

Switching Regression as Robust Estimation against Misclassification in Machine Learning Classification

Aleksandr Michuda

UC Davis Agricultural and Resource Economics

Introduction

- ▶ How does urban labor supply respond to agricultural income shocks?
 - ▶ Rural-Urban linkages
 - ▶ Driving Uber as insurance/income diversification
- ▶ Can machine learning help when location data is unavailable?
 - ▶ Requires less data
 - ▶ But introduces misclassification
 - ▶ **Can we develop an estimator that is robust to that misclassification?**

Data

- ▶ Uber Driver Data
 - ▶ Hours online
 - ▶ Earnings
- ▶ Weather Shocks
 - ▶ Drought Indices (SPI, NDVIA)
- ▶ Predicted Rural Place Origin
 - ▶ SAP Region (4)
 - ▶ FAO Agro-ecological Zones (10)
 - ▶ Distribution of probabilities

Prediction Table

Name	Central	East	North	West
Ahimbisibwe	0.144	0.003	0.001	0.925
Amin	0.149	0.057	0.651	0.140
Auma	0.040	0.267	0.674	0.017
Kadaga	0.148	0.797	0	0.054
Makubuya	0.964	0.018	0	0.017
Museveni	0.164	0.022	0.042	0.769
Oculi	0.015	0.263	0.717	0.003

Misclassification is a problem

- ▶ Predictions might contain misclassification error
 - ▶ For ex: we are classifying “Amin” into North, but it might be that they are actually more connected to the Center.
 - ▶ Not from any systematic bias during machine learning process
 - ▶ We can “imperfectly” break drivers into groups

What if I use OLS?

- ▶ We can estimate with OLS
- ▶ But response to drought will be attenuated.
- ▶ Is there a better way to estimate it?
 - ▶ We can model the misclassification directly
 - ▶ “What’s the probability that I categorize Amin into the North regime, given that they’re truly from the Center?, etc.”

Objectives Today

- ▶ Presenting a Maximum Likelihood Estimator
 - ▶ inspired by switching regression literature
- ▶ How does this estimator perform under varying levels of misclassification compared to OLS?
- ▶ Explore through Monte Carlo Simulations

The Hours Function

- ▶ Each regime $i \in I$ can be expressed as a linear function of SPI^i and $Hours$:

$$Hours = \beta_0^i + \beta_1^i SPI^i + \varepsilon^i$$

- ▶ The goal is to recover β_1^0 and β_1^1

Without Misclassification

- ▶ Suppose there's a true membership indicator, I .
- ▶ The conditional expectation function is then:

$$E(\text{Hours}|I, SPI^i) = 1\{I = 0\}(\beta_0^0 + \beta_1^0 SPI^0) + 1\{I = 1\}(\beta_0^1 + \beta_1^1 SPI^1)$$

- ▶ Without misclassification and with a separation indicator, we can recover β_1^0 and β_1^1 without bias, by using I as a variable in an OLS regression.

Misclassification in Regimes

- ▶ In our case we do not observe I , but we do observe r .
- ▶ r gives us a measure of I with *measurement error*.
- ▶ We can express the measurement error in terms of a matrix of conditional probabilities with $p_i^j = Pr(r = i | I = j)$.

-	$r = 0$	$r = 1$
$I = 0$	p_0^0	p_1^0
$I = 1$	p_0^1	p_1^1

- ▶ If $p_0^1 = p_1^0 = 0$, then there is no misclassification.

Conditional Expectation with Misclassification

- In the case of misclassification, the conditional expectation function is then:

$$\begin{aligned} E(\text{Hours}|r) = & \overbrace{(\beta_0^0 + \beta_1^1 SPI^0)}^{E(\text{Hours}|I=0)} \cdot (1-r) \cdot \overbrace{(1-\lambda)p_0^0}^{Pr(r=0,I=0)} + \\ & \overbrace{(\beta_0^1 + \beta_1^1 SPI^1)}^{E(\text{Hours}|I=1)} \cdot (1-r) \cdot \overbrace{\lambda p_0^1}^{Pr(r=0,I=1)} + \\ & \overbrace{(\beta_0^0 + \beta_1^0 SPI^0)}^{E(\text{Hours}|I=0)} \cdot r \cdot \overbrace{(1-\lambda)p_1^0}^{Pr(r=1,I=0)} + \\ & \overbrace{(\beta_0^1 + \beta_1^1 SPI^1)}^{E(\text{Hours}|I=1)} \cdot r \cdot \overbrace{\lambda p_1^1}^{Pr(r=1,I=1)} \end{aligned}$$

- $Pr(I = 1) = \lambda$

Using OLS with Misclassification

- ▶ If we use the same OLS strategy as before:

$$Hours = 1\{r = 0\} + 1\{r = 1\} + \beta_1^0 SPI^0 \cdot 1\{r = 0\} + \beta_1^1 SPI^1 \cdot 1\{r = 1\} + \varepsilon$$

