

Effect of Electricity Consumption & Pricing on Industrial Growth in India



In fulfilment of the course

ECON F266

Study Oriented Project

Submitted by

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To

Dr. Geetilaxmi Mohapatra

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Abstract

This work examines causality between electricity consumption, electricity price & Index of Industrial Productivity(IIP) using VECM for India using monthly time series data from April 2010 to December 2016. This study finds that IIP is caused by electricity consumption and price, whereas electricity consumption is not caused by IIP as well as electricity price. There is a unidirectional causality from electricity consumption to IIP. So, increasing electricity availability will lead to higher industrial growth due to supply as well as demand factors of industrial goods.

1. Introduction

Energy is a key requirement of industrial development in an economy. Electricity is the most prominent for of energy used in production activities and it should effect industrial productivity of an economy. There has been a lot of research in the past regarding effect of electricity consumption on GDP, employment, income etc. Most of the literature uses long term analysis with annual granularity. Here, the multivariate causality between electricity consumption, Index of Industrial Productivity and electricity price has been studied using very recent monthly data from April 2010 to December 2016. Previous work in this area has been briefly discussed in the next section.

2. Literature Review

Ghosh (2002) [1] studies Granger causality between GDP growth and electricity consumption in India using data from 1950-97. It concludes that there is a unidirectional relationship from GDP to electricity consumption with no feedback effect. So, energy conservation policies can be used successfully without reducing economic growth. VAR model has been used for the econometric analysis of the data. GDP/Capita and KWh/Capita are the variables of interest. GDP/capita has been found to have a chi squared distribution with 1 degree of freedom. GDP growth rate over this period has been 4.5% at an average whereas CAGR Electricity Demand growth was 8%.

The reason for this difference is that with rise in GDP, domestic demand of electricity increases as well as more substitution of conventional sources of energy in industry took place.

Asafu-Adjaye (2002) [2] suggests that either employment or growth can be used as a proxy for growth and time series properties of data involved must be considered. Commercial sector energy use has been considered instead of total energy use. This paper looks at energy consumption as a whole, so the unit has been normalized in terms of oil/capita.

Using Augmented Dickey Fuller test, it was found that real income, energy prices and energy consumption are co integrated of order 1 and thus first differencing is good enough to analyze

the effects. Data used was for 1973-95. It found unidirectional effect from energy consumption to income.

Abbas et. Al [3] found out a bidirectional causality between Agricultural Electricity Consumption and Agricultural GDP. India confirmed to conservation hypothesis at an aggregate level. It recommends using supply enhancing strategies which includes increased production balanced with a managed growth in demand. Data set is from 1972 to 2008 and is relatively new compared to previous two papers. It is clear from the data that demand increased faster than the supply in last 40 years. Agriculture accounts for 1/5th of the total electricity demand in India. Dickey Fuller and Augmented Dickey Fuller were used to check stationarity of data. Error Correction Model (ECM) was used to introduce an error correction term and Granger Causality was modelled. Correlation between total GDP and Total electricity consumption was found to be 0.95. While that between per capita values of same was 0.90. ADF suggests that GDP and Electricity consumption at both per capita and aggregate level are first difference stationary. Causality between GDP growth and Electricity Consumption has been found to be from GDP to Electricity Consumption in both aggregate and per capita levels and long as well as short term.

Ghosh (2009) [4] uses VECM to further establish the long- and short-run Granger causality running from real GDP and electricity supply to employment without any feedback effect implying that the growth in real GDP and electricity supply are responsible for the high level of employment in India.

3. Data Description

The variables used are Electricity consumption in terms of Mega Units, Electricity price in rupees per MWh & the IIP with base year 2004-05. Time period of focus is from April 2010 to December 2016 with monthly granularity.

Availability of electricity has been obtained from Load Generation & Balancing Reports of Central Electricity Authority. It is the sum of transmission losses & realized consumption of electricity and hence a good proxy for consumption. The actual losses are proportional to consumption and the remainder is accounted for by power theft which is basically a form of consumption.

Electricity price has been obtained from Indian Energy Exchange where electricity is traded at a national level & hence acts as a good proxy for average electricity price in India. The Indian Energy Exchange is an electronic system based power trading exchange regulated by the Central Electricity Regulatory Commission.

IIP data is use based as per RBI classification and is obtained from RBI database.

