# Optimality and Sub-optimality of Principal Component Analysis for Spiked Random Matrices

#### Amelia Perry MIT

Joint work with:

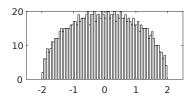
Alex Wein (MIT), Afonso Bandeira (Courant/CDS), Ankur Moitra (MIT)

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 symmetric,  $W_{ij}\sim\mathcal{N}\left(0,1
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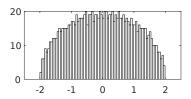


E. P. Wigner, AoM 1958.

V. A. Marchenko, L. A. Pastur, M.S 1967.

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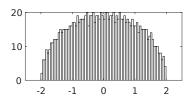
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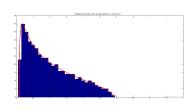
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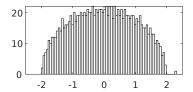
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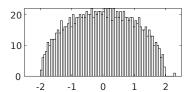
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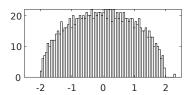
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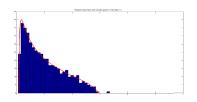
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Visible on the largest eigenvalue when

$$|\beta| > \sqrt{\gamma}$$
,  $\gamma = \frac{n}{N}$ ,  $\beta \in [-1, \infty)$ 

▶ Detection: distinguish reliably (error prob  $\rightarrow$  0)  $Y \sim \frac{1}{\sqrt{n}}W \text{ (Wigner)} \quad \text{vs} \quad Y \sim \frac{1}{\sqrt{n}}W + \lambda xx^T$ 

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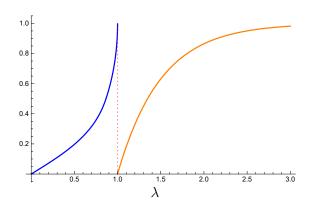
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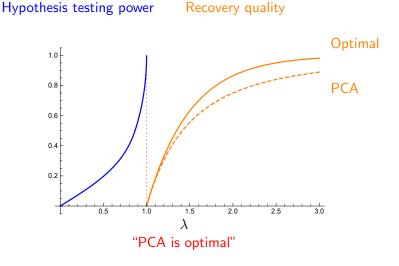
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  - unit sphere
  - ▶ i.i.d. ±1
  - ▶ sparse ±1

Hypothesis testing power Recovery quality





**Optimal** 1.0 0.8 **PCA** 0.6 0.4 0.2 3.0 0.5 1.0 1.5 2.0 2.5 "PCA is sub-optimal"

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This talk: focus on detection threshold (also hypothesis testing bounds, recovery threshold)

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3. Can beat PCA, but only with an inefficient algorithm (e.g. sparse priors, Wishart)

▶ Sequence of distributions  $P_n$  is contiguous to  $Q_n$  if for any sequence of events  $A_n$ ,

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L. Le Cam, 1960.

# Case 1: PCA is optimal

► Taking  $P_n : \frac{1}{\sqrt{n}}W + \lambda xx^T$  with  $x \sim \mathcal{X}$ , and  $Q_n : \frac{1}{\sqrt{n}}W$ 

<sup>[</sup>MRZ15] A. Montanari, D. Reichman, O. Zeitouni, NIPS 2015.

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## Wigner Second Moment

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But what about when we know more about the spike?

# Wigner, Rademacher Prior

$$Y \sim \frac{1}{\sqrt{n}}W$$
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Still Contiguous for  $\lambda < 1$ 

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Contiguity argument goes through for a general class of priors!