- ▶ Leads to biased estimate, proportionate to extent of misclassification

- ▶ $ABias(\beta^0) = \frac{(1-\lambda)p_0^1}{p_0^0 + p_0^1} \cdot (\Sigma_{00}^{-1} \Sigma_{01} \beta^1 - \beta^0)$

- ▶ $\beta^r = [\beta_0^r \ \beta_1^r], \Sigma_{jk} = E(x_j' x_k)$

- ▶ $ABias(\beta^1) = \frac{\lambda p_1^0}{p_1^0 + p_1^1} \cdot (\Sigma_{11}^{-1} \Sigma_{10} \beta^0 - \beta^1)$

- ▶ $\beta^r = [\beta_0^r \ \beta_1^r], \Sigma_{jk} = E(x_j' x_k)$

- ▶ $x_r = [1 \ SPI^r]$

ML Approach

- ▶ Generalizing Lee and Porter (1985) to more than two regimes
 - ▶ Switching Regression with imperfect sample separation
- ▶ Flatten probabilities to a categorical
 - ▶ Take maximum of probabilities as truth, r

original_name	Central	East	North	West	Region Indicator (r)
Ahimbisibwe	0.144	0.003	0.001	0.925	West
Amin	0.149	0.057	0.651	0.140	North
Auma	0.040	0.267	0.674	0.017	North
Kadaga	0.148	0.797	0	0.054	East
Makubuya	0.964	0.018	0	0.017	Central
Museveni	0.164	0.022	0.042	0.769	West
Oculi	0.015	0.263	0.717	0.003	North

A Maximum Likelihood Alternative

- ▶ Each regime is normally distributed with mean $Hours - \beta_0^r - \beta_1^r SPI^r$ and standard deviation σ_r , with density f_r .
- ▶ We can then write the joint density of ε_r and r as:

$$f(\varepsilon_r, r) = f_0(\varepsilon_0) \left[r\lambda p_0^0 + (1-r)\lambda(1-p_0^0) \right] + \\ f_1(\varepsilon_1) \left[r(1-\lambda)(1-p_1^1) + (1-r)(1-\lambda)p_1^1 \right]$$

The Likelihood Function

- ▶ The likelihood function of the estimator is then:

$$\begin{aligned} L(\beta, \sigma, p, \lambda) = & \\ & [f_0(\varepsilon_{i1t})\lambda p_{11} + f_1(\varepsilon_{i1t})(1 - \lambda)p_{10}]^r \\ & \cdot [f_0(\varepsilon_{i0t})\lambda(1 - p_{11}) + f_1(\varepsilon_{i0t})(1 - \lambda)(1 - p_{10})]^{1-r} \end{aligned}$$

- ▶ We can maximize the log-likelihood to find optimal parameters for each of the parameters above.
- ▶ We can run Monte Carlo simulations of the MLE and an OLS analogue to compare the performance of the estimator.

Baseline Values for Simulation

- ▶ Data is modelled as crossection
- ▶ Actual data is panel

Parameter

Simulations in each = 200

Drivers = 275

Time periods = 10

Regimes = 2

$\sigma_0 = \sigma_1 = 1$

$E(SPI^0) = E(SPI^1) = 0$

$Var(SPI^0) = Var(SPI^1) = 1$

$Cov(SPI^0, SPI^1) = 0$

$\beta_0^0 = 20, \beta_0^1 = 35$

$\beta_1^0 = -1, \beta_1^1 = -2$

How is misclassification created?

Misclassification Plots $R = 2$

- Increase severity of misclassification

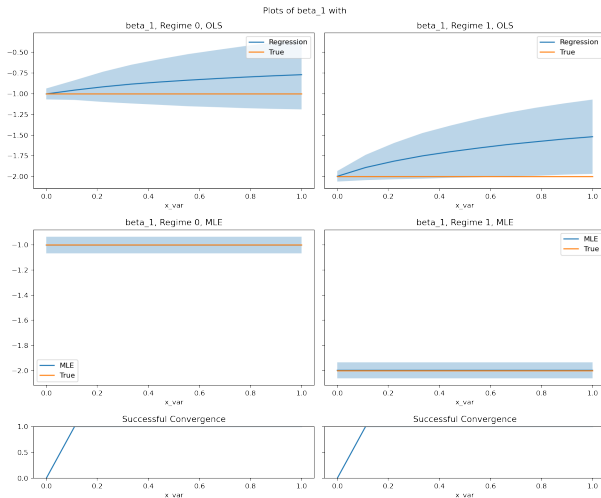


Figure 1: Increasing Misclassification $R = 2$

Generalizing to $R > 2$

- ▶ For $R > 2$, r becomes a categorical variable and we now use the mutually exclusive and exhaustive indicator functions for each regime, $G_i \equiv 1\{r = i\}$
- ▶ There are now $R - 1$, λ parameters
- ▶ The probability matrix will be an $R \times R$ matrix
- ▶ Consider $R=3$:

