In The name of God



University Tehran

Engineering Facility

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Department



Pattern Recognition

Homework #1

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Abstract

In this home work we are going to some calculation and some coding the find the decision boundary of a distribution of different kind of function with one to two dimensions.

And also using the maximum likihood and other tools to make out calculations efficient.

Problem 1:

a)
$$P(x|w_i) = \frac{1}{\pi b} \frac{1}{1 + (\frac{x - a_i}{b})^2}$$
 $i = 1.2$

2, for Normalization 2,
$$\int_{\infty}^{\infty} P(x|\omega_i) dx \cdot 1 \rightarrow \frac{1}{\pi b} \int_{\infty}^{\infty} \frac{dx}{1 + (\frac{x - a_i}{b})^2} = \frac{1}{\pi} f_{g'}(\frac{x - a_i}{b})\Big|_{\infty}^{\infty}$$

$$\chi^{2}$$
 - $2\alpha_{1}\pi + \alpha_{1}^{2} = \chi^{2} - 2\alpha_{2}\pi + \alpha_{2}^{2} - \chi (2\alpha_{1} - 2\alpha_{1}) + \alpha_{2}^{2} - \chi$

$$\chi = \frac{(a_1 - a_1)(a_1 + a_1)}{2(a_1 - a_1)} = \frac{(a_1 + a_1)}{2} \quad \gamma, \quad \chi = \frac{\alpha_1 + \alpha_1}{2}$$

The Matlab code are in the file with the figure

$$\begin{cases} (x|\omega_1)R(\omega_1) \\ q_1 \\ q_2 \\ w \in k_{10} w \quad q_1 > \alpha_1 \ \ \ \ \end{cases} \xrightarrow{\begin{cases} \alpha_1 + \alpha_2 \\ 1 \\ 1 \end{cases}} \begin{cases} (\alpha_1 - \alpha_2) \\ = 1 \quad 1 \quad 1 \quad |\alpha_2 - \alpha_1| \end{cases}$$

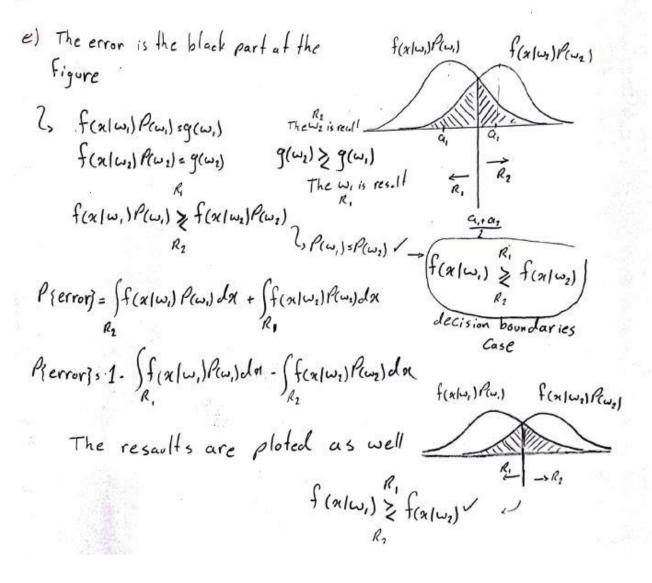
f(xlw,)Pa,)

$$\frac{1}{2\pi} \left(\frac{4}{9} \left(\frac{\alpha_{2} - \alpha_{1}}{2b} \right) + \frac{\pi}{2} + \frac{\pi}{2} - 4 \frac{\pi}{2} \left(\frac{\alpha_{1} - \alpha_{1}}{2b} \right) \right) = \frac{1}{2\pi} \left(\pi + 2 \frac{4}{9} \left(\frac{\alpha_{1} - \alpha_{2}}{2b} \right) \right) = \frac{1}{2} - \frac{1}{\pi} \frac{4}{9} \left(\frac{|\alpha_{1} - \alpha_{1}|}{2b} \right)$$

d) The maximum error acourse when $a_1 = a_2$ witch means.

that the two function are on each other witch means that

they both have the same chance $P\{error] = \frac{1}{2} - \frac{1}{\pi} \frac{1}{g^2} \left| \frac{a_2 - a_1}{2b} \right| \rightarrow a_2 = a_1$ $P\{error\} = \frac{1}{2}$ for this case the chance for w_1 or w_2 is always equal.



f) Risks 2 Sp. dx 2 2 is P(w) f(x/w) V, = 2, P(x/w,) P(w,) + 2,1 P(x/w,) P(w,) 12 = 222 P(x/U2) P(W2) + 212 P(x/W2) P(W2) to chose 2, r. The minimum error accours.

