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Problem 1: Insurance claims

Let X(t) denote number of claims received by an insurance company during a time interval $\tau \in [0, t]$. One can assume that X(t) is a Poisson process with continuous time index $\tau \geq 0$, where τ denotes days from January 1st, 0:00.00.

a) Homogeneous Poisson process

In this task we let the intensity, or rate of the process, be fixed at $\lambda(t) = 3$. The probability that there are more than 325 claims before May 1st (120 days) is

$$P(X(120)_H > 325) = 1 - P(X(120)_H \le 325) = 1 - F(325) = 0.9671,$$

where $P(X(120)_H \le 325) = F(325)$ is the cumulative distribution function of the Poisson process at time t = 120 days.

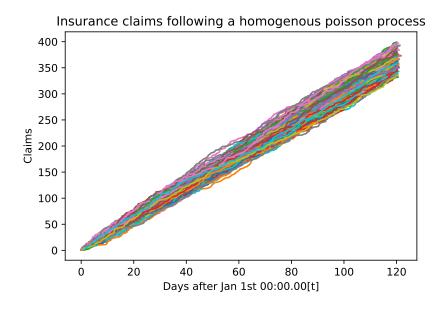


Figure 1: 100 realizations of the homogeneous poisson process with constant $\lambda(t) = 3$

We then simulated the process 1000 times, and counted the amount of claims we received in order to verify the theoretical probability of receiving more than 325 claims. We ended up with receiving more than 325 claims in 976 out of the 1000 simulations, which is pretty much what we expect based on the probability. Visualization shown in figure 3.

b) Inhomogeneous Poisson process

In this section we simulate the number of insurance claims as an inhomogeneous poisson process with intensity $\lambda(t) = 2 + \cos\left(\frac{2\pi(t)}{365}\right)$.

The probability of more than 325 claims before May 1st (120 days) was calculated using the corresponding cumulative distribution in which we calculated numerically.

$$P(X(120)_I > 325) = 1 - P(X(120)_I \le 325) = 1 - F(325) = 0.0235,$$

where $P(X(120)_I \le 325) = F(325)$ is the cumulative distribution function of the Poisson process at time t = 120 days.

Thus, the homogeneous process gives a higher probability of receiving more than 325 claims before May 1st.

By comparing figure 1 and 2, one observe that the rate in the inhomogeneous process changes with time, so the graphs appear less linear.

As well as with the homogeneous process, we made 1000 simulations in order to verify the calculated probability. In this case, we received more than 325 claims in 24 out of the 1000 simulations, which corresponds nicely to the theoretical values. Visualization shown in figure 3.

Insurance claims following an inhomogenous poisson process 300 250 200 150 100 50 0 20 40 60 80 100 120 Days after Jan 1st 00:00.00[t]

Figure 2: 100 realizations of the inhomogeneous poisson process with $\lambda(t)=2+\cos\left(\frac{2\pi(t)}{365}\right)$

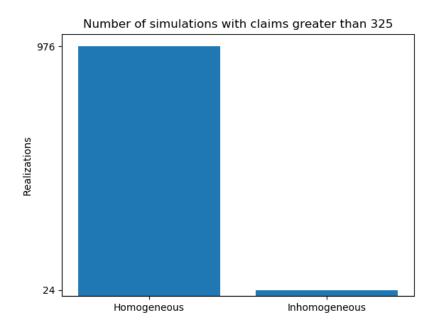


Figure 3: Number of simulations resulting in claims greater than 325 before May 1st $\,$

c) Log-Gaussian distribution

The expected value for the log-gaussian distribution, $E[C_i] = E[e^{Y_i}]$, is calculated by the fact that $Y_i \sim N(\mu, \sigma^2)$.

$$E[C_i] = E[e^{Y_i}] = \int_{\mathbb{R}} e^y \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y-\mu)^2} dy$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \int_{\mathbb{R}} e^{z\sigma + \mu - \frac{1}{2}z^2} \sigma dz = e^\mu \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} e^{\frac{\sigma^2}{2}} e^{-\frac{1}{2}(z-\sigma)^2} dz$$

$$= e^{\mu + \frac{\sigma^2}{2}} \frac{1}{\sqrt{2\pi}} \int_{\mathbb{R}} e^{-\frac{1}{2}(z-\sigma)^2} dz = e^{\mu + \frac{\sigma^2}{2}},$$

using the substitution $z = \frac{y-\mu}{\sigma}$, completing the square in the exponent of the integrand and the gaussian integral $\int_{\mathbb{R}} e^{(z-\sigma)^2} dz = \sqrt{2\pi}$.