	$G_0 = 1$	$G_1 = 1$	$G_2 = 1$
$l = 0$	p_0^0	p_1^0	p_2^0
$l = 1$	p_0^1	p_1^1	p_2^1
$l = 2$	p_0^2	p_1^2	p_2^2

Generalizing to $R > 2$

- The likelihood function now becomes:

$$\begin{aligned} L(\beta, \sigma, \mathbf{p}, \lambda) = & \\ & \left[f_0(\varepsilon_{i0t})\lambda_0 p_0^0 + f_1(\varepsilon_{i0t})\lambda_1 p_0^1 + f_2(\varepsilon_{i0t})(1 - \lambda_0 - \lambda_1)p_0^2 \right]^{G_0} \\ & \cdot \left[f_0(\varepsilon_{i1t})\lambda_0 p_1^0 + f_1(\varepsilon_{i1t})\lambda_1 p_1^1 + f_2(\varepsilon_{i1t})(1 - \lambda_0 - \lambda_1)p_1^2 \right]^{G_1} \\ & \cdot \left[f_0(\varepsilon_{i2t})\lambda_0 p_2^0 + f_1(\varepsilon_{i2t})\lambda_1 p_2^1 + f_2(\varepsilon_{i2t})(1 - \lambda_0 - \lambda_1)p_2^2 \right]^{G_2} \end{aligned}$$

Baseline Values for Simulation ($R = 3$)

- ▶ Unless otherwise stated the values of each parameter in question will be as follows:

Parameter

Simulations in each=200

Drivers =275

Time Periods =10

Regimes =3

$\sigma_0 = \sigma_1 = \sigma_2 = 1$

$E(SPI_0) = E(SPI_1) = E(SPI_2) = 0$

$Var(SPI_0) = Var(SPI_1) = Var(SPI_2) = 1$

$Cov(SPI_j, SPI_k) = 0$

$\beta_0^0 = 10, \beta_0^1 = 20, \beta_0^2 = 35$

$\beta_1^0 = -1, \beta_1^1 = -2, \beta_1^2 = -3$

Misclassification Plot $R = 3$

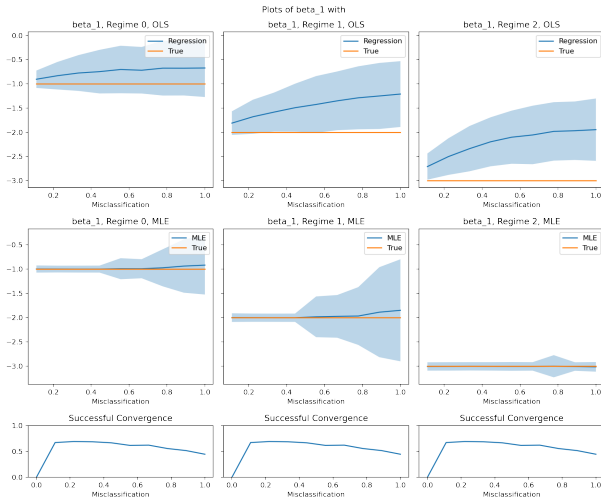


Figure 2: Increasing Misclassification $R = 3$

Conclusion

- ▶ MLE method is robust to misclassification
 - ▶ but converges less often with more regimes
 - ▶ better ways to specify function or calculate standard errors?
- ▶ How best to sell results?
- ▶ Regressions using real data require many regimes
 - ▶ OLS regressions suggest promising results