\[\lambda_{12} P(\delta|\omega_1) P(\omega_1) + \lambda_2 \lambda_1 P(\delta|\omega_1) P(\omega_2) \geq \lambda_2 \lambda_1 P(\delta|\omega_1) P(\omega_1) + \lambda_2 \lambda_1 P(\omega_1) P(\omega_1) P(\omega_1) \geq \lambda_2 \lambda_2 \lambda_1 P(\omega_1) P(\omega_1) + \lambda_2 \lambda_2 P(\omega_1) P(\omeg 2. 22 -> P(x/w2)P(w2) > 2P(x/w.)P(w,) -> P(w.)5P(w2) 222 P(XIW, IP(W) $\frac{1}{\pi b} \frac{1}{1 + \left(\frac{\chi - \alpha_1}{b}\right)^2}, \frac{1}{\pi b} \frac{2}{1 + \left(\frac{\chi - \alpha_1}{b}\right)^2} \longrightarrow$ x2-2a2x+a22+b2=2x2-4a,x1+2a,22b22->x2+(4a1+2a2)x+(2a,2-a2-2b2)=. $\mathcal{X} = (\alpha_2 + 2\alpha_1) + \sqrt{(\alpha_2 - 2\alpha_1)^2 - 4(2\alpha_1^2 - \alpha_1^2 - 2b^2)} = (2\alpha_1 - \alpha_2) + \sqrt{2\alpha_2^2 + 2\alpha_1^2 - 4\alpha_2 \alpha_1 - b^2}$ solve 2, two ans

12721

Risk = 12 f(n/w,) P(w,) dn + f(n/w,) f(n/w) do

Risk = 12 f(n/w,) P(w,) dn + f(n/w) f(n/w) do touch each other of feature Rounders = (1g'(2-a1) + 1g'(2-a1) + 1 (1g'(21-a1) + 1/2 (1/2-a1) + 1/2 (1/2-1g'(21-a1)) Risk = 1 (19"(12-a1) - 19"(12-a1) + 1 (19"(12-a1) + 1) + 1 (19"(12-a1))

Pierror 3 2 f(x/w,)dx + of f(x/w)dx + of f(x/w)dx

 $\frac{1}{2\pi} \left(\frac{1}{g} \left(\frac{\gamma_2 - \alpha_1}{b} \right) - \frac{1}{g} \left(\frac{\gamma_1 - \alpha_1}{b} \right) + \frac{1}{2\pi} \left(\frac{1}{g} \left(\frac{\gamma_1 - \alpha_2}{b} \right) + \frac{\pi}{2} \right) + \frac{1}{2\pi} \left(\frac{\pi}{2} - \frac{1}{g} \left(\frac{\gamma_2 - \alpha_2}{b} \right) \right)$ In the case we minimised if with risk the error

Increases \implies if $\gamma_1 c \gamma_2 \notin \mathbb{R}$ — They don't touch each other

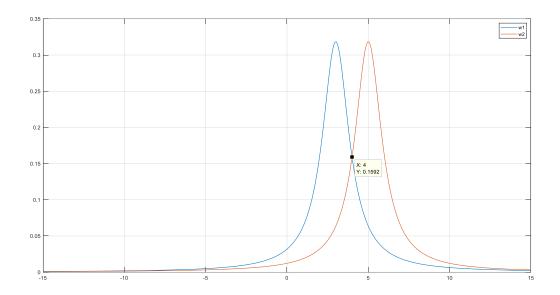


Figure 1

Figure 1 shows that where $x = \frac{a_1 + a_2}{2}$ in that position $P(\omega_1 | x) = P(\omega_2 | x)$ and we have showed that $\frac{3+5}{2} = 4$ witch shows that this is correct.

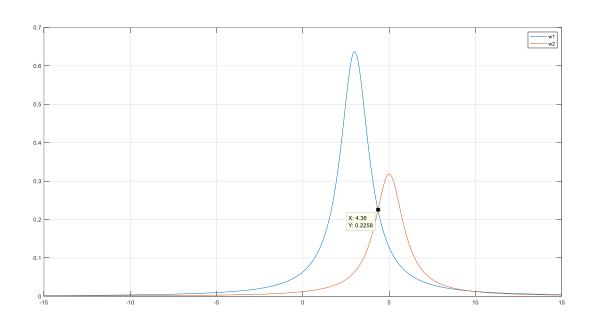


Figure 2

In Figure 2 we can see that when we use rick value as well because they are like a constant variable the will change the border of the 2 regions.