The variance is calculated as

$$Var[C_i] = E[C_i^2] - E[C_i]^2 = e^{\mu + \sigma^2} (e^{\sigma^2} - 1),$$

using
$$E[C_i^2] = E[e^{2Y_i}] = e^{2(\mu + \sigma^2)}$$
.

The total claim amount is denoted by $Z = \sum_{i=1}^{X(t)} C_i$. One can calculate the expected value and variance in total claim amounts using the law of total expectation and -variance, respectively:

$$E[Z] = E\left[\sum_{i=1}^{X(t)} C_i\right] = E\left[E\left[\sum_{i=1}^{X(t)} C_i | X(t)\right]\right]$$

$$= E\left[\sum_{i=1}^{X(t)} E[C_i]\right] = E[X(t)E[C_i]] = E[X(t)e^{\mu + \frac{1}{2}\sigma^2}]$$

$$E[Z] = E[X(t)]e^{\mu + \frac{1}{2}\sigma^2} = \lambda te^{\mu + \frac{1}{2}\sigma^2},$$

where we have used $E[C_i]$ from above and $E[X(t)] = \lambda t$ for the homogeneous poisson process X(t).

$$\begin{split} Var[Z] &= Var\left[\sum_{i=1}^{X(t)}\right] = E\left[Var\left[\sum_{i=1}^{X(t)}C_i|X(t)\right]\right] + Var\left[E\left[\sum_{i=1}^{X(t)}C_i|X(t)\right]\right] \\ &= E\left[\sum_{i=1}^{X(t)}Var[C_i]\right] + Var\left[\sum_{i=1}^{X(t)}E[C_i]\right] \\ &= E[X(T)Var[C_i]] + Var[X(t)E[C_i]] = E[X(t)]Var[C_i] + Var[X(t)]E[C_i] \\ &= E[Z] = \lambda te^{2(\mu + \sigma^2)}. \end{split}$$

	E(.)	Var(.)
C_i	0.22	0.09
Z	80.33	48.72

Table 1: Theoretical variance and expected values for individual and total claim amounts

using $E[C_i]$ and $Var[C_i]$ from above, and that $E[X(t)] = Var[X(t)] = \lambda t$ for the homogeneous poisson process X(t).

The numerical values are listed in the table 1.

Simulations of the claim amounts was simulated for both the homogeneous process with $\lambda=3$ and the inhomogeneous process with

 $\lambda(t) = 3 + \cos(\frac{2\pi t}{365})$. for 100 realizations. Visuals of the simulations are depicted in figure (4-5) below.

As one can see from figure 4 below, the expected value for the total claim amounts fits well to the theoretical value listed in table 1. The observed variance was calculated to Var[Z] = 50.1 MNOK, which also within reach of the theoretical value in table 1.

We observe that the expected total claim amounts is less for the inhomogenous than the homogenous process (see figures 4-5). This is expected, as the rate of the inhomogenous process is always less than or equal to the rate of the homogenous process.

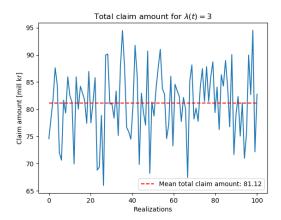


Figure 4: 100 realizations of the inhomogeneous poisson process with $\lambda(t)=3$

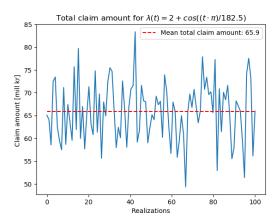


Figure 5: 100 realizations of the inhomogeneous poisson process with $\lambda(t)=2+\cos\left(\frac{2\pi(t)}{365}\right)$

Problem 2: Server jobs

In this problem we are considering the following birth-death process: A computer server can handle at most 32 jobs at a time. Arrivals of jobs are defined to be Poisson distributed with rate $\lambda=25$ per hour. The duration of each job is exponentially distributed, with rate parameter $\mu=1$ per hour. Any arriving job when the servers capacity is full, is forwarded to another server. Thus the current amount of jobs on the server is bounded, $X(t) \in \{0,1,...,32\}$.