Misclassification 2 Beta 0

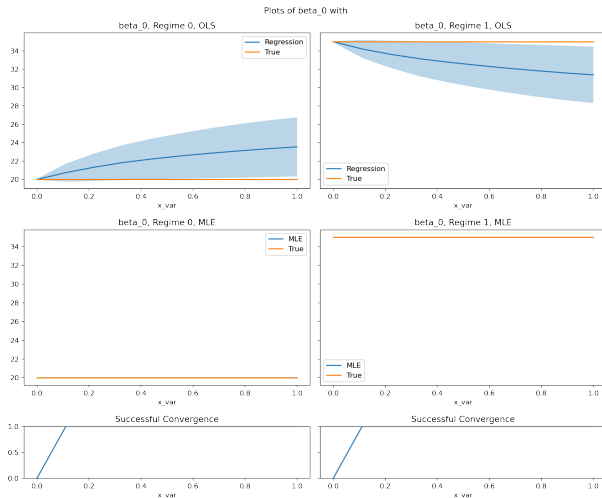


Figure 3: Changing STN of Hours

Misclassification 2 Sigma

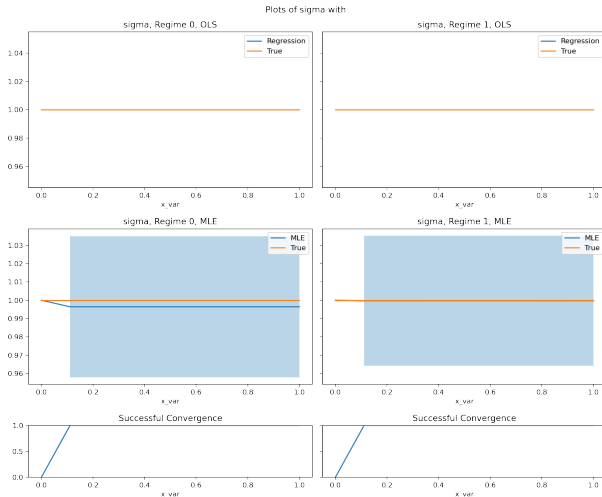


Figure 4: Changing STN of Hours

Misclassification 3 Beta 0

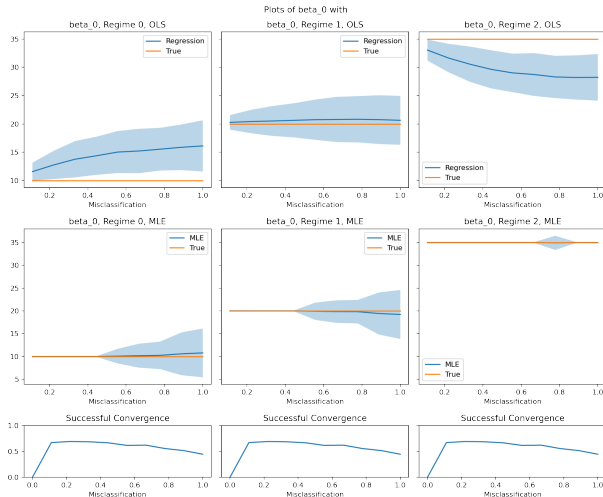


Figure 5: Changing STN of Hours

Misclassification 3 Sigma

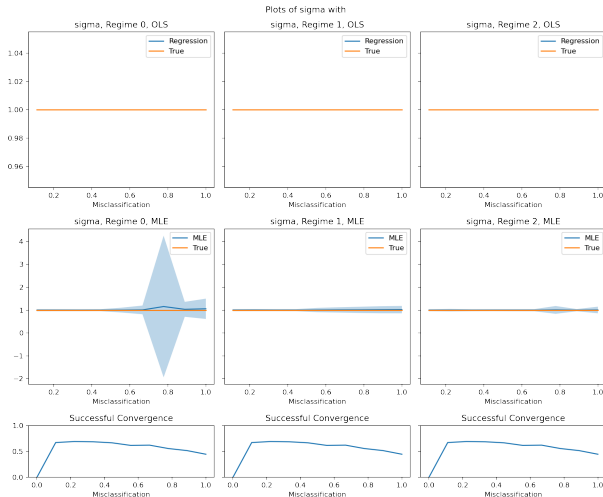
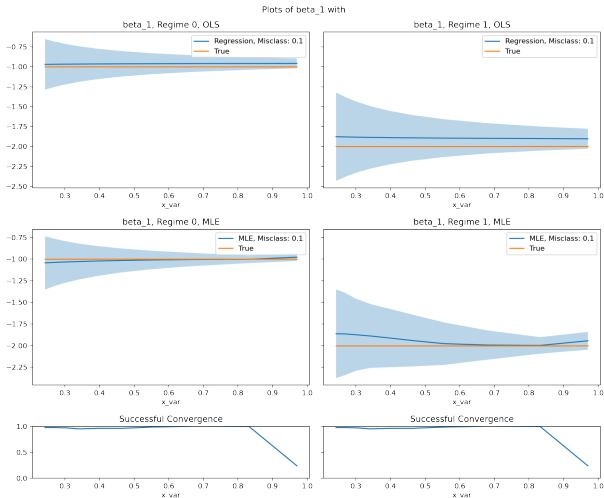


Figure 6: Changing STN of Hours

Noise to Signal Ratio of Hours

- ▶ Focus on increasing the signal to noise ratio symmetrically across the two regimes
- ▶ $STN = \frac{E(y_r)}{\sigma_r}$