Problem 2:

a) Risk,
$$\sum_{i=1}^{C} \sum_{j=1}^{C} \lambda_{ij} P(x|\omega_{i}) P(\omega_{i})$$
 $\delta^{*} = \alpha r_{g} \min_{i=1}^{C} \left(\sum_{i=1}^{C} \lambda_{ij} P(x|\omega_{i}) P(\omega_{i}) \right) \rightarrow C.2$
 $\lambda_{11} P(x|\omega_{i}) P(\omega_{i}) + \lambda_{21} P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{f} \rightarrow compaire$
 $\lambda_{11} P(x|\omega_{i}) P(\omega_{i}) + \lambda_{22} P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{f} \rightarrow compaire$
 $\lambda_{11} P(x|\omega_{i}) P(\omega_{i}) + \lambda_{21} P(x|\omega_{i}) P(\omega_{i}) \xrightarrow{f} \lambda_{12} P(x|\omega_{2}) P(\omega_{2}) + \lambda_{22} P(x|\omega_{2}) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \lambda_{12} P(x|\omega_{2}) P(\omega_{2}) + \lambda_{22} P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{21} \right) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{11} \right) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{11} \right) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{21} \right) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{21} \right) P(x|\omega_{2}) P(\omega_{2}) \xrightarrow{g} \left(\lambda_{21} - \lambda_{21} \right) P(x|\omega_{2}) P(x|\omega_{2})$

b) Liblihood ->
$$\frac{P(x|\omega_1)}{P(x|\omega_2)} \stackrel{R_1}{\underset{R_2}{\nearrow}} \frac{(\lambda_{21} - \lambda_{22})}{(\lambda_{12} - \lambda_{11})} \cdot \frac{P(\omega_1)}{P(\omega_1)} -> \frac{P(x|\omega_2)}{P(x|\omega_1)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{(\lambda_{21} - \lambda_{22})} \cdot \frac{P(\omega_1)}{P(\omega_2)} \stackrel{R_1}{\underset{R_2}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(x|\omega_2)} \cdot \frac{P(x|\omega_1)}{P(x|\omega_1)} \stackrel{R_1}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(x|\omega_2)} \cdot \frac{P(\omega_1)}{P(x|\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(x|\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(x|\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(\omega_1)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(\omega_2)} \stackrel{R_2}{\underset{R_1}{\nearrow}} \frac{(\lambda_{12} - \lambda_{11})}{P(\omega_1)} \stackrel{R_2}{\underset{R_2}{\nearrow}} \frac{(\lambda_{12} - \lambda_{$$

Problem 3:
$$P(x|\omega_i)_s \frac{1}{\pi c} \frac{1}{1+(\frac{x-a_i}{b})} \sum_{i=1}^{n} \frac{x-a_i}{b}$$
 $P(x|\omega_i)_i P(\omega_i)_i \geq P(x|\omega_i)_i P(\omega_i)_i \geq \frac{P(\omega_i)}{1+(\frac{x-a_i}{b})} \sum_{i=1}^{n} \frac{P(\omega_i)}{(\frac{x-a_i}{b})} \frac{P(\omega_i)_i}{1+(\frac{x-a_i}{b})} \frac{P(\omega_i)_i}{1+(\frac{x-a_i}{b})}$

the distance is the corporator for an amount of I -> we can say that

the base of the decision is the distance from

the middle of x vector -> we soil that the

b is constant

but lit we do it completly v

$$\frac{P(\omega_1)}{1 + \left(\frac{\varkappa_1 - \alpha_1}{b}\right)^T \sum_{i=1}^{l-1} \left(\frac{\varkappa_1 - \alpha_1}{b}\right) R_2} \frac{P(\omega_1)}{1 + \left(\frac{\varkappa_2 - \alpha_2}{b}\right)^T \sum_{i=1}^{l-1} \left(\frac{\varkappa_2 - \alpha_2}{b}\right)}$$

$$\frac{P(\omega,1)}{1-\left(\frac{\alpha_{n_1}-a_{n_1}}{b},\frac{\alpha_{n_2}-a_{n_2}}{b}\right)\left[\frac{\gamma_1}{\gamma_2}\frac{\gamma_2}{\gamma_3}\right]\left(\frac{\alpha_{n_1}-a_{n_1}}{\gamma_2}\right)}{\frac{\gamma_1}{\gamma_2}\left(\frac{\alpha_{n_1}-a_{n_2}}{b}\right)} + \frac{\gamma_2}{b}\left(\frac{\alpha_{n_2}-a_{n_2}}{b}\right) + \frac{\gamma_2}{b}\left(\frac{\alpha_{n_2}-a_{n_2}}{b}\right)$$