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Rate function:  $R(u) = \lim_{n \to \infty} \frac{-1}{n} \log \Pr[\langle x, x' \rangle \ge u]$ 

$$\begin{split} \mathbb{E}_{Q_n} \left( \frac{dP_n}{dQ_n} \right)^2 &= \mathbb{E} \exp \left( \frac{\lambda^2 n}{2} \langle x, x' \rangle^2 \right) \quad x, x' \sim \mathcal{X} \\ &= \int_0^\infty \Pr \left[ \exp \left( \frac{\lambda^2 n}{2} \langle x, x' \rangle^2 \right) \geq t \right] dt \\ &= 2 \int_0^\infty \Pr \left[ \langle x, x' \rangle \geq \sqrt{\frac{2 \log t}{\lambda^2 n}} \right] dt \\ &= 2 \int_0^1 \Pr \left[ \langle x, x' \rangle \geq u \right] \exp \left( \frac{\lambda^2 n}{2} u^2 \right) \lambda^2 n \, u \, du \\ &\to 2 \int_0^1 \lambda^2 n \, u \exp \left[ n \left( \frac{\lambda^2}{2} u^2 - R(u) \right) \right] du \end{split}$$

Rate function: 
$$R(u) = \lim_{n \to \infty} \frac{-1}{n} \log \Pr \left[ \langle x, x' \rangle \ge u \right]$$

In other words:  $\Pr[\langle x, x' \rangle \geq u] \approx \exp(-n R(u))$ 

$$\mathbb{E}_{Q_n} \left( \frac{dP_n}{dQ_n} \right)^2 \rightarrow 2 \int_0^1 \lambda^2 n \, u \exp \left[ n \left( \frac{\lambda^2}{2} u^2 - R(u) \right) \right] du$$

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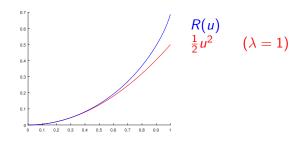
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- ▶ Bounded iff  $\frac{\lambda^2}{2}u^2 R(u) < 0$  for all  $u \in (0,1)$
- How big a parabola can you fit underneath the rate function?
- ▶ E.g. Rademacher prior  $(\pm 1)$  has  $R(u) = \log 2 h\left(\frac{1+u}{2}\right)$  where  $h(p) = -p \log p (1-p) \log(1-p)$



Case 2: PCA can be efficiently beaten

## What if noise is not Gaussian?

$$Y = \frac{1}{\sqrt{n}} \mathbf{W} + \lambda x x^T$$

 $x \sim \text{Unif}\{\mathbb{S}^{n-1}\}, \ W \in \mathbb{R}^{n \times n}$  but  $W_{ii} \sim p(w)$  such that  $\mathbb{E}w = 0, \ \mathbb{E}w^2 = 1.$ 

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**Universality**: spectral properties are unchanged...

D. Feral, S. Péché, CMP 2006.

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## Can you tell which one is which?

For 
$$W_{ij} \sim \text{Unif}(\pm 1)$$
,  $\lambda < 1$ 

$$W + \lambda \sqrt{n} x x^T$$
 vs  $W$ 

## Can you tell which one is which?

For 
$$W_{ij} \sim \text{Unif}(\pm 1)$$
,  $\lambda < 1$ 

$$W + \lambda \sqrt{n} x x^T$$
 vs  $W$ 

```
-1.0000
          1.0000
                    -1.0000
                              -1.0000
                                         1.0000
                                                   -1.0000
1.0000
          1.0000
                    1.0000
                              -1.0000
                                        -1.0000
                                                    1.0000
-1.0000
          1.0000
                    1.0000
                              -1.0000
                                        -1.0000
                                                  -1.0000
-1.0000
          -1.0000
                    -1.0000
                             1.0000
                                        -1.0000
                                                   1.0000
1.0000
          -1.0000
                    -1.0000
                              -1.0000
                                        -1.0000
                                                  1.0000
-1.0000
          1.0000
                    -1.0000
                              1.0000
                                        1.0000
                                                    1.0000
```