a)

By using long-term equality of rates in and out of states we can calculate the long-term probabilities

$$\pi_k = \prod_{i=1}^{k-1} \frac{\lambda^i}{\mu^{i+1}} \frac{\pi_0}{i!}$$

We observe that all long-term probabilities are determined by π_0 . We know

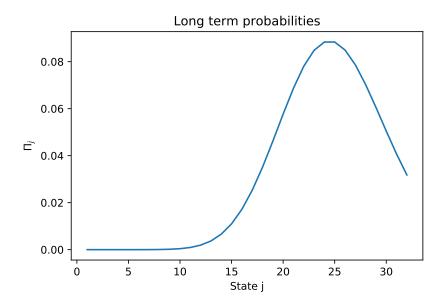


Figure 6: Equilibrium probabilities P(X(t) = n), n = 0,1,...,32.

all probabilities must sum to 1, which yields the following expression for π_0 :

$$\pi_0 = \frac{1}{1 + \sum_{k=1}^{\infty} \prod_{i=0}^{k-1} \frac{\lambda_i}{\mu_{i+1} \cdot i!}}$$

From figure 6 we conclude that the long term probability is greatest for 25 parallel jobs. This makes sense intuitively, because from the Poisson distributed arrivals, we expect 25 new jobs per hour. Simultaneously, we expect 25 jobs to be finished per hour, when each job has an exponential living time of 1 hour. Thus at 25 jobs, we have statistical equilibrium between births and deaths. We expect the process to oscillate around this value.

b)

In this task we are simulating the birth-death process for 168 hours. In figure 7 we have displayed 3 simulations, showing that the process is oscillating around the equilibrium at ≈ 25 . As the process reached 25 parallel jobs, we considered it to be in equilibrium. Based on these simulations, the end of the transient part of the process was estimated at ≈ 2.42 hours, illustrated in figure 7.

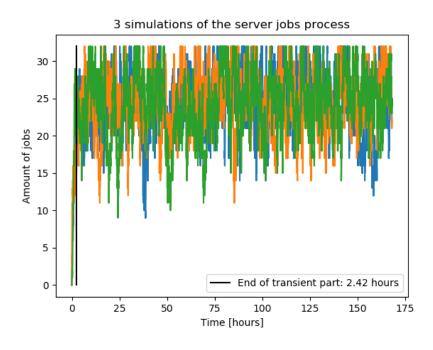


Figure 7: 3 simulations of the server job process

In figure 8 we have compared the theoretical long term probabilities with realizations from 100 simulations, to get as accurate results as possible. Clearly, the simulations correspond pretty well with the analytical calculations, which verifies the probabilities derived in a).

Finally, we also counted the amount of jobs being forwarded to another

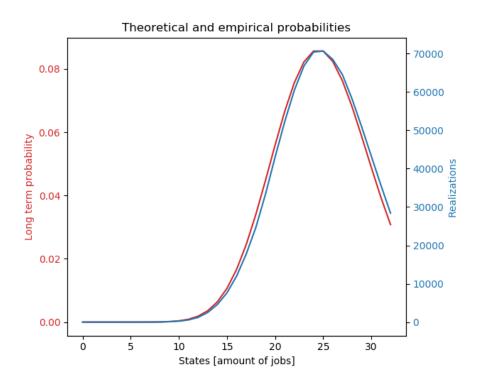


Figure 8: Comparison between the theoretical long term probabilities and the sum of realizations from 100 simulations

server in our simulations, and from 100 simulations the process averaged ≈ 0.75 forwarded jobs per hour.

Appendix

Source code problem 1

Source code for problem 1 are contained in the screen dumps below



Figure 9: Code for simulating the homogeneous Poisson process and its claim amounts



Figure 10: Code for simulating the inhomogeneous Poisson process and its claim amounts

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Average Claims after 120 days) []

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Figure 11: Source code to verify the theoretical probabilities in 1)

Source code problem 2

Source code for solving problem 2 are shown in the screen dumps below

Figure 12: Source code for calculating the long term probabilities in 2a)

Figure 13: Source code for simulating 2b)

Figure 14: Source code for plotting 2b)