4. Analysis

The basic methodology used has the following steps:

1. Deseasonalizing the data using X12 ARIMA

2. Augmented Dickey Fuller Test for unit root
3. AIC Test for lag selection
4. Johansen Cointegration Test
5. VECM test

4.1. Deseasonalization

Data for electricity consumption, price as well as IIP follow a seasonal trend. Electricity consumption is affected by seasonal factors like temperature, harvest season, monsoon etc. IIP is also affected by these factors along with increased production to meet the demands at certain time periods like festive season. Electricity price also varies seasonally since it's a function of demand & supply of electricity, both of which happen to be seasonal.

X12 ARIMA has been used to deseasonalize the data. It is a widely used standard package developed by US Census Department and implemented in several statistical packages. All seasonal data released by US Census Department is deseasonalized whereas in context of India only the actual values are available.

Logarithmic transformations of these deseasonalized variables have been used for further analysis to reduce the effect of any outliers caused by sudden shocks on supply or demand side.

4.2. Augmented Dickey Fuller Test for Unit Root

The Augmented Dickey Fuller Test (ADF) is a unit root test for stationarity. Unit roots can cause unpredictable results in time series analysis. The Augmented Dickey-Fuller test can be used with serial correlation. The ADF test can handle more complex models than the Dickey-Fuller test, and it is also more powerful. [5]

The hypotheses for the test:

H₀: There is a unit root in the time series

H₁: Time series is stationary (or trend-stationary)

There are three models for ADF which are chosen based on the nature of data. These are:

- a. No constant, no trend: $\Delta y_t = \gamma y_{t-1} + v_t$
- b. Constant, no trend: $\Delta y_t = \alpha + \gamma y_{t-1} + v_t$
- c. Constant and trend: $\Delta y_t = \alpha + \gamma y_{t-1} + \lambda t + v_t$

The results of ADF performed for various variables are listed below:

1. Consumption of Electricity (*logconsa*)

Consumption has been found to have a unit root and follows the model with intercept and no trend term, i.e. model b as described above. As a result, first differenced values have been used for subsequent analysis. These have been confirmed to be stationary under all 3 ADF models

```
. dfuller logconsa, regress lags(0)
```

Dickey-Fuller test for unit root Number of obs = 80

	Test Statistic	----- 1% Critical Value	Interpolated Dickey-Fuller 5% Critical Value	----- 10% Critical Value
Z(t)	-0.556	-3.538	-2.906	-2.588

Mackinnon approximate p-value for Z(t) = 0.8806

D.logconsa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
logconsa					
L1.	-.0129163	.0232184	-0.56	0.580	-.0591406 .0333081
_cons	.0659447	.1138589	0.58	0.564	-.160731 .2926203

2. Electricity Price (logpricea)

Price has been found to have unit root following model a. with no intercept term. As a remedy, the first differences have been used for further analysis

```
. dfuller logiipa, noconstant regress lags(0)
```

Dickey-Fuller test for unit root Number of obs = 80

	Test Statistic	----- 1% Critical Value	Interpolated Dickey-Fuller 5% Critical Value	----- 10% Critical Value
Z(t)	0.438	-2.608	-1.950	-1.610

D.logiipa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
logiipa					
L1.	.0002268	.0005178	0.44	0.663	-.0008037 .0012574

3. Index of Industrial Productivity (logiipa)

IIP has unit root without an intercept term. IIP has also been first differenced to prepare VAR model as discussed below.

```
. dfuller logiipa, noconstant regress lags(0)
```

Dickey-Fuller test for unit root Number of obs = 80

----- Interpolated Dickey-Fuller -----				
	Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	0.438	-2.608	-1.950	-1.610

D.logiipa	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
logiipa						
L1.	.0002268	.0005178	0.44	0.663	-.0008037	.0012574

4.3. Akaike's Information Criterion (AIC) test for lag selection

Akaike's information criterion (AIC) compares the quality of a set of statistical models to each other. The AIC will take each model and rank them from best to worst. The "best" model will be the one that neither under-fits nor over-fits. Performing the AIC test on logconsa, logiipa & logiipa, it is observed that suitable number of lags is 2. The results from AIC test are as follows.

```
. varsoc logconsa logpricea logiipa, maxlag(12)
```