VS

```
-0.9988
          1.0011
                    -1.0007
                               -0.9997
                                          0.9990
                                                    -1.0014
 1.0011
          1.0010
                    0.9993
                               -0.9997
                                         -1.0010
                                                     0.9987
-1.0007
          0.9993
                    1.0004
                               -1.0002
                                         -0.9994
                                                    -0.9991
-0.9997
          -0.9997
                    -1.0002
                               1.0001
                                         -1.0002
                                                     0.9997
 0.9990
          -1.0010
                    -0.9994
                               -1.0002
                                         -0.9991
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                              -1.0000
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                                                  -1.0000
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          1.0000
                    1.0000
                              -1.0000
                                        -1.0000
                                                   1.0000
-1.0000
         1.0000
                   1.0000
                              -1.0000
                                        -1.0000
                                                  -1.0000
-1.0000
         -1.0000
                    -1.0000
                             1.0000
                                        -1.0000
                                                  1.0000
1.0000
         -1.0000
                   -1.0000
                              -1.0000
                                        -1.0000
                                                 1.0000
-1.0000
          1.0000
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                             1.0000
                                       1.0000
                                                   1.0000
```

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-0.9988
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                                0.9997
                                         1.0012
                                                    1.0017
```

Let's restrict ourselves to when the density p(w) is smooth.

$$f(Y_{ij}) = f(W_{ij} + \lambda \sqrt{n}x_ix_j)$$

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$$\approx f(W_{ij}) + f'(W_{ij}) \lambda \sqrt{n}x_ix_j$$

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$$\approx f(W_{ij}) + \mathbb{E}[f'(W_{ij})]\lambda \sqrt{n}x_ix_j - (f'(W_{ij}) - \mathbb{E}f'(W_{ij}))\lambda \sqrt{n}x_ix_j$$

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If noise drawn from non-Gaussian p(w): we will beat PCA by applying some function  $f: \mathbb{R} \to \mathbb{R}$  entrywise to our matrix  $Y = W + \lambda \sqrt{n}xx^{\top}$ , followed by PCA.

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▶ It is (close to) a new spiked Wigner matrix with

$$\lambda' = rac{\lambda \mathbb{E} f'\left(W_{ij}\right)}{\sqrt{\mathbb{E} f^2\left(W_{ij}\right)}}.$$

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Calculus of variations gives optimal choice of f:

$$f(w) = \frac{-p'(w)}{p(w)}$$

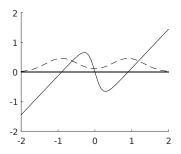
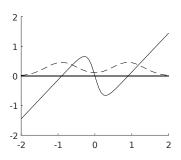
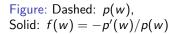
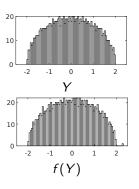


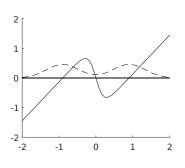
Figure: Dashed: p(w), Solid: f(w) = -p'(w)/p(w)

Related: T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015; F. Krzakala, J. Xu, L. Zdeborová, 2016.









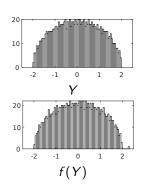
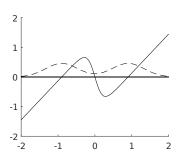


Figure: Dashed: p(w), Solid: f(w) = -p'(w)/p(w)

$$\qquad \text{New threshold at } \lambda = \frac{1}{\sqrt{F_\rho}}, \quad F_\rho = \mathbb{E}_p \left(\frac{p'(w)}{p(w)}\right)^2 \geq 1.$$