$$\times \left(\frac{\frac{\mathcal{H}_{11} - \alpha_{11}}{b}}{\frac{\mathcal{H}_{12} - \alpha_{12}}{b}} \right) = \frac{\mathcal{V}_{1}}{b^{2}} \left(\mathcal{H}_{11} - \alpha_{11} \right)^{2} + \frac{\mathcal{V}_{3}}{b^{2}} \left(\mathcal{H}_{12} - \alpha_{12} \right) \left(\mathcal{H}_{11} - \alpha_{11} \right) + \frac{\mathcal{V}_{2}}{b^{2}} \left(\mathcal{H}_{12} - \alpha_{12} \right) \left(\mathcal{H}_{11} - \alpha_{11} \right) + \frac{\mathcal{V}_{2}}{b^{2}} \left(\mathcal{H}_{12} - \alpha_{12} \right)^{2}$$

$$\frac{\gamma_{1}}{b^{2}}(x_{11}-\alpha_{11})^{2}+\frac{\gamma_{4}}{b^{2}}(x_{12}-a_{12})^{2}+\frac{(\gamma_{3}+\gamma_{2})}{b^{2}}((y_{12}-a_{12})(x_{11}-a_{11})) -$$

rest of Prob 3: $P(\omega_{1})$ $b^{2}+\chi(\chi_{11}-\alpha_{11})^{2}+\gamma_{q}(\chi_{12}-\alpha_{12})^{2}+(\gamma_{3}+\gamma_{2})((\chi_{12}-\alpha_{12})(\chi_{11}-\alpha_{11})) \frac{R_{1}}{R_{2}} \frac{P(\omega_{2})}{b^{2}+N_{1}(\chi_{21}-\alpha_{21})^{2}+N_{q}(\chi_{22}-\alpha_{22})}$ $(N_{3}+N_{2})((\chi_{21}-\alpha_{22})(\chi_{21}-\alpha_{21})) \quad \text{so we showed that in witch case}$ $(N_{3}+N_{2})((\chi_{21}-\alpha_{22})(\chi_{21}-\alpha_{21})) \quad \text{we are going to get the result in ang}$ $Condition \quad \text{we are going to have a result ether } R_{1}, \text{ or } R_{2}$ This could be the view from the top

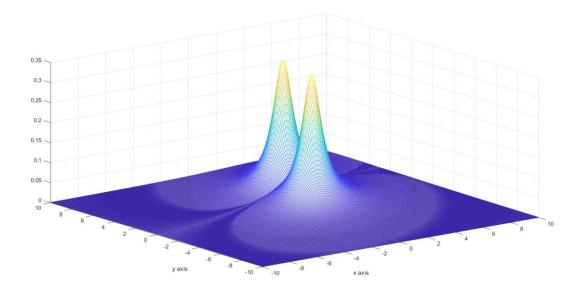


Figure 3

In Figure 3 the
$$a_1=(2.1)$$
 , $a_2=(3.5)$, $b=1, \Sigma=\begin{bmatrix} 1 & 0.5 \\ 0.5 & 1 \end{bmatrix}$

According to Figure 3 me can draw a decision graph like the surface below

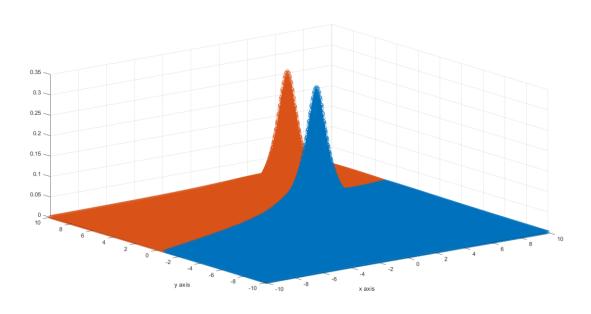


Figure 4

The Figure 4 shows that in every location in the map we can make a decision with higher chance of success.

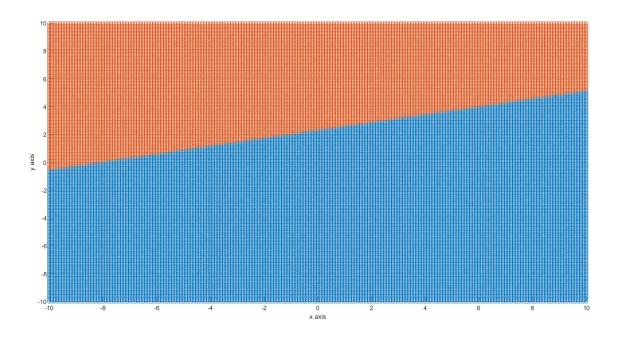


Figure 5

Figure 5 we can see from the top view that with a sample of (X, Y) the location that we choose has superior decision the blue region is for ω_1 and the red region is for region ω_2