Selection-order criteria								
Sample: 1961m2 - 1966m10					Number of obs = 69			
lag	LL	LR	df	p	FPE	AIC	HQIC	SBIC
0	426.946				9.2e-10	-12.2883	-12.2498	-12.1912
1	562.149	270.41	9	0.000	2.4e-11	-15.9463	-15.7922*	-15.5578*
2	572.272	20.247	9	0.016	2.3e-11*	-15.9789*	-15.7092	-15.299
3	576.094	7.6431	9	0.570	2.7e-11	-15.8288	-15.4434	-14.8575
4	580.266	8.3431	9	0.500	3.1e-11	-15.6889	-15.1879	-14.4261
5	584.193	7.8545	9	0.549	3.7e-11	-15.5418	-14.9252	-13.9877
6	590.334	12.282	9	0.198	4.1e-11	-15.459	-14.7268	-13.6134
7	595.801	10.933	9	0.280	4.6e-11	-15.3565	-14.5087	-13.2196
8	604.868	18.135	9	0.034	4.8e-11	-15.3585	-14.3951	-12.9301
9	614.251	18.765	9	0.027	4.9e-11	-15.3696	-14.2906	-12.6498
10	629.126	29.75	9	0.000	4.4e-11	-15.5399	-14.3452	-12.5287
11	636.32	14.388	9	0.109	5.0e-11	-15.4875	-14.1773	-12.1849
12	647.138	21.636*	9	0.010	5.2e-11	-15.5402	-14.1144	-11.9462

Endogenous: logconsa logpricea logiipa
Exogenous: _cons

4.4 Johansen Cointegration Test

Johansen Cointegration Test has been used to find the order of Cointegration in the model. The model is cointegrated of order one according to the results below.

```
. vecrank logconsa logpricea logiipa, trend(constant) max
```

Johansen tests for cointegration					
Trend: constant				Number of obs = 79	
Sample: 1960m4 - 1966m10				Lags = 2	

maximum				trace	5%
rank	parms	LL	eigenvalue	statistic	critical
0	12	634.06601	.	35.2344	29.68
1	17	644.71474	0.23631	13.9370*	15.41
2	20	651.64815	0.16099	0.0702	3.76
3	21	651.68323	0.00089		

maximum				max	5%
rank	parms	LL	eigenvalue	statistic	critical
0	12	634.06601	.	21.2975	20.97
1	17	644.71474	0.23631	13.8668	14.07
2	20	651.64815	0.16099	0.0702	3.76
3	21	651.68323	0.00089		

4.5. Vector Error Correction Model

VECM is used to find causality in cointegrated variables. Three different VECM models have been considered for each of the three variables.

I. Dependent variable as logconsa & explanatory variables as logiipa and logpricea

The VECM model obtained has a non-significant error term and hence there is no long run causality from electricity price and IIP to electricity consumption. Also there is no short run causality from either electricity price or price to consumption as shown below. This can be due to the fact that electricity availability that has been used as a proxy for electricity consumption, depends on government policy and resources available to the country. It is quite possible that developments like coal scam & transition of centre government during period of study might have affected growth in electricity availability.

```
. vec logconsa logpricea logiipa, trend(constant)
```

Vector error-correction model

Sample: 1960m4 - 1966m10 No. of obs = 79

Log likelihood =	644.7147	AIC	= -15.89151
Det(Sigma_ml) =	1.64e-11	HQIC	= -15.68724
		SBIC	= -15.38163

Equation	Parms	RMSE	R-sq	chi2	P>chi2
D_logconsa	5	.010983	0.2582	25.75104	0.0001
D_logpricea	5	.057993	0.1343	11.47621	0.0427
D_logiipa	5	.008421	0.3820	45.73957	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
D_logconsa						
_ce1						
L1.	-.0452799	.0407985	-1.11	0.267	-.1252434	.0346836
logconsa						
LD.	-.4407167	.1230534	-3.58	0.000	-.6818968	-.1995365
logpricea						
LD.	.003665	.0229349	0.16	0.873	-.0412866	.0486165
logiipa						
LD.	-.0791798	.1626831	-0.49	0.626	-.3980327	.2396732
_cons	.0041949	.0013012	3.22	0.001	.0016446	.0067452
D_logpricea						
_ce1						
L1.	-.5828339	.2154159	-2.71	0.007	-1.005041	-.1606264
logconsa						
LD.	-.5618655	.6497219	-0.86	0.387	-1.835297	.711566
logpricea						
LD.	.0662284	.1210962	0.55	0.584	-.1711159	.3035726
logiipa						
LD.	-.1245087	.8589668	-0.14	0.885	-1.808053	1.559035
_cons	-.000223	.0068703	-0.03	0.974	-.0136885	.0132425
D_logiipa						
_ce1						
L1.	.1178796	.0312809	3.77	0.000	.0565701	.1791891
logconsa						
LD.	-.0254508	.0943472	-0.27	0.787	-.210368	.1594664
logpricea						
LD.	.0006252	.0175846	0.04	0.972	-.03384	.0350903
logiipa						