Related: T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015; F. Krzakala, J. Xu, L. Zdeborová, 2016.



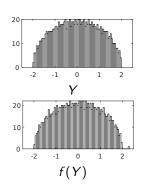
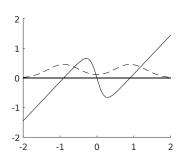


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Gaussian is hardest:  $F_p = 1$ .

Related: T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015: F. Krzakala, J. Xu. L. Zdeborová, 2016.



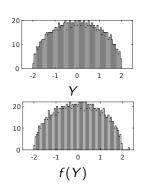


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New threshold at  $\lambda=\frac{1}{\sqrt{F_p}}, \quad F_p=\mathbb{E}_p\left(\frac{p'(w)}{p(w)}\right)^2\geq 1.$  Gaussian is hardest:  $F_p=1.$ 

▶ Theorem: detection is impossible below this threshold!

Related: T. Lesieur, F. Krzakala, L. Zdeborová, Allerton 2015;
F. Krzakala, J. Xu, L. Zdeborová, 2016.

# Case 3: PCA can be inefficiently beaten

# Spiked Wishart model, Rademacher prior

$$\frac{1}{N} \sum_{k=1}^{N} y_k y_k^T$$

$$y_k \sim \mathcal{N}\left(0, I_n\right) \qquad \text{vs} \qquad y_k \sim \mathcal{N}\left(0, I_n + \beta x x^T\right), \ x \sim \text{Unif}\left\{\pm \frac{1}{\sqrt{n}}\right\}^n$$

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- ▶ But...for  $\gamma > 0.698$  a computationally inefficient procedure distinguishes the two models for some  $\beta \in (-\sqrt{\gamma}, 0)$  (below the spectral threshold).

<sup>[</sup>OMH13] A. Onatski, M. J. Moreira, M. Hallin, AoS 2013.

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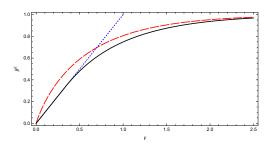
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Is there a computational gap?

## Rademacher Spiked Wishart, Negative $\beta$



- ▶ PCA: succeeds above the line
- ▶ inefficient algorithm: succeeds above the line
- contiguity lower bound: impossible below the line

▶ Goal:  $P_n$  contiguous to  $Q_n$ :  $Q_n(A_n) \to 0 \Rightarrow P_n(A_n) \to 0$ 

J. Banks, C. Moore, J. Neeman, P. Netrapalli, COLT 2016.

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Condition away from rare bad events

J. Banks, C. Moore, J. Neeman, P. Netrapalli, COLT 2016.

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

▶ Sparsity  $\rho \in [0,1]$ 

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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Prior 
$$\mathcal{X}_{\rho}$$
: i.i.d.  $x_i \sim \frac{1}{\sqrt{\rho n}} \left\{ \begin{array}{ll} +1 & \text{w.p. } \rho/2 \\ -1 & \text{w.p. } \rho/2 \\ 0 & \text{w.p. } 1-\rho \end{array} \right.$ 

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

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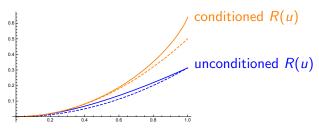
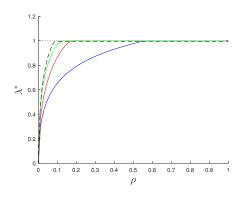


Figure:  $\rho = 0.2$ 

J. Banks, C. Moore, R. Vershynin, J. Xu, 2016.

## Sparse Rademacher: Results



- ▶ unconditioned
- conditioned
- noise-conditioned (upcoming)
- replica prediction (truth)

3 scenarios:

A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Optimality and Sub-optimality of PCA for Spiked Random Matrices and Synchronization," 2016.

A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Statistical limits of spiked tensor models," 2016+

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A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Optimality and Sub-optimality of PCA for Spiked Random Matrices and Synchronization," 2016.

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A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Statistical limits of spiked tensor models," 2016+

- Is it possible to match the replica prediction with a simple method? Or are there fundamental limits to these methods?
- In Case 3, are there true statistical-to-computational gaps, or can some method detect or recover efficiently?
- Can we establish sum-of-squares lower bounds for these computational problems?
- ▶ Is there a connection between SOS lower bounds and replica predictions of hardness?

#### Thanks! Questions?

A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Optimality and Sub-optimality of PCA for Spiked Random Matrices and Synchronization," 2016.

A. Perry, A. S. Wein, A. S. Bandeira, A. Moitra, "Statistical limits of spiked tensor models," 2016+