Problem 4:
$$P(x|\omega_{i}): \begin{cases} \frac{x}{\sigma_{i}^{2}} \exp\left(\frac{-x^{2}}{2\sigma_{i}^{2}}\right) & \text{x.c.} \\ \\ P(x|\omega_{i}): \end{cases} \begin{cases} \frac{x}{\sigma_{i}^{2}} \exp\left(\frac{-x^{2}}{2\sigma_{i}^{2}}\right) & \text{x.c.} \end{cases} \end{cases}$$

$$P(x|\omega_{i}): \begin{cases} \frac{x}{\sigma_{i}^{2}} \exp\left(\frac{-x^{2}}{2\sigma_{i}^{2}}\right) & \text{x.c.} \end{cases} \end{cases}$$

$$P(x|\omega_{i}): P(\omega_{i}) > P(x|\omega_{i}) P(\omega_{2}) \rightarrow \text{Assume equal priors this means} \end{cases}$$

$$P(\omega_{i}): P(\omega_{2}) \rightarrow P(x|\omega_{2}) P(x|\omega_{2}) \rightarrow \text{Assume equal priors this means} \end{cases}$$

$$P(\omega_{i}): P(\omega_{2}) \rightarrow P(x|\omega_{1}) P(x|\omega_{2}) \rightarrow \text{Assume equal priors this means} \end{cases}$$

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$$P(\omega_{i}): P(\omega_{2}) \rightarrow P(x|\omega_{1}) P(x|\omega_{2}) \rightarrow \text{Assume equal priors this means}$$

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$$P(\omega_{i}): P(\omega_{2}) \rightarrow P(x|\omega_{2}) P(x|\omega_{2}) \rightarrow \text{Assume equal priors this means}$$

$$P(\omega_{i}): P(\omega_{i}): P(\omega_{i}): P(\omega_{i}) P(\omega_{i})$$

20,
$$\sigma$$
, $\sqrt{\frac{l_n(\sigma_i) \cdot l_n(\sigma_i)}{\sigma_2^2 - \sigma_i^2}} \gtrsim \chi$ so we can see that χ has a range $\chi \in \mathbb{R}_2$. In this case we can decide $\chi \in \mathbb{R}_1$ or $\chi \in \mathbb{R}_2$.

 $\chi \in \mathbb{R}_1$ or $\chi \in \mathbb{R}_2$.

 $\chi \in \mathbb{R}_1$ or $\chi \in \mathbb{R}_2$.

 $\chi \in \mathbb{R}_1$ the decide $\chi = 2\sigma_2\sigma_1 \sqrt{\frac{l_n(\sigma_1) \cdot l_n(\sigma_1)}{(\sigma_2^2 - \sigma_1^2)}}$.

Store:
$$\mu_{x} = 1.333$$
 $\mu_{y} = 1.6111$

Stor $\begin{bmatrix} \sigma_{x}^{2} & \sigma_{y} \\ \sigma_{x} & \sigma_{y} \end{bmatrix} = \begin{bmatrix} 0,5556 & 0.1852 \\ 0.1852 & 0.9877 \end{bmatrix}$

Lote: $\mu_{x} = -0.0409$
 $\mu_{y} = -0.1818$
 $\begin{bmatrix} \sigma_{x}^{2} & \sigma_{y} \\ \sigma_{y} & \sigma_{y}^{2} \end{bmatrix} = \begin{bmatrix} 1.7677 & -0.0202 \\ -0.0202 & 0.5707 \end{bmatrix}$

$$P(x|w_i) \underset{x \in R_2}{\text{Re}} P(x|w_i)$$
 $\binom{x_i}{y_i} \checkmark -> \text{the calculation}$ $\binom{x_i}{y_i} \checkmark$

c) The drawing has been done in mollab

(x, = 221 P(21w2)P(U2) xeP.

rishir P(alwi)P(wi) -> rizrz -> P(wi) sP(wi) V

 $\lambda_{11}=\alpha \rightarrow P(\alpha 1\omega_2) \gtrsim P(\alpha 1\omega_1) \times 2 \rightarrow \alpha \in R_2$

 $\frac{1}{\sqrt{2\pi}} \frac{1}{|\Sigma_1|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_2)\right) \geq \frac{1}{\sqrt{2\pi}} \frac{1}{|\Sigma_1|^{\frac{1}{2}}} \exp\left(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)\right)$ This case we decide the sample is for R, or R₂/

P(x/w,) & P(x/w) >2 = we reaped part b.