LD.		-.2451641	.1247321	-1.97	0.049	-.4896345	-.0006938
_cons		.0005088	.0009976	0.51	0.610	-.0014465	.0024642

Cointegrating equations

Equation	Parms	chi2	P>chi2
-----	-----	-----	-----
_ce1	2	120.3633	0.0000
-----	-----	-----	-----

Identification: beta is exactly identified

Johansen normalization restriction imposed						
beta		Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
-----	-----	-----	-----	-----	-----	-----
_ce1						
logconsa		1
logpricea		.1159959	.0835405	1.39	0.165	-.0477405 .2797322
logiipa		-4.434261	.4949058	-8.96	0.000	-5.404258 -3.464263
_cons		4.628902

```
. test ([D_logconsa]: LD.logiipa)
```

```
( 1) [D_logconsa]LD.logiipa = 0
```

```
      chi2( 1) =    0.24
      Prob > chi2 =    0.6265
```

```
. test ([D_logconsa]: LD.logpricea)
```

```
( 1) [D_logconsa]LD.logpricea = 0
```

```
      chi2( 1) =    0.03
      Prob > chi2 =    0.8730
```

II. Dependent variable as logiipa & explanatory variables as logpricea and logconsa

There is long run causality from electricity consumption & price to IIP as coefficient of cointegrated equation is negative and significant. It gives the speed of adjustment of IIP to the dependent variables. This indicates a unidirectional long run causality running from electricity consumption to Index of Industrial Productivity. So, improving the availability of power to industries can significantly boost industrial output & growth of new industries.

Also electrification of the country leads to higher demand for industrial goods by stimulating the economic activities & creating new market for industrial goods. Thus, there is both a demand pull & supply push for industrial output due to electricity consumption.

Government must focus on increasing electricity production & improving transmission under the *Make in India* initiative. Electricity is the key input for manufacturing sector and its availability needs to be improved.

It is worth noting that there is no short run causality from either electricity consumption or electricity price to IIP. So, effects of improved electricity production on industry might not be immediate due the delay caused by setting up of new industries and gradual growth in new markets for industrial output.

```
. vec logiipa logconsa logpricea, trend(constant)
```

Vector error-correction model

Sample:	1960m4 - 1966m10	No. of obs	=	79
		AIC	=	-15.89151
Log likelihood	= 644.7147	HQIC	=	-15.68724
Det(Sigma_ml)	= 1.64e-11	SBIC	=	-15.38163

Equation	Parms	RMSE	R-sq	chi2	P>chi2
D_logiipa	5	.008421	0.3820	45.73957	0.0000
D_logconsa	5	.010983	0.2582	25.75104	0.0001
D_logpricea	5	.057993	0.1343	11.47621	0.0427

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
D_logiipa					
_ce1					
L1.	-.5227089	.1387078	-3.77	0.000	-.7945711 -.2508467
logiipa					
LD.	-.2451641	.1247321	-1.97	0.049	-.4896345 -.0006938
logconsa					
LD.	-.0254508	.0943472	-0.27	0.787	-.210368 .1594664
logpricea					
LD.	.0006252	.0175846	0.04	0.972	-.03384 .0350903
_cons	.0005088	.0009976	0.51	0.610	-.0014465 .0024642
D_logconsa					
_ce1					
L1.	.2007828	.180911	1.11	0.267	-.1537963 .5553618
logiipa					

LD.	-.0791798	.1626831	-0.49	0.626	-.3980327	.2396732
logconsa						
LD.	-.4407167	.1230534	-3.58	0.000	-.6818968	-.1995365
logpricea						
LD.	.003665	.0229349	0.16	0.873	-.0412866	.0486165
_cons	.0041949	.0013012	3.22	0.001	.0016446	.0067452

D_logpricea						
_ce1						
L1.	2.584437	.9552103	2.71	0.007	.7122594	4.456615
logiipa						
LD.	-.1245087	.8589668	-0.14	0.885	-1.808053	1.559035
logconsa						
LD.	-.5618655	.6497219	-0.86	0.387	-1.835297	.711566
logpricea						
LD.	.0662284	.1210962	0.55	0.584	-.1711159	.3035726
_cons	-.000223	.0068703	-0.03	0.974	-.0136885	.0132425

Cointegrating equations

Equation	Parms	chi2	P>chi2

_ce1	2	71.91566	0.0000

Identification: beta is exactly identified

Johansen normalization restriction imposed

beta	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	

_ce1						
logiipa	1
logconsa	-.2255167	.0277307	-8.13	0.000	-.279868	-.1711655
logpricea	-.026159	.0185208	-1.41	0.158	-.0624591	.0101411
_cons	-1.043895