Correlation of SPI Shocks

- Increase correlation between SPI variables
- Increase from 0 to 0.9

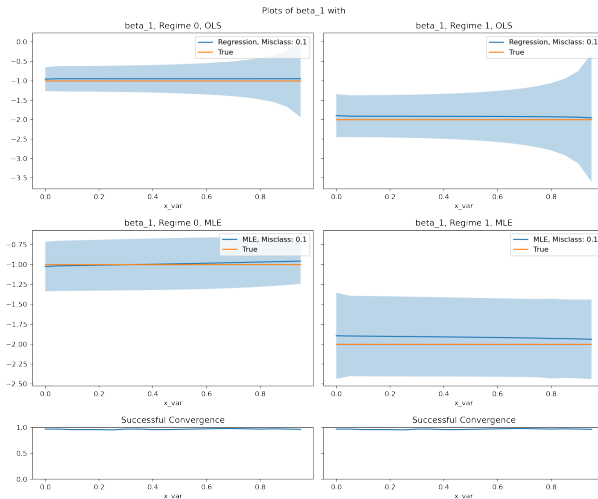


Figure 8: Changing Correlation of Drought Shocks

Difference across Regime Responses

- ▶ $\beta_1^0 = 0$
- ▶ β_1^1 ranges from 0 to 2

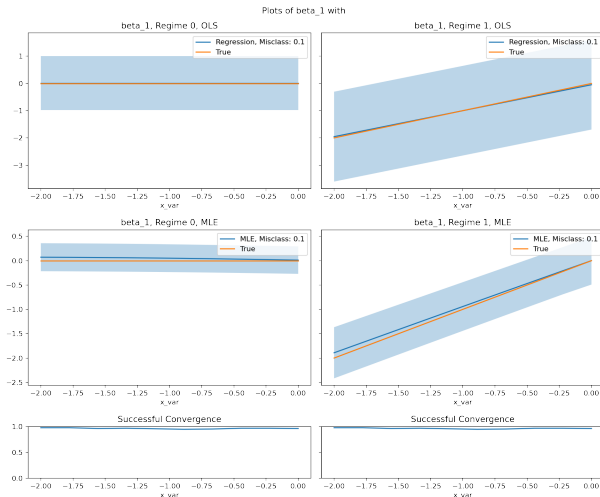


Figure 9: Regime Response Heterogeneity

Noise to Signal Ratio of Hours ($R = 3$)

- Same idea as before, but three regimes now

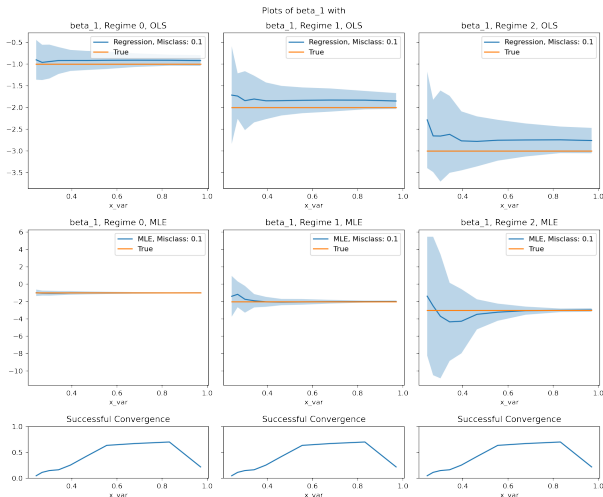


Figure 10: Changing STN of Hours

Correlation of Drought Shocks ($R = 3$)

- Increase correlation between drought variables
- Increase from 0 to 0.9

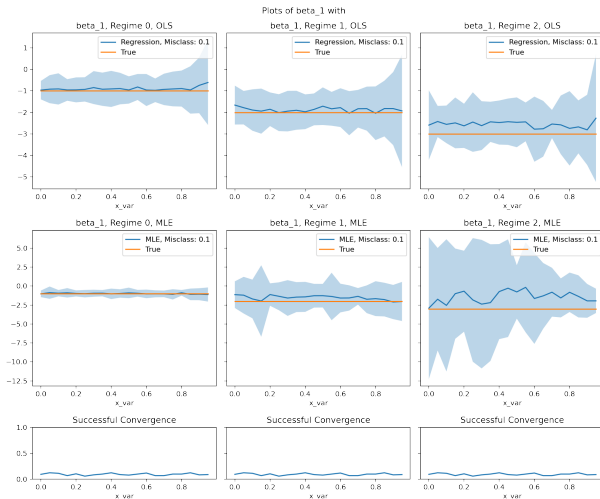


Figure 11: Changing Correlation of Drought Shocks, $R = 3$

Difference across Regime Responses ($R = 3$)

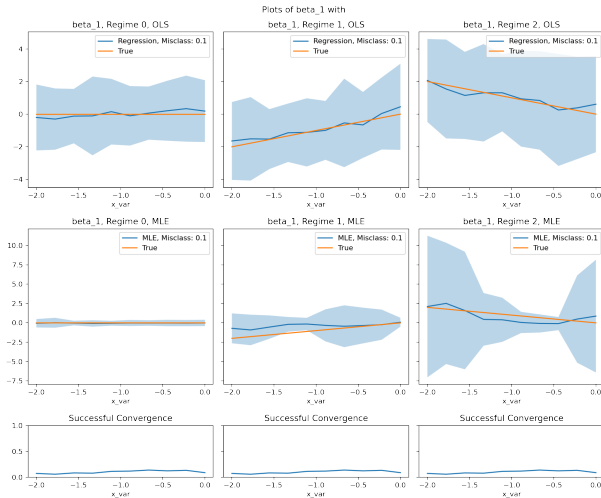
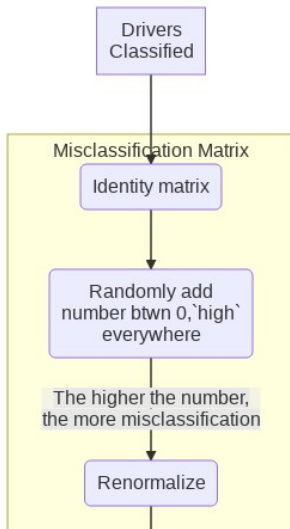


Figure 12: Changing Regime Response Heterogeneity, $R = 3$

Misclassification Procedure

- The misclassification matrix is a “jittered” matrix that introduces misclassification to the drivers after their memberships have already been chosen.