[2,3] -> star -> reality / [1,3] -> star -> centity / Star -> classify

[2.5,2] -> star -> reality / [1.5,2] -> star -> reality/ 2 star -> classify

[0.5,2] -> star -> reality / [2,1] -> star -> reality /

[1,1] 2 star - reality [0,05] 2 star - reality & wrong

[1.5,0] -star - reality [1.5,1] - date - reality /

[1.5, 0.5] - dote-sreality ([0.5, 0.5] - dote-sreality (wrong) Star-sclassify

[-0.5, 0.5] -dote-reality / [-1.5,0]-rdote-reality / 2 dote-sclosefy / [-1.5,0]-rdote-reality

[4.5,-0,5] -> dote -> reality [0,5,-0.5] -> dote -> reality / Z, dote -> clossify Z, dote -> clossify

[1,-1] -> dote -> reality [-1,-1] -> dote -> reality / 2, dote -> classify

[-2,-1] -> dole -> reality 3 wrong answers we

have [15.8] our classifier misclussifies

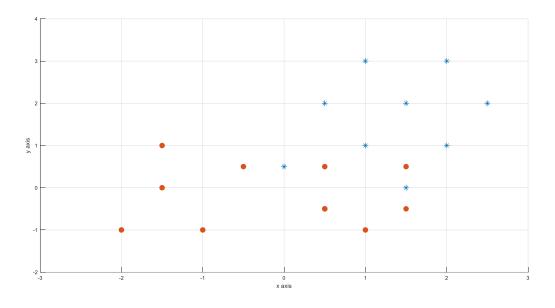


Figure 6

In Figure 6 the samples are showed in the figure and what we are going to find the distribution of the samples.

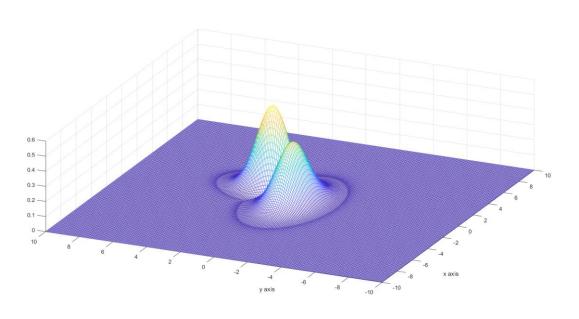


Figure 7

Figure 7 is the 3D view of the Gaussian two dimensional distribution.

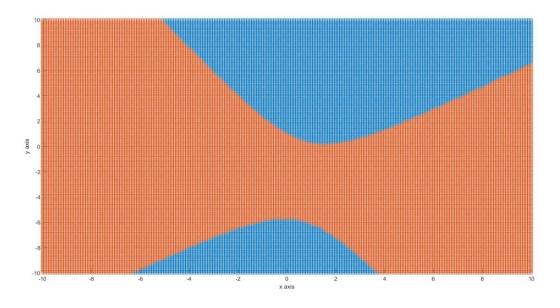


Figure 8

Figure 8 is the top view for the previous distribution we can see the decision boundary of it. (the figure from plot is 3D you should turn it to see this image correctly)

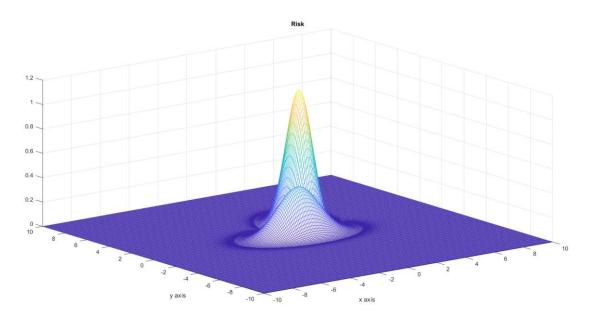


Figure 9

In Figure 9 we can see the rick diagram.

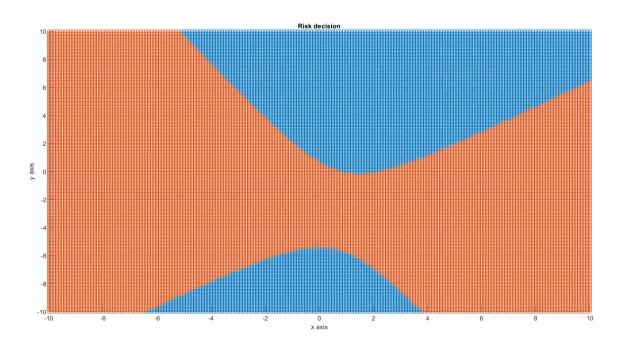


Figure 10

In Figure 10 we can see the decision region that shows in every (X, Y) we can decide either our decision to be ω_1 or ω_2 .