. test ([D_logiipa]: LD.logconsa)

(1) [D_logiipa]LD.logconsa = 0

chi2(1) = 0.07
Prob > chi2 = 0.7873

. test ([D_logiipa]: LD.logpricea)

```
( 1) [D_logiipa]LD.logpricea = 0
```

```
      chi2( 1) =      0.00
Prob > chi2 =      0.9716
```

III. Dependent variable as logpricea & explanatory variables as logiipa and logconsa

Long run causality from IIP & electricity consumption to electricity price is indicated by a negative and significant coefficient of cointegrated equation. This can be due to inflationary causes. But there is no short run causality from either IIP or electricity consumption to price illustrating the relatively fixed nature of electricity prices in short run.

```
. vec logpricea logconsa logiipa, trend(constant)
```

Vector error-correction model

```
Sample: 1960m4 - 1966m10                No. of obs   =          79
                                         AIC           = -15.89151
Log likelihood = 644.7147                HQIC          = -15.68724
Det(Sigma_ml)  = 1.64e-11                SBIC          = -15.38163
```

Equation	Parms	RMSE	R-sq	chi2	P>chi2
D_logpricea	5	.057993	0.1343	11.47621	0.0427
D_logconsa	5	.010983	0.2582	25.75104	0.0001
D_logiipa	5	.008421	0.3820	45.73957	0.0000

	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]
D_logpricea					
_ce1					
L1.	-.0676063	.0249874	-2.71	0.007	-.1165806 -.018632
logpricea					
LD.	.0662284	.1210962	0.55	0.584	-.1711159 .3035726
logconsa					
LD.	-.5618655	.6497219	-0.86	0.387	-1.835297 .711566
logiipa					
LD.	-.1245087	.8589668	-0.14	0.885	-1.808053 1.559035
_cons	-.000223	.0068703	-0.03	0.974	-.0136885 .0132425
D_logconsa					
_ce1					
L1.	-.0052523	.0047325	-1.11	0.267	-.0145277 .0040232
logpricea					
LD.	.003665	.0229349	0.16	0.873	-.0412866 .0486165

logconsa						
LD.	-.4407167	.1230534	-3.58	0.000	-.6818968	-.1995365
logiipa						
LD.	-.0791798	.1626831	-0.49	0.626	-.3980327	.2396732
_cons	.0041949	.0013012	3.22	0.001	.0016446	.0067452

D_logiipa						
_ce1						
L1.	.0136735	.0036285	3.77	0.000	.0065619	.0207852
logpricea						
LD.	.0006252	.0175846	0.04	0.972	-.03384	.0350903
logconsa						
LD.	-.0254508	.0943472	-0.27	0.787	-.210368	.1594664
logiipa						
LD.	-.2451641	.1247321	-1.97	0.049	-.4896345	-.0006938
_cons	.0005088	.0009976	0.51	0.610	-.0014465	.0024642

Cointegrating equations

Equation	Parms	chi2	P>chi2

_ce1	2	17.96452	0.0001

Identification: beta is exactly identified

Johansen normalization restriction imposed

beta	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
<hr/>						
_ce1						
logpricea	1
logconsa	8.620996	2.2813	3.78	0.000	4.149731	13.09226
logiipa	-38.22775	9.026229	-4.24	0.000	-55.91883	-20.53666
_cons	39.90575
<hr/>						

. test ([D_logpricea]: LD.logconsa)

(1) [D_logpricea]LD.logconsa = 0

chi2(1) = 0.75
Prob > chi2 = 0.3872

. test ([D_logpricea]: LD.logiipa)

(1) [D_logpricea]LD.logiipa = 0

chi2(1) = 0.02
Prob > chi2 = 0.8847

5. Conclusion

From the analysis above it can be inferred that in the period from 2010-2016, there is a unidirectional causation from Electricity Consumption & Price to Index of Industrial Production. Further, electricity price does not affect its consumption indicating inelastic demand schedule for electricity.

Electricity consumption leads to higher industrial productivity owing to increased power consumption in industry as well as high household demand for industrial goods which require electricity to operate. This can be seen as a supply push effect of electricity consumption on IIP.

Increasing production as well as improving distribution networks can put India on path to growth. Electrification of the entire country will lead to higher demand for industrial products as well as start of new industries at various scales. Thus, policy makers must focus on increasing electricity availability to the country. Also, energy conservation measures like use of LEDs for lighting should be encouraged as it increases productivity per unit electricity.

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