Two Regimes STN Sigma

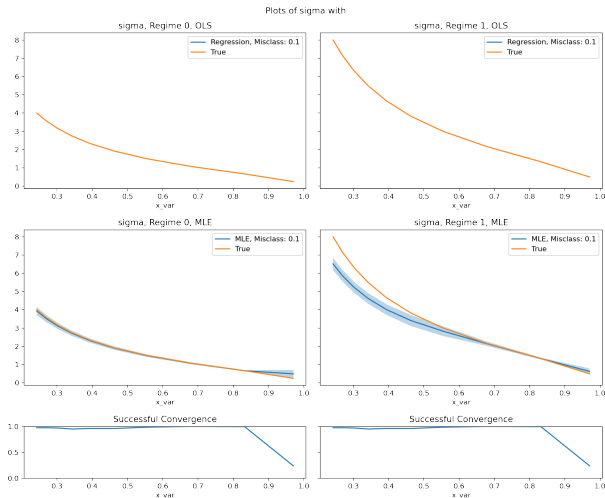


Figure 14: Two Regimes STN Sigma

Two Regimes STN Beta 0

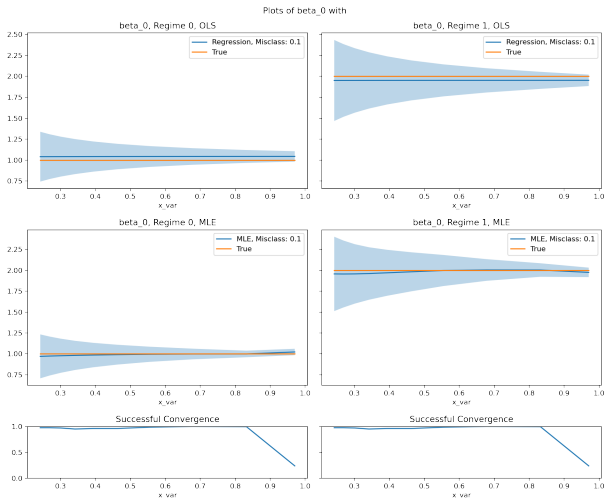


Figure 15: Two Regimes STN β_0

Two Regimes Drought Correlation Sigma

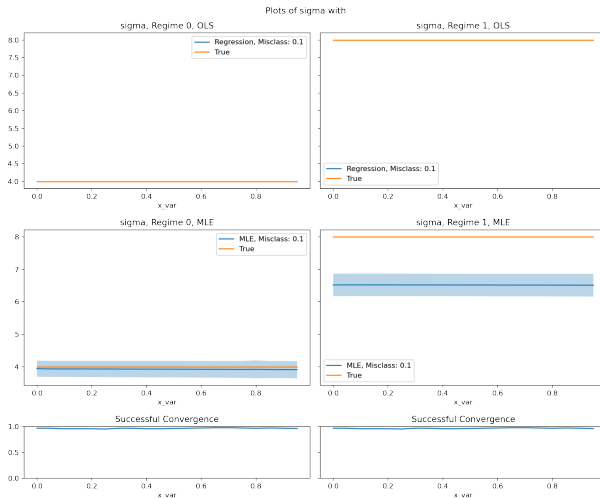


Figure 16: Two Regimes Drought Correlation Sigma

Two Regimes Drought Correlation Beta 0

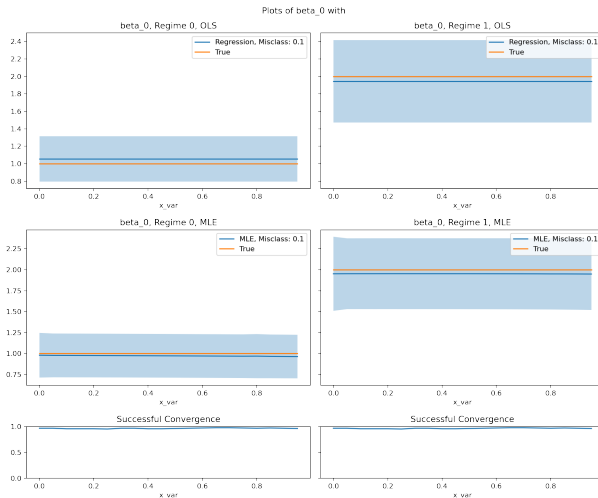


Figure 17: Two Regimes Drought Correlation Beta 0

Two Regimes Response Sigma

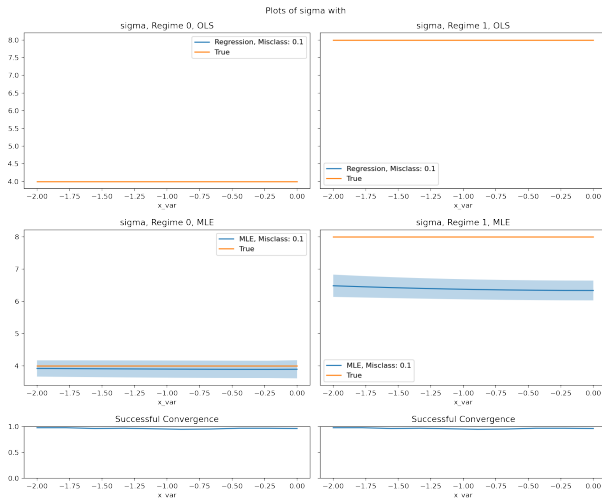


Figure 18: Two Regimes Response Sigma

Two Regimes Response Beta 0