If we calculate the decision boundary when $P(\omega_1) = \frac{1}{3}$ and $P(\omega_2) = \frac{2}{3}$ it's like the when we found the region when we had ricks but the difference is the amplitude of the probability

Problem 6: 2 = (21, cx 2 c... cx d) + (3 = Prob (21 = 1/w)

- a) It means the probability of the ith parameter from x vector to become 1 if we know that we are in the wij space and we has accord in that case, and we know the we have c space for wij and the x vector has all parameter
- b) \rightarrow we know for deciding $g(n) = f(x|w_j) P(w_j) \rightarrow$ for showing $P(x_j) = P(x_j) = P(x_j) P(w_j) \rightarrow P(x_j) = P(x_j) =$

9 (21) = ln (f(21wi)) + ln (P(wi)) -> ln (f(21wi)) = [ln(Pi) + [ln(Pi)] + [ln

9,(2) = 2 2; ln(Pij) + [ln(1-Pij) + P(wi))

Problem 7: Ex is the probility to get wrong answere for wis space Es = I PENERIW 3 = 1-PENERIWS. (, E, = 1 - P{xER, [w,] s P{xER, [w,] E2=1-P{xER2/W2}=P{xER; 1W2} \$\lambda_{21} PK XER, and XEW_3 + \$\lambda_{22} PixER_2 and XEW_3 + \$\lambda_{12} PixER_2 and XER,} PEXER; and REW; 3, PEXER, IW, 3 P(W) PEXER, IW, 3 = 1-E, PIXER, W. ? SE. Y = 2,1 P(xER, |w.)P(w.) + 2,1 P(xER, |w.) P(w.) P{x = n1 | w 2 } 1 - E2 + 222 Piner21 w2) P(w) + 212 Pixer2 wy) P(w) PEXERILUZZEZ = 2,1 (1-E,) P(w,) + 2,1 E, P(w,) + 2,2 (1-E,2) P(w,) + 2,2 E, P(w,) r=p(w,) 2,+ P(w,) 2,2+p(w,)(2,2-2,1)E,+p(w,)(2,1-2,2)E2

Problem 8: a) Risk =
$$\sum_{i=1}^{c+1} \sum_{j=1}^{c} \lambda_{ij} P\{x|\omega_{j}, \lambda_{\omega_{i}}\} = \sum_{i=1}^{c+1} \sum_{j=1}^{c} \lambda_{ij} P\{x|\omega_{j}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\} P\{\omega_{i}\}\} P\{\omega_{i}\} P\{\omega_{i}$$

$$i = \arg\min_{i \neq 1} g_i(x) - 2g(x) = \sum_{i \neq 1} \lambda_s (1 - P(w_i|x_i)) + \lambda_r P(w_{c+1}|x_i)$$

$$g_i(x_i) = \begin{cases} \lambda_s (1 - P(w_i|x_i)) & i \neq c+1 \\ \lambda_r \{P(w_{c+1}|x_i)\} & i = c+1 \end{cases}$$
the ith sentence

if As P(we+1/2) > 2r(1-P(w; 12)) Vi izctiv

arg ming. (n) = ary max (P(wilm)) -> In the critical condition

P(witing) > P(wila) & arrivering > 28(1-P(wila))

 $\frac{\lambda_r}{\lambda_s} P(\omega_{cn}|\underline{n}) \geq (1 - P(\omega_i|\underline{n})) \rightarrow P(\omega_i|\underline{n}) \geq 1 - \frac{\lambda_r}{\lambda_s} P(\omega_{c+1}|\underline{n}) \rightarrow P(\omega_i|\underline{n}) \geq 1 - \frac{\lambda_r}{\lambda_s} P(\omega_{c+1}|\underline{n}) \rightarrow P(\omega_i|\underline{n}) \geq 1 - \frac{\lambda_r}{\lambda_s}$

Second way:

$$R(\omega; |\underline{x}) = \sum_{d=1}^{C} \lambda_{ij} P(\omega_{i}|\underline{x}) = \sum_{d=1}^{C} \lambda_{s} P(\omega_{i}|\underline{x}) = \lambda_{s} (1 - R(\omega_{i}|\underline{x}))$$

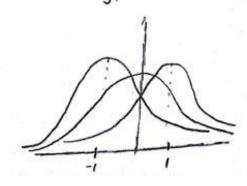
$$(ritical condition $\lambda_{s} R(\omega_{c+1}|\underline{x}) = \lambda_{s} (1 - P(\omega_{i}|\underline{x})) \rightarrow$$$

$$\lambda_r \geq \lambda_s \left(1 - \rho(\omega_i | \underline{x})\right) \rightarrow \frac{\lambda_r}{\lambda_s} \geq 1 - \rho(\omega_i | \underline{x}) \rightarrow \rho(\omega_i | \underline{x}) \geq 1 - \frac{\lambda_r}{\lambda_s}$$

if areo wil means that there is no threat against rejecting so it's better to reject anyway and there is no risk in rejecting.

