0.0115)	0.0092 (0.0109)	0.0659 (0.0317)	0.1280 (0.1439)	0.0101 (0.0365)	14.68
RAT	ODGM	0.0573 (0.0544)	0.0621 (0.0507)	0.0912 (0.0246)	0.5059 (0.1253)	0.0712 (0.0687)	70.01
KAZ	MLE	-0.0000 (0.0000)	0.0000 (0.0000)	0.0181 (0.0067)	0.0068 (0.0155)	0.0000 (0.0000)	26.09
KAZ	MEAN	-0.0000 (0.0000)	0.0000 (0.0000)	0.0182 (0.0069)	0.0052 (0.0140)	0.0000 (0.0000)	25.57
KAZ	ODGM	0.0000 (0.0000)	0.0000 (0.0000)	0.0181 (0.0066)	0.0106 (0.0217)	0.0015 (0.0075)	25.24

**Table 2:** Bias and accuracy

We first consider E1 and E2. Judging from Table 2, for five databases (apart from KAZ) ODGM drastically overestimated  $p$ -values. In contrast, out-of-bag methods seem relatively unbiased. Absolute errors were also clearly improved when out-of-bag methods were applied. The test results in Figure 1 confirm that all pairs of out-of-bag methods methods against ODGM were significantly different, with the single exception of absolute error (E2) for database INT. For database KAZ, E1 and E2 are too low to be apparent from Table 2. However, the differences are statistically significant, also for KAZ. The tests also indicate that MEAN has advantages over MLE in bias (E1). For absolute error (E2), the situation is less clear.

We now consider E3, E4, and E5 in Table 2. For E3, the relative support value error, OOB methods were always superior to ODGM, except for database KAZ. The improvements are also statistically significant, apart from INT. MLE performed always better than MEAN, and significantly better in most of the cases.

For E4, the error on binary significance estimation, ODGM errors were extreme: In four out of six cases, the majority of subgraphs were falsely recognized as significant. MLE and MEAN estimation provided a significant improvement for all databases, except for KAZ. Moreover, MLE performed better than MEAN in all cases except for KAZ (three times with significant differences).

Finally for E5, the class bias error, both MLE and MEAN drastically improved on ODGM. The differences are statis-

tically significant for all databases, except KAZ. There were virtually no differences between MLE and MEAN.

Table 2 gives the mean number of subgraphs generated in the last column. Much less subgraphs were generated by OOB methods, except for KAZ. There is a general trend for lower fractions of subgraphs with shrinking database sizes, compared to ODGM.

### 4.3 Discussion

We investigate the interpretable error measures E3, E4, and E5 again in the charts in Figure 2, where their values are plotted across increasing database size. The latter was logarithmically transformed, to visually better separate similarly-sized databases MOU, MCC, and RAT.

Across the three smallest databases, error measures vary. However, for all databases apart from KAZ, the difference to ODGM is always pronounced, and performance of methods varies similarly across these databases. For the KAZ database, there is hardly any advantage for out-of-bag methods.

The estimation of the prominent class for each subgraph, captured by E5, is the easiest task. Except for MUL, out-of-bag methods have E5 values close to zero. The corresponding values for ODGM are disquietingly high.

E3 and E4 are related, since the support values (analyzed in E3) provide the input for calculating  $p$ -values (analyzed in E4). In fact, the improvements made by out-of-bag methods in estimating the former helped approximating the latter substantially better, as the slopes of the lines are similar, but

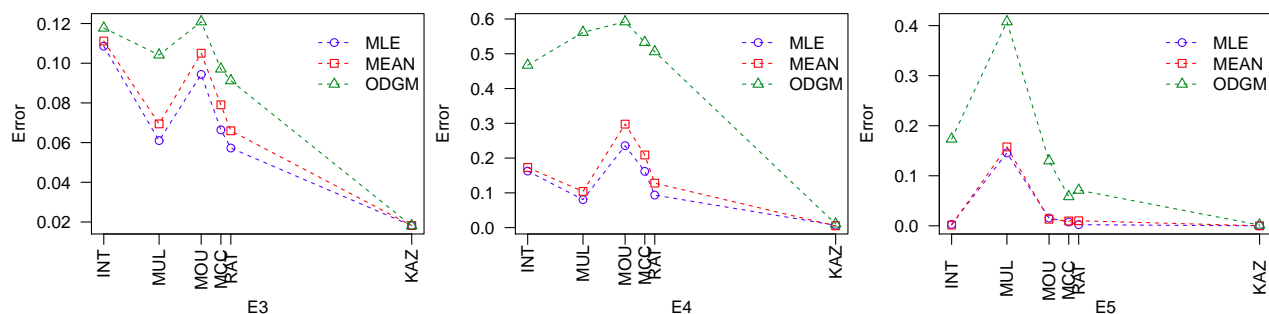


Figure 2: Error measures across increasing logarithmic database size

the difference to ODGM is much larger for E4 than for E3. Comparing out-of-bag methods, MLE is significantly better than MEAN in the majority of cases for these measures.

For larger databases, there is a trend towards increasing performance for all methods, as indicated by dashed lines. We attribute this to the increasing precision of ODGM (which is wrapped by the out-of-bag methods) in the presence of a larger number of learning data. This is also supported by the fact that the number of subgraphs generated for KAZ by ODGM is close to that of the other methods. In other words, a lot of confounding subgraphs have been removed for the smaller databases by out-of-bag estimation, but for a database this large, ODGM is nearly as precise as the other methods. However, the fraction of circa 50 % of subgraphs wrongly determined by ODGM in E4 even for the second largest database (RAT) shows clearly that mining significant subgraphs from training data in a single pass can not be judged reliable, unless for really large training databases: only for the KAZ database, E4 drops to levels similar to out-of-bag methods.

## 5. CONCLUSIONS

The proposed out-of-bag methods are able to significantly improve on support values and derived statistical quantities, especially on binary class significance (E4), and allow to remove a large fraction of subgraphs that are wrongly determined as significant by ordinary discriminative graph mining. The effect is particularly pronounced for small databases, which are typical in the field of toxicology, but less pronounced for larger ones. However, given the discussed results, it may not be possible to tell beforehand if a database is “large enough”. We suggest that, to reliably estimate the described properties, it is not sufficient to process databases in a single pass, as in ordinary discriminative graph mining, where even the assessment of class bias is not reliable. The same holds for the more sophisticated relative support values and binary class significance. Instead, sampling methods may be necessary, which can be conveniently combined with out-of-bag estimation.

Support values (class-specific occurrences) and binary class significance are probably the most important measures for experts analyzing the subgraphs, or fitting computational models using subgraph descriptors, for example toxicologists working in chemical risk assessment. For these quantities, the experiments have shown that, given a subgraph, using information from related subgraphs is beneficial, as the maximum likelihood method significantly outperforms mean estimation in most cases.

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