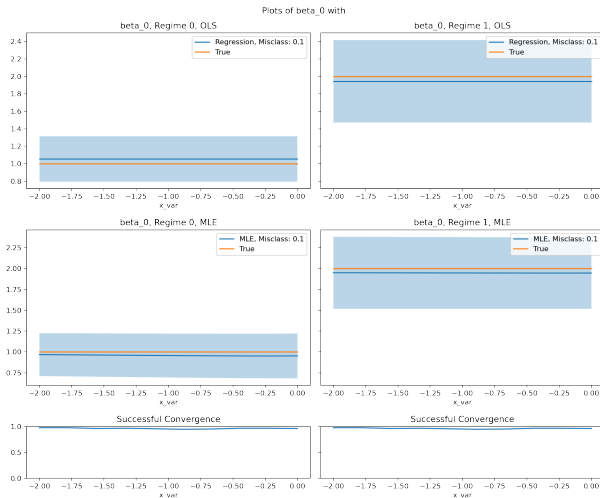


Figure 19: Two Regimes Response Sigma

Three Regimes STN Sigma

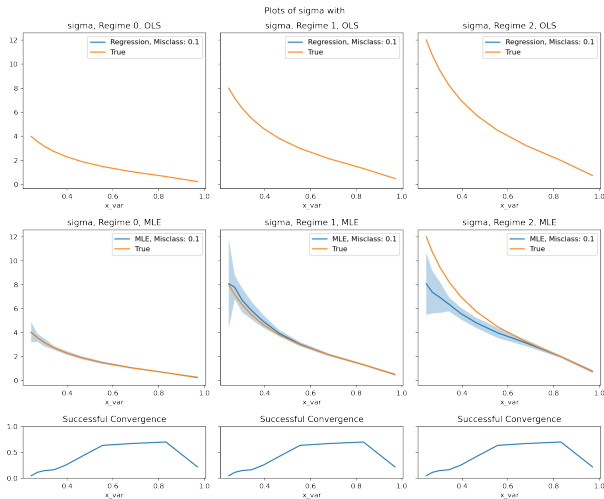


Figure 20: Three Regimes STN Sigma

Three Regimes STN Beta 0

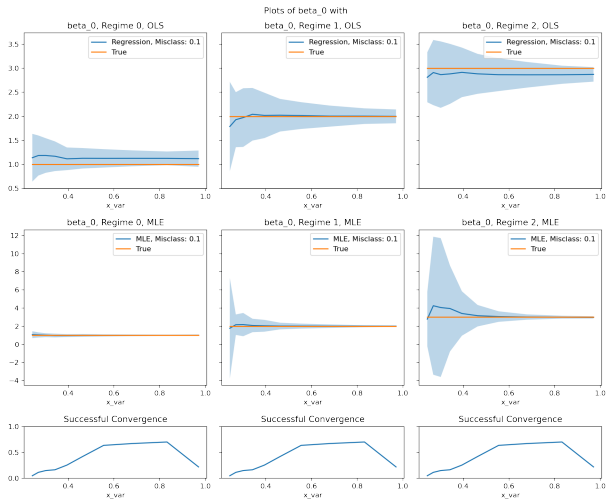
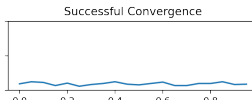
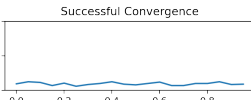
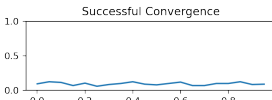
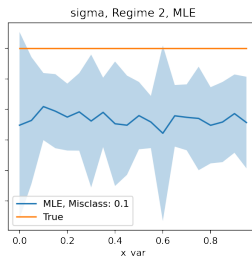
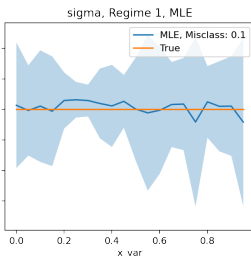
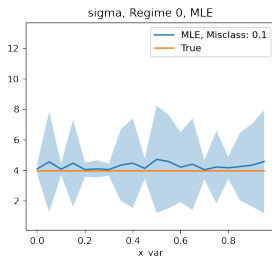
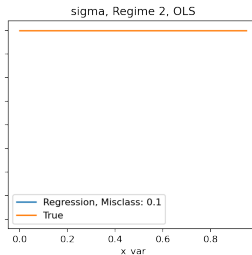
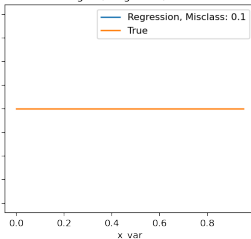
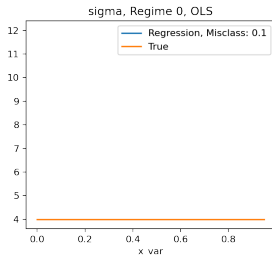