if 2r=2s It means the risk of rejecting is quit higher then the error so we must never rejeting In this case we should alway choose and never reject b) y(n): { f(x(w)) P(w) ist, ... c \[\frac{\lambda_{5} - \lambda_{r}}{\lambda_{5}} \frac{\cappa_{6}}{\lambda_{7}} \frac{\cappa_{6}}{\lambda_{8}} \frac{\cappa_{6}}{\lambda_{7}} \frac{\cappa_{6}}{\lambda_{7}} \frac{\cappa_{6}}{\lambda_{7}} \frac{\cappa_{7}}{\lambda_{8}} \frac{\cappa_{8}}{\lambda_{8}} \frac{\cappa_{7}}{\lambda_{8}} \frace{\cappa_{7}}{\lambda_{8}} \frac{\cappa_{7}}{\lambda_{8}} \frace{ 9:(0)>9:(0)> f(0)(0)> f(0)(0) > f(0)(0) > f(0)(0) > f(0)(0) f(x1w;1P(w;) > f(x1w;)P(w;) -> (f(x1w;)P(w;) > f(x1w;)P(w;) from part a -> P(w) (2) > 1- 2r xp(m) P(w) (x) P(m) x1-2r) P(m) P(x/w;)P(v;) > (1-2r)P(x) - P(x) = [P(x/w;)P(w;) P(2/w;)P(w;) > (1-25) [P(x/w;)P(w) -> i=C+1 => $q_{i}(n)$ = $\begin{cases} f(x_{i}|u_{i}|P(u_{i})) & i=1,1,...,c \\ (1-\frac{\lambda_{i}}{\lambda_{i}}) \sum_{j} f(x_{j}|u_{j}|P(u_{j})) & i=C+1 \end{cases}$ P(w,) = P(w2) = { (x/w,) = N(1,1) fx(x/w2) ~ N(-1,1) $g_{1}(x)$, $f(x|\omega_{1})P(\omega_{1})$, $\frac{1}{\sqrt{2\pi}}e^{-\frac{((x-1)^{2})}{2\sqrt{2\pi}}}e$ is3-) 93(x1) 2, (1-4) = f(x/w;)P(w;) = 3 x 1 (e - (x-1)2 + e 2)

How does $\frac{\lambda r}{\lambda s}$, change it starts at $\frac{\lambda r}{\lambda s}$ so \rightarrow It means that rejecting has no risk and cost It means that the gen case is bigger the the other g. witch means it's the best case \rightarrow therfor in this case it's better to reject while the $\frac{\lambda r}{\lambda s}$ increases to $1 \rightarrow$ in some coses it would be better to not take a risk and in some case $g_i(x)$ might be bigger then $g_{cu}(x) \rightarrow if \frac{\lambda r}{\lambda s} in this case <math>g_{cu}(x)$ are all smaller than $g_i(x)$ so we should take any risk afall



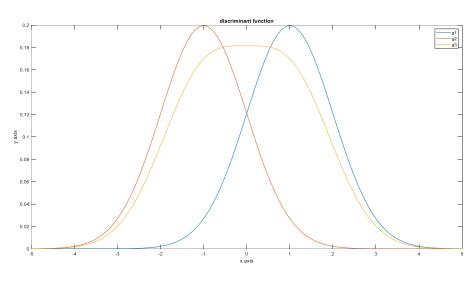


Figure 11

The Figure 11 shows all the $g_i(x)$ from $g_1(x)$ to $g_3(x)$ and the calculation are mentioned in the previous page

Problem 9.

a) Risk =
$$\frac{1}{|\alpha|} \int_{R_1}^{\infty} dx \int_{A_1}^{\infty} f(x_1|\alpha_1) f(x_1|\alpha_2) dx + \int_{R_2}^{\infty} \lambda_{21} f(\alpha_1) f(\alpha_1|\alpha_1) dx$$

Risk + $\int_{R_1}^{\infty} \lambda_{11} f(\alpha_1) f(\alpha_1|\alpha_2) dx + \int_{R_2}^{\infty} \lambda_{21} f(\alpha_1) f(\alpha_1|\alpha_1) dx$

- $\int_{R_2}^{\infty} f(\alpha_1) f(\alpha_1|\alpha_2) d\alpha_1 \longrightarrow \int_{A_1}^{\infty} f(\alpha_1|\alpha_2) d\alpha_1 + \int_{R_2}^{\infty} \lambda_{12} f(\alpha_1) f(\alpha_1|\alpha_2) d\alpha_2$

- $\int_{R_2}^{\infty} f(\alpha_1|\alpha_2) f(\alpha_1|\alpha_2) d\alpha_1 \longrightarrow \int_{R_2}^{\infty} f(\alpha_1|\alpha_2) d\alpha_2 + \int_{R_2}^{\infty} f(\alpha_1|\alpha_2) d\alpha_2 \longrightarrow \int_{R_2}^{\infty} f(\alpha_$

b) no if the two distribution where not the or the chance for $P(w_i) \neq P(w_i)$ -> we couldn't get the same result so if we don't have the conditions from the past part statuet would not satisfy the desigion regions.

Reference

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