Figure 21: Three Regimes STN β_0

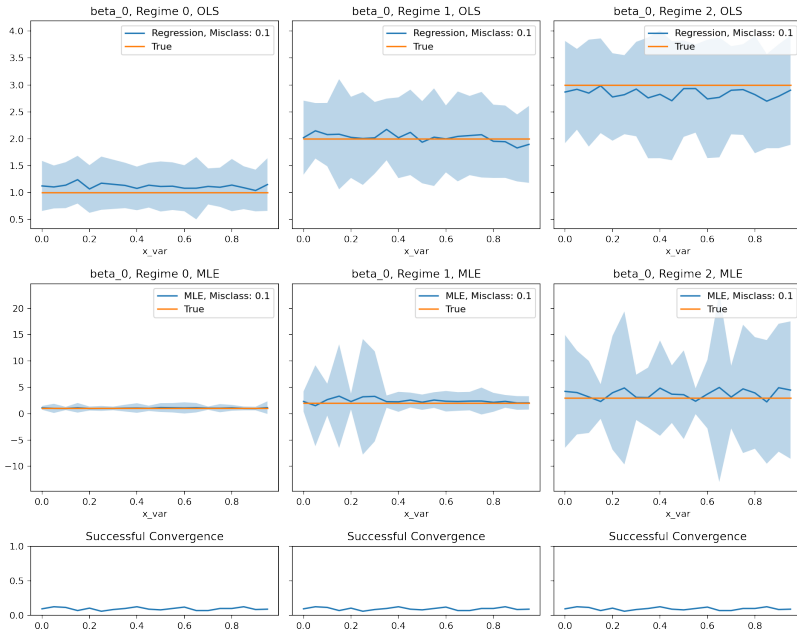
Three Regimes Drought Correlation Sigma

Plots of sigma with
sigma, Regime 1, OLS



Three Regimes Drought Correlation Beta 0

Plots of β_0 with



Three Regimes Response Sigma



Figure 24: Three Regimes Response Sigma

Three Regimes Response Beta 0

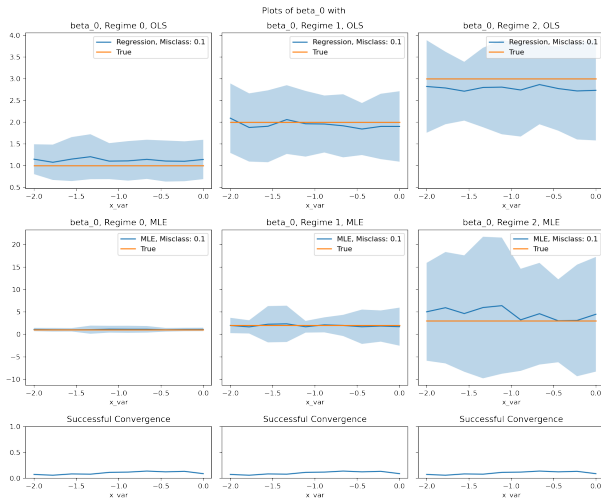


Figure 25: Three Regimes Response Sigma

OLS Regressions on Region

	Parameter Estimates					
	Parameter	Std. Err.	T-stat	P-value	Lower CI	Upper CI
region_class_Central	74.024	0.7525	98.372	0.0000	72.549	75.499
region_class_East	69.941	0.8790	79.570	0.0000	68.219	71.664
region_class_North	63.988	1.1560	55.354	0.0000	61.723	66.254
region_class_West	74.381	0.7921	93.902	0.0000	72.828	75.933
region_class_Central:lagged_Central	0.1891	0.0584	3.2357	0.0012	0.0746	0.3037
region_class_East:lagged_East	0.1619	0.0529	3.0634	0.0022	0.0583	0.2655
region_class_North:lagged_North	0.1621	0.0927	1.7485	0.0804	-0.0196	0.3437
region_class_West:lagged_West	0.1224	0.0551	2.2221	0.0263	0.0144	0.2303

Figure 26: Regression Results using SAP Region

MLE Estimates on Region

- ▶ Unavailable as MLE does not converge.